

Chesapeake Bay Finfish Investigations

US FWS FEDERAL AID PROJECT<br>F-61-R-8<br>2011-2012<br>

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# UNITED STATES DEPARTMENT OF THE INTERIOR <br> FISH AND WILDLIFE SERVICE PERFORMANCE REPORT 

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## Executive Summary

The primary objective of the Chesapeake Bay Finfish Investigations Survey was to monitor and biologically characterize resident and migratory finfish species in the Maryland portion of the Chesapeake Bay. This Survey provides information regarding relative abundance, age and size structure, recruitment, growth, mortality, and migration patterns of finfish populations in Maryland's Chesapeake B ay. The da ta generated are utilized in both intrastate and interstate management processes and provides a reference point for future fisheries management considerations.

The Head-of-Bay (HOB) channel catfish population was assessed with a surplus production model covering the years, 1980-2011, and the Choptank River channel catfish stock was assessed with a catch survey a nalysis (CSA; 1993 -- 2011). T rends in channel catfish populations from Nanticoke, Patuxent, and Potomac rivers were described from available relative abundance indices.

The HOB channel catfish assessment utilized a fishery dependent relative abundance index, and two fishery independent relative abundance indices, a gill net survey, and a trawl survey. The model fit the data well, but as usual, biomass and mortality ratios were more precisely estimated than absolute biomass and mortality estimates. $\mathrm{B}: \mathrm{B}_{\mathrm{msy}}$ ratio in the final year was 1.55 , which indicates
that the stock is not overfished, and $F: F_{\text {msy }}$ ratio in 2011 was 0.86 , which indicates that overfishing was not occurring. However, fishing mortality has trended upward since 2009.

The CSA model fit the population data moderately well. Pre-recruit (channel catfish < 355 mm ) population abundance generally tracked the increase in the survey's relative abundance values, with relatively low pre-recruit abundance during 1995-2004, followed by relatively high pre-recruit abundance for the remainder of the time series. Post-recruit channel catfish ( $>355 \mathrm{~mm}$ ) abundance varied between 200,000 a nd 400,000 c hannel catfish from 1993 - 2007. A fter 2007, r ecruited channel catfish abundance accelerated quite swiftly with the recruited population increasing from an estimated 664,000 fish in 2008 to 1.06 million fish in 2011. Instantaneous fishing mortality ( F ) was generally low, varying between $\mathrm{F}=0.04$ and 0.15 for most of the assessment period. Average F for the entire time series was $\mathrm{F}=0.13$ and F in the final year of the assessment was 0.11 . Model outputs and survey results strongly suggest that fishing mortality at recent levels is not impacting population growth.

Relative abundance indices from ot her river s ystems w ere 1 argely i nconclusive, but populations appear to be stable. Nanticoke River commercial fishery CPUE's were quite variable and exhibited no discernable trend. Young-of-year production, as determined from a seine survey was also not definitive, but production was more consistent during 1989-1997 than in recent years. Patuxent River channel catfish landings have been trendless throughout the past 25 years. Only the fish pot relative abundance index provided a complete enough time series to warrant investigation, and it has been trending downward since 2006. Young-of-year production, as determined from a seine survey indicated that the last years of high juvenile abundance were in 2001 and 2003. The Potomac River drift gill net survey indicated that the biomass index was below the $75^{\text {th }}$ percentile since 2005. Y oung-of-year production, as determined from a seine survey indicated low a nd intermittent juvenile production since 1985. C ommercial landings have been relatively stable at lower levels since 2002.

Populations of American shad in Maryland continue to be impacted by predation, by-catch and $t$ urbine $m$ ortality. $T$ he $s$ urplus pr oduction $m$ odel popul ation e stimate of A merican shad abundance in the lower Susquehanna River exhibited no significant trend over the times series (1986-2012), but population abundance has been steadily increasing since 2007. Estimates of hook and line GM CPUE vary without trend over the time series in the lower Susquehanna River (19842012) and the Potomac River (1996-2012). In the Nanticoke River, GM CPUE was the highest in the time series (1988-2012). The percentage of repeat spawners continues to increase in the lower Susquehanna and Nanticoke rivers. Juvenile American shad indices have improved in the Potomac River, but generally remain low in Maryland tributaries.

The age structure of hickory shad in a Susquehanna River tributary remains consistent, with a wide range of ages and a high percentage of older fish. The arcsine-transformed proportion of these repeat spawners (sexes combined) has not changed significantly over the time series (2004-2012; $r^{2}$ $=0.028, P=0.67$; Figure 18), although the total percentage of repeat spawners in $2012(64.0 \%)$ was the lowest total percent of repeat spawners of the time series (2004-2012).

According to the most recent ASMFC stock assessment, the coast wide meta-complex of river herring stocks on the U.S. Atlantic coast is depleted to near historic lows. River herring age structure in the Nanticoke River appears to be truncating, especially for blueback herring. Observed declines in length-at-age generally occur toward the end of the time series. The GM CPUE for juvenile al ewife and blueback he rring de creased in 2012 i n all M aryland tributaries. D ue to Amendment 2 to the ASMFC FMP for American shad and river herring, it is not legal to harvest river herring within the jurisdiction of Maryland. This moratorium on river herring should promote an increased spawning stock, leading to increased production of juvenile river herring.

Weakfish have experienced a sharp decline in abundance coast wide. Recreational catch estimates by the NMFS for Maryland declined from 475,348 fish in 2000 to 237 fish in 2011. Maryland's commercial weakfish harvest declined to 378 pounds in 2011, and was the lowest catch on record. The 2012 mean length for weakfish from the onboard pound net survey was 284 mm TL. The 2011 length frequency distribution and RSD analysis indicate that only smaller weakfish were available in Maryland waters. The charter boat CPUE has significantly declined from 1993-2011.

Summer flounder mean length from the pound net survey was 338 mm TL in 2012, the ninth lowest mean value in the 20 year survey. Relative stock densities in the 2012 onboard pound net survey indicated a slight decrease in the stock and quality categories with a corresponding increase in the preferred category compared to 2011. Charter boat CPUE has declined from 1993-2003, but has been relatively stable for the past eight years. The NMFS 2011 coast wide stock assessment concluded that summer flounder stocks were not overfished, overfishing was not occurring and the rebuilding target has been met as of 2010 .

Mean length of bluefish from the pound net survey in 2012 was 298 mm TL, less than the time series mean. Length distribution and RSD analysis indicated a slight shift to larger bluefish in 2012. Recreational and commercial bluefish harvest increased in 2011, but still remained below the long te rm me an. The 2011 coast $w$ ide s tock a ssessment upda te i ndicated the stock was not overfished and overfishing is not occurring.

The mean length of Atlantic croaker examined from the pound net survey in 2012 was 274 mm TL; this was the third lowest value of the 20 year time series. For Atlantic croaker from the onboard pound net survey the $\mathrm{RSD}_{\text {perferred }}$ category decreased, with a corresponding slight increase occurring in the remaining RSD categories. Maryland Atlantic croaker total commercial harvest increased in 2011 to 704,019 pounds; while the 2011 recreational harvest estimated of 554,206 fish decreased compared to 2010. Compared to 2010, the 2011 charter boat geometric catch per angler decreased to 4.66 fish per angler, but was still above the long term mean.

Spot mean length decreased in 2012 and was the lowest value on record. The spot juvenile index spiked to the time series high in 2010, declined to near the time series low in 2011, but rose in 2012 to the eighth highest value in the 24 year time series. Commercial harvests increased sharply in 2009 and remained high through 2011, while the recreational estimate dropped well bellow the time series mean. The charter boat geometric mean catch per angler increased in 2011, but was still the fourth lowest value of the 19 year time series.

Resident / premigratory striped bass sampled in the Chesapeake Bay during the summer - fall 2011 pound net and hook and line commercial fisheries ranged from 1 to 13 years of age. Three year old (2008 year-class), four year old (2007 year-class), five year old (2006 year-class), six year old (2005 year-class) and seven year old (2004 year-class) striped bass dominated biological samples taken from pound nets. These five year-classes com prised $88 \%$ of the sample. C heck station sampling de termined that the majority of the pound ne $t$ and hook-and-line fishery harvest was composed of four to seven year old individuals from the 2004 through 2007 year-classes.

The 2011-2012 commercial striped bass drift gill net fishery harvest was comprised primarily of fish 4, 5, 6 and 7 years old from the 2005 through 2008 year-classes. Striped bass from the 2007 year-class (five year old fish) composed $47 \%$ of the total drift gill net harvest. The 2007 and 2006 (ages 5 and 6) cohorts accounted for $71 \%$ of the total harvest while age groups 8 to 11 year-old fish contributed $2 \%$ to the total. Striped bass present in commercial drift gill net samples collected from check stations ranged in age from age 3 to 11 (2001 to 2009 year-classes).

Fish harvested during the 2011-2012 Atlantic coast fishing season ranged from age 4 (2008 year-class) to age 21 ( 1991 year-class). Fourteen year-classes were represented in the sampled harvest. Approximately $72 \%$ of striped bass harvested were ages 7 through 10 . Striped bass were recruited into the Atlantic coast fishery as young as age 4, but due to the 24 inch minimum size limit, few fish younger than age 6 w ere harvested, which is similar to previous years. B ased on the estimated catch-at-age, the most common age harvested during the 2011-2012 Atlantic coast harvest was age 9 (2003 year-class), which represented $34 \%$ of the fishery. Large contributions were also made by the 2004 year-class (age 8) and the 2005 year- class (age 7), which represented $16 \%$ and $13 \%$ of the fishery, respectively.

The spring, 2012 spawning stock survey indicated that there were 18 age-classes of striped bass present on the Potomac River and Upper Bay spawning grounds. These fish ranged in age from 2 to 19 years old. Male striped bass ranged in age from 2 to 15 years old, with age 8 and age 9 fish (2004 and 2003 year-classes) being the most abundant component of the male striped bass spawning stock. The majority of females were ages 9 to 14, with most female collected at age 9 (2003 yearclass). During the spring, 2012 spawning season, the contribution of age 8 and older females to the female spawning stock increased to $80 \%$.

The 2012 striped bass juvenile index, the annual measure of striped bass spawning success in Chesapeake Bay, was 0.9 , the lowest index measured in survey's 59 -year history. T his is significantly lower than the long-term average juvenile index of 12.0. Highly variable spawning success is a hallmark of striped bass populations. Typically, several years of average reproduction are i nterspersed with occasional 1 arge and small year-classes. S pawning s uccess is he avily influenced by e nvironmental conditions such as flow rates and water temperature. In 2011, biologists documented one of the most successful striped bass spawns on record and these 1-year-old fish are very abundant. The successful spawning years of 1989, 1996, and 2001 were also followed by below-average or poor years.

Several ot her s pecies of a nadromous fish, such as white perch, yellow perch, and river herring, experienced low r eproductive success in 2012, poi nting to 1 arge-scale envi ronmental conditions as the probable cause because warm winters and dry springs are unfavorable spawning conditions for anadromous fish. However, the survey documented increased reproduction of species that spawn offshore or in higher salinity bay water, like Atlantic croaker and bay anchovies. During this year's survey, biologists identified and counted more than 31,000 fish of 54 different species. DNR biologists have monitored the r eproductive success of striped ba ss a nd ot her s pecies in Maryland's portion of the Chesapeake Bay annually since 1954

During the 2012 trophy season, biologists intercepted 209 fishing trips, interviewed 447 anglers, and examined 130 striped bass. The average total length of striped bass sampled was 863 mm total length ( mm TL) ( 34.0 inches). The average weight was 6.7 kg ( 14.7 lbs ). Striped bass sampled from the trophy fishery ranged in age from 5 to 17 years old. The 2003 year-class (age 9) and 2004 year-class (age 8 ) were the most frequently observed cohorts. Average catch rate based on angler interviews was 0.2 fish per hour.

Maryland Department of Natural Resources biologists continued to tag and release striped bass in 2012 in support of the US FWS coordinated interstate, coastal population study for growth and mortality. A total of 688 striped bass were tagged and released with USFWS internal anchor tags. Of this sample, 682 were tagged in the Potomac River and the upper Chesapeake Bay area during the spring spawning stock assessment survey. A total of 6 striped bass were tagged during an abbreviated cooperative USFWS / SEAMAP Atlantic Ocean tagging cruise.

## APPROVAL

Michael Luisi, Assistant Director Estuarine \& Marine Fisheries Division<br>Maryland Fisheries Service<br>Maryland Department of Natural Resources

## ACKNOWLEDGEMENTS

The M aryland Department of N atural R esources (MD DNR) would like to thank the Maryland Watermen's Association commercial captains and their crews who allowed us to sample their commercial catches. We also wish to thank RMC Environmental Services personnel for their aid in acquiring tag returns and catch data from the fish lifts at Conowingo Dam. Appreciation is also extended to MD DNR Hatchery personnel, Brian Richardson and staff for otolith analysis of juvenile a nd a dult A merican s had and to C onnie L ewis, Fisheries $S$ tatistics, for pr oviding commercial landings. We would also like to thank Captain John Collier and crew of the R/VLaidly, for their assistance during the winter trawl survey.

Striped bass were collected for portions of this study from commercial pound nets owned and operated by Maryland Watermen's Association commercial captains and their crews. Striped bass were collected from the Atlantic Ocean trawl and gill net fisheries by Gary Tyler and Steve Doctor. Experimental drift gill nets were operated by Robert Boarman, on the Potomac River, and Michael Cannan, on the Upper Chesapeake Bay.

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## PROJECT NO. 1

JOB NO. 1

# POPULATION VITAL RATES OF RESIDENT FINFISH IN SELECTED TIDAL AREAS OF MARYLAND'S CHESAPEAKE BAY 

Prepared by Paul G. Piavis and Edward Webb, III

## INTRODUCTION

The primary objective of Job 1 was to provide data and analysis from routine monitoring of $t$ he $f$ ollowing $r$ esident $s$ pecies: $w$ hite pe rch (Morone americana), yellow p erch ( Perca flavescens), channel c atfish (Ictalurus punctatus) and white cat fish (Ameiurus catus) from selected tributaries in the Maryland portion of the Chesapeake Bay. In order to update finfish population assessments and management plans, data on popul ation vital rates should be current and c learly de fined. P opulation vi tal r ates include g rowth, m ortality, a nd r ecruitment. Efficiency is of ten 1 acking when upd ating or initiating as sessments be cause da ta a re r arely compiled a nd s ynopsized in one convenient s ource. D ata collected in a n a ntecedent s urvey (MULTIFISH, F-54-R) have proved invaluable in compiling technical reports and providing the basis for sound $m$ anagement $r$ ecommendations for $t$ hese $s$ pecies. $T$ his $j$ ob will e nhance $t$ his efficiency by detailing current results of routine monitoring.

## METHODS

## I. Field Operations

## Upper Chesapeake Bay Winter Trawl

The uppe r C hesapeake Bay winter bot tom trawl s urvey is designed to collect fisheryindependent data for the assessment of population trends of white perch, yellow perch, channel catfish, and white c atfish. Upper Chesapeake B ay was di vided into f our s ampling a reas; Sassafras River (S AS), Elk River (EB), upper Chesapeake B ay (UB), and middle Chesapeake

Bay (MB). E ighteen sampling stations, e ach a pproximately 2.6 km ( 1.5 miles ) in 1 ength and variable in width, were created throughout the study area (Figure 1). Each sampling station was divided into west/north or east/south halves by drawing a line parallel to the shipping channel. Sampling depth was divided into two strata; shallow water ( $<6 \mathrm{~m}$ ) and deep water ( $>6 \mathrm{~m}$ ). Each site vi sit w as t hen r andomized f or de pth s trata a $n d \mathrm{t}$ he nor $\mathrm{th} /$ south or e ast/west di rectional components.

The winter trawl s urvey employed a 7.6 m wide bottom t rawl c onsisting of 7.6 c m stretch-mesh in the wings and body, 1.9 c m stretch-mesh in the cod end and a 1.3 cm stretchmesh liner. Following the 10 -minute tow at approximately 3 knots, the trawl was retrieved into the boat by winch and the catch emptied into either a cul ling board or large tub if catches were large. A minimum of 50 fish per species were sexed and measured. Non-random samples of
 determination. A ll s pecies caught were identified and counted. If c atches were prohibitively large $t$ o pr ocess, t otal num bers w ere extrapolated f rom vol umetric c ounts. V olumetric subsamples were taken from the top of the tub, the middle of the tub, and the bottom of the tub. Six sampling rounds were scheduled from early December 2011 through February 2012.

Trawl sites ha ve be en c onsistent t hroughout t he s urvey, but w eather a nd ope rational issues caus ed incomplete s ampling in some years. The 2003 s urvey was ha mpered byice conditions such that only one of six rounds was completed. Retirement of the captain of the R/V Laidly dur ing 20041 ed to no $r$ ounds be ing completed. O nly $1-1 / 2$ rounds of the scheduled six rounds were completed in 2005 be cause of catastrophic engine failure. Ice-cover prevented the final two rounds of the 2007 s urvey and on e round of the 2009 from being completed. Ice conditions also affected the 2010 and 2011 sample years where only 56 and 66 of the scheduled 108 trawls w ere com pleted, respectively. During 2012, 107 of the scheduled 108 ha uls w ere completed.

## Choptank River Fishery Independent Sampling

In 2012 , six experimental fyke nets were set in the Choptank River to sample the four resident species from this system. Nets were set at river kilometers 63.6, 65.4, 66.6, 72.5, 74.4 and 78.1 and were fished two to three times per week from 23 February through 2 April (Figure 2). These nets contained a 64 mm stretch-mesh body and 76 mm stretch-mesh in the wings ( 7.6 m long) and leads ( 30.5 m long). Nets were set perpendicular to the shore with the wings at $45^{\circ}$ angles.

Net hoops were brought aboard first to ensure that all fish were retained. Fish were then removed and placed into a tub and identified. All yellow perch and a subsample of up to 30 fish of each target s pecies were s exed and measured. A 11 non -target s pecies w ere count ed and released. O toliths f rom a s ubsample ofw hite a nd yellow pe rch were r emoved f or a ge determination.

## Upper Chesapeake Bay Fishery Dependent Sampling

Commercial fyke net catches were sampled for yellow perch on 4 February 2012 in and around Middle River, 16 February 2012 in Sassafras River, and 20 February 2012 in Northeast River (Figures 3, 4, 5). A ll yellow pe rch were measured and sexed (unculled) ex cept when catches $w$ ere p rohibitively $l$ arge. A $s$ ubsample $w$ as pur chased $f$ or otolith e xtraction a nd subsequent age determination.

## Nanticoke River Fishery Dependent Sampling

From 22 February 2012 to 30 April 2012, resident species were sampled from pound nets set by commercial fishermen on the N anticoke River. Previous years have included fyke n et samples. This segment of the survey was completed in coordination with Project 2, Job 1 of this grant. Nets were set from B arren C reek ( 35.7 rkm ) dow nstream to Monday's Gut (30.4 rkm; Figure 6). Net sites and dates fished were at the di scretion of the com mercial fishermen. All
yellow perch caught were sexed, measured for total length and a non-random sample of otoliths removed for age determination. Thirty randomly selected white perch from the fyke nets were sexed and measured and a subsample was processed for age determination (otoliths). A bushel of unculled, mixed catfish species was randomly selected, identified as channel or white catfish and total lengths measured.

The 2011 sampling s eason $w$ as severely truncated due to s now and ice conditions. A s such, the yellow perch run had finished before sampling was initiated. In addition, sample sizes for channel catfish and white catfish were also very low.

## II. Data compilation

## Population Age Structures

Population a ge s tructures were determined for yellow perch and white perch from the Choptank and Nanticoke rivers and the upper Chesapeake Bay (trawl and commercial sampling separately). A ge-at-length ke ys for yellow perch and white perch ( separated by sex) from the Choptank River, Nanticoke River, and upper Bay commercial fyke net surveys were constructed by determining the proportion-at-age per $20-\mathrm{mm}$ length group and applying that proportion to the total number-at-length. For the upper Bay trawl survey, an age-length key was constructed in 10 mm increments and the age-at-length key was applied to individual hauls. Total number by sex were added together to get total numbers at age.

## Length-frequency

Relative $s$ tock density (RSD) $w$ as us ed to describe 1 ength structures for $w$ hite $p$ erch, yellow p erch, c hannel c atfish, a nd white c atfish. G ablehouse (1984) advocated i ncremental RSD's to characterize fish length distributions. This method groups fish into five broad length categories: stock, qua lity, pr eferred, $m$ emorable a nd $t$ rophy. $T$ he $m$ inimum 1 ength of e ach category is based on all-tackle world records such that the minimum stock length is $20-26 \%$ of the w orld record length (WRL), mini mum qua lity le ngth is $36-41 \%$ of the W RL, m inimum
preferred length is $45-55 \%$ of the WRL, minimum memorable length is $59-64 \%$ of the WRL and minimum trophy length is $74-80 \%$ of the WRL. M inimum 1 engths w ere a ssigned from either the cut-offs listed by Gablehouse et al (1984) or were derived from world record lengths as recorded by the International Game Fish Association. Current length-frequency histograms were produced for all target species encountered.

## Growth

Growth in length over time and weight in relation to length were described with standard fishery equations. T he allometric growth equation (weight $(\mathrm{g})=\alpha^{*}$ length $(\mathrm{mmTL})^{3}$ ) described weight change as a function of length, and the vonBertalanffy growth equation (Length $=\mathrm{L}_{\infty}\left(1-e^{-}\right.$ $\left.\begin{array}{c}\mathrm{K}\left(\mathrm{t} \mathrm{t}_{0}\right)\end{array}\right)$ described change in length with respect to age. Both equations were fit for white perch and yellow perch males, females, and sexes combined with SAS nonlinear procedures, Excel Solver (Microsoft Corporation 1993), or Evolver genetic tree algorithms (Palisades Corporation 2001). Growth data for target species encountered in the trawl survey were not compiled due to the size selectivity of the gear.

## Mortality

Catch c urves for C hoptank R iver, Nanticoke R iver, a nd uppe r C hesapeake B ay white perch were based on $\log _{e}$ transformed catch-per-unit-effort (CPUE) data for ages 6-10 for males and females. The slope of the line was -Z and M was assumed to be 0.20 . Instantaneous fishing mortality (F) was Z-M.

Choptank River yellow perch mortality was estimated with a ratio method to determine survivorship (S), where $S=($ CPUE ages $4-10+$ in year $t) /($ CPUE ages $3-10+$ in year $t-1)$. Total instantaneous mortality ( Z ) was $-\log _{e}(\mathrm{~S})$, a nd $\mathrm{F}=\mathrm{Z}-\mathrm{M}$ where M was a ssumed to be 0.25 . T he only ex ception to this method was the 2002 e stimate where all a ge-classes w ere us ed for the survivorship estimate. Current Nanticoke River yellow perch rates were not estimated because of unequal recruitment rates, varying annual sample sizes, and an inability to assign associated effort
data to catches. Instantaneous $m$ ortality $r$ ates for yellow $p$ erch $f$ rom u pper Bay commercial samples were calculated with a statistical catch-at-age model (see Project 1, Job2).

## Recruitment

Recruitment data were p rovided from age $1+$ abundance in the winter trawl survey and young-of-year relative abundance from the Estuarine Juvenile Finfish Survey (EJFS; see Project 2, Job2, Task 3 of this report). Cohort splitting was used to determine $1+$ abundance in the winter trawl survey. A ny yellow perch $<130 \mathrm{~mm}$, white perch $<110 \mathrm{~mm}$, and channel catfish $<135$ mm were assumed $1+$. Since white catfish abundance was not well represented in the upper Bay trawl catches, data were not compiled for this species.

Previous yellow perch assessments indicated a suite of selected head-of-bay sites from the EJFS which provided a g ood index of j uvenile a bundance. T herefore, onl y t he H owell Pt ., Ordinary Pt., Tim's Creek, Elk Neck Park, Parlor Pt., and Welch Pt. permanent sites were used to determine the yellow perch juvenile relative abundance index. H owever, the Ordinary Pt. seine site was lost because of bulkhead construction and the replacement site was not included in the index. This index is reported as an average $\log _{e}($ catch +1$)$ index. White perch and channel catfish juvenile $r$ elative abundance was $t$ he geometric $m$ ean (GM) abundance $f$ rom al $l$ ba ywide permanent sites. Sites and methodology are reported in Project 2 Job 3 Task 3 of this report.

## Relative Abundance

Relative abundance of target species was determined as the grand mean abundance from all surveys where reliable effort data were available. For white perch and yellow perch, relative abundance as CPUE at age was determined from the catch-at-age matrices. F yke net effort for yellow pe rch was defined as the a mount of effort ne eded to collect $95 \%$ of each year's c atch.

This is necessary to ameliorate the effects of effort expended to catch white perch after the main yellow perch spawning run. T he CPUE at age matrix included all yellow perch encountered. Prior to 1993, a ll s ampling b egan 1 M arch, but the s tart d ate h as va ried s ince 1993 (usually beginning mid-February). In order to standardize data, CPUE from 1 March to the $95 \%$ catch end time was utilized for time-trend analysis.

## RESULTS

Data are summarized either in tables or figures or ganized by datat ype (age structure, length structure, etc.), species, and survey. Data summaries are provided in these locations:

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White perch
Yellow perch
Channel catfish
White catfish

## Growth

White perch
Yellow perch

## Mortality

White perch
Yellow perch

## Recruitment

White perch
Yellow perch
Channel catfish

## Relative Abundance

White perch
Yellow perch Channel catfish White catfish

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Table 4. Yellow perch catch at age matrix from upper Chesapeake Bay winter trawl survey, 2000-2012.

Table 5. Yellow perch catch at age matrix from Choptank River fyke net survey, 1988 2012.

Table 6. Yellow perch catch at age matrix from upper Chesapeake Bay commercial fyke net survey, 1999 - 2012.

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$5^{3}$

Table 1. White perch catch-at-age matrix from upper Chesapeake Bay winter trawl survey, 2000 2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 2000 | 1,321 | 9,382 | 4,256 | 2,751 | 1,034 | 616 | 845 | 93 | 88 | 55 |
| 2001 | 2,796 | 5,375 | 8,628 | 1,658 | 2,519 | 547 | 1,321 | 1,402 | 324 | 199 |
| 2002 | 17,571 | 150 | 3,670 | 1,516 | 2,359 | 1,006 | 1,947 | 1,067 | 277 | 638 |
| 2003 | 1,655 | 3,123 | 573 | 263 | 365 | 419 | 1,479 | 33 |  | 197 |
| 2004 | NOT SAMPLED |  |  |  |  |  |  |  |  |  |
| 2005 | 973 | 1,684 | 460 | 846 | 216 | 77 | 25 | 242 | 28 | 12 |
| 2006 | 9,597 | 3,172 | 7,589 | 2,283 | 1,680 | 469 | 285 | 281 | 65 | 130 |
| 2007 | 2,521 | 1,699 | 1,229 | 2,408 | 1,387 | 335 | 381 | 30 | 26 | 133 |
| 2008 | 16,173 | 2,715 | 6,995 | 5,269 | 1,654 | 571 | 229 | 252 | 93 | 93 |
| 2009 | 5,838 | 16,227 | 686 | 2,969 | 5,588 | 4,716 | 113 | 1,628 | 344 | 67 |
| 2010 | 4,943 | 2,679 | 4,591 | 159 | 3,205 | 1,184 | 1,963 | 154 | 252 | 388 |
| 2011 | 2,569 | 3,044 | 2,164 | 2,916 | 710 | 1,614 | 884 | 896 | 50 | 153 |
| 2012 | 10,231 | 3,532 | 1,713 | 840 | 873 | 938 | 1,695 | 756 | 1,016 | 304 |

Table 2. White perch catch-at-age matrix from Choptank River fyke net survey, 2000-2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 2000 | 0 | 1 | 1,573 | 9,923 | 9,671 | 1,709 | 6,212 | 576 | 404 | 0 |
| 2001 | 0 | 2,177 | 4,947 | 14,849 | 11,090 | 8,135 | 1,305 | 3,399 | 474 | 0 |
| 2002 | 0 | 650 | 2,390 | 8,708 | 5,007 | 5,626 | 1,065 | 1,883 | 818 | 30 |
| 2003 | 0 | 572 | 9,594 | 8,773 | 8,684 | 364 | 7,217 | 1,881 | 835 | 834 |
| 2004 | 0 | 98 | 9,118 | 3,083 | 3,531 | 4,310 | 325 | 2,401 | 863 | 559 |
| 2005 | 0 | 801 | 3,759 | 12,029 | 7,543 | 4,687 | 1,682 | 397 | 2,531 | 116 |
| 2006 | 0 | 402 | 16,863 | 816 | 8,175 | 4,051 | 440 | 515 | 305 | 4,013 |
| 2007 | 0 | 258 | 1,931 | 25,125 | 2,719 | 11,741 | 4,194 | 1,655 | 1,834 | 1,452 |
| 2008 | 0 | 95 | 5,643 | 4,387 | 13,435 | 1,153 | 4,592 | 2,610 | 478 | 1,048 |
| 2009 | 0 | 369 | 149 | 5,220 | 1,427 | 9,501 | 1,150 | 1,793 | 1,021 | 650 |
| 2010 | 0 | 246 | 4,691 | 730 | 12,145 | 4,258 | 13,037 | 1,617 | 2,170 | 1,155 |
| 2011 | 0 | 21 | 247 | 5,313 | 844 | 5,080 | 3,115 | 3,824 | 553 | 1,027 |
| 2012 | 0 | 25 | 1,190 | 595 | 2,412 | 1,053 | 1,394 | 572 | 1,075 | 289 |

Table 3. White perch catch-at-age matrix from Nanticoke River fyke and pound net survey, 2000 - 2012. 2007 -- 2009 include Marshyhope River data.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 2000 | 0 | 42 | 593 | 6,074 | 6,471 | 2,813 | 1,942 | 365 | 81 | 0 |
| 2001 | 0 | 0 | 681 | 796 | 3,262 | 1,822 | 689 | 785 | 94 | 38 |
| 2002 | 0 | 5 | 1,469 | 1,927 | 504 | 2,124 | 1,132 | 632 | 244 | 135 |
| 2003 | 0 | 97 | 318 | 2,559 | 1,567 | 446 | 994 | 652 | 180 | 175 |
| 2004 | 0 | 6,930 | 3,892 | 12,215 | 3,259 | 1,835 | 1,297 | 1,361 | 443 | 886 |
| 2005 | 0 | 826 | 1,302 | 5,847 | 3,903 | 5,288 | 2,400 | 1,237 | 1,497 | 2,582 |
| 2006 | 0 | 0 | 5,759 | 3,280 | 5,298 | 3,488 | 3,590 | 1,287 | 861 | 799 |
| 2007 | 0 | 497 | 1,948 | 12,876 | 727 | 6,236 | 2,260 | 2,716 | 977 | 1,573 |
| 2008 | 0 | 33 | 902 | 1,188 | 2,780 | 824 | 1,457 | 665 | 593 | 496 |
| 2009 | 0 | 70 | 1,351 | 4,135 | 2,117 | 6,216 | 1,188 | 1,651 | 889 | 1,470 |
| 2010 | 0 | 101 | 273 | 155 | 414 | 315 | 1,113 | 88 | 143 | 166 |
| 2011 | 0 | 933 | 1,625 | 7,817 | 1,167 | 4,433 | 1,750 | 5,133 | 1.050 | 3,034 |
| 2012 | 4 | 134 | 387 | 176 | 539 | 214 | 330 | 57 | 276 | 85 |

Table 4. Yellow perch catch at age matrix from upper Chesapeake Bay winter trawl survey, 2000 - 2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 2000 | 44 | 77 | 13 | 85 | 3 | 15 | 4 | 0 | 0 | 5 |
| 2001 | 669 | 43 | 78 | 12 | 44 | 3 | 0 | 3 | 0 | 0 |
| 2002 | 1,170 | 847 | 83 | 178 | 14 | 86 | 0 | 8 | 4 | 0 |
| 2003 | 343 | 985 | 3,050 | 327 | 437 | 28 | 175 | 0 | 14 | 0 |
| 2004 | NOT SAMPLED |  |  |  |  |  |  |  |  |  |
| 2005 | 446 | 320 | 0 | 70 | 9 | 0 | 0 | 0 | 0 | 0 |
| 2006 | 1,580 | 1,738 | 738 | 0 | 146 | 18 | 0 | 15 | 0 | 0 |
| 2007 | 167 | 150 | 385 | 112 | 71 | 26 | 2 | 0 | 0 | 0 |
| 2008 | 1,053 | 256 | 572 | 504 | 131 | 0 | 0 | 0 | 0 | 0 |
| 2009 | 215 | 1,051 | 54 | 117 | 105 | 23 | 1 | 0 | 0 | 0 |
| 2010 | 862 | 101 | 260 | 18 | 28 | 11 | 6 | 0 | 2 | 0 |
| 2011 | 51 | 185 | 29 | 118 | 0 | 15 | 6 | 0 | 0 | 0 |
| 2012 | 1,138 | 464 | 156 | 6 | 9 | 5 | 0 | 45 | 0 | 0 |

Table 5. Yellow perch catch at age matrix from Choptank River fyke net survey, 1988 - 2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 1988 | 0 | 9 | 268 | 9 | 2 | 21 | 19 | 1 | 1 | 5 |
| 1989 | 0 | 0 | 80 | 234 | 81 | 41 | 8 | 2 | 2 | 0 |
| 1990 | 0 | 22 | 179 | 82 | 273 | 53 | 10 | 8 | 5 | 1 |
| 1991 | 0 | 7 | 41 | 53 | 18 | 44 | 9 | 2 | 2 | 0 |
| 1992 | 0 | 1 | 8 | 14 | 15 | 7 | 6 | 0 | 0 | 0 |
| 1993 | 0 | 3 | 75 | 150 | 98 | 109 | 37 | 7 | 4 | 0 |
| 1994 | 0 | 42 | 158 | 25 | 81 | 87 | 78 | 64 | 5 | 18 |
| 1995 | 0 | 79 | 258 | 23 | 68 | 67 | 42 | 37 | 5 | 21 |
| 1996 | 0 | 857 | 343 | 267 | 35 | 81 | 47 | 27 | 43 | 9 |
| 1997 | 0 | 14 | 641 | 99 | 86 | 0 | 19 | 24 | 8 | 0 |
| 1998 | 0 | 142 | 77 | 583 | 26 | 31 | 0 | 8 | 3 | 17 |
| 1999 | 0 | 306 | 8,514 | 86 | 3,148 | 32 | 9 | 8 | 0 | 6 |
| 2000 | 0 | 329 | 92 | 1,378 | 27 | 140 | 0 | 7 | 0 | 0 |
| 2001 | 0 | 878 | 1,986 | 102 | 1,139 | 19 | 72 | 2 | 0 | 0 |
| 2002 | 0 | 334 | 1,336 | 1,169 | 38 | 430 | 104 | 51 | 3 | 0 |
| 2003 | 0 | 369 | 440 | 922 | 333 | 34 | 226 | 35 | 32 | 2 |
| 2004 | 0 | 60 | 504 | 177 | 120 | 103 | 0 | 61 | 0 | 7 |
| 2005 | 0 | 1,667 | 137 | 416 | 134 | 55 | 140 | 23 | 52 | 15 |
| 2006 | 0 | 173 | 1,858 | 176 | 395 | 64 | 66 | 42 | 0 | 7 |
| 2007 | 0 | 1,512 | 737 | 1,560 | 33 | 182 | 109 | 28 | 10 | 12 |
| 2008 | 0 | 39 | 1,303 | 130 | 326 | 13 | 49 | 20 | 0 | 0 |
| 2009 | 0 | 0 | 866 | 2,119 | 140 | 127 | 23 | 3 | 0 | 6 |
| 2010 | 0 | 48 | 104 | 1,045 | 2,410 | 52 | 162 | 0 | 9 | 0 |
| 2011 | 0 | 193 | 0 | 40 | 721 | 882 | 53 | 109 | 0 | 0 |
| 2012 | 50 | 255 | 1088 | 20 | 0 | 259 | 578 | 5 | 12 | 0 |

Table 6. Yellow perch catch at age matrix from upper Chesapeake Bay commercial fyke net survey, 1999-2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 1999 | 0 | 0 | 1,621 | 33 | 337 | 408 | 28 | 0 | 2 | 0 |
| 2000 | 0 | 35 | 138 | 2937 | 129 | 369 | 211 | 0 | 0 | 0 |
| 2001 | 0 | 0 | 83 | 90 | 432 | 17 | 9 | 17 | 0 | 0 |
| 2002 | 0 | 52 | 117 | 528 | 56 | 1,000 | 14 | 39 | 53 | 0 |
| 2003 | 0 | 27 | 565 | 78 | 361 | 45 | 418 | 6 | 15 | 25 |
| 2004 | 0 | 4 | 473 | 499 | 62 | 50 | 3 | 43 | 2 | 2 |
| 2005 | 0 | 18 | 27 | 1,320 | 414 | 73 | 37 | 0 | 26 | 5 |
| 2006 | 0 | 32 | 476 | 9 | 848 | 245 | 0 | 1 | 10 | 0 |
| 2007 | 0 | 2 | 290 | 1,400 | 23 | 548 | 168 | 3 | 0 | 14 |
| 2008 | 0 | 70 | 3,855 | 3,782 | 4,820 | 75 | 789 | 149 | 14 | 2 |
| 2009 | 0 | 87 | 128 | 663 | 490 | 648 | 5 | 80 | 35 | 0 |
| 2010 | 0 | 3 | 356 | 125 | 274 | 281 | 260 | 0 | 23 | 0 |
| 2011 | 0 | 41 | 56 | 703 | 152 | 355 | 183 | 102 | 0 | 0 |
| 2012 | 0 | 19 | 462 | 38 | 548 | 14 | 244 | 99 | 54 | 35 |

Table 7. Yellow perch catch at age matrix from Nanticoke River fyke and pound net survey, 1999 - 2012. 2007 -- 2009 include Marshyhope River data.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |
| 1999 | 0 | 10 | 1,072 | 323 | 295 | 22 | 0 | 4 | 14 | 22 |
| 2000 | 0 | 0 | 16 | 561 | 78 | 83 | 7 | 0 | 0 | 0 |
| 2001 | 0 | 2 | 36 | 114 | 737 | 48 | 36 | 3 | 0 | 0 |
| 2002 | 0 | 128 | 9 | 60 | 36 | 940 | 39 | 24 | 6 | 0 |
| 2003 | 0 | 17 | 123 | 2 | 49 | 2 | 45 | 1 | 2 | 0 |
| 2004 | 0 | 7 | 58 | 93 | 0 | 1 | 10 | 21 | 1 | 0 |
| 2005 | 0 | 59 | 6 | 34 | 35 | 0 | 1 | 0 | 4 | 0 |
| 2006 | 0 | 56 | 381 | 18 | 34 | 50 | 4 | 3 | 6 | 5 |
| 2007 | 0 | 38 | 244 | 291 | 37 | 32 | 16 | 0 | 0 | 2 |
| 2008 | 0 | 36 | 238 | 144 | 148 | 25 | 9 | 4 | 2 | 7 |
| 2009 | 0 | 37 | 374 | 660 | 336 | 126 | 9 | 0 | 11 | 0 |
| 2010 | 0 | 0 | 0 | 3 | 6 | 5 | 0 | 0 | 0 | 0 |
| 2011 | 0 | 2 | 6 | 31 | 22 | 20 | 10 | 2 | 0 | 0 |
| 2012 | 0 | 28 | 12 | 8 | 11 | 15 | 14 | 4 | 1 | 0 |

Table 8. Relative stock densities (RSD's) of white perch from the upper Chesapeake Bay winter trawl survey, 2000 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(125 \mathrm{~mm})$ | Quality <br> $(200 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(305 \mathrm{~mm})$ | Trophy <br> $(380 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2000 | 76.9 | 22.1 | 0.9 | 0.1 | 0.0 |
| 2001 | 89.8 | 9.9 | 0.3 | 0.0 | 0.0 |
| 2002 | 87.1 | 12.0 | 0.8 | 0.0 | 0.0 |
| 2003 | 83.6 | 14.3 | 1.2 | 0.5 | 0.0 |
| 2004 | NOT SAMPLED |  |  |  |  |
| 2005 | 83.9 | 16.1 | 0.0 | 0.0 | 0.0 |
| 2006 | 88.4 | 10.8 | 0.1 | $<0.1$ | 0.0 |
| 2007 | 92.3 | 7.0 | 0.7 | 0.0 | 0.0 |
| 2008 | 91.2 | 8.2 | 0.6 | 0.0 | 0.0 |
| 2009 | 92.0 | 7.3 | 0.6 | 0.0 | 0.0 |
| 2010 | 89.6 | 9.7 | 0.7 | 0.0 | 0.0 |
| 2011 | 87.2 | 11.6 | 1.2 | 0.0 | 0.0 |
| 2012 | 86.4 | 12.7 | 0.9 | 0.0 | $<0.1$ |

Figure 7. White perch length-frequency from 2012 upper Chesapeake Bay winter trawl survey.


Table 9. Relative stock densities (RSD's) of white perch from the Choptank River fyke net survey, 1993-2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(125 \mathrm{~mm})$ | Quality <br> $(200 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(305 \mathrm{~mm})$ | Trophy <br> $(380 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 72.5 | 25.0 | 2.4 | 0.1 | 0.0 |
| 1994 | 76.8 | 21.3 | 1.8 | 0.1 | 0.0 |
| 1995 | 84.3 | 14.9 | 0.8 | 0.0 | 0.0 |
| 1996 | 86.4 | 13.1 | 0.5 | 0.0 | 0.0 |
| 1997 | 80.0 | 19.1 | 0.8 | 0.1 | 0.0 |
| 1998 | 71.9 | 26.2 | 1.8 | $<0.1$ | 0.0 |
| 1999 | 80.2 | 18.7 | 1.1 | $<0.1$ | 0.0 |
| 2000 | 72.0 | 25.9 | 2.1 | 0.0 | 0.0 |
| 2001 | 84.6 | 14.4 | 1.0 | 0.0 | 0.0 |
| 2002 | 71.6 | 26.6 | 1.7 | 0.1 | 0.0 |
| 2003 | 76.4 | 22.2 | 1.3 | 0.1 | 0.0 |
| 2004 | 75.6 | 23.6 | 1.0 | 0.1 | 0.0 |
| 2005 | 78.5 | 19.9 | 1.5 | 0.1 | 0.0 |
| 2006 | 70.5 | 26.7 | 2.7 | $<0.1$ | 0.0 |
| 2007 | 76.5 | 21.7 | 1.7 | 0.0 | 0.0 |
| 2008 | 73.8 | 24.9 | 1.2 | $<0.1$ | 0.0 |
| 2009 | 73.0 | 25.5 | 1.4 | 0.1 | 0.0 |
| 2010 | 62.3 | 35.0 | 2.7 | $<0.1$ | 0.0 |
| 2011 | 63.0 | 33.5 | 3.2 | 0.3 | 0.0 |
| 2012 | 51.9 | 42.9 | 4.9 | 0.2 | 0.0 |

Figure 8. White perch length-frequency from 2012 Choptank River fyke net survey.


Table 10. Relative stock densities (RSD's) of white perch from the Nanticoke River fyke and pound net survey, 1995 - 2012. Minimum length cut-offs in parentheses. 2007 -- 2009 include Marshyhope River data.

| Year | Stock <br> $(125 \mathrm{~mm})$ | Quality <br> $(200 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(305 \mathrm{~mm})$ | Trophy <br> $(380 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | 56.3 | 35.4 | 5.2 | 3.0 | 0.0 |
| 1996 | 37.8 | 54.2 | 7.3 | 0.7 | 0.0 |
| 1997 | 37.5 | 58.4 | 4.0 | $<0.1$ | 0.0 |
| 1998 | 30.4 | 63.1 | 6.4 | $<0.1$ | 0.0 |
| 1999 | 37.2 | 57.7 | 5.0 | $<0.1$ | 0.0 |
| 2000 | 31.3 | 58.9 | 9.7 | $<0.1$ | 0.0 |
| 2001 | 26.2 | 60.7 | 12.5 | 0.6 | 0.0 |
| 2002 | 32.4 | 52.9 | 14.3 | 0.4 | 0.0 |
| 2003 | 26.4 | 60.6 | 11.9 | 1.1 | 0.0 |
| 2004 | 23.0 | 61.0 | 14.0 | 2.0 | 0.0 |
| 2005 | 25.3 | 52.8 | 19.3 | 2.6 | 0.0 |
| 2006 | 26.1 | 56.7 | 16.3 | $<0.1$ | 0.0 |
| 2007 | 36.3 | 52.4 | 10.0 | 1.4 | 0.0 |
| 2008 | 36.2 | 50.9 | 12.2 | 0.7 | 0.0 |
| 2009 | 33.6 | 53.2 | 12.2 | 1.0 | 0.0 |
| 2010 | 22.0 | 53.6 | 23.1 | 1.1 | 0.2 |
| 2011 | 25.1 | 53.0 | 19.1 | 2.7 | 0.0 |
| 2012 | 30.4 | 47.7 | 19.9 | 2.0 | 0.0 |

Figure 9. White perch length-frequency from 2012 Nanticoke River fyke and pound net survey.


Table 11. Relative stock densities (RSD's) of yellow perch from the upper Chesapeake Bay winter trawl survey, 2000 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(140 \mathrm{~mm})$ | Quality <br> $(216 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(318 \mathrm{~mm})$ | Trophy <br> $(405 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2000 | 84.2 | 14.3 | 1.5 | 0.0 | 0.0 |
| 2001 | 90.6 | 7.9 | 1.4 | 0.0 | 0.0 |
| 2002 | 87.8 | 10.7 | 1.5 | 0.0 | 0.0 |
| 2003 | 87.5 | 9.9 | 1.9 | 0.0 | 0.0 |
| 2004 | NOT SAMPLED |  |  |  |  |
| 2005 | 98.6 | 1.4 | 0.0 | 0.0 | 0.0 |
| 2006 | 97.7 | 1.7 | 0.5 | 0.0 | 0.0 |
| 2007 | 98.7 | 0.4 | 0.8 | 0.0 | 0.0 |
| 2008 | 94.2 | 4.6 | 1.2 | 0.0 | 0.0 |
| 2009 | 93.4 | 4.6 | 2.0 | 0.0 | 0.0 |
| 2010 | 80.7 | 16.7 | 2.6 | 0.0 | 0.0 |
| 2011 | 83.7 | 12.8 | 3.5 | 0.0 | 0.0 |
| 2012 | 92.6 | 5.9 | 1.5 | 0.0 | 0.0 |

Figure 10. Yellow perch length-frequency from the 2012 upper Chesapeake Bay winter trawl survey.


Table 12. Relative stock densities (RSD's) of yellow perch from the Choptank River fyke net survey, 1989 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(140 \mathrm{~mm})$ | Quality <br> $(216 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(318 \mathrm{~mm})$ | Trophy <br> $(405 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1989 | 66.7 | 24.4 | 8.2 | 0.7 | 0.0 |
| 1990 | 64.8 | 27.3 | 7.8 | 0.0 | 0.0 |
| 1991 | 58.7 | 23.4 | 18.0 | 0.0 | 0.0 |
| 1992 | 45.3 | 26.4 | 24.5 | 3.8 | 0.0 |
| 1993 | 34.6 | 31.7 | 30.3 | 3.3 | 0.0 |
| 1994 | 23.4 | 33.6 | 36.6 | 6.4 | 0.0 |
| 1995 | 45.5 | 28.1 | 23.1 | 3.3 | 0.0 |
| 1996 | 74.1 | 18.2 | 7.2 | 0.5 | 0.0 |
| 1997 | 57.5 | 29.3 | 12.9 | 0.3 | 0.0 |
| 1998 | 10.5 | 72.9 | 16 | 0.6 | 0.0 |
| 1999 | 86.0 | 12.4 | 2.4 | $<0.1$ | 0.0 |
| 2000 | 71.6 | 19.0 | 9.1 | 0.2 | 0.0 |
| 2001 | 83.6 | 13.0 | 3.3 | $<0.1$ | 0.0 |
| 2002 | 59.8 | 33.1 | 6.9 | 0.2 | 0.0 |
| 2003 | 67.0 | 27.4 | 5.4 | 0.2 | 0.0 |
| 2004 | 54.2 | 34.6 | 10.7 | 0.4 | 0.0 |
| 2005 | 75.1 | 17.2 | 7.4 | 0.2 | 0.0 |
| 2006 | 53.5 | 32.1 | 13.8 | 0.6 | 0.0 |
| 2007 | 74.9 | 15.0 | 9.9 | 0.2 | 0.0 |
| 2008 | 76.4 | 16.1 | 7.3 | 0.2 | 0.0 |
| 2009 | 77.3 | 17.4 | 5.1 | $<0.1$ | 0.0 |
| 2010 | 64.3 | 25.6 | 10.0 | 0.1 | 0.0 |
| 2011 | 50.1 | 32.6 | 16.9 | 0.3 | 0.0 |
| 2012 | 51.5 | 30.8 | 16.7 | 1.0 | 0.0 |

Figure 11. Yellow perch length-frequency from the 2012 Choptank River fyke net survey.


Table 13. Relative stock densities (RSD's) of yellow perch from the upper Chesapeake Bay commercial fyke net survey, 1988, 1990, 1998-2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(140 \mathrm{~mm})$ | Quality <br> $(216 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(318 \mathrm{~mm})$ | Trophy <br> $(405 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1988 | 71.8 | 25.3 | 3.1 | 0.0 | 0.0 |
| 1990 | 6.7 | 71.7 | 21 | 0.1 | 0.0 |
| 1998 | 24.2 | 51.0 | 24.7 | $<0.1$ | 0.0 |
| 1999 | 40.2 | 52.3 | 7.3 | 0.2 | 0.0 |
| 2000 | 55.1 | 37.2 | 7.6 | $<0.1$ | 0.0 |
| 2001 | 27.1 | 48.8 | 24.0 | 0.0 | 0.0 |
| 2002 | 17.8 | 63.1 | 18.9 | 0.2 | 0.0 |
| 2003 | 19.5 | 54.6 | 24.6 | 1.3 | 0.0 |
| 2004 | 9.6 | 66.3 | 23.8 | 0.3 | 0.0 |
| 2005 | 45.2 | 42.2 | 12.1 | 0.5 | 0.0 |
| 2006 | 35.0 | 52.8 | 12.0 | 0.2 | 0.0 |
| 2007 | 40.1 | 47.9 | 11.5 | 0.5 | 0.0 |
| 2008 | 31.6 | 55.3 | 13.0 | 0.1 | 0.0 |
| 2009 | 30.6 | 47.6 | 21.4 | 0.4 | 0.0 |
| 2010 | 20.9 | 60.3 | 18.2 | 0.6 | 0.0 |
| 2011 | 27.0 | 50.2 | 22.4 | 0.4 | 0.0 |
| 2012 | 22.1 | 54.5 | 22.6 | 0.7 | 0.0 |

Figure 12. Yellow perch length frequency from the 2012 upper Chesapeake commercial fyke net survey.


Table 14. Relative stock densities (RSD's) of yellow perch from the Nanticoke River fyke and pound net survey, 1999 - 2012. Minimum length cut-offs in parentheses; 2007-- 2009 includes Marshyhope River data.

| Year | Stock <br> $(140 \mathrm{~mm})$ | Quality <br> $(216 \mathrm{~mm})$ | Preferred <br> $(255 \mathrm{~mm})$ | Memorable <br> $(318 \mathrm{~mm})$ | Trophy <br> $(405 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1999 | 12.4 | 28.8 | 55.6 | 3.2 | 0.0 |
| 2000 | 3.1 | 19.5 | 72 | 5.2 | 0.0 |
| 2001 | 2.4 | 22.2 | 66.6 | 8.9 | 0.0 |
| 2002 | 2.9 | 18.9 | 62.5 | 15.7 | 0.0 |
| 2003 | 10.9 | 46.6 | 36.3 | 6.2 | 0.0 |
| 2004 | 1.6 | 27.2 | 60.7 | 10.5 | 0.0 |
| 2005 | 16.2 | 33.8 | 38.7 | 11.3 | 0.0 |
| 2006 | 4.1 | 34.1 | 57.1 | 4.7 | 0.0 |
| 2007 | 15.7 | 21.8 | 57.1 | 5.4 | 0.0 |
| 2008 | 27.4 | 25.0 | 42.1 | 5.5 | 0.0 |
| 2009 | 9.0 | 28.0 | 53.9 | 9.0 | 0.0 |
| 2010 | 0.0 | 14.3 | 78.6 | 7.1 | 0.0 |
| 2011 | 2.2 | 15.0 | 75.3 | 7.5 | 0.0 |
| 2012 | 24.7 | 16.1 | 44.1 | 15.0 | 0.0 |

Figure 13. Yellow perch length frequency from the 2012 Nanticoke River survey fyke and pound net survey.


Table 15. Relative stock densities (RSD's) of channel catfish from the upper Chesapeake Bay winter trawl survey, 2000 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(255 \mathrm{~mm})$ | Quality <br> $(460 \mathrm{~mm})$ | Preferred <br> $(510 \mathrm{~mm})$ | Memorable <br> $(710 \mathrm{~mm})$ | Trophy <br> $(890 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2000 | 88.5 | 4.5 | 6.4 | 0.6 | 0.0 |
| 2001 | 92.7 | 2.5 | 4.7 | 0.0 | 0.0 |
| 2002 | 89.4 | 7.3 | 3.2 | 0.0 | 0.0 |
| 2003 | 89.5 | 5.3 | 5.3 | 0.0 | 0.0 |
| 2004 | NOT SAMPLED |  |  |  |  |
| 2005 | 73.8 | 10.0 | 16.2 | 0.0 | 0.0 |
| 2006 | 96.4 | 2.0 | 1.6 | 0.0 | 0.0 |
| 2007 | 95.6 | 2.2 | 2.2 | 0.0 | 0.0 |
| 2008 | 91.4 | 3.7 | 4.9 | 0.0 | 0.0 |
| 2009 | 94.1 | 2.1 | 3.8 | 0.0 | 0.0 |
| 2010 | 84.6 | 9.2 | 5.8 | 0.4 | 0.0 |
| 2011 | 76.3 | 14.0 | 9.7 | 0.0 | 0.0 |
| 2012 | 88.5 | 5.9 | 5.1 | 0.4 | 0.0 |

Figure 14. Length frequency of channel catfish from the 2012 upper Chesapeake Bay winter trawl survey.


Table 16. Relative stock densities (RSD's) of channel catfish from the Choptank River fyke net survey, 1993 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(255 \mathrm{~mm})$ | Quality <br> $(460 \mathrm{~mm})$ | Preferred <br> $(510 \mathrm{~mm})$ | Memorable <br> $(710 \mathrm{~mm})$ | Trophy <br> $(890 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 53.4 | 24.0 | 22.6 | 0.0 | 0.0 |
| 1994 | 61.9 | 15.8 | 22.2 | 0.0 | 0.0 |
| 1995 | 21.0 | 20.4 | 58.6 | 0.0 | 0.0 |
| 1996 | 40.8 | 14.1 | 35.6 | 0.0 | 0.0 |
| 1997 | 19.8 | 16.4 | 63.8 | 0.0 | 0.0 |
| 1998 | 33.3 | 9.2 | 57.5 | 0.0 | 0.0 |
| 1999 | 31.3 | 10.6 | 58.1 | 0.0 | 0.0 |
| 2000 | 63.7 | 8.4 | 27.9 | 0.0 | 0.0 |
| 2001 | 53.2 | 6.7 | 40.1 | 0.0 | 0.0 |
| 2002 | 19.8 | 14.3 | 65.9 | 0.0 | 0.0 |
| 2003 | 84.2 | 5.8 | 9.9 | 0.0 | 0.0 |
| 2004 | 58.8 | 10.0 | 31.2 | 0.0 | 0.0 |
| 2005 | 79.2 | 9.3 | 11.5 | 0.0 | 0.0 |
| 2006 | 72.3 | 12.6 | 15.1 | 0.0 | 0.0 |
| 2007 | 84.9 | 7.1 | 8.0 | 0.0 | 0.0 |
| 2008 | 79.6 | 8.1 | 12.3 | 0.0 | 0.0 |
| 2009 | 74.3 | 8.2 | 27.0 | 0.0 | 0.0 |
| 2010 | 69.0 | 12.0 | 18.9 | 0.0 | 0.0 |
| 2011 | 73.4 | 13.4 | 13.2 | 0.0 | 0.0 |
| 2012 | 14.1 | 7.0 | 78.5 | 0.2 | 0.1 |

Figure 15. Channel catfish length frequency from the 2012 Choptank River fyke net survey.


Table 17. Relative stock densities (RSD's) of channel catfish from Nanticoke River fyke and pound net survey, 1995 - 2012. 2007 -- 2009 include Marshyhope River fyke net data. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(255 \mathrm{~mm})$ | Quality <br> $(460 \mathrm{~mm})$ | Preferred <br> $(510 \mathrm{~mm})$ | Memorable <br> $(710 \mathrm{~mm})$ | Trophy <br> $(890 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | 72.3 | 19.4 | 8.2 | 0.0 | 0.0 |
| 1996 | 65.8 | 23.8 | 10.4 | 0.0 | 0.0 |
| 1997 | 62.2 | 27.5 | 10.2 | 0.0 | 0.0 |
| 1998 | 60.3 | 27.7 | 12.0 | 0.0 | 0.0 |
| 1999 | 80.6 | 14.6 | 4.7 | 0.0 | 0.0 |
| 2000 | 70.9 | 22.1 | 7.1 | 0.0 | 0.0 |
| 2001 | 70.2 | 22.9 | 6.9 | 0.0 | 0.0 |
| 2002 | 56.4 | 31.1 | 12.5 | 0.0 | 0.0 |
| 2003 | 52.3 | 29.2 | 18.4 | 0.0 | 0.0 |
| 2004 | 60.8 | 27.8 | 11.5 | 0.0 | 0.0 |
| 2005 | 48.8 | 30.6 | 20.6 | 0.0 | 0.0 |
| 2006 | 63.7 | 23.0 | 13.3 | 0.0 | 0.0 |
| 2007 | 67.4 | 22.8 | 9.8 | 0.0 | 0.0 |
| 2008 | 69.4 | 17.8 | 12.6 | 0.3 | 0.0 |
| 2009 | 66.5 | 18.4 | 15.1 | 0.0 | 0.0 |
| 2010 | 45.0 | 23.3 | 30.0 | 1.7 | 0.0 |
| 2011 | 74.1 | 13.0 | 13.0 | 0.0 | 0.0 |
| 2012 | 22.5 | 30.2 | 47.3 | 0.0 | 0.0 |

Figure 16. Channel catfish length frequency from the 2012 Nanticoke River fyke and pound net survey.


Table 18. Relative stock densities (RSD's) of white catfish from the upper Chesapeake Bay winter trawl survey, 2000 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(165 \mathrm{~mm})$ | Quality <br> $(255 \mathrm{~mm})$ | Preferred <br> $(350 \mathrm{~mm})$ | Memorable <br> $(405 \mathrm{~mm})$ | Trophy <br> $(508 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2000 | NONE COLLECTED |  |  |  |  |
| 2001 | 41.9 | 54.8 | 3.2 | 0.0 | 0.0 |
| 2002 | 57.1 | 42.9 | 0.0 | 0.0 | 0.0 |
| 2003 | 85.0 | 15.0 | 0.0 | 0.0 | 0.0 |
| 2004 | NOT SAMPLED |  |  |  |  |
| 2005 | 96.6 | 3.4 | 0.0 | 0.0 | 0.0 |
| 2006 | 90.0 | 10.0 | 0.0 | 0.0 | 0.0 |
| 2007 | 85.7 | 14.3 | 0.0 | 0.0 | 0.0 |
| 2008 | 85.7 | 14.3 | 0.0 | 0.0 | 0.0 |
| 2009 | 83.0 | 17.0 | 0.0 | 0.0 | 0.0 |
| 2010 | 87.0 | 10.9 | 2.2 | 0.0 | 0.0 |
| 2011 | 81.9 | 17.3 | 0.8 | 0.0 | 0.0 |
| 2012 | 70.2 | 26.9 | 3.0 | 0.0 | 0.0 |

Figure 17. White catfish length frequency from the 2012 upper Chesapeake Bay winter trawl survey.


Table 19. Relative stock densities (RSD's) of white catfish from the Choptank River fyke net survey, 1993 - 2012. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(165 \mathrm{~mm})$ | Quality <br> $(255 \mathrm{~mm})$ | Preferred <br> $(350 \mathrm{~mm})$ | Memorable <br> $(405 \mathrm{~mm})$ | Trophy <br> $(508 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 45.6 | 19.4 | 4.9 | 27.2 | 2.9 |
| 1994 | 42.2 | 28.9 | 10.2 | 18.8 | 0.0 |
| 1995 | 19.3 | 47.8 | 8.9 | 23.1 | 0.9 |
| 1996 | 45.6 | 22.1 | 6.1 | 24.4 | 1.5 |
| 1997 | 29.7 | 48.5 | 6.9 | 12.9 | 2.0 |
| 1998 | 42.6 | 44.1 | 2.9 | 10.3 | 0.5 |
| 1999 | 44.8 | 38.6 | 5.9 | 10.8 | 0.0 |
| 2000 | 50.6 | 29.2 | 7.6 | 12.4 | 0.3 |
| 2001 | 44.8 | 29.5 | 4.8 | 20.0 | 1.0 |
| 2002 | 7.8 | 38.9 | 15.4 | 35.5 | 2.4 |
| 2003 | 25.2 | 35.8 | 11.9 | 26.5 | 0.4 |
| 2004 | 15.2 | 54.8 | 20.9 | 9.5 | 0.0 |
| 2005 | 37.4 | 41.0 | 15.5 | 6.0 | 0.0 |
| 2006 | 29.1 | 45.4 | 13.3 | 12.0 | 0.2 |
| 2007 | 49.6 | 39.1 | 7.5 | 3.8 | 0.0 |
| 2008 | 26.1 | 44.4 | 13.8 | 15.5 | 0.3 |
| 2009 | 25.3 | 48.6 | 9.9 | 15.8 | 0.5 |
| 2010 | 19.6 | 52.5 | 11.3 | 16.2 | 0.4 |
| 2011 | 23.5 | 33.5 | 9.7 | 33.1 | 0.2 |
| 2012 | 12.5 | 50.6 | 13.3 | 22.9 | 0.8 |

Figure 18. White catfish length frequency from the 2012 Choptank River fyke net survey.


Table 20. Relative stock densities (RSD's) of white catfish from the Nanticoke River fyke and pound net survey, 1995 - 2012. 2007 -- 2009 include Marshyhope River fyke net data. Minimum length cut-offs in parentheses.

| Year | Stock <br> $(165 \mathrm{~mm})$ | Quality <br> $(255 \mathrm{~mm})$ | Preferred <br> $(350 \mathrm{~mm})$ | Memorable <br> $(405 \mathrm{~mm})$ | Trophy <br> $(508 \mathrm{~mm})$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | 35.7 | 32.8 | 14.3 | 16.6 | 0.6 |
| 1996 | 42.4 | 36.9 | 10.5 | 9.6 | 0.6 |
| 1997 | 42.1 | 37.4 | 10.9 | 8.2 | 1.4 |
| 1998 | 27.9 | 48.2 | 17.4 | 6.0 | 0.0 |
| 1999 | 41.0 | 34.5 | 14.4 | 10.1 | 0.0 |
| 2000 | 39.9 | 42.1 | 12.0 | 6.0 | 0.0 |
| 2001 | 46.2 | 28.2 | 16.0 | 9.0 | 0.6 |
| 2002 | 37.0 | 34.6 | 15.2 | 12.8 | 0.5 |
| 2003 | 17.6 | 32.4 | 23.5 | 25.0 | 1.5 |
| 2004 | 13.2 | 45.3 | 34.9 | 6.6 | 0.0 |
| 2005 | 47.0 | 30.3 | 13.6 | 9.1 | 0.0 |
| 2006 | 70.0 | 21.1 | 4.3 | 4.6 | 0.0 |
| 2007 | 40.0 | 37.3 | 14.7 | 8.0 | 0.0 |
| 2008 | 62.5 | 24.1 | 8.5 | 4.6 | 0.3 |
| 2009 | 55.8 | 21.8 | 10.5 | 10.5 | 1.4 |
| 2010 | 21.4 | 25.0 | 14.3 | 28.6 | 10.7 |
| 2011 | 43.7 | 43.7 | 5.7 | 5.7 | 6.9 |
| 2012 | 11.9 | 25.8 | 29.6 | 30.5 | 2.2 |

Figure 19. White catfish length frequency from the 2012 Nanticoke River fyke and pound net survey.


Table 21. White perch growth parameters from Choptank River for males, females, and sexes combined. NA=data not available $\mathrm{NSF}=$ no solution found or small sample size.

| Sample Year | Sex | Allometry |  | von Bertalanffy |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | alpha | beta | L-inf | K | $\mathrm{t}_{0}$ |
| 2004 | F | $6.4 \times 10^{-6}$ | 3.17 |  | NSF |  |
|  | M | NSF |  |  | NSF |  |
|  | Combined | $4.5 \times 10^{-6}$ | 3.23 |  | NSF |  |
| 2005 | F | $4.8 \times 10^{-6}$ | 3.23 | 288 | 0.36 | 0.00 |
|  | M | $4.8 \times 10^{-6}$ | 3.22 | 374 | 0.10 | -2.10 |
|  | Combined | $3.8 \times 10^{-6}$ | 3.27 | 304 | 0.25 | -1.60 |
| 2006 | F | NSF | 285 | 0.36 | 0.40 |  |
|  | M | NSF | 275 | 0.42 | 0.60 |  |
|  | Combined | $7.8 \times 10^{-5}$ | 2.69 | 273 | 0.4 | 0.60 |
| 2007 | F | $1.6 \times 10^{-5}$ | 3.00 | 269 | 0.33 | 0.28 |
|  | M | $5.8 \times 10^{-5}$ | 2.74 | 247 | 0.32 | 0.06 |
|  | Combined | $1.9 \times 10^{-5}$ | 2.96 | 265 | 0.31 | 0.15 |
| 2008 | F | $3.0 \times 10^{-6}$ | 3.29 | 317 | 0.23 | -1.44 |
|  | M | $3.7 \times 10^{-6}$ | 3.25 | 227 | 0.32 | -1.98 |
|  | Combined | $2.2 \times 10^{-6}$ | 3.35 | 284 | 0.28 | -0.89 |
| 2009 | F | $2.8 \times 10^{-6}$ | 3.32 | 338 | 0.20 | -1.33 |
|  | M | $2.5 \times 10^{-6}$ | 3.32 | 225 | 0.49 | -0.77 |
|  | Combined | $1.9 \times 10^{-6}$ | 3.38 | 281 | 0.32 | -0.17 |
| 2010 | F | $4.0 \times 10^{-6}$ | 3.26 | 312 | 0.18 | -1.38 |
|  | M | $4.2 \times 10^{-6}$ | 3.23 |  | NSF |  |
|  | Combined | $2.6 \times 10^{-6}$ | 3.33 |  | NSF |  |
| 2011 | F | $2.3 \times 10^{-6}$ | 3.35 |  | NSF |  |
|  | M | $2.4 \times 10^{-6}$ | 3.34 | 217 | 0.49 | 0.44 |
|  | Combined | $2.0 \times 10^{-6}$ | 3.38 |  | NSF |  |
| 2012 | F | $6.9 \times 10^{-6}$ | 3.17 | 264 | 0.47 | 0.81 |
|  | M | $4.5 \times 10^{-6}$ | 3.23 | 227 | 0.39 | -0.21 |
|  | Combined | $3.1 \times 10^{-6}$ | 3.31 | 251 | 0.46 | 0.68 |
| 2000-2012 | F | $4.5 \times 10^{-6}$ | 3.23 | 303 | 0.20 | -1.41 |
|  | M | $5.7 \times 10^{-6}$ | 3.18 | 241 | 0.26 | -1.24 |
|  | Combined | $3.2 \times 10^{-6}$ | 3.29 | 288 | 0.21 | -1.25 |

Table 22. White perch growth parameters from Nanticoke River for males, females, and sexes combined. NA=data not available NSF=no solution found or small sample size.

| Sample Year | Sex | (allometry) | (von Bertalanffy) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | alpha | beta | L-inf | K | $\mathrm{t}_{0}$ |
| 2003 | F | NA |  | 386 | 0.11 | -2.90 |
|  | M |  |  | 263 | 0.30 | -0.21 |
|  | Combined |  |  | 329 | 0.16 | -1.90 |
| 2004 | F | $5.3 \times 10^{-6}$ | 3.22 | 322 | 0.25 | -0.30 |
|  | M | $2.4 \times 10^{-6}$ | 3.35 | 288 | 0.21 | -1.50 |
|  | Combined | $2.6 \times 10^{-6}$ | 3.35 | 335 | 0.18 | -1.20 |
| 2005 | F | $2.3 \times 10^{-6}$ | 3.36 | 313 | 0.23 | -0.53 |
|  | M | NSF |  | 313 | 0.14 | -2.65 |
|  | Combined | $1.50 \times 10^{-6}$ | 3.44 | 321 | 0.17 | -1.60 |
| 2006 | F | NA |  | 311 | 0.22 | -1.41 |
|  | M |  |  | 279 | 0.19 | -2.54 |
|  | Combined |  |  | 321 | 0.16 | -2.60 |
| 2007 | F | $6.2 \times 10^{-6}$ | 2.76 | 299 | 0.23 | -0.81 |
|  | M | $1.0 \times 10^{-6}$ | 3.08 | 282 | 0.24 | -0.79 |
|  | Combined | $3.4 \times 10^{-6}$ | 2.87 | 297 | 0.23 | -0.70 |
| 2008 | F | $4.1 \times 10^{-6}$ | 3.25 | 295 | 0.35 | 0.23 |
|  | M | $8.0 \times 10^{-6}$ | 3.12 | 254 | 0.38 | -0.20 |
|  | Combined | $3.6 \times 10^{-6}$ | 3.27 | 288 | 0.32 | -0.16 |
| 2009 | F | $3.4 \times 10^{-6}$ | 3.28 | 285 | 0.33 | 0.47 |
|  | M | $1.4 \times 10^{-4}$ | 2.58 | 273 | 0.18 | -1.70 |
|  | Combined | $5.9 \times 10^{-6}$ | 3.18 | 284 | 0.25 | -0.33 |
| 2010 | F | $1.7 \times 10^{-6}$ | 3.41 | 345 | 0.16 | -1.36 |
|  | M | $3.4 \times 10^{-5}$ | 2.85 | 275 | 0.25 | -0.46 |
|  | Combined | $2.7 \times 10^{-6}$ | 3.32 | 318 | 0.18 | -1.03 |
| 2011 | F | $1.6 \times 10^{-6}$ | 3.42 | 313 | 0.25 | -0.20 |
|  | M | $7.8 \times 10^{-6}$ | 3.13 | 265 | 0.26 | -0.31 |
|  | Combined | $1.5 \times 10^{-6}$ | 3.43 | 293 | 0.24 | -0.39 |
| 2012 | F | $4.5 \times 10^{-6}$ | 3.25 | NSF |  |  |
|  | M | $1.0 \times 10^{-5}$ | 3.08 | 318 | 0.16 | -1.56 |
|  | Combined | $2.9 \times 10^{-6}$ | 3.32 | 344 | 0.14 | -1.83 |

Table 23. Yellow perch growth parameters from Choptank River for males, females, and sexes combined. NA=data not available $\mathrm{NSF}=$ no solution found or small sample size.

| Sample Year | Sex | allometry |  | von Bertalanffy |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | alpha | beta | L-inf | K | $\mathrm{t}_{0}$ |
| 2003 | F | NA |  | 264 | 0.82 | 0.36 |
|  | M | NA |  | 263 | 0.35 | -0.8 |
|  | Combined | NA |  | 255 | 0.5 | -0.7 |
| 2004 | F | NA |  | 306 | 0.41 | -0.4 |
|  | M | NA |  | 253 | 0.34 | -1.2 |
|  | Combined | NA |  | 259 | 0.51 | -0.5 |
| 2005 | F | NA |  | 293 | 0.64 | -0.5 |
|  | M | NA |  | 244 | 0.63 | 0.1 |
|  | Combined | NA |  | 258 | 0.45 | -1.6 |
| 2006 | F | NA |  | 297 | . 36 | -1.05 |
|  | M | NA |  | 291 | . 24 | -1.09 |
|  | Combined | NA |  | 290 | . 26 | -2.00 |
| 2007 | F | $2.3 \times 10^{-5}$ | 2.88 | 308 | 0.52 | 0.19 |
|  | M | $1.3 \times 10^{-5}$ | 2.97 | 279 | 0.29 | -1.40 |
|  | Combined | $1.1 \times 10^{-5}$ | 3.02 | 277 | 0.54 | -0.01 |
| 2008 | F | $5.8 \times 10^{-6}$ | 3.12 | 322 | 0.43 | -0.12 |
|  | M | $1.1 \times 10^{-5}$ | 3.00 | 253 | 0.26 | -2.82 |
|  | Combined | $8.1 \times 10^{-6}$ | 3.06 | 289 | 0.40 | -0.59 |
| 2009 | F | $8.7 \times 10^{-6}$ | 3.06 | 315 | 0.40 | -0.63 |
|  | M | $2.8 \times 10^{-6}$ | 3.26 | 288 | 0.35 | -0.24 |
|  | Combined | $4.4 \times 10^{-6}$ | 2.18 | 308 | 0.29 | -1.71 |
| 2010 | F | $1.3 \times 10^{-5}$ | 2.97 |  | NSF |  |
|  | M | $4.7 \times 10^{-6}$ | 3.16 |  | NSF |  |
|  | Combined | $9.9 \times 10^{-6}$ | 3.02 |  | NSF |  |
| 2011 | F | $1.2 \times 10^{-6}$ | 3.02 |  | NSF |  |
|  | M | $4.7 \times 10^{-6}$ | 3.17 |  | NSF |  |
|  | Combined | $3.2 \times 10^{-6}$ | 3.25 |  | NSF |  |
| 2012 | F | $7.0 \times 10^{-6}$ | 3.08 | 374 | 0.18 | -2.22 |
|  | M | $1.5 \times 10^{-6}$ | 3.37 | 257 | 0.29 | -2.62 |
|  | Combined | $6.7 \times 10^{-6}$ | 3.09 | 295 | 0.32 | -1.38 |
| 2000-2012 | F | $7.6 \times 10^{-6}$ | 3.09 | 350 | 0.28 | -1.29 |
|  | M | $2.9 \times 10^{-6}$ | 3.25 | 296 | 0.16 | -3.36 |
|  | Combined | $5.1 \times 10^{-6}$ | 3.15 | 271 | 0.35 | -1.38 |

Table 24. Yellow perch growth parameters from upper Chesapeake Bay fyke nets for males, females, and sexes combined. NA=data not available NSF=no solution found.

| Sample Year | Sex | allometry |  | von Bertalanffy |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | alpha | beta | L-inf | K | $\mathrm{t}_{0}$ |
| 2003 | F | $6.68 \times 10^{-7}$ | 3.53 | 298 | 0.47 | 0.03 |
|  | M | NSF |  | 246 | 0.44 | -1.1 |
|  | Combined | $4.14 \times 10^{-7}$ | 3.61 | 275 | 0.53 | -0.1 |
| 2004 | F | $1.18 \times 10^{-6}$ | 3.43 | 297 | 0.75 | 1.14 |
|  | M | NSF |  | 256 | 0.37 | -2.5 |
|  | Combined | $7.08 \times 10^{-7}$ | 3.52 | 273 | 1.04 | 1.35 |
| 2005 | F | $4.40 \times 10^{-7}$ | 3.62 | 358 | 0.25 | -0.7 |
|  | M | $5.61 \times 10^{-7}$ | 3.55 | 244 | 0.41 | -0.5 |
|  | Combined | $1.69 \times 10^{-7}$ | 3.79 | 256 | 0.64 | 0.32 |
| 2006 | F | $5.15 \times 10^{-5}$ | 2.75 | 288 | 0.34 | -2 |
|  | M | $4.75 \times 10^{-5}$ | 2.73 | 240 | 0.41 | -2 |
|  | Combined | $4.72 \times 10^{-5}$ | 2.75 | 244 | 0.6 | -2 |
| 2007 | F | $1.96 \times 10^{-6}$ | 3.35 | 325 | 0.34 | -0.09 |
|  | M | $4.38 \times 10^{-6}$ | 3.18 | 240 | 0.61 | 0.61 |
|  | Combined | $6.68 \times 10^{-7}$ | 3.54 | 267 | 0.64 | 0.55 |
| 2008 | F | $7.83 \times 10^{-6}$ | 3.11 | 339 | 0.26 | -2.14 |
|  | M | $3.32 \times 10^{-6}$ | 3.24 |  | NSF |  |
|  | Combined | $3.89 \times 10^{-6}$ | 3.23 | 275 | 0.41 | -1.97 |
| 2009 | F | $1.30 \times 10^{-6}$ | 3.43 | 294 | 0.43 | -0.78 |
|  | M | $6.09 \times 10^{-6}$ | 3.13 | 220 | 0.97 | -0.14 |
|  | Combined | $6.23 \times 10^{-6}$ | 3.56 | 245 | 0.90 | 0.13 |
| 2010 | F | $1.62 \times 10^{-4}$ | 2.57 | 392 | 0.51 | 0.04 |
|  | M | $1.92 \times 10^{-6}$ | 3.34 | 247 | 0.88 | 0.99 |
|  | Combined | $3.40 \times 10^{-5}$ | 2.84 | 296 | 0.66 | 0.40 |
| 2011 | F | $3.1 \times 10^{-8}$ | 4.10 |  | NSF |  |
|  | M | $9.4 \times 10^{-7}$ | 3.47 |  | NSF |  |
|  | Combined | $9.1 \times 10^{-6}$ | 3.90 | 245 | 0.66 | -1.93 |
| 2012 | F | $1.4 \times 10^{-6}$ | 3.39 | 294 | 0.44 | -0.31 |
|  | M | $7.8 \times 10^{-6}$ | 3.06 | 253 | 0.89 | 1.22 |
|  | Combined | $7.7 \times 10^{-6}$ | 3.50 | 269 | 0.73 | 0.53 |
| 1998-2012 | F | $4.5 \times 10^{-6}$ | 3.20 | 305 | 0.30 | -1.28 |
|  | M | $3.5 \times 10^{-6}$ | 3.22 | 244 | 0.36 | -2.28 |
|  | Combined | $2.1 \times 10^{-6}$ | 3.33 | 262 | 0.54 | -0.36 |

Table 25. Yellow perch growth parameters from upper Nanticoke River for males, females, and sexes combined. NA=data not available NSF=no solution found or small sample size.

| Sample Year | Sex | Allometry |  | von Bertalanffy |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | alpha | beta | L-inf | K | $\mathrm{T}_{0}$ |
| 2003 | F |  |  | 324 | 0.49 | -0.3 |
|  | M | NA | 273 | 0.38 | -1.4 |  |
|  | Combined |  |  | 298 | 0.56 | -0.6 |
| 2004 | F |  |  | 326 | 0.43 | -1.1 |
|  | M | NA | 284 | 0.32 | -3.4 |  |
|  | Combined |  |  | 290 | 0.68 | -0.5 |
| 2005 | F | NSF | 332 | 0.56 | -0.1 |  |
|  | M | $3.40 \times 10^{-5}$ | 2.84 | 286 | 0.68 | 0.1 |
|  | Combined | NSF | 342 | 0.35 | -1.1 |  |
| 2006 | F | NA | 313 | 0.73 | 0.3 |  |
|  | M |  |  | 297 | 0.57 | -0.1 |
|  | Combined |  |  | 301 | 0.78 | 0.4 |
| 2007 | F | $1.80 \times 10^{-6}$ | 3.38 | 346 | 0.35 | -0.8 |
|  | M | $7.37 \times 10^{-6}$ | 3.10 |  | NSF |  |
|  | Combined | $1.18 \times 10^{-6}$ | 3.45 | 308 | 0.42 | -0.8 |
| 2008 | F | $3.37 \times 10^{-6}$ | 3.26 | 325 | 0.63 | 0.28 |
|  | M | $6.79 \times 10^{-6}$ | 3.10 | 259 | 0.92 | 0.45 |
|  | Combined | $9.96 \times 10^{-7}$ | 3.46 | 285 | 0.90 | 0.55 |
| 2009 | F | $3.0 \times 10^{-5}$ | 2.87 | NSF |  |  |
|  | M | $7.5 \times 10^{-5}$ | 2.67 | 292 | 0.40 | -0.01 |
|  | Combined | $1.1 \times 10^{-5}$ | 3.05 | 317 | 0.32 | -1.10 |
| 2010 | F | NSF |  |  | NSF |  |
|  | M | NSF |  |  | NSF |  |
|  | Combined | NSF |  |  | NSF |  |
| 2011 | F | $5.4 \times 10^{-5}$ | 2.74 |  | NSF |  |
|  | M | $3.3 \times 10^{-6}$ | 3.23 |  | NSF |  |
|  | Combined | $1.6 \times 10^{-5}$ | 2.95 |  | NSF |  |
| 2012 | F |  | 2.93 |  | . 053 | 0.08 |
|  | M | $1.8 \times 10^{-6}$ | 3.34 | 311 | . 034 | -0.41 |
|  | Combined | $8.6 \times 10^{-6}$ | 3.07 | 312 | . 063 | 0.43 |
| 2000-2012 | F | $9.42 \times 10^{-6}$ | 3.07 | 347 | 0.30 | -1.20 |
|  | M | $1.1 \times 10^{-5}$ | 3.01 | 294 | 0.34 | -1.11 |
|  | Combined | $3.7 \times 10^{-6}$ | 3.23 | 307 | 0.40 | -0.84 |

Table 26. Estimated instantaneous fishing mortality rates (F) for white perch. Based on catch curve analysis of ages $6-10+$. NR= not reliable; NA=not available; MIN= minimal, at or near M estimate.

|  | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Choptank | 0.46 | 0.1 | 0.58 | 0.58 | 0.40 | MIN | 0.35 | 0.99 | 0.29 | 0.08 |
| Nanticoke | 0.31 | NR | NR | 0.22 | 0.18 | 0.16 | 0.12 | 0.66 | NR | NR |
| Upper Bay trawl | 0.13 | NA | 0.50 | 0.12 | 0.19 | 0.26 | 0.54 | 0.76 | 0.51 | 0.08 |

Table 27. Estimated instantaneous fishing mortality rates ( F ) for yellow perch. $\mathrm{NR}=$ not reliable; $\mathrm{MIN}=$ minimal, at or near M estimate.

|  | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Choptank $^{1}$ | 0.05 | NR | 0.08 | MIN | 0 | NR | 0.17 | MIN | 0.56 | 0.12 |
| ${\text { Upper } \mathrm{Bay}^{2}}^{2}$ | 0.30 | 0.30 | 0.31 | 0.10 | 0.14 | 0.02 | 0.14 | 0.19 | 0.24 | 0.28 |

${ }^{1}$ Based on ratio of CPUE of ages 4-10+ (year t) to CPUE of ages 3-10+ (year t-1)
except 2002 estimate where all available ages were used, and 2009 estimate where ratio of ages 5-10 and 4-10 were used.
${ }^{2} \mathrm{~N}$-weighted population F from Piavis and Webb in publ.
Figure 20. Baywide young-of-year relative abundance index for white perch, 1962 - 2012, based on EJFS data. Bold horizontal line=time series average. Error bars indicate 95\% CI's.


Figure 21. Age 1 white perch relative abundance from upper Chesapeake Bay winter trawl survey. Not sampled in 2004, small sample sizes 2003 and 2005.


Figure 22. Head-of-Bay young-of-year relative abundance index for yellow perch, 1979 - 2012, based on Estuarine Juvenile Finfish Survey data. Horizontal line=time series average. Error bars indicate $95 \%$ confidence interval.


Figure 23. Age 1 yellow perch relative abundance from upper Chesapeake Bay winter trawl survey. Horizontal line=time series average. Not sampled in 2004, small sample sizes 2003 and 2005.


Figure 24. Age 1 channel catfish relative abundance from upper Chesapeake Bay winter trawl survey. Not sampled in 2004, small sample sizes 2003 and 2005.


Table 28. White perch relative abundance (N/tow) and total effort from the upper Chesapeake Bay winter trawl survey, 2000-2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ | $\begin{aligned} & \text { Sum } \\ & \text { CPE } \end{aligned}$ | Total effort |
| 2000 | 16.7 | 118.8 | 53.9 | 34.8 | 13.1 | 7.8 | 10.7 | 1.2 | 1.1 | 0.7 | 258.7 | 79 |
| 2001 | 24.5 | 47.1 | 75.7 | 14.5 | 22.1 | 4.8 | 11.6 | 12.3 | 2.5 | 1.7 | 217.3 | 114 |
| 2002 | 159.7 | 1.4 | 33.4 | 13.8 | 21.4 | 9.1 | 17.7 | 9.7 | 2.5 | 5.8 | 274.6 | 110 |
| 2003 | 83.3 | 156.1 | 28.7 | 13.1 | 18.2 | 20.9 | 73.9 | 1.7 | 0.0 | 9.9 | 405.8 | 20 |
| 2004 | NOT SAMPLED |  |  |  |  |  |  |  |  |  |  |  |
| 2005 | 22.6 | 39.2 | 10.7 | 19.7 | 5.0 | 1.8 | 0.6 | 5.6 | 0.6 | 0.3 | 106.1 | 43 |
| 2006 | 88.9 | 29.4 | 70.3 | 21.1 | 15.6 | 4.3 | 2.6 | 2.6 | 0.6 | 1.2 | 236.6 | 108 |
| 2007 | 35.5 | 23.9 | 17.3 | 33.9 | 19.5 | 4.7 | 5.4 | 0.4 | 0.4 | 1.9 | 142.9 | 71 |
| 2008 | 149.8 | 25.1 | 64.8 | 48.8 | 15.3 | 5.3 | 2.1 | 2.3 | 0.9 | 0.9 | 315.2 | 108 |
| 2009 | 64.9 | 180.3 | 7.6 | 33.0 | 62.1 | 52.4 | 1.3 | 18.1 | 3.8 | 0.7 | 424.2 | 90 |
| 2010 | 88.3 | 69.8 | 82.0 | 2.8 | 26.5 | 21.2 | 35.1 | 2.8 | 4.5 | 6.9 | 339.9 | 56 |
| 2011 | 32.9 | 39.0 | 27.7 | 37.4 | 9.1 | 20.7 | 11.3 | 11.5 | 0.6 | 2.0 | 192.3 | 66 |
| 2012 | 71.5 | 24.7 | 12.0 | 5.9 | 6.1 | 6.6 | 11.8 | 5.3 | 7.1 | 2.1 | 153.1 | 143 |

Table 29. White perch relative abundance (N/net day) and total effort from the Choptank River fyke net survey, 2000-2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | $10+$ | Sum <br> CPE | Total <br> effort |  |  |  |  |
| 2000 | 0.0 | 0.0 | 5.1 | 32.0 | 31.2 | 5.5 | 20.0 | 1.9 | 1.3 | 0.0 | 97.0 | 310 |  |  |  |  |
| 2001 | 0.0 | 7.0 | 16.0 | 47.9 | 35.8 | 26.2 | 4.2 | 11.0 | 1.5 | 0.0 | 149.6 | 310 |  |  |  |  |
| 2002 | 0.0 | 2.1 | 7.8 | 28.5 | 16.4 | 18.4 | 3.5 | 6.2 | 2.7 | 0.1 | 85.5 | 306 |  |  |  |  |
| 2003 | 0.0 | 2.2 | 36.8 | 33.6 | 33.3 | 1.4 | 27.7 | 7.2 | 3.2 | 3.2 | 148.5 | 261 |  |  |  |  |
| 2004 | 0.0 | 0.4 | 36.3 | 12.3 | 14.1 | 17.2 | 1.3 | 9.6 | 3.4 | 2.2 | 96.8 | 251 |  |  |  |  |
| 2005 | 0.0 | 3.4 | 16.0 | 51.2 | 32.1 | 19.9 | 7.2 | 1.7 | 10.8 | 0.5 | 142.7 | 235 |  |  |  |  |
| 2006 | 0.0 | 1.7 | 71.5 | 3.5 | 34.6 | 17.2 | 1.9 | 2.2 | 1.3 | 17.0 | 150.8 | 236 |  |  |  |  |
| 2007 | 0.0 | 1.3 | 9.5 | 123.8 | 13.4 | 57.8 | 20.7 | 8.2 | 9.0 | 7.2 | 250.8 | 203 |  |  |  |  |
| 2008 | 0.0 | 0.4 | 22.8 | 17.7 | 54.2 | 4.6 | 18.5 | 10.5 | 1.9 | 4.2 | 134.8 | 248 |  |  |  |  |
| 2009 | 0.0 | 1.8 | 0.7 | 24.9 | 6.8 | 45.2 | 5.5 | 8.5 | 4.9 | 3.1 | 101.3 | 210 |  |  |  |  |
| 2010 | 0.0 | 1.7 | 32.6 | 5.1 | 84.3 | 29.6 | 90.5 | 11.2 | 15.1 | 8.0 | 195.5 | 223 |  |  |  |  |
| 2011 | 0.0 | 0.1 | 1.0 | 22.0 | 3.5 | 21.0 | 12.9 | 15.8 | 2.3 | 4.2 | 82.7 | 242 |  |  |  |  |
| 2012 | 0.0 | 0.1 | 5.4 | 2.7 | 11.0 | 4.8 | 6.4 | 2.6 | 4.6 | 1.4 | 62.0 | 220 |  |  |  |  |

Table 30. Yellow perch relative abundance ( $\mathrm{N} /$ tow) and total effort from the upper Chesapeake Bay winter trawl survey, 2000-2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ | Sum <br> CPE | Total effort |
| 2000 | 0.6 | 1.0 | 0.2 | 1.1 | 0.0 | 0.2 | 0.1 | 0.0 | 0.0 | 0.1 | 3.1 | 79 |
| 2001 | 5.9 | 0.4 | 0.7 | 0.1 | 0.4 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 7.5 | 114 |
| 2002 | 10.6 | 7.7 | 0.8 | 1.6 | 0.1 | 0.8 | 0.0 | 0.1 | 0.0 | 0.0 | 21.7 | 110 |
| 2003 | 17.2 | 49.2 | 152.5 | 16.4 | 21.8 | 1.4 | 8.8 | 0.0 | 0.7 | 0.0 | 268.0 | 20 |
| 2004 | NOT SAMPLED |  |  |  |  |  |  |  |  |  |  |  |
| 2005 | 10.4 | 7.4 | 0.0 | 1.6 | 0.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 19.7 | 43 |
| 2006 | 14.1 | 16.1 | 6.8 | 0.0 | 1.4 | 0.2 | 0.0 | 0.1 | 0.0 | 0.0 | 38.6 | 108 |
| 2007 | 2.4 | 2.1 | 5.4 | 1.6 | 1.0 | 0.4 | 0.0 | 0.0 | 0.0 | 0.0 | 12.9 | 71 |
| 2008 | 9.8 | 2.4 | 5.3 | 4.7 | 1.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 23.3 | 108 |
| 2009 | 2.4 | 11.7 | 0.6 | 1.3 | 1.2 | 0.3 | 0.0 | 0.0 | 0.0 | 0.0 | 17.4 | 90 |
| 2010 | 15.4 | 1.8 | 4.6 | 0.3 | 0.5 | 0.2 | 0.1 | 0.0 | $<0.1$ | 0.0 | 22.9 | 56 |
| 2011 | 0.9 | 3.1 | 0.5 | 2.0 | 0.0 | 0.3 | 0.1 | 0.0 | 0.0 | 0.0 | 6.9 | 66 |
| 2012 | 10.6 | 4.3 | 1.5 | 0.1 | 0.1 | $<0.1$ | 0.0 | 0.4 | 0.0 | 0.0 | 17.1 | 107 |

Table 31. Yellow perch relative abundance (N/net day) and total effort from the Choptank River fyke net survey, 1988-2012.

| YEAR | AGE |  |  |  |  |  |  |  |  |  | Sum Total <br> CPE effort |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10+ |  |  |
| 1988 | 0.0 | 0.2 | 4.5 | 0.2 | 0.0 | 0.4 | 0.3 | 0.0 | 0.0 | 0.1 | 5.7 | 59 |
| 1989 | 0.0 | 0.0 | 1.2 | 3.4 | 1.2 | 0.6 | 0.1 | 0.0 | 0.0 | 0.0 | 6.6 | 68 |
| 1990 | 0.0 | 0.3 | 2.6 | 1.2 | 4.0 | 0.8 | 0.1 | 0.1 | 0.1 | 0.0 | 9.3 | 68 |
| 1991 | 0.0 | 0.1 | 0.6 | 0.8 | 0.3 | 0.6 | 0.1 | 0.0 | 0.0 | 0.0 | 2.5 | 70 |
| 1992 | 0.0 | 0.0 | 0.1 | 0.1 | 0.1 | 0.1 | 0.1 | 0.0 | 0.0 | 0.0 | 0.5 | 113 |
| 1993 | 0.0 | 0.0 | 0.6 | 1.3 | 0.8 | 0.9 | 0.3 | 0.1 | 0.0 | 0.0 | 4.0 | 120 |
| 1994 | 0.0 | 0.4 | 1.4 | 0.2 | 0.7 | 0.8 | 0.7 | 0.6 | 0.0 | 0.2 | 4.9 | 114 |
| 1995 | 0.0 | 0.7 | 2.1 | 0.2 | 0.6 | 0.6 | 0.3 | 0.3 | 0.0 | 0.2 | 5.0 | 121 |
| 1996 | 0.0 | 6.1 | 2.5 | 1.9 | 0.3 | 0.6 | 0.3 | 0.2 | 0.3 | 0.1 | 12.2 | 140 |
| 1997 | 0.0 | 0.1 | 4.2 | 0.6 | 0.6 | 0.0 | 0.1 | 0.2 | 0.1 | 0.0 | 5.8 | 153 |
| 1998 | 0.0 | 0.9 | 0.5 | 3.8 | 0.2 | 0.2 | 0.0 | 0.1 | 0.0 | 0.1 | 5.8 | 154 |
| 1999 | 0.0 | 1.7 | 47.8 | 0.5 | 17.7 | 0.2 | 0.1 | 0.0 | 0.0 | 0.0 | 68.0 | 178 |
| 2000 | 0.0 | 2.0 | 0.6 | 8.4 | 0.2 | 0.9 | 0.0 | 0.0 | 0.0 | 0.0 | 12.0 | 164 |
| 2001 | 0.0 | 5.3 | 11.9 | 0.6 | 6.8 | 0.1 | 0.4 | 0.0 | 0.0 | 0.0 | 25.1 | 167 |
| 2002 | 0.0 | 1.9 | 7.5 | 6.6 | 0.2 | 2.4 | 0.6 | 0.3 | 0.0 | 0.0 | 19.5 | 178 |
| 2003 | 0.0 | 3.1 | 3.6 | 7.6 | 2.8 | 0.3 | 1.9 | 0.3 | 0.3 | 0.0 | 19.8 | 121 |
| 2004 | 0.0 | 0.4 | 3.2 | 1.1 | 0.8 | 0.7 | 0.0 | 0.4 | 0.0 | 0.0 | 6.6 | 156 |
| 2005 | 0.0 | 9.0 | 0.7 | 2.2 | 0.7 | 0.3 | 0.8 | 0.1 | 0.3 | 0.1 | 14.2 | 186 |
| 2006 | 0.0 | 1.1 | 11.8 | 1.1 | 2.5 | 0.4 | 0.4 | 0.3 | 0.0 | 0.0 | 17.6 | 158 |
| 2007 | 0.0 | 10.8 | 5.3 | 11.1 | 0.2 | 1.3 | 0.8 | 0.2 | 0.1 | 0.1 | 29.9 | 140 |
| 2008 | 0.0 | 0.2 | 7.8 | 0.8 | 2.0 | 0.1 | 0.3 | 0.1 | 0.0 | 0.0 | 11.3 | 166 |
| 2009 | 0.0 | 0.0 | 6.1 | 14.8 | 1.0 | 0.9 | 0.2 | 0.0 | 0.0 | 0.0 | 23.0 | 143 |
| 2010 | 0.0 | 0.4 | 0.8 | 7.9 | 18.3 | 0.4 | 1.2 | 0.0 | 0.1 | 0.0 | 26.3 | 144 |
| 2011 | 0.0 | 1.2 | 0.0 | 0.2 | 4.6 | 5.6 | 0.3 | 0.7 | 0.0 | 0.0 | 12.6 | 158 |
| 2012 | 0.4 | 2.3 | 9.8 | 0.2 | 0.0 | 2.3 | 5.2 | $<0.1$ | 0.1 | 0.0 | 20.5 | 111 |

Figure 25. Choptank River yellow perch relative abundance from fyke nets, 1988-2012. Effort standardized from 1 March - $95 \%$ total catch date. Log-transformed trendline statistically significant at $\mathrm{P}=0.01$.

- Observed CPUE ——Predicted CPUE


Figure 26. Channel catfish relative abundance (N/tow) from the upper Chesapeake Bay winter trawl survey, 2000-2012. Not surveyed in 2004, small sample sizes in 2003 and 2005.


Figure 27. Channel catfish relative abundance (N/net day) from the Choptank River fyke net survey, 2000 - 2012. Horizontal line indicates time series average relative abundance.


Figure 28. White catfish relative abundance (N/net day) from the Choptank River fyke net survey, 2000-2012. Horizontal line indicates time series average relative abundance.


## PROJECT NO. 1

JOB NO. 2

# POPULATION ASSESSMENT OF CHANNEL CATFISH IN MARYLAND WITH SPECIAL EMPHASIS ON HEAD-OF-BAY STOCKS 

Prepared by Paul G. Piavis and Edward Webb, III

## INTRODUCTION

The objective of Job 2 was to assess channel catfish (Ictaluras punctatus) stock size, describe trends in recruitment, and relate current and historical mortality estimates to various biological reference points. Channel catfish were introduced into Maryland waters as early as the late 1800 's. Since those introductions, channel catfish have become self-sustaining, expanded their range, and are considered a naturalized species (Sauls et al 1998).

Channel catfish inhabit fresh or brackish waters in the Chesapeake Bay and its tributaries. Currently, recreational and commercial channel catfish fisheries are unregulated in tidal waters in Maryland (no minimum size limit, creel limit or seasonal closures). The Potomac River Fisheries Commission (PRFC) manages channel catfish in the Potomac River mainstem. The minimum size limit in the Potomac River is 203 mm ( 8 inches; TL) for commercial and recreational fisheries with no closed season or catch limits.

Channel catfish are important to recreational and commercial fishers throughout Maryland's portion of the Chesapeake Bay. The Marine Recreational Information Program (MRIP) produces estimates of recreational catch with fair precision (National Oceanic and Atmospheric Administration, personal communication, January 10, 2013). Estimated channel catfish recreational harvest (MRIP) averaged 240,600 lbs during 1982

- 2011; for the five year period, 2007-2011, average recreational catfish harvest was $417,100 \mathrm{lbs}$ ( $73 \%$ above the long term average). In 2011, channel catfish was the third largest recreational harvest in Maryland (by weight), trailing only striped bass and white perch.

In addition to MRIP estimates, recreational harvest estimates are available from geographically and temporally limited surveys. A Maryland Department of Natural Resources (MD DNR) creel survey conducted during the spring of 1985 in the lower Susquehanna River estimated that recreational fishers harvested 25,894 channel catfish (Weinrich et al. 1986). The estimated Susquehanna recreational harvest in 1985 was four times higher than any other year of the survey (1980-1984). Commercial harvest in the Susquehanna River and upper Chesapeake Bay region mimicked the increased recreational harvest over that same period.

Commercial channel catfish harvest peaked in 1996 at 2.45 million lbs and declined to $723,000 \mathrm{lbs}$ by 2007 before rising to near record levels of 2.17 million lbs in 2011. Channel catfish commercial landings (by weight) were second only to Atlantic menhaden during 2011. Areas above the Chesapeake Bay bridges accounted for $64 \%$ of the total Maryland commercial harvest in 2011, and averaged $60 \%$ of total landings during the five year period, 2007-2011.

Channel catfish populations were last assessed in 2009 (Piavis and Webb 2010). This Job is an update of the 2009 assessment. The 2009 assessment attempted to describe population dynamics in 3 systems, the Head-of-Bay (HOB; areas north of the Preston Lane Memorial Bridges), Choptank River, and Potomac River. However, the one-way trip nature of the Potomac River catfish indices made fitting population models
unreliable. For this report, channel catfish populations were modeled with a surplus production model for the HOB, and a Catch-Survey Analysis (CSA) for the Choptank River. For other systems, indices of relative abundance (fishery dependent and fishery independent, when available) were utilized to illustrate trends in population abundance.

## METHODS

## Bay-wide Landings

Maryland commercial fishery landings were available from the 1920's, but fishers were only required to report catch as general catfish landings (mixed species, predominately bullheads (Ameiurus spp.), channel catfish, and white catfish (Ameiurus catus) until 1996. Beginning in 1996, commercial fishers were required to report catfish landings as general, channel, or white catfish. The amount of channel catfish reported in the general category for the years 1996 - 2011 was calculated by determining the proportion of channel catfish in the combined white and channel catfish landings. This proportion was then multiplied by the amount of general catfish landed. The estimated annual landings of channel catfish in the general category were then added to the declared channel catfish landings for an estimated total commercial removal. To determine commercial channel catfish landing prior to 1996, the general catfish landings were multiplied by the average proportion of channel catfish of the total declared catfish landings by species for the years 1996 - 2011. Bullheads were considered an insignificant portion of landings prior to 1996.

Recreational landings, as estimated by the MRIP were fairly precise, but several years contained estimates where the proportional standard error (PSE) was $>40 \%$. A regression of estimated recreational harvests with PSE's $<0.40$ versus commercial landings was highly significant $\left(\mathrm{R}^{2}=0.88 \mathrm{P}<0.001\right)$. Therefore, estimated harvest from years with $\mathrm{PSE}<40 \%$ were compared to commercial landings to determine the average percentage of recreational landings to commercial landings. The average percentage was then applied to annual commercial harvest of years when PSE's of the recreational estimate exceeded $40 \%$.

## Head-of-Bay Surplus Production Model

Surplus production models fit biomass estimates to the equation

$$
\begin{equation*}
B_{t+1}=B_{t}+r B_{t}\left(1-B_{t} / K\right)-C_{t} \tag{1}
\end{equation*}
$$

where $r$ is the intrinsic rate of increase, $K$ is carrying capacity and $C_{t}$ is total removals in year $t$.

The model took the form of the Haddon (2001) implementation where a series of biomass estimates were generated to maximize a log-likelihood function by solving for r , K , and initial biomass $\left(B_{0}\right)$. An estimated index was derived from the equation

$$
\begin{equation*}
I=q\left[\left(B_{t+1}+B_{t}\right) / 2\right] e^{\mathrm{E}} \tag{2}
\end{equation*}
$$

where $I$ is the estimated index, $q$ is catchability and $e^{\varepsilon}$ is the lognormal residual error. This form simplified the solution by not having to solve for a catchability parameter for each index. In this closed form, average catchability for each index was $e^{(1 / n) \Sigma \ln \left(I_{t} / B\right)}{ }_{t}$.

The log function to be maximized was simply the sum of all log-likelihoods multiplied by a weighting factor.

The log-likelihood function for an individual index is

$$
\begin{equation*}
L L=-n / 2[\ln (2 \pi)+2 \ln (\sigma)+1] \tag{3}
\end{equation*}
$$

where $\sigma=\sqrt{\sum\left(\ln I_{t}-\ln I_{t}, \text { exp }\right)^{2} \sqrt{ } / n}$, and $n$ is the number of data points in the series. This assessment utilized an equal weighting scheme.

All runs were performed in an Excel spreadsheet using the Solver algorithm to estimate biomass and solve for the 3 unknown parameters ( $B_{0}, r, K$ ). Reference points and fishing mortality were estimated from standard relationships (Prager 1994; Haddon 2001):

Maximum Sustainable Yield $=r K / 4$
$\mathrm{B}_{\text {msy }}=K / 2$
$\mathrm{F}_{\mathrm{msy}}=r / 2$
Instantaneous fishing mortality $(F)=-\ln \left[1-C_{t} /\left(B_{t}+B_{t+1}\right) / 2\right]$.

## Model Inputs

There were five available indices of relative abundance available for modeling purposes. There were three fishery dependent indices (commercial CPUE's from the fyke net, pound net, and fish pot fisheries), and two fishery independent indices [Striped Bass Spawning Stock Survey (SBSSS), Project 2, Job 3, Task 2; and the Upper Bay winter trawl survey, Project 1 Job 1]. Positively correlated indices were identified, and a
final run was completed using the commercial fyke net CPUE index, the fishery independent drift gill net survey, and the bottom trawl survey.

The fishery dependent commercial fyke net index was derived from MD DNR Fisheries Service commercial landings database. Effort data for these gear types were available from 1980-1984, 1990, and 1992-2011. An index of effort was constructed to standardize landings because commercial catch reporting was completed monthly and not on a per trip basis. The index was nominal fishing effort, or simply the total number of nets declared by fishers in any month. Only fishers that reported catfish harvest > 500 lbs were used for relative abundance estimates. This eliminated fishers that were not targeting channel catfish. The final annual index was total pounds harvested divided by total nominal effort.

Fishery independent data from the experimental SBSSS in the HOB were compiled and included in the surplus production model (Figure 1). Since the model is a weight-based model, indices based on numbers were transformed to weight-based indices. Channel catfish weight per gill net set was estimated by determining average channel catfish length per mesh size per gill net set and applying a length-weight formula from the Susquehanna Flats area of the HOB (Fewlass 1980):

$$
\log _{10}(W)=3.09684 \times \log _{10}(L)-2.1622
$$

where W is weight $(\mathrm{g})$ and L is total length $(\mathrm{cm})$. The average weight per gill net set and mesh size was then multiplied by the total number captured per mesh size and net set. The final index was the geometric mean weight per net set standardized to 1000 -gill net yards $\times$ hours.

The fishery independent HOB winter trawl survey provided channel catfish relative abundance for the HOB (Figure 2). Species count data from this survey (20002002; 2006-2011) were transformed to biomass per tow with the same allometric equation utilized in the drift gill net index formulation. The index was geometric mean channel catfish biomass per tow for channel catfish greater than 355 mm . Observation of commercial fishing practices suggested that fish $<355 \mathrm{~mm}$ are not marketable.

Total removals by the commercial and recreational fisheries were estimated on a regional basis. Removals from HOB were easily obtained from the commercial landings data base because fishermen are required to submit landings by system. Recreational landings from HOB were estimated as the proportion of inland recreational landings (MRIP data) to bay-wide commercial landings, for all years pooled, multiplied by annual HOB commercial landings.

## Uncertainty

Bootstrapping, or resampling residuals and adding them to the natural logarithm of the expected indices, and re-exponentiating the values was used to quantify model uncertainty ( $n=1,000$ trials). Mean, median, standard deviation and coefficient of variation were calculated for all fitted parameters and each estimate of annual biomass and F. Confidence intervals ( $80 \% \mathrm{CI}$ ) were determined from cumulative percent distributions of the bootstrapped parameter estimates.

## Choptank River Catch-Survey Analysis (CSA)

The CSA relates pre-recruit relative abundance to post-recruit relative abundance in the following year, such that:

$$
\begin{equation*}
R_{t+1}=\left(R_{t}+P_{t}\right) e^{-M_{t}}-C_{t} e^{-M_{t}\left(1-T_{t}\right)} \tag{4}
\end{equation*}
$$

where $R$ is the post-recruit abundance, $P$ is the pre-recruit abundance, $M$ is instantaneous natural mortality, $C$ is harvest, and $T$ is the fraction of time between the survey and the harvest.

The model assumes survey catch $r$ and $p$ for post-recruits and pre-recruits, respectively, relate to abundance by a survey catchability $(q)$ such that:

$$
\begin{equation*}
r_{t}=R_{t} q \tag{5}
\end{equation*}
$$

and,

$$
\begin{equation*}
p_{t}=P_{t} q \Phi \tag{6}
\end{equation*}
$$

where $\Phi$ is a scalar relating the pre-recruit catchability to post-recruit catchability.
Substituting [5] and [6] into equation [4] yields

$$
\begin{equation*}
r_{t+1}=\left(r_{t}+p_{t} / \Phi\right) e^{-M}-q C_{t} e^{-M_{t}\left(1-T_{t}\right)} \tag{7}
\end{equation*}
$$

Adding a process error term ( $\varepsilon$ ) into [7] yields

$$
\begin{equation*}
r_{t+1}=\left(r_{t}+p_{t} / \Phi\right) e^{-M}-q C_{t} e^{-M_{t}\left(1-T_{t}\right) \mathrm{E}} \tag{8}
\end{equation*}
$$

Measurement error ( $\eta$ and $\delta$ ) is similarly incorporated into [5] and [6]

$$
\begin{gather*}
p_{t}=P_{t} q e \eta  \tag{9}\\
r_{t}=R_{t} q \Phi e^{\delta} \tag{10}
\end{gather*}
$$

Collie and Kruse (1998) advocated using a single error model structure. The allobservation error structure produced similar results to the mixed error model and was less likely to be over parameterized (Collie and Kruse 1998). This approach produced the objective function to be minimized:

$$
\begin{equation*}
\mathrm{SSQ}=\lambda_{\eta} \Sigma \eta^{2}+\lambda_{\delta} \Sigma \delta^{2} \tag{11}
\end{equation*}
$$

This yields $i+1$ parameters to be estimated with $i-2 \mathrm{df}$. The model was compiled in a Microsoft Excel spreadsheet and the Solver routine was used to fit the model.

Population size of fully recruited fish $\left(\mathrm{R}_{t}\right)$ for the Choptank River was estimated as $\mathrm{r}_{t} / q$ and the population size of pre-recruits $\left(\mathrm{P}_{t}\right)$ was $\mathrm{p}_{t} / \Phi q$. Harvest rate $h$ was estimated as

$$
\begin{equation*}
h_{t}=C_{t} /\left[\left(P_{t}+R_{t}\right) e^{-M_{t} T_{t}} .\right. \tag{12}
\end{equation*}
$$

Total instantaneous fishing mortality (F) was

$$
\begin{equation*}
F_{t}=-\log _{e}\left(1-h_{t}\right) . \tag{13}
\end{equation*}
$$

## Model Inputs

Pre-recruit and post-recruit indices of abundance were determined from MD DNR Fisheries Service fyke net catches (Figure 3; Project 1 Job 1). Pre-recruits were those channel catfish $<356 \mathrm{~mm}$. Post-recruit channel catfish were those fish $>355 \mathrm{~mm}$ TL. Numbers of pre-recruit and post-recruit channel catfish were determined for each fyke net visit by applying the proportion of pre-recruit and post-recruit channel catfish from the length subsample to the total catch. Numbers of pre- and post-recruit channel catfish from each net lift were divided by gear soak time. The final indices were the arithmetic mean of each net CPUE.

Harvest estimates were determined for the commercial and recreational fisheries. Numbers of commercially harvested channel catfish were determined by dividing pounds harvested (by gear type) by estimated average weight of legal channel catfish. Average legal weight was determined from our fyke net catches. The same allometric equation used for the HOB analysis was used to transform average length to average weight.

Recreational channel catfish harvest for the Choptank River was estimated from total inland harvest estimates from the MRIP (National Marine Fisheries Service, personal communication, January 10, 2013). The proportion of recreational to commercial landings was determined by dividing total recreational inland landings by bay-wide commercial landings. That proportion was applied to Choptank River commercial landings to estimate recreational landings in this system. Negligible release losses were assumed for all fisheries.

Relative catchability of pre-recruits $(\Phi)$ was set at 1.0 because length-frequencies indicated that channel catfish were recruited to the survey gear. Natural mortality ( $M$ ) was 0.20 . An initial catchability for the runs was set at $5.0 \times 10^{-6}$. The fraction of year that the survey preceded the fishery $(\mathrm{T})$ was 0.5 .

## Uncertainty

The model was bootstrapped 5,000 times by resampling residuals and adding them to the natural logarithm of the expected index values, and then re-exponentiating the values. Mean, median, standard deviation and CV's were calculated for $q$ and each estimate of $\mathrm{P}_{t}$ and $\mathrm{R}_{t}$, exclusive of the terminal year. Confidence intervals ( $80 \%$ ) were determined from cumulative percent distributions of the bootstrapped parameter estimates.

## Other Areas

Previous attempts to fit population models to other areas have failed, largely due to lack of fishery independent surveys (Piavis and Webb 2010). Qualitative methods to describe population trends in Nanticoke, Patuxent, and Potomac rivers were employed.

## Landings

Channel catfish landings were determined from MD DNR commercial landings database for the Nanticoke and Patuxent rivers. Adjustments due to changes in the species reporting requirements were identical to the bay-wide landings discussed above. The Potomac River Fisheries Commission (PRFC) provided commercial landings from the Potomac River (Potomac River Fisheries Commission, personal communication, February 20, 2013). Catfish landings were identified to species from 2003-2012. From 1985 - 2002, catfish were coded as mixed (white catfish and channel catfish) and bullhead species. Channel catfish landings for the period 1985 - 2002 were estimated as mixed catfish landings $\times$ proportion of channel catfish of total catfish landings during the nearest 5 year period, 2003 - 2007 (0.85). From 1964 - 1984, catfish landings were reported as mixed bullhead and catfish species. Channel catfish landings for the period 1964 - 1984 were estimated as catfish landings $\times$ proportion of channel catfish of total landings during the period $1985-2002$.

## Fishery Dependent Relative Abundance Indices

Area specific relative abundance indices were determined from the fishery dependent commercial landings database. The indices were computed in the same manner as detailed in the Model Inputs section above for the HOB surplus production model. Gear specific indices were constructed for the fyke net, pound net, and fish pot fisheries.

## Fishery Independent Relative Abundance Indices

A gill net survey designed to estimate spawning stock biomass of striped bass in Potomac River (SBSSS) was utilized to describe population trends (Figure 1). This
survey is analogous to the drift gill net survey in HOB that was included in the HOB surplus production model. However, the Potomac index was included as a numbers based index instead of transforming to a biomass index as required by the surplus production model. Data encompassed the time period 1984 - 2012.

Channel catfish juvenile recruitment was determined from the Estuarine Juvenile Finfish Survey (EJFS; Project 2, Job 3, Task 3). The EJFS is designed to estimate young-of-year striped bass (Morone saxatilis) relative abundance, but it has proved valuable in determining year-class strength of other species as well. Relative juvenile abundance indices were available for the Nanticoke, Potomac, and Patuxent rivers (Figure 4).

## RESULTS

## Landings

Baywide commercial landings generally varied between 400,000 pounds and 700,000 pounds from 1929 through the mid-1970's (Figure 5). Landings increased rapidly from 1976 through 1996 to a time series maximum of 2.4 million pounds. Since 1996, landings decreased to a recent low in 2007, and then increased to a near time series high in 2011. The 2011 harvest was 2.1 million pounds. Baywide recreational landings estimates have varied greatly over the period 1983-2011 (Figure 6). A time series low was estimated in 1988, but recreational landings trended upward through 1996, corresponding to the rise in commercial landings. Recreational landings during the period 1997 - 2007 were notably low, but a general rebound occurred during 2007 2011.

## Head-of-Bay Surplus Production Model

Total estimated fishery removals from HOB, by weight, exhibited a dome-shaped pattern for much of the assessment time-period (1980 - 2011). However, landings increased from 0.4 million pounds to 1.7 million pounds over the period $2005-2011$ (commercial and recreational combined; Figure 7). The model included three biomass indices, a fishery dependent fyke net index (1980-1984, 1990, 1991-2011), and two fishery independent indices (the gill net survey, 1985-2011; and the winter trawl survey, 2000 - 2011). The fyke net index exhibited a bimodal pattern with one peak in 1990 and a broader peak covering the years 2006 - 2009 (Figure 8). The fishery independent gill net survey indicated relatively high index values during 1985-1987, a time period where no fyke net index was available. The gill net index corroborated the higher fyke net index during the last three years of the time series (Figure 9). The winter trawl survey also validated the increased biomass over the last 4 years of the assessment period (Figure 10), but this index suggested that biomass was at its highest in 2011 whereas the fyke net index and gill net index suggested some decline over the period 2010-2011.

The model fit the data well. Estimated parameters $r, K$, and $B_{0}$ were $0.68,8.7$ million pounds, and 2.2 million pounds, respectively. Biomass increased from 2.2 million pounds in 1980 to 7.5 million pounds in 1989. Channel catfish biomass then trended lower to 3.8 million pounds in 2000 , but nearly doubled to 7.6 million pounds in 2010. The final year biomass estimate (2011) was 6.8 million pounds (Figure 11). Instantaneous fishing mortality (F) peaked from 1996-1999, but then fell to low levels during 2004 - 2010. Instantaneous fishing mortality in the final year of the assessment (2011) was estimated to be 0.29 (Figure 11). Over the course of the assessment, F
averaged 0.24 . Biomass at maximum sustainable yield ( $\mathrm{B}_{\mathrm{msy}}$ ) was estimated as $1 / 2 \mathrm{~K}$ or 4.4 million pounds. $\mathrm{F}_{\text {msy }}$ was estimated as $1 / 2 r$ or 0.34 . Maximum sustainable yield was estimated $r K / 4$, or 1.5 million pounds.

Previous studies have indicated that the absolute values for biomass and fishing mortality from surplus production models may not be particularly precise, but the ratios of $\mathrm{B}: \mathrm{B}_{\mathrm{msy}}$ and $\mathrm{F}: \mathrm{F}_{\mathrm{msy}}$ are particularly robust (Prager 1994). Ratios of $\mathrm{B}: \mathrm{B}_{\mathrm{msy}}$ and $\mathrm{F}: \mathrm{F}_{\mathrm{msy}}$ indicated a period of increasing surplus biomass and moderate F between 1983 and the mid 1990's. Fishing mortality then rose to unsustainable levels for six of nine years during $1995-2003$, that is, the $\mathrm{F}: \mathrm{F}_{\text {msy }}$ ratio was greater than 1.0 (Figure 12). After 2003, the $\mathrm{F}: \mathrm{F}_{\text {msy }}$ ratio declined and the $\mathrm{B}: \mathrm{B}_{\text {msy }}$ ratio increased. The $\mathrm{B}: \mathrm{B}_{\text {msy }}$ and $\mathrm{F}: \mathrm{F}_{\text {msy }}$ ratios in the final year of the assessment were 1.55 and 0.86 , respectively. Based on these point estimates, the HOB channel catfish stock is not overfished and overfishing is not occurring.

Bootstrapping provided estimates of uncertainty for this model (Table 1). The bootstrap procedure produced 983 valid trials out of 1,000 attempts ( $98.3 \%$ ). The intrinsic rate of increase ( $r$ ) was precisely estimated ( $\mathrm{CV}=29 \%$ ). Estimates of $K$ and $B_{0}$ were less precisely estimated with CV's equal to $35 \%$ and $40 \%$, respectively. Initial biomass $\left(B_{0}\right)$ is generally regarded as a nuisance parameter that has lower importance than $r$ and $K$ in model outputs and subsequent management advice. Coefficients of variation of annual biomass estimates ranged from $19 \%-37 \%$. In contrast, the ratio $B: B_{\text {msy }}$ was very precisely estimated in all years (CV range $=6 \%-19 \%$ ). Comparisons of the confidence intervals also demonstrate the increased precision of the ratio estimates (Figures 13 and 14). Coefficients of variation of annual fishing mortality estimates
ranged from $19 \%-48 \%$. In contrast, the ratio $\mathrm{F}: \mathrm{F}_{\mathrm{msy}}$ was precisely estimated in all years $(\mathrm{CV}$ range $=12 \%-28 \%)$. Comparisons of the confidence intervals also demonstrate the increased precision of the ratio estimates (Figures 15 and 16). In the final year of the assessment (2011), there was only a $0.7 \%$ chance that channel catfish biomass was below $\mathrm{B}_{\mathrm{msy}}$, and a $5.9 \%$ chance that overfishing was occurring (i.e., $\mathrm{F}: \mathrm{F}_{\mathrm{msy}}$ $>1.0$ ).

## Choptank River Catch-Survey Analysis (CSA)

Total channel catfish removal the from Choptank River, in numbers, was estimated for the assessment time period 1993-2011. Commercial and recreational harvest was generally low during 1993 - 2004, ranging from 20,000 - 50,000 fish, except for the nearly 100,000 fish estimated for 1999. After 2004, harvest increased substantially, ending in 2011 at a time series high (Figure 17). The model included two indices from a MD DNR Fisheries Service fishery independent fyke net survey. One index was a pre-recruit relative abundance index and the other was a post-recruit relative abundance index. The pre-recruit index remained in a low range, relative to the entire time series, from 1995 - 2006. The pre-recruit index increased after 2006, more than doubling the previous high relative abundance value (Figure 18). The post-recruit index indicated a similar pattern, but the higher relative abundance of the recruited fish did not begin until 2008 and ended the time series with the four highest relative abundance values in the last five years of the of the survey (Figure 19).

The CSA model fit the population data moderately well. Catchability of the survey $(q)$ was estimated as $1.85 \times 10^{-6}$. Pre-recruit population abundance generally
tracked the increase in the survey's relative abundance values, with relatively low prerecruit abundance during 1995 - 2004, followed by relatively high pre-recruit abundance for the remainder of the time series (Figure 20). Post-recruit channel catfish abundance varied between 200,000 and 400,000 channel catfish from 1993 - 2007 (Figure 20). After 2007, recruited channel catfish abundance accelerated quite swiftly with the recruited population increasing from an estimated 664,000 fish in 2008 to 1.06 million fish in 2011. Instantaneous fishing mortality (F) was generally low, varying between 0.04 and 0.15 for most of the assessment period (Figure 21). Average F for the entire time series was 0.13 and F in the final year of the assessment was 0.11 . No F-based, biomass-based, or abundance-based biological reference points have been determined for Chesapeake Bay area channel catfish stocks. Therefore, no conclusions may be definitively drawn regarding overfishing or overfished status for Choptank River channel catfish stocks. Model outputs and survey results strongly suggest that fishing mortality at recent levels is not impacting population growth.

Bootstrapping provided estimates of uncertainty for this model (5,000 trials; Table 2). Survey catchability (q) was precisely estimated (CV=22\%). Coefficients of variation for pre-recruit abundance estimates ranged from $34 \%-41 \%$ with some of the highest CV's in the last 5 years of the assessment. Coefficients of variation for postrecruit abundance were more variable than the pre-recruit abundances. Coefficients of variation ranged from $29 \%-52 \%$. Again, a temporal trend is evidenced where the higher CV's occur in the latter years of the assessment. Total population size (pre-recruit and post-recruit abundances) provided a better fit, with CV's ranging from $28 \%-41 \%$. Fishing mortality estimates were also precisely estimated with CV's ranging from 23\% -
$36 \%$. Graphs of confidence intervals for population estimates and F estimates indicate that in general, population estimates may be biased high, and F may be biased low (Figures $22-25$ ). In addition these graphs also depict the temporal uncertainty in the population estimates in the latter part of the time series.

## Other Areas

Nanticoke River channel catfish data included commercial fishery landings, fishery dependent relative abundance, and a fishery independent seine survey. Commercial landings from 1987 - 2011 were variable ranging form just under 20,000 pounds to 145,0000 pounds (Figure 26). Commercial fishery CPUE's were quite variable and exhibited no discernable trend (Figures 27, 28). Young-of-year production, as determined from the EJFS seine survey is also not definitive, but production was more consistent during 1989-1997 than in recent years (Figure 29).

Patuxent River channel catfish data included commercial fishery landings, fishery dependent relative abundance, and a fishery independent seine survey. Patuxent River landings have been trendless throughout the past 25 years (Figure 30). Only the fish pot relative abundance index provided a complete enough time series to warrant investigation. Relative abundance values have been trending downward since 2006 (Figure 31). Young-of-year production, as determined from the EJFS seine survey indicated that the last years of high juvenile abundance were in 2001 and 2003 (Figure 32). No juvenile channel catfish were encountered in 2006, 2008, 2010, or 2012.

Potomac River channel catfish landings, as report to Potomac River Fishery Commission, had to be adjusted for difference in reporting requirements similar to
landings from the MD DNR commercial database. Estimated landings of channel catfish from Potomac River showed a peak in 1964, a relatively broad peak in the mid 1980's followed by a rapid decline with landings generally less than 100,000 pounds since 1991 (Figure 33). The Potomac River drift gill net survey indicated that the biomass index was below the $75^{\text {th }}$ percentile since 2005 (Figure 34). Young-of-year production, as determined from the EJFS seine survey, indicated low and intermittent juvenile production since 1985 (Figure 35).

## DISCUSSION

Channel catfish provide valuable recreational and commercial fisheries while occupying an important ecological niche among brackish-tidal fresh ecosystems in Maryland's portion of the Chesapeake Bay. The primary objective of this Job was to describe trends in channel catfish abundance throughout the Bay region. Model runs proved informative for HOB and Choptank River channel catfish populations. These areas accounted for $77 \%$ of total MD commercial channel catfish harvest. Channel catfish populations in Nanticoke, Patuxent, and Potomac rivers were assessed through qualitative examination of available relative abundance data.

The HOB surplus production model indicated a period of population increase from 1980 - 1989 followed by a decline through 2000 (estimated as 3.8 million pounds). Since 2000, population biomass increased to an average of 7.1 million pounds over the last 5 year period, 2007 - 2011. These results generally mimic the original model run (Piavis and Webb 2010).

Maximum sustainable yield (MSY) was identified as 1.48 million pounds. Total estimated removals were above MSY in only 2 years during the expansion/plateau phase
of channel catfish abundance (1980 - 1993). Total estimated removals exceeded MSY in each year except 2000 during the period when the population contracted (1994-2000). Recently, harvest (commercial and recreational) was above MSY in 2003 and 2011. The population biomass during 2011 was $55 \%$ higher than $\mathrm{B}_{\text {msy }}\left(\mathrm{B}_{\text {msy }}=\right.$ the population biomass that can sustain harvest at MSY), given that the $\mathrm{B}: \mathrm{B}_{\text {msy }}$ ratio for 2011 was 1.55 . $\mathrm{AB}: \mathrm{B}_{\text {msy }}$ ratio greater than 1.0 indicates that the stock is not overfished. This metric has proven more robust than absolute biomass values from surplus production modeling (Prager 1994). The robustness of the ratio estimates becomes evident with the inspection of uncertainty parameters. Confidence intervals are tighter around B ratio estimates (CV range: $6 \%-19 \%$ ) than absolute B estimates (CV range: $19 \%-37 \%$ ).

Inspection of the trajectories of F moved opposite that of biomass. As F increased, the population biomass stabilized until F increased beyond $\mathrm{F}_{\text {msy }}$, at which point population biomass contracted. Conversely, the period beginning in 2000 had F rates below $\mathrm{F}_{\mathrm{msy}}$ and population biomass expanded. In the final year of the assessment, the $\mathrm{F}: \mathrm{F}_{\mathrm{msy}}$ ratio was $0.86 . \mathrm{F}: \mathrm{F}_{\text {msy }}$ ratios less than 1.0 indicate that overfishing is not occurring. Similar to the $\mathrm{B}: \mathrm{B}_{\text {msy }}$ ratio, the F ratio is a more robust estimate of the status of F than absolute values (Prager 1994).

The winter trawl survey (Project 1 Job 1) has limited temporal coverage, but the trawl survey results indicated strong year-classes for the 2004, 2006, 2008, and 2011 cohorts. Given expected growth rates, the increased biomass in recent years is attributed to the higher juvenile production of the 2004 and 2006 year-classes. The 2008 and 2011 year-classes should sustain population expansion for future years if the commercial and recreational fisheries remain stable.

The Choptank River assessment was the first assessment of channel catfish using a CSA. Population trajectories indicated an expanding population which closely tracked our experimental fyke net indices. No biological reference points exist to determined overfished or overfishing status, but given that populations are estimated at time series highs, overfishing is unlikely to have occurred for extended periods of time in the Choptank River channel catfish fishery.

Uncertainty analysis indicated that abundance parameters were only moderately estimated. Relative abundance indices for both pre- and post-recruit fish were at baseline levels compared to values later in the time series. The abundance increase over the last few years provides the only contrast in population size. Magnusson and Hilborn (2007) investigated what population trajectories and models provided informative fishery management advice. Although the authors did not investigate CSA type models, results indicated that fishery population models that performed the best did so when there were sustained contrasting periods of population abundance. Given that our results show a relatively recent increase in abundance, a full population cycle may help increase precision.

Channel catfish relative abundance trends in Nanticoke, Patuxent, and Potomac rivers are largely uninformative, but general trends are evident. Fishery dependent CPUE's in Nanticoke River have been hovering around the $75^{\text {th }}$ percentile for several years indicating a generally stable population. Patuxent River fishery dependent CPUE's and Potomac River fishery independent CPUE's have been below the $75^{\text {th }}$ percentile for some time. Meaningful population increases are unlikely, given that juvenile indices in all three systems have been weak over the last several years. However, juvenile indices
reported from seine catches may not be the best indicator of juvenile production (Piavis and Webb 2010).

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Table 1. Uncertainty parameters for Head-of-Bay channel catfish surplus production model.

| Parameter/Year | Estimate | Mean | Median | Std Dev | C.V. |
| :--- | ---: | ---: | ---: | ---: | ---: |
| $r$ | 0.679 | 0.655 | 0.678 | 0.195 | 0.288 |
| $K_{1}$ | $8,731,750$ | $9,664,473$ | $9,153,815$ | $3,367,350$ | 0.348 |
| $B_{0}$ | $2,249,834$ | $2,619,064$ | $2,428,093$ | $1,038,700$ | 0.397 |
|  |  |  |  |  |  |
| $B_{1981}$ | $2,821,114$ | $3,231,434$ | $3,046,373$ | $1,136,196$ | 0.352 |
| $B_{1982}$ | $3,652,119$ | $4,096,747$ | $3,883,035$ | $1,218,638$ | 0.297 |
| $B_{1983}$ | $4,654,158$ | $5,105,782$ | $4,884,121$ | $1,268,241$ | 0.248 |
| $B_{1984}$ | $5,373,523$ | $5,779,725$ | $5,571,109$ | $1,287,020$ | 0.223 |
| $B_{1985}$ | $6,077,590$ | $6,424,085$ | $6,216,390$ | $1,302,695$ | 0.203 |
| $B_{1986}$ | $6,259,858$ | $6,560,064$ | $6,366,858$ | $1,339,972$ | 0.204 |
| $B_{1987}$ | $6,627,386$ | $6,922,635$ | $6,730,080$ | $1,370,500$ | 0.198 |
| $B_{1988}$ | $7,115,460$ | $7,423,628$ | $7,242,488$ | $1,424,805$ | 0.192 |
| $B_{1989}$ | $7,474,514$ | $7,815,591$ | $7,653,198$ | $1,516,023$ | 0.194 |
| $B_{1990}$ | $6,874,531$ | $7,269,466$ | $7,107,412$ | $1,623,843$ | 0.223 |
| $B_{1991}$ | $6,422,310$ | $6,869,635$ | $6,685,288$ | $1,631,443$ | 0.237 |
| $B_{1992}$ | $6,600,940$ | $7,080,139$ | $6,878,927$ | $1,618,894$ | 0.229 |
| $B_{1993}$ | $6,419,896$ | $6,918,528$ | $6,707,599$ | $1,638,142$ | 0.237 |
| $B_{1994}$ | $6,479,235$ | $6,994,993$ | $6,767,826$ | $1,643,919$ | 0.235 |
| $B_{1995}$ | $6,022,445$ | $6,549,684$ | $6,312,031$ | $1,663,016$ | 0.254 |
| $B_{1996}$ | $5,533,883$ | $6,074,234$ | $5,828,579$ | $1,650,997$ | 0.272 |
| $B_{1997}$ | $4,431,561$ | $4,982,211$ | $4,728,953$ | $1,628,196$ | 0.327 |
| $B_{1998}$ | $4,396,059$ | $4,961,798$ | $4,680,998$ | $1,616,503$ | 0.326 |
| $B_{1999}$ | $3,918,258$ | $4,494,430$ | $4,207,879$ | $1,609,432$ | 0.358 |
| $B_{2000}$ | $3,814,214$ | $4,404,857$ | $4,073,018$ | $1,629,558$ | 0.370 |
| $B_{2001}$ | $4,382,048$ | $4,981,695$ | $4,664,610$ | $1,662,089$ | 0.334 |
| $B_{2002}$ | $4,707,414$ | $5,285,731$ | $4,985,547$ | $1,665,146$ | 0.315 |
| $B_{2003}$ | $5,389,065$ | $5,931,705$ | $5,641,595$ | $1,659,903$ | 0.280 |
| $B_{2004}$ | $4,966,816$ | $5,453,140$ | $5,170,563$ | $1,650,481$ | 0.303 |
| $B_{2005}$ | $5,562,602$ | $6,026,080$ | $5,751,897$ | $1,644,692$ | 0.273 |
| $B_{2006}$ | $6,536,262$ | $6,961,828$ | $6,667,953$ | $1,643,980$ | 0.236 |
| $B_{2007}$ | $6,923,565$ | $7,312,886$ | $7,050,980$ | $1,692,339$ | 0.231 |
| $B_{2008}$ | $7,350,280$ | $7,756,405$ | $7,499,247$ | $1,749,995$ | 0.226 |
| $B_{2009}$ | $7,161,607$ | $7,609,118$ | $7,359,249$ | $1,847,713$ | 0.243 |
| $B_{2010}$ | $7,210,083$ | $7,711,917$ | $7,460,331$ | $1,898,155$ | 0.246 |
| $B_{2011}$ | $6,780,867$ | $7,325,968$ | $7,070,692$ | $1,961,292$ | 0.268 |
|  |  |  |  |  |  |

Table 1. (Continued)

| Parameter/Year | Estimate | Mean | Median | Std Dev | C.V. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{F}_{1980}$ | 0.288 | 0.291 | 0.264 | 0.140 | 0.481 |
| $\mathrm{F}_{1981}$ | 0.181 | 0.177 | 0.166 | 0.065 | 0.370 |
| $\mathrm{F}_{1982}$ | 0.129 | 0.124 | 0.121 | 0.037 | 0.296 |
| $\mathrm{F}_{1983}$ | 0.178 | 0.171 | 0.168 | 0.043 | 0.249 |
| $\mathrm{F}_{1984}$ | 0.140 | 0.136 | 0.134 | 0.029 | 0.215 |
| $\mathrm{F}_{1985}$ | 0.194 | 0.191 | 0.189 | 0.038 | 0.199 |
| $\mathrm{F}_{1986}$ | 0.143 | 0.142 | 0.141 | 0.029 | 0.201 |
| $\mathrm{F}_{1987}$ | 0.094 | 0.094 | 0.093 | 0.018 | 0.190 |
| $\mathrm{F}_{1988}$ | 0.078 | 0.078 | 0.077 | 0.014 | 0.187 |
| $\mathrm{F}_{1989}$ | 0.196 | 0.195 | 0.191 | 0.041 | 0.209 |
| $\mathrm{F}_{1990}$ | 0.236 | 0.236 | 0.227 | 0.062 | 0.262 |
| $\mathrm{F}_{1991}$ | 0.165 | 0.163 | 0.158 | 0.041 | 0.254 |
| $\mathrm{F}_{1992}$ | 0.215 | 0.210 | 0.205 | 0.051 | 0.240 |
| $\mathrm{F}_{1993}$ | 0.187 | 0.183 | 0.178 | 0.045 | 0.245 |
| $\mathrm{F}_{1994}$ | 0.282 | 0.275 | 0.268 | 0.069 | 0.251 |
| $\mathrm{F}_{1995}$ | 0.345 | 0.338 | 0.326 | 0.097 | 0.287 |
| $\mathrm{F}_{1996}$ | 0.594 | 0.591 | 0.554 | 0.223 | 0.377 |
| $\mathrm{F}_{1997}$ | 0.419 | 0.417 | 0.387 | 0.165 | 0.396 |
| $\mathrm{F}_{1998}$ | 0.591 | 0.590 | 0.543 | 0.256 | 0.434 |
| $\mathrm{F}_{1999}$ | 0.513 | 0.510 | 0.468 | 0.221 | 0.433 |
| $\mathrm{F}_{2000}$ | 0.266 | 0.259 | 0.247 | 0.094 | 0.363 |
| $\mathrm{F}_{2001}$ | 0.307 | 0.297 | 0.285 | 0.097 | 0.328 |
| $\mathrm{F}_{2002}$ | 0.184 | 0.178 | 0.173 | 0.051 | 0.287 |
| $\mathrm{F}_{2003}$ | 0.413 | 0.402 | 0.391 | 0.114 | 0.284 |
| $\mathrm{F}_{2004}$ | 0.190 | 0.186 | 0.182 | 0.051 | 0.276 |
| $\mathrm{F}_{2005}$ | 0.074 | 0.073 | 0.072 | 0.017 | 0.230 |
| $\mathrm{F}_{2006}$ | 0.118 | 0.116 | 0.116 | 0.024 | 0.203 |
| $\mathrm{F}_{2007}$ | 0.082 | 0.081 | 0.081 | 0.017 | 0.205 |
| $\mathrm{F}_{2008}$ | 0.143 | 0.141 | 0.140 | 0.029 | 0.208 |
| $\mathrm{F}_{2009}$ | 0.123 | 0.121 | 0.119 | 0.028 | 0.232 |
| $\mathrm{F}_{2010}$ | 0.196 | 0.193 | 0.189 | 0.046 | 0.238 |
| $\mathrm{F}_{2011}$ | 0.293 | 0.289 | 0.279 | 0.081 | 0.280 |

Table 1. (Continued).

| Parameter/Year | Estimate | Mean | Median | Std Dev | C.V. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (B/B MSY$)_{1980}$ | 0.515 | 0.536 | 0.527 | 0.100 | 0.186 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1981}$ | 0.646 | 0.671 | 0.665 | 0.123 | 0.184 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1982}$ | 0.837 | 0.864 | 0.853 | 0.161 | 0.187 |
| $\left(\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}\right)_{1983}$ | 1.066 | 1.091 | 1.077 | 0.209 | 0.192 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1984}$ | 1.231 | 1.240 | 1.238 | 0.221 | 0.178 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1985}$ | 1.392 | 1.380 | 1.393 | 0.221 | 0.160 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1986}$ | 1.434 | 1.404 | 1.436 | 0.195 | 0.139 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1987}$ | 1.518 | 1.481 | 1.517 | 0.189 | 0.128 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1988}$ | 1.630 | 1.587 | 1.625 | 0.185 | 0.117 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1989}$ | 1.712 | 1.667 | 1.709 | 0.173 | 0.104 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1990}$ | 1.575 | 1.537 | 1.564 | 0.133 | 0.087 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1991}$ | 1.471 | 1.448 | 1.463 | 0.118 | 0.081 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1992}$ | 1.512 | 1.496 | 1.513 | 0.122 | 0.081 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1993}$ | 1.470 | 1.459 | 1.473 | 0.113 | 0.077 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1994}$ | 1.484 | 1.476 | 1.488 | 0.113 | 0.077 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1995}$ | 1.379 | 1.376 | 1.381 | 0.100 | 0.073 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1996}$ | 1.268 | 1.271 | 1.269 | 0.096 | 0.075 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1997}$ | 1.015 | 1.031 | 1.021 | 0.103 | 0.100 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1998}$ | 1.007 | 1.029 | 1.014 | 0.114 | 0.111 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{1999}$ | 0.897 | 0.928 | 0.910 | 0.131 | 0.142 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2000}$ | 0.874 | 0.910 | 0.891 | 0.153 | 0.168 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2001}$ | 1.004 | 1.040 | 1.021 | 0.176 | 0.169 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2002}$ | 1.078 | 1.109 | 1.097 | 0.186 | 0.168 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2003}$ | 1.234 | 1.254 | 1.256 | 0.199 | 0.158 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2004}$ | 1.138 | 1.146 | 1.159 | 0.174 | 0.152 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2005}$ | 1.274 | 1.275 | 1.290 | 0.192 | 0.150 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2006}$ | 1.497 | 1.482 | 1.499 | 0.206 | 0.139 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2007}$ | 1.586 | 1.552 | 1.585 | 0.176 | 0.114 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2008}$ | 1.684 | 1.645 | 1.683 | 0.161 | 0.098 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2009}$ | 1.640 | 1.605 | 1.640 | 0.126 | 0.079 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2010}$ | 1.651 | 1.624 | 1.652 | 0.115 | 0.071 |
| $\left(\mathrm{B} / \mathrm{B}_{\text {MSY }}\right)_{2011}$ | 1.553 | 1.535 | 1.549 | 0.092 | 0.060 |

Table 1. (Continued).

| Parameter/Year | Estimate | Mean | Median | Std Dev | C.V. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\left({\left.\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1980}}^{\text {l }}\right.$ | 0.848 | 0.840 | 0.820 | 0.210 | 0.250 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{1981}$ | 0.532 | 0.525 | 0.515 | 0.126 | 0.240 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{1982}$ | 0.379 | 0.377 | 0.368 | 0.096 | 0.255 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1983}$ | 0.523 | 0.527 | 0.511 | 0.146 | 0.277 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1984}$ | 0.411 | 0.421 | 0.404 | 0.120 | 0.284 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1985}$ | 0.572 | 0.592 | 0.564 | 0.167 | 0.282 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1986}$ | 0.422 | 0.440 | 0.418 | 0.118 | 0.267 |
| $\left(\text { ( } / F_{\text {MSY }}\right)_{1987}$ | 0.278 | 0.289 | 0.276 | 0.074 | 0.257 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1988}$ | 0.230 | 0.239 | 0.229 | 0.058 | 0.241 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1989}$ | 0.577 | 0.596 | 0.574 | 0.128 | 0.214 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1990}$ | 0.695 | 0.711 | 0.698 | 0.127 | 0.179 |
| $\left(\mathrm{F} / \mathrm{F}_{\mathrm{MSY}}\right)_{1991}$ | 0.485 | 0.491 | 0.485 | 0.082 | 0.167 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{1992}$ | 0.632 | 0.635 | 0.628 | 0.104 | 0.164 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{1993}$ | 0.551 | 0.551 | 0.546 | 0.085 | 0.155 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1994}$ | 0.830 | 0.828 | 0.824 | 0.122 | 0.148 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1995}$ | 1.016 | 1.009 | 1.014 | 0.140 | 0.139 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1996}$ | 1.749 | 1.733 | 1.749 | 0.270 | 0.156 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{1997}$ | 1.235 | 1.216 | 1.228 | 0.212 | 0.175 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{1998}$ | 1.739 | 1.715 | 1.717 | 0.364 | 0.212 |
| $\left(\text { ( } / F_{\text {MSY }}\right)_{1999}$ | 1.509 | 1.486 | 1.478 | 0.356 | 0.240 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2000}$ | 0.784 | 0.768 | 0.764 | 0.180 | 0.235 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2001}$ | 0.903 | 0.890 | 0.882 | 0.218 | 0.245 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{2002}$ | 0.543 | 0.539 | 0.529 | 0.132 | 0.245 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2003}$ | 1.216 | 1.222 | 1.183 | 0.313 | 0.256 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{2004}$ | 0.559 | 0.566 | 0.544 | 0.139 | 0.245 |
| $\left(\text { ( } / \mathrm{F}_{\text {MSY }}\right)_{2005}$ | 0.219 | 0.223 | 0.214 | 0.054 | 0.245 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2006}$ | 0.348 | 0.357 | 0.344 | 0.084 | 0.236 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2007}$ | 0.242 | 0.249 | 0.241 | 0.052 | 0.208 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2008}$ | 0.421 | 0.431 | 0.418 | 0.079 | 0.185 |
| $\left(\mathrm{F} / \mathrm{F}_{\text {MSY }}\right)_{2009}$ | 0.361 | 0.367 | 0.359 | 0.057 | 0.156 |
| $\left(\text { ( } / F_{\text {MSY }}\right)_{2010}$ | 0.577 | 0.581 | 0.573 | 0.082 | 0.140 |
| $\left(\right.$ (F/F $\mathrm{MSY}^{\text {\% }}$ 2011 | 0.863 | 0.863 | 0.863 | 0.107 | 0.123 |

Table 2. Uncertainty parameters for Choptank River channel catfish catch survey analysis model.

| Parameter/Year | Estimate | Mean | Median | Std Dev | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
| q | $1.85 \mathrm{E}-06$ | 2.16E-06 | 2.19E-06 | 4.77E-07 | 0.221 |
| Pre-Recruit 1993 | 101,595 | 100,559 | 92,171 | 40,577 | 0.404 |
| Pre-Recruit 1994 | 316,238 | 297,941 | 279,696 | 100,228 | 0.336 |
| Pre-Recruit 1995 | 65,634 | 64,680 | 59,944 | 23,545 | 0.364 |
| Pre-Recruit 1996 | 78,966 | 78,659 | 72,997 | 29,971 | 0.381 |
| Pre-Recruit 1997 | 20,696 | 20,680 | 19,172 | 7,821 | 0.378 |
| Pre-Recruit 1998 | 36,260 | 36,603 | 33,993 | 13,636 | 0.373 |
| Pre-Recruit 1999 | 96,360 | 100,482 | 93,262 | 35,625 | 0.355 |
| Pre-Recruit 2000 | 178,617 | 166,470 | 153,832 | 62,212 | 0.374 |
| Pre-Recruit 2001 | 66,616 | 64,523 | 59,737 | 24,312 | 0.377 |
| Pre-Recruit 2002 | 52,342 | 51,391 | 47,666 | 19,694 | 0.383 |
| Pre-Recruit 2003 | 187,214 | 179,899 | 166,454 | 70,751 | 0.393 |
| Pre-Recruit 2004 | 63,482 | 62,327 | 57,941 | 23,994 | 0.385 |
| Pre-Recruit 2005 | 224,734 | 221,814 | 204,670 | 83,745 | 0.378 |
| Pre-Recruit 2006 | 276,295 | 270,202 | 249,740 | 105,216 | 0.389 |
| Pre-Recruit 2007 | 520,434 | 491,965 | 452,239 | 191,903 | 0.390 |
| Pre-Recruit 2008 | 658,633 | 629,197 | 578,751 | 248,905 | 0.396 |
| Pre-Recruit 2009 | 574,665 | 563,424 | 516,412 | 229,798 | 0.408 |
| Pre-Recruit 2010 | 371,201 | 367,107 | 339,080 | 146,601 | 0.399 |
| Pre-Recruit 2011 | 737,275 | 733,489 | 670,468 | 295,111 | 0.402 |
| Post-Recruit 1993 | 218,647 | 186,474 | 172,077 | 85,420 | 0.458 |
| Post-Recruit 1994 | 244,226 | 217,038 | 202,840 | 80,856 | 0.373 |
| Post-Recruit 1995 | 441,531 | 404,291 | 379,300 | 116,201 | 0.287 |
| Post-Recruit 1996 | 398,701 | 367,430 | 343,261 | 107,117 | 0.292 |
| Post-Recruit 1997 | 361,367 | 335,513 | 310,513 | 102,812 | 0.306 |
| Post-Recruit 1998 | 294,084 | 272,904 | 250,830 | 88,529 | 0.324 |
| Post-Recruit 1999 | 234,328 | 217,268 | 196,348 | 80,092 | 0.369 |
| Post-Recruit 2000 | 181,047 | 170,454 | 145,268 | 84,153 | 0.494 |
| Post-Recruit 2001 | 278,624 | 260,007 | 233,493 | 107,228 | 0.412 |
| Post-Recruit 2002 | 272,933 | 255,977 | 230,623 | 102,200 | 0.399 |
| Post-Recruit 2003 | 241,904 | 227,243 | 203,796 | 95,496 | 0.420 |
| Post-Recruit 2004 | 308,347 | 290,354 | 260,212 | 122,668 | 0.422 |
| Post-Recruit 2005 | 271,513 | 255,836 | 228,206 | 114,740 | 0.448 |
| Post-Recruit 2006 | 302,012 | 286,785 | 253,707 | 146,053 | 0.509 |
| Post-Recruit 2007 | 369,851 | 352,396 | 311,445 | 184,875 | 0.525 |
| Post-Recruit 2008 | 664,130 | 626,531 | 564,530 | 275,783 | 0.440 |
| Post-Recruit 2009 | 950,114 | 895,230 | 814,146 | 383,669 | 0.429 |
| Post-Recruit 2010 | 1,109,139 | 1,055,001 | 955,081 | 454,308 | 0.431 |
| Post-Recruit 2011 | 1,061,724 | 1,014,047 | 911,008 | 458,810 | 0.452 |

Table 2. (Continued).

| Parameter/Year | Estimate | Mean | Median |  | Std Dev |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Total N 1993 | 320,241 | 287,033 | 269,692 | 98,759 | 0.344 |
| Total N 1994 | 560,464 | 514,979 | 483,455 | 141,928 | 0.276 |
| Total N 1995 | 507,165 | 468,971 | 49,451 | 130,834 | 0.279 |
| Total N 1996 | 477,667 | 446,089 | 415,554 | 125,575 | 0.282 |
| Total N 1997 | 382,062 | 356,193 | 329,231 | 108,130 | 0.304 |
| Total N 1998 | 330,344 | 309,507 | 283,956 | 97,825 | 0.316 |
| Total N 1999 | 330,688 | 317,750 | 286,987 | 102,785 | 0.323 |
| Total N 2000 | 359,664 | 336,924 | 304,541 | 130,968 | 0.389 |
| Total N 2001 | 345,240 | 324,529 | 293,562 | 124,827 | 0.385 |
| Total N 2002 | 325,275 | 307,368 | 278,729 | 116,640 | 0.379 |
| Total N 2003 | 429,119 | 407,142 | 370,326 | 149,827 | 0.368 |
| Total N 2004 | 371,829 | 352,680 | 318,933 | 140,143 | 0.397 |
| Total N 2005 | 496,247 | 477,650 | 437,249 | 178,390 | 0.373 |
| Total N 2006 | 578,307 | 556,987 | 506,969 | 225,807 | 0.405 |
| Total N 2007 | 890,285 | 844,362 | 768,634 | 336,842 | 0.399 |
| Total N 2008 | $1,322,763$ | $1,255,727$ | $1,156,690$ | 468,615 | 0.373 |
| Total N 2009 | $1,524,779$ | $1,458,654$ | $1,336,612$ | 554,893 | 0.380 |
| Total N 2010 | $1,480,340$ | $1,422,108$ | $1,296,256$ | 560,392 | 0.394 |
| Total N 2011 | $1,798,999$ | $1,747,536$ | $1,599,135$ | 686,374 | 0.393 |
|  |  |  |  |  |  |
| F 1993 | 0.071 | 0.089 | 0.085 | 0.031 | 0.343 |
| F 1994 | 0.039 | 0.045 | 0.045 | 0.010 | 0.230 |
| F 1995 | 0.041 | 0.047 | 0.047 | 0.011 | 0.228 |
| F 1996 | 0.079 | 0.090 | 0.091 | 0.021 | 0.229 |
| F 1997 | 0.062 | 0.071 | 0.072 | 0.017 | 0.243 |
| F 1998 | 0.143 | 0.167 | 0.169 | 0.044 | 0.261 |
| F 1999 | 0.402 | 0.470 | 0.481 | 0.130 | 0.277 |
| F 2000 | 0.055 | 0.066 | 0.066 | 0.020 | 0.302 |
| F 2001 | 0.035 | 0.042 | 0.041 | 0.012 | 0.296 |
| F 2002 | 0.096 | 0.114 | 0.113 | 0.035 | 0.303 |
| F 2003 | 0.131 | 0.155 | 0.153 | 0.048 | 0.310 |
| F 2004 | 0.114 | 0.138 | 0.135 | 0.046 | 0.336 |
| F 2005 | 0.297 | 0.359 | 0.344 | 0.133 | 0.370 |
| F 2006 | 0.247 | 0.303 | 0.287 | 0.120 | 0.396 |
| F 2007 | 0.093 | 0.113 | 0.109 | 0.040 | 0.355 |
| F 2008 | 0.131 | 0.157 | 0.151 | 0.053 | 0.335 |
| F 2009 | 0.118 | 0.140 | 0.136 | 0.047 | 0.335 |
| F 2010 | 0.132 | 0.158 | 0.153 | 0.056 | 0.354 |
| F 2011 | 0.109 | 0.129 | 0.124 | 0.046 | 0.355 |
|  |  |  |  |  |  |

Figure 1. Head-of-Bay and Potomac River fishery independent drift gill net sampling locations, 1985 -- 2011.


Figure 2. Head-of-Bay winter trawl sites, 1999 -- 2012 (triangles=main bay sites, squares=Elk River sites, circles=Sassafras River sites).


Figure 3. Choptank River fyke net locations, 2011. Circles indicate sites.

1


Figure 4. Estuarine Juvenile Finfish Survey seine site locations, 1962 -- 2012.


Figure 5. Adjusted Maryland commercial channel catfish landings, 1929-2011.


Figure 6. Estimated channel catfish landings from the recreational fishery, 1983-2011. Error bars $=1$ standard error.


Figure 7. Head-of Bay channel catfish removals from commercial and recreational fisheries, 1980-2011.
$\square$ Recreational $\square$ Commercial


Figure 8. Observed and expected HOB commercial fyke net index, 1980-2011.


Figure 9. Observed and expected biomass index from HOB gill net survey, 1985 - 2011.
$\square$ Observed ■ Expected


Figure 10. Observed and expected channel catfish biomass index from upper Bay winter trawl survey, 2000-2002 and 2006-2011.


Figure 11. Biomass and fishing mortality estimates from Head-of-Bay channel catfish surplus production model.


Figure 12. Biomass and fishing mortality ratios from Head-of-Bay channel catfish surplus production model.

- B/BMSY - -F/FMSY


Figure 13. Biomass estimate and $80 \%$ confidence intervals from Head-of-Bay channel catfish surplus production model.


Figure 14. Biomass ratio and $80 \%$ confidence intervals from Head-of-Bay channel catfish surplus production model.


Figure 15. Fishing mortality and $80 \%$ confidence intervals from Head-of-Bay channel catfish surplus production model.


Figure 16. Fishing mortality ratio and $80 \%$ confidence intervals from Head-of-Bay channel catfish surplus production model.


Figure 17. Choptank River channel catfish removals from commercial and recreational fisheries, 1993 - 2011
-Commercial $\mathbf{\Delta}$ Recreational


Figure 18. Observed and expected pre-recruit channel catfish index from Choptank River catch survey analysis.
—OBSERVED • EXPECTED


Figure 19. Observed and expected post-recruit channel catfish index from Choptank River catch survey analysis.
-OBSERVED - EXPECTED


Figure 20. Estimated pre-recruit and post-recruit channel catfish abundance from Choptank River catch survey analysis.
$\square$ pre rec pop'n est ( N ) $\square$ post rec pop'n est ( N )


Figure 21. Estimated fishing mortality for Choptank River channel catfish from a catch survey analysis.


Figure 22. Choptank River channel catfish pre-recruit abundance with $80 \%$ confidence intervals from catch survey analysis.
——LOWER CI ——UPPER CI - - ESTIMATE


Figure 23. Choptank River channel catfish post-recruit abundance with $80 \%$ confidence intervals from catch survey analysis.


Figure 24. Total channel catfish population abundance estimates and $80 \%$ confidence intervals from Choptank River catch survey analysis.


Figure 25. Estimated fishing mortality and $80 \%$ confidence intervals for Choptank River channel catfish from catch survey analysis.


Figure 26. Nanticoke River channel catfish commercial landings, 1987-2011.


Figure 27. Nanticoke River commercial fish pot channel catfish relative abundance and $75^{\text {th }}$ percentile, 1980 - 2011. Anomalous 2008 value truncated for scale $(1,141)$.


Figure 28. Nanticoke River commercial fish fyke net channel catfish relative abundance and $75^{\text {th }}$ percentile, $1980-2011$. Anomalous 1999 value truncated for scale $(1,056)$.


Figure 29. Nanticoke River channel catfish young-of-year from Estuarine Juvenile Finfish Survey, 1975 - 2012.


Figure 30. Patuxent River channel catfish commercial landings, 1987-2011.


Figure 31. Patuxent River commercial fish pot channel catfish relative abundance and $75^{\text {th }}$ percentile, $1981-2011$.


Figure 32. Patuxent River channel catfish young-of-year from Estuarine Juvenile Finfish Survey, 1983-2012.


Figure 33. Potomac River channel catfish commercial landings, 1964-2012. Data from Potomac River Fishery Commission.


Figure 34. Channel catfish biomass index from Potomac River gill net survey, 1985 2011.
$\square$ Geometric Mean Biomass - 75TH PERCENTILE


Figure 35. Potomac River channel catfish young-of-year from Estuarine Juvenile Finfish Survey, 1975-2012.


# PROJECT NO. 2 

JOB NO. 1

# STOCK ASSESSMENT OF ADULT AND JUVENILE ALOSINE SPECIES IN THE CHESAPEAKE BAY AND SELECTED TRIBUTARIES 

Prepared by
Karen M. Capossela and Anthony A. Jarzynski

## INTRODUCTION

The pr imary objective of Project 2 , Job 1 w as to a ssess t rends in the stock s tatus of American shad, hickory shad and river herring (i.e., alewife and blueback herring) in Maryland's portion of the Chesapeake B ay and s elected $\mathrm{tributaries}$. species and their subsequent s pawning success in Maryland tributaries was col lected for this project by $t$ he Maryland D epartment of N atural R esources (MDNR) using bot h fishery dependent a nd independent s ampling gear. On t he N anticoke R iver, biologists worked w ith commercial fishermen to collect sex, age and stock composition data and to estimate relative abundance of adult American s had, hi ckory s had and river he rring. Survey bi ologists also independently sampled ichthyoplankton. Similar data were collected for adult American shad in the lower S usquehanna River be low the Conowingo D am, a nd hi ckory s had a bundance w as assessed in a tributary tot he $S$ usquehanna $R$ iver (Deer $C$ reek). Summer $s$ ampling ta rgeted juvenile alosines in the Chester River.

The data collected during this study were used to prepare and update stock assessments and fishery management plans for the A tlantic States Marine Fisheries Commission (ASMFC), Susquehanna River A nadromous F ish R estoration C ooperative ( SRAFRC), Chesapeake B ay Program's Living R esources C ommittee and Maryland Sea Grant E cosystem-Based Fisheries Management Program.

## METHODS

## Data Collection

## Susquehanna River

Adult American shad were angled by MDNR staff from the Conowingo Dam tailrace on the lower S usquehanna River two t imes per week from 4 April 2012 through 30 M ay 2012 (Figure 1). Two rods were fished simultaneously; each rod was rigged with two shad darts and lead weight was added when required to achieve proper depth. All American shad were sexed (by expression of $g$ onadal pr oducts), $t$ otal 1 ength ( mm TL ) a nd fork 1 ength ( mm FL ) w ere measured and scales were removed below the insertion of the dorsal fin for ageing and spawning history a nalysis. Fish ing ood ph ysical condition (including unspent orripe females) were tagged with Floy tags (color-coded to identify the year tagged) and released. A MDNR hat was given to fishers as a reward for returned tags.

Scales collected from all rivers for A merican shad, hickory shad and river herring were aged using Cating's method (Cating 1953). A minimum of four scales per sample were cleaned, mounted be tween $t$ wo glass s lides a nd $r$ ead for a ge a nd s pawning hi story us ing a Bell a nd Howell M T-609 microfiche reader. The s cale edge was count ed as a year-mark due to the assumption that each fish had completed a full year's growth at the time of capture. Ages were not assigned to regenerated scales or to scales that were difficult to read. Two readers aged all scales separately, and then jointly re-aged any s cales t hat w ere as signed different ages. If agreement about an age could not be reached, the scales were not included in the age structure analysis. Hickory shad scales from the S usquehanna R iver were a ged by the Restoration and Enhancement P rogram. R epeat s pawning m arks w ere count ed on all al osine s cales dur ing ageing.

Normandeau Associates, Inc. was responsible for obs erving and/or collecting American shad at the Conowingo Dam fish lifts. American shad collected in the East Fish Lift (EFL) were deposited into a trough, di rected p ast a $4^{\prime} \mathrm{x} 10^{\prime} \mathrm{c}$ ounting w indow, i dentified t o s pecies and counted by experienced technicians. A merican shad captured from the West Fish Lift (WFL) were counted and either used for experiments (e.g., hatchery brood stock, oxytetracycline [OTC] analysis, sacrificed for o tolith extraction) or returned to the tailrace. For both lifts, tags were used to identify American shad tagged in the current year and in previous years.

Recreational data from a non-random roving creel survey were collected from anglers in the C onowingo Dam tailrace dur ing the MDNR A merican shad hook and line survey. In this survey, stream ba nk an glers were i nterviewed about American shad catch a nd hour s s pent fishing. A vol untary logbook survey also provided location, catch a nd hours spent fishing for American shad in the lower Susquehanna River (including the C onowingo tailrace a nd D eer Creek) for each participating angler. The same information was collected for hickory shad in the Susquehanna River (including the Conowingo tailrace and Deer Creek), North East Creek and Big Elk Creek.

Due tothe low num ber of hi ckory shad typically obs erved bythis project, MDNR's Susquehanna Restoration a nd E nhancement P rogram provided a dditional hickory s had da ta (2004-2012) from their brood stock collection. Hickory shad were col lected in Deer Creek (a Susquehanna R iver tributary) for ha tchery brood stock a nd were s ubsampled for age, repeat spawning marks, sex, length and weight. In 2004 and 2005, fish were collected using hook and line fishing; fish have been collected using electrofishing gear from 2006 to the present.

## Nanticoke River

Four commercial pound nets were surveyed for American shad, hickory shad and river herring between 22 February and 30 April 2012 (Figure 2). Cooperating commercial watermen did not use fyke nets in 2012. Fish captured from these nets were sorted according to species and transferred to the s urvey boat for processing. All nets were sampled two da ys per week during the survey period. Fish were sexed (by expression of gonadal products), measured (TL and FL), and scales were removed below the insertion of the dorsal fin for ageing and spawning history a nalysis. Otoliths f rom d ead adult A merican shad were removed a nd s ent t o the Delaware Division of Fish and Wildlife (DE DFW) for OTC analysis.

Ichthyoplankton $s$ amples $w$ ere conducted in cooperation w ith $t$ he Fish Habitat \& Ecosystem Program (Federal A id Grant F-63-R, Segment 2, J ob 1, S ection 3) twice per week from 2 April to 30 April 2012 in the Nanticoke River. The presence/absence of alosine eggs or larvae was noted (time and field conditions prevented species identification of alosine eggs or larvae). These samples were collected following historical methodology: the river was divided into eighteen one-mile cells and ten of these cells were randomly selected during each sampling day (Figure 3). The ichthyoplankton net was constructed of $500 \mu \mathrm{~m}$ mesh net with a 500 mm metal r ing opening. T he ne t was towed for t wo minutes at approximately t wo knots. A t the conclusion of the tow, the contents were flushed down into a masonry jar for presence/absence determination.

## Chester River

Juvenile American shad, hickory shad and river herring were sampled once eve ry two weeks in the Chester River from 12 July to 20 September 2012 with a $30.5 \mathrm{~m} \times 1.2 \mathrm{~m} \times 6.4 \mathrm{~mm}$ mesh ha ul seine and a 4.9 m s emi-balloon ot ter trawl. The t rawl w as cons tructed of t reated
nylon mesh netting measuring 38 mm stretch-mesh in the body and 33 mm stretch-mesh in the codend, with a n unt reated 12 mm stretch-mesh knot less me sh line r . T he $16^{\prime}$ headrope was equipped with floats and the footrope was equipped with a 3.2 mm chain. The net used 0.61 m long by 0.30 m high trawl doors attached to a 6.1 m bridle leading to a 24.4 m towrope. Trawls were tow ed in the same direction as the tide. Each seine site was located on a beach di rectly across from a $t$ rawl site. The six paired seine and trawl sites were loc ated a mini mum of 0.5 miles apart on the C hester R iver (Figure 4). Sites were selected based on the ava ilability of seinable beaches and historical spawning importance. A 11 collected alosines were counted and measured (FL and TL).

## Potomac River

The Striped Bass S pawning S tock Survey (SBSSS; Project 2, J ob 3, T ask 2) provided American shad scales from the Potomac River to compare age structure and repeat spawning of fish in this river with fish sampled in the Susquehanna and Nanticoke Rivers. American shad were captured in gill nets targeting striped bass from 26 M arch to 7 M ay 2012. All A merican shad were sexed, measured (TL and FL), and scales were removed below the insertion of the dorsal fin for ageing and spawning history analysis.

## Data Analysis

## Ichthyoplankton

The pe rcent of pos itive t ows (i.e., t hose c ontaining alosine e ggs or 1 arvae) was determined as the number of tows with eggs and/or larvae divided by the total number of tows. These data have been reported since 2005 .

## Sex, Age and Stock Composition

Male-female ratios were derived for American shad angled at the Conowingo Dam in the Susquehanna River. Male-female ratios were also derived for A merican shad, alewife he rring and blueback herring captured by pound a nd fyke nets in the Nanticoke River. Due to the low number of hickory shad captured in the Nanticoke River survey, hickory shad male-female ratios were derived from data provided by the MDNR Restoration and Enhancement Program's brood stock collection on the Susquehanna River.

Age determination from scales was at tempted for all A merican shad and river herring samples col lected from $t$ he Susquehanna, Nanticoke and $P$ otomac Rivers. All r eadable American shad scales w ere a ged. In 2012, we increased the number of a lewife a nd blueback herring measured at each site by $50 \%$. Approximately $30 \%$ of measured river herring were aged, and a n a ge-length ke y was us ed t o covert from num ber at l ength t o num ber at a ge. The percentages ofr epeat spawners by species a nd system (sexes com bined) w ere arcsinetransformed (in degrees) before 1 ooking f or linear trends over t ime. For a 1 l statistics, significance was determined at $\alpha=0.05$.

All hatchery produced juvenile A merican shad stocked in Maryland, Delaware and the Susquehanna ba sin ha ve uni quef luorescent O TC m arks. O tolithe xamination byt he Pennsylvania Fish and Boat Commission (PFBC) and the DE DFW indicated the percent of nonhatchery fish present from American shad collected in the WFL and Maryland's portion of the Nanticoke River, respectively.

## Adult Relative Abundance

Catch-per-unit-effort (CPUE) from the C onowingo Dam tailrace was calculated as the number of adult fish captured per boat hour. We computed a combined lift CPUE as the total number of adult fish lifted per hour of lifting at the EFL and WFL. The geometric mean (GM) of adult A merican shad CPUE for both the tailrace ar ea and the lifts $w$ as then calculated as the average $\mathrm{LN}(\mathrm{CPUE}+1)$ for each fishing/lifting day, transformed back to the original scale. In addition, the relative abundance (GM CPUE) of American shad, alewife herring and blueback herring in the Nanticoke River was calculated as the average LN (CPUE + 1) for each net day by gear type, transformed back to the original scale. No CPUE was calculated for hickory shad in the Nanticoke River due to the low number encountered by both gear types over the time series; instead, the number of hickory shad captured by gear type is reported. In the Potomac River, the SBSSS calculated CPUE as the number of fish caught per $1,000 \mathrm{~s}$ quare yards of experimental drift gill ne t per hour fished. C atch-per-angler-hour (CPAH) for American shad and hickory shad in the Susquehanna River were also calculated from the shad logbooks. The roving creel survey was used to calculate a CPAH for American shad.

Historically, CPUE for American shad from the Nanticoke was only calculated with data from one pound ne $t \mathrm{t}$ hat was mostc onsistently sampled over the time s eries (Mill C reek). Similarly, alewife and blueback herring CPUE were only calculated with fyke net data because pound nets were not consistently set in ideal habitat for river herring. This report follows these historical protocols.

Chapman's $m$ odification of the P etersen statistic w as us ed to estimate abunda nce of American shad in the Conowingo Dam tailrace (Chapman 1951):

$$
N=(C+1)(M+1) /(R+1)
$$

where $N$ is the relative population estimate, $C$ is the number of fish examined for tags at the EFL and WFL, $M$ is the number of fish tagged minus $3 \%$ tag loss, and $R$ is the number of tagged fish recaptured.

Overestimation of abunda nce $b$ yt he $P$ etersen $s$ tatistic (due $t o$ low recapture $r$ ates) necessitated the additional use of a bi omass surplus pr oduction model (SPM; M acall 2002, Weinrich et al. 2008):

$$
N_{t}=N_{t-1}+\left[r N_{t-1}\left(\left(1-N_{t-1}\right) / K\right)\right]-C_{t-1}
$$

where $N_{t}$ is the population (numbers) in year $t, N_{t-1}$ is the population (numbers) in the previous year, $r$ is the intrinsic rate of population increase, $K$ is the maximum population size, and $C_{t-1}$ is losses a ssociated w ith upstream a nd do wnstream $f$ ish pa ssage and bycatch mortality in the Atlantic herring fishery in the previous year (equivalent to catch in a surplus production model). The dynamics of this population are governed by the logistic growth curve. Model parameters were e stimated using a non -equilibrium a pproach $t$ hat $f$ ollows an ob servation-error fi tting method (i.e., assumes that all er rors oc cur in the relationship be tween true stock size a nd the index used to measure it). Assumptions include proportional bycatch of American shad in the Atlantic herring fishery and accurate adult American shad turbine mortality estimates. The SPM required starting values for the initial population in 1985 (set as 7,876 by the Petersen statistic for this year; calculation described above), a c aring capacity estimate (set as $3,040,551 \mathrm{f}$ ish, which was three times the highest Petersen estimate of the time series), and an estimate of the intrinsic rate of growth (set as 0.50 ). These starting values were adjusted by the model during the fitting procedure.

## Mortality

Catch curve an alysis w as us ed to estimate total i nstantaneous mortality ( $Z$ ) of a dult American shad and river he rring in the N anticoke River. Additionally, Z w as cal culated for American and hickory shad in the Susquehanna River. The number of repeat spawning marks was used in this estimation instead of age because ageing techniques for American shad scales are tenuous (McBride et al. 2005). Therefore, the Z calculated for these fish represents mortality associated with repeat spawning. Assuming $t$ hat c onsecutive s pawning o ccurred, the $\ln$ transformed spawning group frequency was plotted against the cor responding number of times spawned:

$$
\ln \left(S_{f x}+1\right)=a+Z * W_{f x}
$$

where $S_{f x}$ is number of fish with $1,2, \ldots f$ spawning marks in year $x, a$ is the $y$-intercept, and $W_{f x}$ is frequency of spawning marks $(1,2, \ldots f)$ in year $x$. Using $Z$, annual mortality (A) was obtained from a table of exponential functions and derivatives (Ricker 1975).

Natural and fishing $m$ ortality $w$ ere not es timated for an $y$ al osine $s$ pecies be cause American shad, hickory shad and river herring fisheries are closed in Maryland. Commercial landings, commercial and recreational bycatch, and EFL and WFL mortalities were cons idered when estimating the minimum total losses of adult American shad in Maryland waters,

## Juvenile Abundance

CPUE for seine and trawl surveys on the Chester River were not calculated for juvenile alosine species due to historically low catches of $t$ hese $s$ pecies in $t$ his river. However, the
numbers of American shad, hickory shad and river he rring captured by these gear types are reported. The MDNR Estuarine Juvenile Finfish Seine Survey (EJFS; Project 2, J ob 3, Task 3) provided juvenile i ndices (geometric mean catch per ha ul) for al ewife herring and blueback herring from fixed stations within the Nanticoke River and the upper Chesapeake Bay, and for American shad in the Nanticoke and P otomac rivers, uppe $\mathrm{r} C$ hesapeake B ay a nd b aywide. Hickory shad data are not reported by the EJFS due to small sample sizes.

## RESULTS

## Ichthyoplankton

Fertilized clupeid eggs and/or larvae were not found in any of the ichthyoplankton tows in $2012(\mathrm{n}=86$; Table 1). Salinity at tow stations ranged from 0.1 to 4.8 ppt . An absence of observed fertilized eggs and/or larvae also occurred from 2006-2008. The available data indicate that clupeid egg and/or larvae presence was highest in 2010 (43\%; 2005-2012).

## American Shad

## Sex, Age and Stock Composition

The $m$ ale-female $r$ atio of adult A merican shad ca ptured $b$ y hook a nd 1 ine from $t$ he Conowingo tailrace was $1: 0.72$. Of the 191 fish sampled by this gear, 177 were successfully scale-aged (Table 2). Males were p resent in age groups 3-6 and females w ere found in age groups 3-8. The 2007 year-class (age 5) and the 2006 year-class (age 6 ) were the most abundant for males and females, respectively, accounting for $45 \%$ of males and $46 \%$ of females (Table 2). Thirty-four percent of males and $73 \%$ of females w ere repeat s pawners. Past percentages of repeat spawners for both males and females were low, particularly be fore 1997 (Figure 5), but the arcsine-transformed proportion of these repeat spawners (sexes combined) has significantly
increased over the time series (1984-2012; $r^{2}=0.45, P<0.001$; Figure 6). Of the 129 readable adult A merican shad otoliths collected from the WFL at Conowingo Dam in 2012, 71\% were classified as non-hatchery fish (M. Hendricks PA Fish and Boat Comm., pers. comm. 2012).

The m ale-female r atio for adul A merican shad capt ured in the N anticoke R iver w as 1:0.5. Of the 178 American shad collected from the Nanticoke pound and fyke nets in 2012, 172 were s ubsequently a ged (Table 2). Males w ere present in age groups 3-7 and females w ere found in age groups 4-7. The most abundant year-classes by sex were the 2007 year-class (age 5) for both males ( $40 \%$ ) and females ( $46 \%$; Table 2). Forty percent of males and $56 \%$ of females were repeat spawners. The arcsine-transformed proportion of Nanticoke River repeat spawning American shad (sexes combined) has significantly increased over the time series, (1988-2012; $r^{2}$ $=0.35, P=0.00$ 2; Figure 7). Fifty-two adult A merican shad ot oliths collected from the Nanticoke River were sent to DE DFW for OTC analysis in 2012. Forty-nine of the 52 scales were readable, and results indicated that $55 \%$ were non-hatchery fish (M. Stangl, pers. comm. 2012).

The m ale-female r atio for adul t A merican shad capt ured in the P otomac R iver w as 1:1.22. Of the 71 American shad collected, 67 were successfully aged (Table 2). Males were present in age groups 4-7 and females were present in age groups 5-8. The most abundant yearclasses by sex were the 2007 year-class (age 5) for males ( $47 \%$ ) and the 2006 year class (age 6) for females (34\%). Thirty-four percent males and $60 \%$ of females were repeat spawners. The arcsine-transformed pr oportion of Potomac R iver repeat s pawning American shad (sexes combined) showed no significant trend over the time series (2002-2012; $r^{2}=0.054, P=0.49$; Figure 8).

## Adult Relative Abundance

Sampling at the C onowingo D am occurred for 18 days in 2012. A total of 226 adult American shad were encountered by the gear; 217 of these fish were captured by MDNR staff from a boat and the remaining 9 were captured by shore anglers. MDNR staff tagged 190 (84\%) of the sampled fish. To remain consistent with historical calculations, only the 217 fish captured from the boa $t$ were used to calculate $t$ he hook a nd 1 ine C PUE. No tagged American shad recaptures were reported from either commercial fishermen or recreational anglers.

The E FL ope rated for 62 days be tween 2 April and 5 June. The 2012 s eason was the third longest season of EFL operation and had the highest number of lifts since the EFL became operational in 1991. Of the 22,143 American shad that passed at the EFL, 39\% (8,665 fish) passed between 22 April and 11 May. Peak passage was on 24 April; 1,710 American shad were recorded on this date. Twenty-four of the American shad counted at the EFL counting windows were identified as being tagged in 2012; only 2 fish passed that were tagged in 2011 (Table 3).

The Conowingo WFL operated for 37 days between 23 A pril and 1 June. The 1,486 captured American shad w ere retained for hatchery op erations, sacrificed for cha racterization data collection, or returned alive to the tailrace. Peak capture from the WFL was on 5 May when 135 American shad were collected. The four tagged American shad recaptured by the WFL in 2012 were fish tagged in 2012 (Table 3).

The Petersen statistic estimated 150,743 American shad in the Conowingo Dam tailrace in 201 2, a nd the S PM e stimated a popul ation of 111,500 fish. Despite differences in yearly estimates, the ove rall p opulation t rends de rived f rom e ach m ethod a re similar (Figure 9). Specifically, SPM es timates declined from 2001 to 2007 a nd increased from 2008 t o 2012 .

Petersen estimates follow a similar pattern if the high levels of uncertainty in 2004 and 2008 (due to low recapture rates) are considered.

Estimates of hook a nd line GM C PUE vary without trend over the time series (19842012; $r^{2}=0.11, P=0.07$ ). Abundance is particularly variable from 200 7-2012 and remains below the high indices observed from 1999 to 2002 (Figure 10). The Conowingo Dam combined lift GM CPUE significantly increased over the time series (1980-2012; $r^{2}=0.33, P<0.001$ ); the GM CPUE decreased steadily from 2002 to 2008 , i ncreased from 2009 through 2011 , a nd decreased in 2012 (Figure 11).

Fifty-eight interviews were conducted ove $\mathrm{r} f$ ive da ys dur ing the creel s urvey at t he Conowingo Dam Tailrace. The CPAH in 2012 was the third lowest since the start of the survey in 2001 (Table 4), and CPAH has de creased over the time series (2001-2012; $r^{2}=0.46, P=$ 0.02). Five anglers returned logbooks in 2012; four logbooks contained information from fishing trips in the lower S usquehanna R iver. Although A merican shad C PAH c alculated from shad logbook da ta d ecreased s ignificantly ove $\mathrm{r} t$ he time s eries $\left(1999-2012 ; r^{2}=0.35, P=0.03\right.$ ), CPAH has steadily increased since 2009 (Table 5).

The 2012 Nanticoke River pound net GM CPUE was the highest it has been since the start of the survey in 1988. The GM CPUE significantly increased over the time series (19882012; $r^{2}=0.24, P=0.07$, Figure 12). The Potomac River CPUE increased significantly over the time series (1996-2012; $r^{2}=0.23, P=0.053$ ), although CPUE in each of the past four years has been lower than the CPUE in 2007 and 2008 (Figure 13).

## Mortality

The C onowingo Dam tailrace total ins tantaneous mortality e stimate from catch curve analysis (using repeat spawning instead of age) resulted in $Z=0.61(A=45.7 \%)$. The Nanticoke

River mortality estimate was $\mathrm{Z}=0.82(\mathrm{~A}=56.0 \%)$. E stimated American shad mortalities (in numbers) from Maryland waters are presented in Table 6.

## Juvenile Abundance

No juvenile American shad were captured in seines or trawls in the Chester River in 2012 (Table 7). Data provided by the EJFS indicated that juvenile American shad indices decreased in 2012 baywide, in the up per Chesapeake Bay, and in the Nanticoke River (Figures 14-16). In contrast, the Potomac River index increased in 2012 and remains above the time series mean (Figure 17). Juvenile indices were not corrected for hatchery contribution.

## Hickory Shad

Sex, Age and Stock Composition
The number of hickory shad captured from the Nanticoke River ( $\mathrm{n}=22$ ) was not large enough to draw meaningful conclusions a bout sex and age com position. However, 1,014 hickory shad were s ampled by the brood stock collection s urvey in Deer Creek. The malefemale ratio was 2.06:1. Of the total fish captured by this survey, 200 were successfully aged. Males were present in a ge groups 3-6 and females were found in age groups 3-7. The most abundant y ear-classes by sex were the 2008 y ear-class (age 4) for both males ( $42.6 \%$ ) a nd females (33.8\%; Table 8). Hickory shad sampled from 2004 to 2012 ranged from 2 to 9 years of age, with ages 3 through 8 present every year except for 2012 (Table 9). The 2012 s ampling year was the only year of the times series where only ages 3 to 7 were present. The arcsinetransformed proportion of these repeat spawners (sexes combined) has not changed significantly over the time series (2004-2012; $r^{2}=0.028, P=0.67$; Figure 18). However, the total percent of repeat spawners in 2012 ( $64.0 \%$ ) is the lowest of the time series (2004-2012; Table 10).

## Relative Abundance

Shad logbook data indicated that hickory shad CPAH did not vary significantly over the time series (1998-2012; $\left.r^{2}=0.13, P=0.18\right)$; however, hickory shad CPAH decreased in 2012 (Table 11). On the Nanticoke River, only 22 fish were captured by pound nets.

## Mortality

Total instantaneous mortality in the Susquehanna River (Deer Creek) was estimated as Z $=0.68$. T his e stimate is less than the 2010 Z estimate $(\mathrm{Z}=0.74)$ but similar to the 2011 Z estimate $(Z=0.67)$. Annual mortality in 2012 was estimated as $A=49.3 \%$.

## Juvenile Abundance

During the 2012 sampling in the Chester River, no juvenile hickory shad were collected in the seine or the trawl (Table 7). The last time this survey encountered no hi ckory shad in either gear was 2008 (2007-2012). The 2011 c atch remains the highest for both seines $(\mathrm{n}=6)$ and trawls $(\mathrm{n}=9)$ from 2007-2012.

## Alewife and Blueback Herring

## Sex, Age and Stock Composition

The 2012 male-female ratio for Nanticoke River alewife herring was 1:1.7. Of the 533 alewives sampled, 166 were subsequently aged. Age groups 3-7 were present and the 2007 yearclass (age 5, sexes combined) was the most abundant, accounting for $33.3 \%$ of the total catch. Females were most abundant at age 5 and males at age 4 (Table 12). The 2012 male-female ratio
for N anticoke R iver bl ueback herring was 1:0.78. Of the 403 blueback he rring sampled, 136 were subsequently aged. B lueback herring were present from ages 2-7 and the 2008 year-class (age 4 , sexes combined) was the most abundant, accounting for $42.9 \%$ of the sample (Table 12).

For the Nanticoke River, $40.8 \%$ of alewife herring and $23.7 \%$ of blueback herring were repeat s pawners (sexes combined; Table 12). There was no $t$ rend in the ar csine-transformed proportion of alewife herring repeat spawners over the time series (1989-2012; $r^{2}<0.007 P=$ 0.70 ); however, blueback herring exhibited a decreasing trend over the same time series (19892012; $r^{2}=0.61, P<0.001$; Figure 19). For male alewife and blueback herring, $75.3 \%$ and $77.4 \%$ were first time spawners, respectively; $49.7 \%$ of female alewife and $74.9 \%$ of female blueback herring were first time spawners.

Mean length-at-age was cal culated for aged fish only. Mean length-at-age for female alewife herring from the Nanticoke River is greater than the corresponding mean length-at-age for males (Table 13 ). Female bl ueback herring $m$ ean length-at-age is also greater $t$ han the corresponding $m$ ale $m$ ean 1 ength-at-age ( Table 14). Age $s$ tructure app ears to be truncating, especially for bl ueback herring, a nd o bserved declines in mean length-at-age generally oc cur toward the end of the time series (Tables 13 a nd 14). The lengths of female alewife herring at ages 4 to 8 and male al ewife he rring at a ges 4 to 7 have de creased significantly since 1989 (Table 15). The lengths of female blueback herring at ages 3 to 7 and male blueback herring at ages 3 to 7 have significantly decreased since 1989 (Table 15).

## Adult Relative Abundance

Fyke nets were not fished in the Nanticoke River in 2012 and no data are available for this year. Our protocol has been to only calculate alewife and blueback herring CPUE from fyke net data because pound nets were not consistently set in ideal habitat for river herring. As of

2011, the GM CPUE for Nanticoke River alewife herring captured in fyke nets varied without trend over the time series (1990-2011; $r^{2}=0.14, P=0.09$; Figure 20); in contrast, the GM CPUE for blueback herring decreased over the time series (1989-2011; $r^{2}=0.64, P<0.001$; Figure 20). As of 30 May 2012, 290 pounds of river herring were reported landed, despite the closure of the fishery (there was no differentiation between species in the com mercial river he rring fishery). Total commercial landings for river he rring in Maryland waters were at multi-decadenal lows before the closure of the fishery (Figure 21).

## Mortality

Total instantaneous mortality for Nanticoke River alewife herring (sexes combined) was estimated as $\mathrm{Z}=1.10(\mathrm{~A}=66.7 \%)$. Total instantaneous mortality for Nanticoke River blueback herring (sexes combined) was $\mathrm{Z}=1.43(\mathrm{~A}=76.1 \%)$. No estimates of M and F were calculated for 2012 because the fishery for river herring closed on 26 December 2011.

## Juvenile Abundance

Juvenile seining in the Chester River produced no juvenile alewife or blueback herring. (Table 7). Data provided by the EJFS indicated that the GM CPUE for juvenile alewife and blueback herring in the Nanticoke River and upper Bay decreased in 2012 (Figures 22-23). This contrasts with the increase observed in blueback herring indices in both the Nanticoke River and upper Bay in 2011.

## DISCUSSION

## American Shad

American shad are historically one of the most important exploited fish species in North America, but the stock has dr astically de clined due to the loss of habitat, ove rfishing, ocean bycatch, stream blockages and pollution. American shad restoration in the uppe r Chesapeake Bay began in the 1970s with the building of fish lifts and the stocking of juvenile American shad. Maryland closed the commercial and recreational American shad fisheries in 1980, and the ocean intercept fishery closed in 2005. The American shad adult stock has shown some improvement since the inception of restoration efforts, although the 2007 A SMFC stock assessment indicated that stocks were still declining in most river systems along the east coast (ASMFC 2007).

The popul ation s ize of A merican $s$ had do es appear $t$ obe increasing int he 1 ower Susquehanna, pa rticularly s ince 2007 ( SPM e stimate). This follows a p eriod ( 2002 to 2007) when calculated indices of abundance generally decreased (including the hook a nd line CPUE, logbook C PAH a nd c reel C PAH). D espite $t$ his $t$ rend in a bundance, $t$ he 2012 hook a nd 1 ine CPUE was the lowest it has been since 1986 and there is no significant trend in CPUE over time. Gizzard s had a re increasing in a bundance in $t$ he $S$ usquehanna dr ainage a nd $m$ ay $r$ educe $t$ he number of lifted American shad by using the lifts the mselves, thus a ffecting lift CPUE. The Potomac R iver C PUE is increasing (1996-2012); how ever, the CPUE in $t$ he N anticoke R iver shows no significant trend (1988-2011), which suggests uneven area-wide recovery.

The Petersen estimate and the SPM are both useful techniques for providing estimates of American shad abundance at the Conowingo Dam. The SPM likely underestimates American shad abundance. For example, the estimated Conowingo Dam lift efficiency (defined as annual number of American shad lifted at Conowingo Dam divided by population estimate) was as high
as $98.7 \%$ in 2004 , a nd it is unl ikely $t$ hat $t$ he da $m$ pa ssed ne arly $100 \%$ of $t$ he $f$ ish in $t$ he Conowingo Dam tailrace. The Petersen statistic likely overestimates the population, especially in years of low recapture of tagged fish. However, the trends (rather than the actual numbers) produced bythe estimate/model should be emphasized when assessing the population at the Conowingo Dam in the Susquehanna River.

Scales are the only validated ageing structures for determining the age of American shad (Judy 1960, McBride et al. 2005). However, Cating's method of using transverse grooves is no longer recommended: c omparisons of A merican shad scales from different populations show different groove frequencies to the freshwater zone and first three annuli (Duffy et al. 2011). We will $r$ emain consistent $w$ ith historical a geing methods $u$ ntil a lternative a geing structures a re investigated.

The percent of repeat spawning American shad below the Conowingo Dam has increased over time, particularly since the truck and transport to locations above Safe Harbor Dam ceased in 1997 when the EFL was automated. The percent of repeat spawners was generally less than $10 \%$ in the early 1980s in the Conowingo Dam tailrace (Weinrich et al. 1982). In contrast, $50 \%$ of aged American shad at the Conowingo Dam were repeat spawners in 2012, a nd, on a verage, $27 \%$ of a ged fish were repeat spawners over the past five years. T urbine mortality for dams above the Conowingo Dam is considered to be $100 \%$, and the end of truck and transport in 1997 may have resulted in more fish surviving to return in following years. The same trend occurs in the P otomac R iver, w here t here is no hi story of t ruck a nd t ransport and da $\mathrm{ms}: \mathrm{t}$ he ave rage percent of repeat spawners was $17 \%$ in the 1950 s (Walburg and Sykes 1957), and is currently 48\%. Increased repeat spawning in both river systems may indicate increased survival of adult fish. This could be due to decreased harvest in Atlantic Ocean fisheries, increased abundance leading to more fish reaching older ages, and/or reductions in natural mortality.

The 2012 c alculated Z for A merican shad in the Conowingo Dam tailrace ( $\mathrm{Z}=0.61$ ) is below the $Z_{30}$ established for rivers in neighboring states (range $=0.62-0.76$ ), with the exception of the Hudson River $\left(Z_{30}=0.54\right.$; ASMFC 2007). The 2012 calculated $Z$ for American shad in the Nanticoke River ( $\mathrm{Z}=0.82$ ) is greater than the $\mathrm{Z}_{30}$ established for all rivers in neighboring states (range $=0.54-0.76$; A SMFC 2007). The $Z_{30}$ established for nor thern $r$ ivers (North C arolina to Maine; $\mathrm{Z}_{30}=1.93$ ) is greater than the 2012 Z for the Conowingo Dam tailrace and the Nanticoke River (ASMFC 2007). These calculated mortality e stimates may be maximum rates be cause repeat s pawning m arks are assessed during t he s pawning s eason after fish have returned to freshwater but before developing a new spawning mark.

No juvenile American shad have been captured in the Chester River trawls or seines since 2005. Baywide juvenile American shad indices decreased in 2012, as did juvenile indices in the upper Chesapeake Bay and the Nanticoke River. Only the juvenile index in the Potomac River increased in 2012. Other juvenile surveys in the Chesapeake Bay tributaries (from Maryland to Virginia [Virginia Institute of Marine Science, pers. comm.]) observed low numbers of a variety of juvenile a nadromous species in 2012, s uggesting poor recruitment. This low reproductive success is 1 ikely due $t$ o na tural va riability in $w$ eather c onditions. Fish lifted a bove t he Conowingo D am may reduce the num ber of po tential spawners due to turbine $m$ ortality, a nd inefficient lift facilities above the Conowingo Dam may also prevent spawners from reaching optimal s pawning ha bitat a bove $t$ he $Y$ ork $H$ aven $D$ am, $t$ hus a ffecting $j$ uvenile pr oduction. Predation by apex predators, particularly striped bass and the recently introduced flathead and blue catfish, may also affect juvenile survival.

## Hickory Shad

Hickory s had stocks ha ve dr astically de clined due to the loss of ha bitat, ove rfishing, stream bl ockages and pollution. A statewide moratorium on $t$ he $h$ arvest of hi ckory $s$ had in Maryland waters was implemented in 1981 and is still in effect today.

Adult hickory shad are difficult to c apture due to their a version to fishery independent (fish lifts) a nd de pendent (pound a nd fyke net) gears. Very few hi ckory shad are hi storically observed using the EFL in the Susquehanna River. A notable exception was in 2011 when 20 hickory shad were counted at the EFL counting window. No hickory shad were observed in the EFL in 2012. Despite $t$ he $t$ raditionally 1 ow $n$ umber of hi ckory shad obs erved pa ssing $t$ he Conowingo Dam, Deer Creek (a tributary to the Susquehanna River) has the greatest densities of hickory shad in Maryland (Richardson et al. 2009). Catch rates exceed four fish per hour for all years except 2009 a nd 2010 according to shad logbook data collected from Deer Creek anglers (1998-2012). Hickory shad are s ensitive to light a nd generally s trike a rtificial lur es mor e frequently when flows are somewhat elevated and the water is slightly turbid. Consequently, the low C PAH for hi ckory shad in 2009 m ay be di rectly related to the low flow and clear water conditions encountered by Deer Creek anglers and observed by MDNR staff during that spring season.

Hickory shad age structure has remained relatively consistent, with a wide range of ages and a high percentage of older fish. Ninety percent of hickory shad from the upper Chesapeake Bay spawn by age four, and this stock generally consists of few virgin fish (Richardson et. al 2004). Repeat spawning has remained relatively consistent over the 2004-2012 time series, with the percent of repeat spawners ranging between 64-89\%.

Hickory shad relative a bundance $m$ etrics in the $N$ anticoke $R$ iver (pound and fyke ne $t$ CPUE) a re $t$ enuous, pr esumably be cause of gear a voidance. $T$ herefore, $r$ elative a bundance
analysis f or hi ckory s had int he N anticoke River was di scontinued. Extensive spring electrofishing conducted in conjunction with Maryland stocking efforts in the Nanticoke River watershed concluded that stocks increased in this system from 2002 to 2009 (Richardson 2009). Maryland stocking and sampling of American shad in the Nanticoke River ended in 2009.

Estimates of $Z$ are attributable solely to $M$ because only a catch and release fishery exists for hickory shad in Maryland. The high percent of repeat spawners is also indicative of very low bycatch mortality. Hickory shad ocean bycatch is minimized compared to the ot her alosines because both mature adults and immature sub-adults migrate and overwinter closer to the coast (ASMFC 2009). This is confirmed by the fact that few hickory shad are observed portside as bycatch int he oc ean small-mesh fisheries (Matthew C ieri, Maine Dep. Marine R es., pe rs. comm.).

Hickory shad adults may spawn up to six weeks before American shad (late March to late April versus late April to early June), and juvenile hickory shad reach a larger size earlier in the summer. Because of their larger size, ability to a void gear, and preference for d eeper water, sampling for j uvenile hi ckory s had from mid-summer through fall is generally unsuccessful (Richardson et al. 2009). These juveniles also exhibit the same sensitivity to light as the adults, migrating $t$ o deeper, darker $w$ ater aw ay $f$ rom $t$ he $s$ hallow be aches $s$ ampled by hal $s$ eines. Sampling would need to be initiated prior to 1 J une in order to accurately assess hickory shad juvenile production.

## Alewife and Blueback Herring

Alewife and blueback herring num bers $h$ ave dr astically de clined for the $s$ ame reasons discussed pr eviously for American a nd hi ckory shad. A ccording tothe most recent stock assessment, the coa stwide meta-complex of $r$ iver he rring s tocks on $t$ he $U$.S. A tlantic coast is
depleted to near historic lows, and declines in the mean length of at least one age were observed in most rivers examined (ASMFC 2012). The depleted status indicates that there was evidence for declines in abundance due to a variety of factors, but the relative importance of these factors in stock reduction could not be de termined (ASMFC 2012). R iver he rring w ere a lso de emed depleted in the N anticoke R iver (ASMFC 2012). This assessment cor responds with the low commercial river herring landings observed in previous years in both the Nanticoke River and the entire state of Maryland. Specifically, the truncating age structure for river herring may be a sign of excessive mortality rates.

Juvenile a lewife a nd bl ueback pr oduction in the N anticoke R iver and upper B ay has generally been erratic, with frequent declines in abundance to very low levels. In 2012, alewife and blueback herring CPUE decreased for juveniles in both of these regions. Juvenile alewife and blueback herring indices de creased in all regions in 2012, a ccording to Maryland's E FJS survey; no river herring were encountered in the Patuxent River (Project 2, Job 3, Task 3).

Because river herring landings along the east coast have decreased significantly, ASMFC passed Amendment 2 of the A SMFC Interstate Fishery Management $P$ lan for A merican $S$ had and River Herring. T his a mendment required states to de velop a nd i mplement a s ustainable fishery pl an for jurisdictions wishing to maintain an open commercial or recreational fishery. Due to the decline in and persistently low levels of river herring in Maryland, a moratorium on the possession of river herring went into effect on 26 December 2011. It is no longer legal to possess river herring within the jurisdiction of Maryland unless the possessor has a bill of sale identifying the river herring as legally caught in waters not under Maryland jurisdiction. The expectation is that the new moratorium on river he rring will lead to increased production of juvenile river herring, and (in three to five years) an increase in the spawning stock.

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Figure 23. Upper bay juvenile alewife and blueback herring geometric mean CPUE (catch per haul), 1959-2012.

Table 1. P ercentage of sites with c lupeid eggs or larvae and num ber of sites sampled in the Nanticoke River (2005-2012).

| Year | Total <br> Sites | Percent of Sites <br> with Clupeid <br> Eggs/Larvae |
| ---: | ---: | ---: |
| 2005 | 80 | 5.0 |
| 2006 | 80 | 0.0 |
| 2007 | 78 | 0.0 |
| 2008 | 109 | 0.0 |
| 2009 | 97 | 8.2 |
| 2010 | 70 | 42.9 |
| 2011 | 73 | 32.9 |
| 2012 | 86 | 0.0 |

Table 2. Number of adult American shad and repeat spawners by sex and age sampled from the Conowingo Dam tailrace (hook and line), Nanticoke River (gears combined) and Potomac River in 2012.

Conowingo Dam Tailrace

| AGE | Male |  | Female |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Repeats | N | Repeats | N | Repeats |
| 3 | 4 | 0 | 1 | 0 | 5 | 0 |
| 4 | 31 | 5 | 1 | 0 | 32 | 5 |
| 5 | 46 | 15 | 14 | 5 | 60 | 20 |
| 6 | 22 | 15 | 34 | 25 | 56 | 40 |
| 7 | 0 | 0 | 23 | 23 | 23 | 23 |
| 8 | 0 | 0 | 1 | 1 | 1 | 1 |
| Totals | 103 | 35 | 74 | 54 | 177 | 89 |
| Percent <br> Repeats | $34.0 \%$ |  | $73.0 \%$ |  | $50.3 \%$ |  |

Nanticoke River

| AGE | Male |  | Female |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Repeats | N | Repeats | N | Repeats |
| 3 | 5 | 0 | 0 | 0 | 5 | 0 |
| 4 | 38 | 3 | 4 | 0 | 42 | 3 |
| 5 | 44 | 22 | 26 | 10 | 70 | 32 |
| 6 | 23 | 18 | 21 | 16 | 44 | 34 |
| 7 | 1 | 1 | 6 | 6 | 7 | 7 |
| 8 | 0 | 0 | 0 | 0 | 0 | 0 |
| Totals | 111 | 44 | 57 | 32 | 170 | 76 |
| Percent <br> Repeats | $39.6 \%$ |  | $56.1 \%$ |  | $45.2 \%$ |  |

Potomac River

| AGE | Male |  | Female |  | Total |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Repeats | N | Repeats | N | Repeats |  |  |  |
| 3 | 0 | 0 | 0 | 0 | 0 | 0 |  |  |  |
| 4 | 4 | 0 | 0 | 0 | 4 | 0 |  |  |  |
| 5 | 14 | 4 | 11 | 4 | 25 | 8 |  |  |  |
| 6 | 9 | 3 | 12 | 7 | 21 | 10 |  |  |  |
| 7 | 3 | 3 | 9 | 7 | 12 | 10 |  |  |  |
| 8 | 0 | 0 | 3 | 3 | 3 | 3 |  |  |  |
| Totals | 30 | 10 | 35 | 21 | 65 | 31 |  |  |  |
| Percent | $33.3 \%$ |  | $60.0 \%$ |  | $47.7 \%$ |  |  |  |  |
| Repeats |  |  |  |  |  |  |  |  |  |

Table 3. Number of recaptured American shad in 2012 at the Conowingo Dam East and West Fish Lifts by tag color and year.

| East Lift |  |  |
| :---: | :---: | :---: |
| Tag Color | Year Tagged | Number Recaptured |
| Orange | 2012 | 24 |
| Green | 2011 | 2 |
| West Lift |  |  |
| Tag Color | Year Tagged | Number Recaptured |
| Orange | 2012 | 4 |

Table 4. Catch ( numbers), e ffort (hours fished) a nd c atch-per-angler-hour (C PAH) from the recreational creel survey in the Susquehanna River below Conowingo Dam, 2001-2012. Due to sampling limitations, no data were available for 2011.

| Year | Number of <br> Interviews | Hours <br> Fished for <br> American <br> Shad | American <br> Shad Catch | American <br> Shad <br> CPAH |
| :---: | :---: | :---: | :---: | :---: |
| 2001 | 90 | 202.9 | 991 | 4.88 |
| 2002 | 52 | 85.3 | 291 | 3.41 |
| 2003 | 65 | 148.2 | 818 | 5.52 |
| 2004 | 97 | 193.3 | 233 | 1.21 |
| 2005 | 29 | 128.8 | 63 | 0.49 |
| 2006 | 78 | 227.3 | 305 | 1.34 |
| 2007 | 30 | 107.5 | 128 | 1.19 |
| 2008 | 16 | 32.5 | 24 | 0.74 |
| 2009 | 40 | 85.0 | 120 | 1.41 |
| 2010 | 36 | 64.0 | 114 | 1.78 |
| 2011 |  |  |  |  |
| 2012 | 58 | 189.0 | 146 | 0.77 |

Table 5. Catch (numbers), effort (hours fished) and catch-per-angler-hour (CPAH) from spring logbooks for American shad, 1999-2012.

| Year | Number <br> of <br> Returned <br> Logbooks | Hours <br> Fished for <br> American <br> Shad | American <br> Shad <br> Catch | American <br> Shad <br> CPAH |
| :---: | :---: | :---: | :---: | :---: |
| 1999 | 7 | 160.5 | 463 | 2.88 |
| 2000 | 10 | 404.0 | 3,137 | 7.76 |
| 2001 | 8 | 272.5 | 1,647 | 6.04 |
| 2002 | 8 | 331.5 | 1,799 | 5.43 |
| 2003 | 9 | 530.0 | 1,222 | 2.31 |
| 2004 | 15 | 291.0 | 1035 | 3.56 |
| 2005 | 12 | 258.5 | 533 | 2.06 |
| 2006 | 16 | 639.0 | 747 | 1.17 |
| 2007 | 10 | 242.0 | 873 | 3.61 |
| 2008 | 14 | 559.5 | 1,269 | 2.27 |
| 2009 | 15 | 378.0 | 967 | 2.56 |
| 2010 | 16 | 429.5 | 857 | 2.00 |
| 2011 | 9 | 174.0 | 413 | 2.37 |
| 2012 | 4 | 177.5 | 491 | 2.77 |

Table 6. Estimated adult American shad mortalities (in numbers) in Maryland waters (1997-2012). Lower Susquehanna River (below the Conowingo Dam) abundance estimates are derived from the surplus production model (SPM). W est Fish Lift mortality includes mortality due to day-to-day operations.

| Year | Total <br> Commercial <br> Landings in <br> Maryland's <br> Cortion of <br> Chesapeake <br> Bay | Conowingo <br> Dam East <br> Fish Lift <br> Mortality | Conowingo <br> Dam West <br> Fish Lift <br> Mortality | Estimated <br> Commercial <br> Chesapeake <br> Bay Bycatch <br> Mortality ${ }^{2}$ | Recreational <br> Bycatch <br> Mortality | Ocean <br> Commercial <br> Landings ${ }^{3}$ | Minimum <br> Total <br> Losses | Conowingo <br> Dam <br> Tailrace <br> Abundance <br> Estimate |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1997 | 0 | 43,790 | 2,274 | 4,200 | Unknown | 24,859 | 75,123 | 159,878 |
| 1998 | 0 | 16,152 | 1,300 | 4,200 | Unknown | 18,526 | 39,908 | 161,430 |
| 1999 | 0 | 43,455 | 3,136 | 4,200 | Unknown | 13,623 | 64,414 | 193,920 |
| 2000 | 0 | 60,452 | 3,102 | 4,200 | Unknown | 4,834 | 72,588 | 207,028 |
| 2001 | 0 | 130,876 | 2,607 | 4,200 | Unknown | 2,347 | 140,030 | 205,924 |
| 2002 | 0 | 40,142 | 2,837 | 4,200 | Unknown | 1,882 | 49,061 | 134,373 |
| 2003 | 0 | 50,224 | 2,160 | 4,200 | Unknown | 621 | 57,205 | 129,196 |
| 2004 | 0 | 29,911 | 1,218 | 4,200 | Unknown | 220 | 35,549 | 111,931 |
| 2005 | 0 | 42,873 | 1,412 | 4,200 | Unknown | 0 | 48,485 | 109,654 |
| 2006 | 0 | 41,201 | 1,696 | 4,200 | Unknown | 0 | 95,582 | 94,790 |
| 2007 | 0 | 14,120 | 1,737 | 4,200 | Unknown | 0 | 20,057 | 77,166 |
| 2008 | 0 | 7,075 | 1,477 | 4,200 | Unknown | 0 | 12,752 | 80,208 |
| 2009 | 0 | 15,490 | 1,566 | 4,200 | Unknown | 0 | 21,256 | 90,989 |
| 2010 | 0 | 21,793 | 1,219 | 4,200 | Unknown | 0 | 27,212 | 98,743 |
| 2011 | 0 | 5,159 | 1,038 | 4,200 | Unknown | 0 | 10,397 | 103,500 |
| 2012 | 0 | 8,714 | 710 | 4,200 | Unknown | 0 | 13,952 | 111,550 |

1 Estimated to be $100 \%$ of fish passing above Holtwood Dam and $25 \%$ turbine mortality of fish passing back through Conowingo Dam.
2 Extrapolated from American shad observed mortalities from pound nets in the upper Chesapeake Bay.
3 Reported numbers were calculated by multiplying total pounds by an estimated four pounds per fish.

Table 7. Number of juvenile a losines captured by species in seines and trawls on t he Chester River, 2007-2012.

Seine

|  | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| American Shad | 0 | 0 | 0 | 0 | 0 | 0 |
| Hickory Shad | 0 | 0 | 0 | 5 | 9 | 0 |
| Alewife | 1 | 1 | 18 | 2 | 19 | 0 |
| Blueback | 334 | 36 | 19 | 28 | 1,214 | 0 |

Trawl

|  | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| American Shad | 0 | 0 | 0 | 0 | 0 | 0 |
| Hickory Shad | 3 | 0 | 1 | 0 | 6 | 0 |
| Alewife | 33 | 12 | 27 | 11 | 6 | 0 |
| Blueback | 1 | 0 | 5 | 0 | 0 | 0 |

Table 8. N umber of adult hickory shad and repeat spawners by sex and age sampled from the brood stock collection survey in Deer Creek (Susquehanna River tributary) in 2012.

| AGE | Male |  | Female |  | Total |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Repeats | N | Repeats | N | Repeats |  |  |  |  |  |  |  |  |  |  |
| 3 | 40 | 0 | 13 | 0 | 53 | 0 |  |  |  |  |  |  |  |  |  |  |
| 4 | 55 | 42 | 24 | 18 | 79 | 60 |  |  |  |  |  |  |  |  |  |  |
| 5 | 28 | 28 | 21 | 21 | 49 | 49 |  |  |  |  |  |  |  |  |  |  |
| 6 | 6 | 6 | 9 | 9 | 15 | 15 |  |  |  |  |  |  |  |  |  |  |
| 7 | 0 | 0 | 4 | 4 | 4 | 4 |  |  |  |  |  |  |  |  |  |  |
| 8 | 0 | 0 | 2 | 2 | 0 | 0 |  |  |  |  |  |  |  |  |  |  |
| Totals | 129 | 76 | 87 | 66 | 200 | 128 |  |  |  |  |  |  |  |  |  |  |
| Percent <br> Repeats | $58.9 \%$ |  |  |  |  |  |  | $73.2 \%$ |  |  |  |  |  |  |  | $64.0 \%$ |

Table 9. Percent of hickory shad by age and number sampled from the brood stock collection survey in Deer Creek (Susquehanna River tributary) by year, 2004-2012.

| Year | N | Age 2 | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 | Age 8 | Age 9 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2004 | 80 |  | 7.5 | 23.8 | 27.5 | 18.8 | 18.8 | 3.8 |  |
| 2005 | 80 |  | 6.3 | 17.5 | 28.8 | 33.8 | 11.3 | 1.3 | 1.3 |
| 2006 | 178 | 0.6 | 9 | 31.5 | 29.8 | 20.2 | 7.3 | 1.7 |  |
| 2007 | 139 |  | 6.5 | 23.7 | 33.8 | 20.9 | 12.2 | 2.2 | 0.7 |
| 2008 | 149 |  | 9.4 | 29.5 | 33.6 | 20.1 | 5.4 | 2 |  |
| 2009 | 118 |  | 7.6 | 16.9 | 44.9 | 19.5 | 10.2 | 0.8 |  |
| 2010 | 240 |  | 12.5 | 37.9 | 31.3 | 11.3 | 6.7 | 0.4 |  |
| 2011 | 216 |  | 30.1 | 30.1 | 27.3 | 8.8 | 2.78 | 0.93 |  |
| 2012 | 200 |  | 26.5 | 39.5 | 24.5 | 7.5 | 2.0 |  |  |

Table 10. Percent repeat spawning hickory shad (sexes combined) by year from the brood stock collection survey in Deer Creek (Susquehanna River tributary), 2004-2012.

| Year | N | Percent <br> Repeats |
| :---: | :---: | :---: |
| 2004 | 80 | 68.8 |
| 2005 | 80 | 82.5 |
| 2006 | 178 | 67.4 |
| 2007 | 139 | 79.1 |
| 2008 | 149 | 83.9 |
| 2009 | 118 | 89.0 |
| 2010 | 240 | 75.4 |
| 2011 | 216 | 68.5 |
| 2012 | 200 | 64.0 |

Table 11. Catch (numbers), effort (hours fished) and catch-per-angler-hour (CPAH) from spring logbooks for hickory shad, 1998-2012.

| Year | Number of <br> Returned <br> Logbooks | Hours <br> Fished for <br> Shad | Hickory <br> Shad Catch | Hickory <br> Shad <br> CPAH |
| :---: | :---: | :---: | :---: | :---: |
| 1998 | 19 | 600.0 | 4,980 | 8.30 |
| 1999 | 15 | 817.0 | 5,115 | 6.26 |
| 2000 | 14 | 655.0 | 3,171 | 14.8 |
| 2001 | 13 | 533.0 | 2,515 | 4.72 |
| 2002 | 11 | 476.0 | 2,433 | 5.11 |
| 2003 | 14 | 635.0 | 3,143 | 4.95 |
| 2004 | 18 | 750.0 | 3,225 | 4.30 |
| 2005 | 19 | 474.0 | 2,094 | 4.42 |
| 2006 | 20 | 766.0 | 4,902 | 6.40 |
| 2007 | 17 | 401.0 | 3,357 | 8.37 |
| 2008 | 22 | 942.0 | 5,465 | 5.80 |
| 2009 | 15 | 561.0 | 2,022 | 3.60 |
| 2010 | 16 | 552.0 | 1,956 | 3.54 |
| 2011 | 9 | 224.3 | 1,802 | 8.03 |
| 2012 | 5 | 190.0 | 866 | 4.56 |

Table 12. Catch-at-age and repeat spawners by sex and age for adult alewife and blueback herring sampled from the Nanticoke River in 2012. Approximately $30 \%$ of measured river he rring w ere aged, and an age-length key was used to covert from num ber at length to number at age.

Alewife Herring

| AGE | Male |  | Female |  | Total |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Repeats | N | Repeats | N | Repeats |  |  |
| 3 | 33 | 0 | 36 | 0 | 70 | 0 |  |  |
| 4 | 76 | 4 | 85 | 0 | 160 | 4 |  |  |
| 5 | 65 | 24 | 111 | 74 | 175 | 94 |  |  |
| 6 | 20 | 20 | 74 | 67 | 94 | 87 |  |  |
| 7 | 0 | 0 | 26 | 26 | 26 | 26 |  |  |
| 8 | 0 | 0 | 0 | 0 | 0 | 0 |  |  |
| 9 | 0 | 0 | 0 | 0 | 0 | 0 |  |  |
| Totals | 194 | 47 | 332 | 167 | 526 | 214 |  |  |
| Percent <br> Repeats | 5 |  |  |  |  |  |  |  |

Blueback Herring

| AGE | Male |  | Female |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Repeats | N | Repeats | N | Repeats |
| 2 | 4 | 0 | 0 | 0 | 4 | 0 |
| 3 | 44 | 0 | 24 | 0 | 68 | 0 |
| 4 | 101 | 7 | 71 | 9 | 172 | 16 |
| 5 | 63 | 30 | 64 | 28 | 127 | 58 |
| 6 | 14 | 14 | 13 | 4 | 27 | 18 |
| 7 | 0 | 0 | 3 | 3 | 3 | 3 |
| 8 | 0 | 0 | 0 | 0 | 0 | 0 |
| Totals | 226 | 51 | 175 | 44 | 401 | 95 |
| Percent Repeats | 22.6\% |  | 25.1\% |  | 23.7\% |  |

Table 13. Mean length-at-age by sex for al ewife he rring sampled from the N anticoke R iver, 1989-2012.

| Year | Males |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2 | 3 | 4 | 5 | 6 | 7 |  |  |  |
| 1989 |  | 230 | 236 | 243 | 256 | 261 |  |  |  |
| 1990 |  | 221 | 231 | 244 | 250 | 263 | 264 |  |  |
| 1991 |  | 224 | 234 | 240 | 251 | 260 | 243 |  |  |
| 1992 |  | 216 | 228 | 238 | 247 | 254 |  |  |  |
| 1993 |  | 208 | 225 | 239 | 246 | 248 | 246 |  |  |
| 1994 |  | 207 | 219 | 231 | 239 | 246 |  |  |  |
| 1995 |  | 214 | 226 | 238 | 246 | 251 | 244 |  |  |
| 1996 | 212 | 219 | 228 | 238 | 242 | 263 |  |  |  |
| 1997 |  | 213 | 228 | 233 | 240 |  | 252 |  |  |
| 1998 |  | 217 | 225 | 238 | 243 | 254 |  |  |  |
| 1999 |  | 211 | 222 | 233 | 238 | 244 |  |  |  |
| 2000 |  | 220 | 228 | 238 | 258 |  |  |  |  |
| 2001 |  | 225 | 234 | 240 | 247 |  |  |  |  |
| 2002 |  | 225 | 233 | 241 | 244 | 248 |  |  |  |
| 2003 |  | 228 | 239 | 245 | 251 |  |  |  |  |
| 2004 |  | 228 | 242 | 251 | 250 |  |  |  |  |
| 2005 |  | 214 | 226 | 236 | 252 | 252 |  |  |  |
| 2006 |  | 219 | 223 | 235 | 242 |  |  |  |  |
| 2007 |  | 219 | 227 | 235 | 248 |  |  |  |  |
| 2008 |  | 216 | 217 | 229 | 235 | 278 |  |  |  |
| 2009 |  | 221 | 224 | 231 | 241 |  |  |  |  |
| 2010 |  | 221 | 224 | 232 | 248 |  |  |  |  |
| 2011 |  |  | 215 | 229 | 233 | 244 |  |  |  |
| 2012 |  |  |  |  |  |  |  |  |  |

Females

| Year |  |  | Age |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 1989 |  | 229 | 244 | 253 | 267 | 277 | 286 |  |  |
| 1990 |  | 225 | 238 | 253 | 261 | 274 | 283 | 286 |  |
| 1991 |  | 227 | 243 | 251 | 263 | 270 | 273 | 286 |  |
| 1992 |  | 223 | 240 | 248 | 256 | 265 | 276 | 279 |  |
| 1993 |  | 225 | 233 | 247 | 256 | 265 | 277 |  |  |
| 1994 |  | 219 | 228 | 243 | 254 | 258 | 270 |  |  |
| 1995 |  | 221 | 235 | 252 | 263 | 268 | 274 |  | 280 |
| 1996 |  | 219 | 231 | 250 | 257 | 267 | 268 | 260 |  |
| 1997 |  | 228 | 234 | 242 | 253 | 267 | 271 |  |  |
| 1998 |  | 224 | 235 | 245 | 255 | 264 |  | 277 |  |
| 1999 |  | 220 | 229 | 242 | 250 | 260 | 272 |  |  |
| 2000 |  | 237 | 237 | 250 | 257 | 270 |  |  |  |
| 2001 |  | 239 | 243 | 249 | 256 | 266 | 270 |  |  |
| 2002 |  | 226 | 238 | 248 | 255 | 260 | 263 |  |  |
| 2003 |  | 235 | 239 | 250 | 260 | 263 |  |  |  |
| 2004 |  |  | 233 | 259 | 262 | 270 |  |  |  |
| 2005 |  | 228 | 240 | 243 | 257 | 267 | 272 |  |  |
| 2006 |  | 220 | 236 | 247 | 256 | 256 | 264 | 277 |  |
| 2007 |  | 217 | 231 | 238 | 248 | 256 | 276 | 279 |  |
| 2008 |  | 215 | 231 | 242 | 252 | 261 |  |  |  |
| 2009 |  |  | 234 | 245 | 257 | 251 |  |  |  |
| 2010 |  |  |  | 226 | 236 | 247 | 256 | 268 | 275 |
| 2011 |  |  | 218 | 233 | 249 | 260 | 263 |  |  |
| 2012 |  |  |  |  |  |  |  |  |  |

Table 14. Mean length-at-age by sex for blueback herring sampled from the Nanticoke River, 1989-2012.

| Year | Age |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 1989 |  | 218 | 227 | 234 | 245 | 259 | 262 | 279 |  |
| 1990 |  | 218 | 232 | 239 | 249 | 258 | 263 | 270 |  |
| 1991 |  | 217 | 229 | 237 | 247 | 258 | 260 | 273 |  |
| 1992 |  | 212 | 224 | 235 | 245 | 251 | 260 | 256 |  |
| 1993 |  | 205 | 224 | 237 | 247 | 256 | 262 | 261 |  |
| 1994 |  | 213 | 223 | 238 | 250 | 256 |  |  |  |
| 1995 |  | 220 | 226 | 233 | 247 | 256 |  |  |  |
| 1996 | 205 | 219 | 230 | 240 | 244 | 270 | 261 |  |  |
| 1997 |  | 212 | 225 | 238 | 241 | 247 | 257 |  |  |
| 1998 |  | 212 | 225 | 233 | 245 | 253 |  |  |  |
| 1999 |  | 200 | 222 | 232 | 239 | 251 |  |  |  |
| 2000 |  | 219 | 225 | 235 | 246 | 249 |  |  |  |
| 2001 |  | 218 | 231 | 235 | 250 |  |  |  |  |
| 2002 |  | 217 | 229 | 234 | 243 |  |  |  |  |
| 2003 | 215 | 230 | 240 | 238 |  |  |  |  |  |
| 2004 | 216 | 231 | 234 | 245 | 250 |  |  |  |  |
| 2005 |  | 222 | 226 | 238 |  |  |  |  |  |
| 2006 |  | 209 | 224 | 235 | 236 | 270 |  |  |  |
| 2007 |  | 207 | 221 | 227 | 266 |  |  |  |  |
| 2008 |  | 206 | 216 | 220 |  |  |  |  |  |
| 2009 |  | 214 | 219 | 231 |  |  |  |  |  |
| 2010 |  | 219 | 227 |  | 228 |  |  |  |  |
| 2011 |  | 206 | 220 | 226 | 234 |  |  |  |  |
| 2012 | 212 | 207 | 217 | 229 | 229 |  |  |  |  |

Females

| Year | Age |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 1989 |  | 227 | 236 | 244 | 257 | 271 | 279 | 297 |  |
| 1990 |  |  | 241 | 252 | 262 | 271 | 281 | 286 | 291 |
| 1991 |  | 228 | 238 | 251 | 260 | 264 | 273 | 285 |  |
| 1992 |  | 230 | 230 | 250 | 260 | 264 | 272 | 281 |  |
| 1993 |  | 220 | 236 | 246 | 259 | 269 | 277 | 290 | 296 |
| 1994 |  | 215 | 226 | 245 | 260 | 272 | 282 | 277 |  |
| 1995 |  | 228 | 235 | 248 | 260 | 264 | 270 |  |  |
| 1996 |  | 218 | 238 | 249 | 257 | 275 | 278 |  |  |
| 1997 |  | 226 | 242 | 247 | 254 | 268 | 276 | 290 |  |
| 1998 |  |  | 233 | 246 | 257 | 265 | 281 |  |  |
| 1999 |  | 219 | 236 | 244 | 253 | 273 |  |  |  |
| 2000 |  | 227 | 231 | 243 | 260 | 269 | 275 |  |  |
| 2001 |  | 219 | 242 | 248 | 260 | 273 |  |  |  |
| 2002 |  | 220 | 235 | 246 | 257 | 260 |  |  |  |
| 2003 | 224 | 235 | 248 | 252 | 264 | 283 |  |  |  |
| 2004 |  | 236 | 245 | 254 | 262 | 262 |  |  |  |
| 2005 |  | 241 | 236 | 248 | 264 |  |  |  |  |
| 2006 |  | 204 | 235 | 242 | 246 |  |  |  |  |
| 2007 |  | 217 | 221 | 246 | 247 | 266 |  |  |  |
| 2008 |  | 213 | 227 | 234 | 252 | 251 | 261 |  |  |
| 2009 |  | 227 | 232 | 242 | 260 | 278 |  |  |  |
| 2010 |  |  | 243 | 238 | 247 |  |  |  |  |
| 2011 |  | 201 | 240 | 238 | 251 | 262 |  |  |  |
| 2012 |  | 213 | 230 | 244 | 249 | 256 |  |  |  |

Table 15. Regression statistics for length by age and sex over time for alewife and blueback herring (1989-2012). Only ages with consistent representation over time were considered. Bolded values indicate significant changes in length-at-age over time.

| Alewife | Males |  |  |  |  | Females |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | N | Slope | $r^{2}$ | $P$ | N | Slope | $r^{2}$ | $P$ |  |  |
| 3 | 391 | -0.115 | 0.00432 | 0.194 | 122 | -0.21 | 0.0199 | 0.121 |  |  |
| 4 | 1389 | -0.386 | 0.0573 | $<\mathbf{0 . 0 0 1}$ | 1276 | -0.313 | 0.0429 | $<\mathbf{0 . 0 0 1}$ |  |  |
| 5 | 1137 | -0.39 | 0.0562 | $<\mathbf{0 . 0 0 1}$ | 1711 | -0.27 | 0.0341 | $<\mathbf{0 . 0 0 1}$ |  |  |
| 6 | 473 | -0.393 | 0.0608 | $<\mathbf{0 . 0 0 1}$ | 1092 | -0.261 | 0.034 | $<\mathbf{0 . 0 0 1}$ |  |  |
| 7 | 70 | -0.937 | 0.178 | $<\mathbf{0 . 0 0 1}$ | 353 | -0.377 | 0.0772 | $<\mathbf{0 . 0 0 1}$ |  |  |
| 8 |  |  |  |  | 96 | -0.518 | 0.0735 | $\mathbf{0 . 0 0 8}$ |  |  |


| Blueback | Males |  |  |  |  | Females |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | N | Slope | $r^{2}$ | $P$ | N | Slope | $r^{2}$ | $P$ |  |
| 3 | 216 | -0.32 | 0.069 | $<\mathbf{0 . 0 0 1}$ | 57 | -0.496 | 0.156 | $\mathbf{0 . 0 0 2}$ |  |
| 4 | 905 | -0.337 | 0.0558 | $<\mathbf{0 . 0 0 1}$ | 790 | -0.13 | 0.00556 | 0.036 |  |
| 5 | 966 | -0.277 | 0.0263 | $<\mathbf{0 . 0 0 1}$ | 955 | -0.275 | 0.0322 | $<\mathbf{0 . 0 0 1}$ |  |
| 6 | 657 | -0.583 | 0.0681 | $<\mathbf{0 . 0 0 1}$ | 699 | -0.448 | 0.0399 | $<\mathbf{0 . 0 0 1}$ |  |
| 7 | 281 | -0.602 | 0.03 | $\mathbf{0 . 0 0 4}$ | 341 | -0.406 | 0.0349 | $<\mathbf{0 . 0 0 1}$ |  |
| 8 | 90 | -0.259 | 0.00247 | 0.641 | 111 | -0.43 | 0.0198 | 0.141 |  |
| 9 | 21 | -4.561 | 0.258 | $\mathbf{0 . 0 1 9}$ | 33 | -0.0055 | $<0.0001$ | 0.996 |  |

Figure 1. Conowingo Dam Tailrace (Susquehanna River) hook a nd line sampling location for American shad in 2012.


Figure 2. Nanticoke River pound net sites for adult alosine sampling in 2012. The Mill Creek pound net site used for calculating American shad CPUE is identified.


Figure 3. Nanticoke River sites for alosine ichthyoplankton sampling in 2012.


Figure 4. Chester River sampling sites for juvenile alosine species in 2012. Each black circle indicates the approximate location of a paired seine and trawl site.


Figure 5. Percent of American shad repeat spawners by sex collected in the Conowingo Dam tailrace (1982-2012).


Figure 6. Arcsine-transformed percentages of repeat spawning American shad (sexes combined) collected from the Conowingo Dam tailrace, 1984-2012.


Figure 7. Arcsine-transformed percentages of repeat spawning American shad (sexes combined) collected from the Nanticoke River, 1988-2012.


Figure 8. Trends in arcsine-transformed percentages of repeat spawning American shad (sexes combined) collected from the Potomac River, 2002-2012.


Figure 9. Conowingo Dam tailrace adult American shad abundance estimates from the Petersen statistic and the surplus production model (SPM), 1986-2012.


Figure 10. American shad geometric mean CPUE (fish per boat hour) from the Conowingo Dam tailrace hook and line sampling, 1984-2012.


Figure 11. American shad geometric mean CPUE (fish per lift hour) from the East and West Fish Lifts at the Conowingo Dam, 1980-2012.


Figure 12. American shad geometric mean CPUE (fish per net day) from the Mill Creek pound net in the Nanticoke River, 1988-2012. No pound nets were fished in 2004.


Figure 13. American shad geometric mean CPUE (fish per 1000 s quare yards of ex perimental drift gill net per hour fished) from the Potomac River, 1996-2012.


Figure 14. Baywide juvenile American shad geometric mean CPUE (catch per haul), 1959-2012.


Figure 15. Upper Chesapeake Bay juvenile A merican shad geometric mean CPUE (catch per haul), 1959-2012.


Figure 16. Nanticoke R iver juvenile A merican shad geometric mean CPUE (catch per haul), 1959-2012.


Figure 17. P otomac R iver juvenile A merican shad geometric m ean CPUE ( catch per ha ul), 1959-2012.


Figure 18. Arcsine-transformed percentages of repeat spawning hickory shad (sexes combined) collected from Deer Creek (Susquehanna River tributary), 2004-2012.


Figure 19. Arcsine-transformed percentages of repeat spawning alewife and blueback he rring (sexes and gears combined) from the Nanticoke River, 1989-2012.


Figure 20. Geometric mean CPUE (catch per net day) of adult a lewife and blueback herring from Nanticoke River fyke nets, 1989-2011. No fyke nets were fished in 2012. The CPUE for blueback herring is significantly declining over the time series.


Figure 21. Maryland's commercial river herring landings, 1929-2012.


Figure 22. Nanticoke River juvenile alewife and blueback herring geometric mean CPUE (catch per haul), 1959-2012.


Figure 23. Upper Bay juvenile alewife and blueback herring geometric mean CPUE (catch per haul), 1959-2012.


# PROJECT NO. 2 

JOB NO. 2

# STOCK ASSESSMENT OF SELECTED RECREATIONALLY IMPORTANT ADULT MIGRATORY FINFISH IN MARYLAND'S CHESAPEAKE BAY 

Prepared by Harry W. Rickabaugh Jr. and Katherine M. Messer

## INTRODUCTION

The pr imary obj ective of Project 2 Job 2 w as t o characterize r ecreationally important migratory finfish stocks in Maryland's Chesapeake Bay by age, length, weight, growth and sex. Weakfish (Cynoscion regalis), bluefish (Pomatomus saltatrix), Atlantic croaker (Micropogonias undulates), summer flounder (Paralichthys dentatus) a nd spot (Leiostomus xanthurus) a re ve ry i mportant s port fish in M aryland's C hesapeake B ay. Red dr um (Sciaenops ocellatus), bl ack dr um ( Pogonias cromis), s potted s eatrout (Cynoscion nebulosus) and S panish m ackerel (Scomberomorus maculates) ar el ess popular in M aryland be cause of 1 ower abundance, but ar et argeted by ang lers when available (Chesapeake Bay Program 1993). Atlantic menhaden (Brevoortia tyrannus) are a key component to the Bay's food chain as forage for predatory sport fish (Hartman and Brandt 1995, Overton et al 2000).

The M aryland Department of N atural R esources ( MD DNR) ha s co nducted summer pound ne t sampling for these species since 1993. T he data collected from this effort pr ovide information $f$ or $t$ he pr eparation and upda ting of $s$ tock a ssessments a nd fishery management plans for the Chesapeake Bay, the Atlantic States Marine Fisheries Commission (ASMFC) and the Mid-Atlantic Fishery Management Council (MAFMC).

This inf ormation is a lso utilized by the M D DNR in managing $t$ he state's va luable migratory finfish resources through the regulatory/statutory process.

## METHODS

## Data Collection

The onboard pound $n$ et $s$ urvey $r$ elies on vol untary cooperation of po und ne $t$ fishermen. Pound nets from the lower Chesapeake B ay and P otomac River have been consistently m onitored throughout the 20 years of this survey (1993-2012). H owever, since no c ooperating fishermen could be located on the lower Potomac River, sampling was not conducted in this area for 2009, but sampling resumed in 2010. Five commercial pound nets were sampled at the Potomac River and Chesapeake Bay between cove Point and Point No Point in 2012 (Figure 1). Each site was sampled once every two weeks, weather and fisherman's s chedule p ermitting. The c ommercial fishermen set all n ets sampled as part of their regular fishing routine. Net soak time and manner in which they were fished were consistent w ith the f isherman's da y-to-day ope rations. In 2012 additional data was gathered from a commercial gill net fisherman in Fishing Bay, this supplemented the pound net data. Two circle gillnets ( $3 \frac{1}{4}$ inch mesh, 1000 ft in length, and 6 ft deep) were sampled on July 1st. Only Atlantic croaker and spotted seatrout were caught.

All targeted species were measured from each net when possible. When it was not practical to measure all fish, a random sample of each species was measured and the remaining individuals enumerated if possible. All measurements were to the nearest mm total length (TL) except for Spanish mackerel, which were measured to the nearest mm
fork length (FL). Fifty randomly selected menhaden were measured to the ne arest mm FL each day, when available, and scale samples were taken from 25 of the measured fish. Menhaden s cales w ere aged b y t wo M D D NR bi ologists. Water t emperature $\left({ }^{\circ} \mathrm{C}\right)$, salinity (ppt), GPS coordinates (NAD 83), date and hours fished were also recorded at each net.

Otoliths, weight (g), TL (mm) and sex were de termined from a sub sample of weakfish, spot and Atlantic croaker. Prior to 2011 Atlantic croaker and weakfish otoliths were processed a nd aged by the S outh C arolina D epartment of N atural R esources (SC DNR). Otoliths from 2011 and 2012 w ere aged by MD DNR biologists. 2010 Atlantic croaker ages from SC DNR were compared to the MD DNR ages to evaluate consistency between agencies. Left ot oliths w ere s ectioned and read by S C D NR a nd the $r$ ight otoliths were sectioned and read by MD DNR, meaning an y obs erved differences were for t he e ntire a ging pr ocess, not j ust di fferences in r eader i nterpretation. Forty six otoliths were compared and there was $96 \%$ agreement in Atlantic croaker ages. All spot otoliths from 2012 were processed and aged by MD DNR, as in previous years. For all three $s$ pecies, the left otolith from e ach specimen was $m$ ounted to a gl ass s lide for sectioning. Otoliths were mounted in Crystalbond 509 and were sectioned with a Buehler IsoMet ${ }^{\circledR}$ Low Speed Saw using two blades separated by a 0.4 mm spacer. The Buehler 15 HC diamond wafering blades are 101.6 mm in diameter and 0.3048 mm thick. The 0.4 mm sections were then mounted on microscope slides and viewed under a microscope at 5 X to 6 X to determine the number of annuli. All age structures were read by two readers. If readers did not agree, both readers reviewed the structures together, and if a greement
still c ould not be reached the sample was not assigned an a ge. If the left otolith was damaged, missing or miss cut the right otolith was substituted.

Juvenile indices were calculated for weakfish, Atlantic croaker and spot from the MD DNR Blue Crab Trawl Survey data. This survey utilizes a 4.9 m semi-balloon otter trawl with a body and cod end of $25-\mathrm{mm}$-stretch-mesh and a 13 -mm-stretch-mesh cod end liner towed for 6 m in at $4.0-4.8 \mathrm{~km} / \mathrm{h}$. T he s ystems sampled included the Chester River, E astern B ay, C hoptank R iver a nd P atuxent R iver ( six fixed s ampling s tations each), T angier S ound (five fixed stations) and P ocomoke S ound (eight fixed stations). Each station was sampled once a month from May - October. Juvenile croaker, spot and weakfish collected by this survey have be en e numerated, and e ntered into a computer database since 1989 (Davis et al.1995).

## Analytical Procedures

Commercial and recreational harvest for the $t$ arget s pecies $w$ ere ex amined utilizing Maryland's mandatory com mercial $r$ eporting $s$ ystem and the Marine Recreational Information Program (MRIP; National M arine F isheries S ervice, personal communication), respectively. MRIP data was downloaded on January 17, 2013. Since these data sets are not finalized until the spring of the following year, harvest data for this report are through 2011. Harvest from M aryland's commercial reporting system was divided by area i nto Chesapeake Bay, A tlantic Ocean (including coastal ba ys) and unknown areas.

Beginning in 1993, Maryland has required charter boatc aptains to submit log books indicating the number of trips, number of anglers and number of fish harvested and
released by species. Trips in which a species was targeted but not caught could not be distinguished in the log books since no indication of target species is given. Chesapeake Bay geometric mean cat ch per an gler (CPA) indices were de rived for eight of the ten target species. No indices w ere cal culated for red drum due to small sample size, or menhaden, since it is not recreationally harvested. Log (catch / angler trip) compared to year $w$ as analyzed using line ar $r$ egression to identify $s$ ignificant tr ends in relative abundance. The statewide MRIP estimates include all anglers (private and for hire) and all areas (Chesapeake Bay, Coastal Bays and Atlantic Ocean). All Maryland charter boat data was from Chesapeake Bay for the target species. The for hire inland only estimates do not include the Atlantic Ocean and are only for anglers that paid another individual to take them fishing, and may be more comparable to the charter boat log data. Numbers of fish harvested by charter boa ts for each species w as com pared to statewide MRIP recreational catch estimates (numbers), MRIP inland only for hire estimates (numbers), and r eported C hesapeake B ay c ommercial 1 andings ( pounds), us ing linear regression, with P values of 0.01 or less were considered significant. Since the 2012 charter log book data had not been finalized, only data through 2011 was utilized for analysis.

Instantaneous total mortality $r$ ates $f$ or $w$ eakfish and Atlantic cr oaker w ere calculated using the Ssentongo and Larkin (1973) length based method,

$$
\mathrm{Z}=\left\{\mathrm{K} /\left(\mathrm{y}_{\mathrm{bar}}-\mathrm{y}_{\mathrm{c}}\right)\right\}
$$

where lengths are converted: $\mathrm{y}=-\log _{e}\left(1-\mathrm{L} / \mathrm{L}_{\infty}\right)$, and $\mathrm{y}_{\mathrm{c}}=-\log _{e}\left(1-\mathrm{L}_{\mathrm{c}} / \mathrm{L}_{\infty}\right), \mathrm{L}=$ total length, $L_{c}=$ length of first recruitment to the fisheries, $K=$ growth coefficient and $L_{\infty}=$ length that an average fish would achieve if it continued to grow. Von Bertalanffy parameters ( K and $\mathrm{L}_{\infty}$ ) for weakfish for all years were estimated from otolith ages collected during
the 1999 Chesapeake B ay pound ne ts urvey (Jarzynski et al 2000). V on B ertalanffy parameters for croaker mortality estimates were derived from pooled ages (otoliths; $\mathrm{n}=$ $2,284)$ de termined from 200 3-2012 Chesapeake B ay pound ne $t \mathrm{~s}$ urvey data, a nd June through $S$ eptember 2003-2012 measurements of a ge $z$ ero croaker ( $\mathrm{n}=197$ ) from MD DNR Blue C rab Trawl S urvey Tangier S ound samples (Chris Walstrum M D D NR personnel communication 2008). Trawl data were included to provide age zero fish that had not recruited to the pound net gear, and represented samples taken from the same time period and region as the pound net samples. Parameters for weakfish were $L_{\infty}=840$ mm TL and $\mathrm{K}=0.08$. $\mathrm{L}_{\mathrm{c}}$ was 305 m m TL . Parameters for A tlantic croaker es timates from 2003-2012 were $\mathrm{L}_{\infty}=413.7 \mathrm{~mm}$ TL and $\mathrm{K}=0.321$, while $\mathrm{L}_{\mathrm{c}}$ for Atlantic croaker was 229 mm TL.

Relative stock de nsity (RSD) w as us ed to characterize 1 ength di stributions for weakfish, s ummer f lounder, bl uefish a nd A tlantic croaker ( Gablehouse 1984). Only onboard pound net sampling was utilized for this analysis. Incremental RSD's group fish into five br oad descriptive length categories: stock, quality, preferred, memorable and trophy. The minimum length of each category is based on all-tackle world records such that the mini mum s tock le ngth is $20-26 \%$, minimum qua lity 1 ength is $36-41 \%$, minimum preferred length is $45-55 \%$, minimum m emorable length is $59-64 \%$ a nd minimum trophy length is $74-80 \%$ of the world record lengths. M inimum lengths for the target species were assigned from either the cut-offs listed by Gablehouse (1984) or derived from world record lengths recorded by the International Game Fish Association (Table 1).

Length frequency di stributions were constructed for s ummer flounder, Atlantic croaker and s pot, ut ilizing onboard sampling length d ata divided into 20 mml ength groups. In order to detect differences gill net caught fish a nd pound net c aught fish, length frequency distributions were calculated separately. Only A tlantic croaker sample size was adequate to construct length frequency distributions for both gears.

Length-at-age keys were constructed for weakfish and Atlantic croaker using age samples through 2012. Age and length data were assigned to 20 mm TL groups for each species and then the le ngth-at-age key was applied to the length frequency by ye ar to determine the proportion at age for croaker in 2000 and 2002 through 2012 and weakfish from 2003 through 2012. Age length keys for spot were constructed for 2007 t hrough 2012. A ge and length data were assigned to 10 mm T L groups for spot and then the length-at-age key was applied to the length frequency to determine the proportion at age by ye ar. It was n ecessary to supplement M D D NR s pot a ges with V irginia $M$ arine Recourses Commission (VMRC) spot age data for a small number of fish greater than 27 cm in the 2007, 2011 and 2012 samples.

Chesapeake B ay j uvenile i ndices w ere calculated as the geometric me an (GM) catch per tow. Since juvenile weakfish have be en consistently caught only in Tangier and Pocomoke sounds, only these areas were utilized in this analysis to minimize zeros that may represent uns uitable ha bitat r ather t han relative abundance. Similarly t he Atlantic croaker index was limited to Tangier Sound, Pocomoke Sound and the Patuxent River. All sites and ar eas were used for the spot index. Indices and $95 \%$ confidence intervals were derived using SAS ${ }^{\circledR}$ software (SAS 2006).

## RESULTS and DISCUSSION

The Potomac River and the Cheasapeake Bay were sampled from May 22, 2012 through September 11, 2012 (Table 2). All target species, and twenty non-target species (Table 3) were encountered during this time period.

## Weakfish

Ninety three weakfish were sampled in the 2012 pound net s urvey, the eighth lowest catch of the 20 year time series. Weakfish mean length in 2012 was 284 mm TL , an increase from the 2011 mean length of 236 mm TL , just below the time series annual mean 1 ength of 295 m m T L (Table 4 ). Weakfish RSD results for 2012 were $11 \%$ $\mathrm{RSD}_{\text {quality, }} 2 \% \mathrm{RSD}_{\text {preferred }}$ and $\mathrm{RSD}_{\text {stock }}$ accounting remaining catch (Table 5 ). This follows $t$ hree consecutive years in which all sampled weakfish were in $t$ he $\mathrm{RSD}_{\text {stock }}$ category. The 2012 onboard pound net survey length frequency distribution corroborated the shift to larger sizes with $65 \%$ of sampled weakfish in the 230 to 270 mm TL groups (Figure 2).

Chesapeake Bay weakfish length-frequencies were truncated during 1993-1998, while those for 1999 and 2000 c ontained considerably more weakfish greater than 380 mm TL. However, this trend reversed from 2001 to 2011, with far fewer large weakfish being encountered. All of the weakfish sampled in the 2011 pound ne $t$ s urvey were below the recreational size limit of 331 mm TL ( 13 inches) and the commercial size limit of 305 m m TL ( 12 inches). This trend ended in 2012, with 14 and 22 of 93 w eakfish above the recreational and commercial size limits, respectively.

In 2012, females accounted for $74 \%$ of fish sampled from the pound net survey ( $\mathrm{n}=52$ ). Female mean TL and mean weight were 289 mm TL and 250 g , respectively, while males averaged 282 mm TL and 212 g . In 2011, females averaged 242 mm TL and 147 g and accounted for $65 \%$ of fish sampled ( $\mathrm{n}=23$ ), while male mean length and weight were 233 mm TL and 127 g , respectively.

Total Maryland commercial weakfish harvest (Chesapeake B ay and Atlantic Ocean combined) in 2011 declined to 378 pounds, with the C hesapeake B ay portion decreasing from 40 pounds in 2010 to 24 pounds in 2011 (Figures 3 and 4). The 2011 total harvest was the lowest of the 82 year time series and was well be low Maryland's average of 620,020 pounds pe $r$ year. $M$ aryland $r$ ecreational a nglers ha rvested an estimated 237 weakfish $(\operatorname{PSE}=91)$ during 2011, with an estimated weight of 134 pounds $(\operatorname{PSE}=89.3 ;$ Figure 5$)$. The number of weakfish harvested by the recreational fishery in 2011 represented a 20 fold decrease compared to the 2010 estimate $(4,784)$, and was the lowest of the 1981-2011 time series. According to the MRIP estimates, Maryland anglers released 18,500 $(\mathrm{PSE}=46.8)$ weakfish in 2011, a more than 8 fold decrease from 2010 $(162,733, \operatorname{PSE}=46.8)$. Estimated recreational harvest decreased steadily from 475,348 fish in 2000 to ne ar zero in 2006, and recovered slightly in 2007 and 2010 before dropping back to ne ar z ero in 2011. Both the recreational ha rvest es timates and the reported commercial landings in 2010 and 2011 may have been affected by a regulation change that took place in April 2010. The new regulation reduced the bag limit from 3 fish to 1 fish per angler per day, and the commercial harvest was limited to a bycatch only fishery, with daily catch limits of 50 pounds in the Chesapeake Bay and 100 pounds in the Atlantic Ocean.

The reported harvest from Maryland charter boat captains has ranged from 2,042 to 75,154 weakfish from 1993 to 2011 (Figure 6), with a dramatic decline occurring in 2003. The reported charter boat harvest had the same trend as the reported commercial harvest $\left(\mathrm{R}^{2}=0.64, \mathrm{P}<0.001\right)$ and the statewide MRIP estimate $\left(\mathrm{R}^{2}=0.81, \mathrm{P}<0.001\right)$, but not the inland for hire only MRIP estimate. Of the 27,734 entries reported, only one was not included in this a nalysis since the CPA e xceeded 200. The 2011 geometric mean of 0.58 weakfish per angler was the third lowest mean of the time series (Figure 7). The geometric mean CPA has declined significantly from $1993-2011\left(\mathrm{R}^{2}=0.81, \mathrm{P}<\right.$ 0.001).

The 2012 weakfish j uvenile G M decreased after increasing for t hree s traight years, and was the $2^{\text {nd }}$ lowest va lue inthe 24 yeartime series (Figure 8). Weakfish juvenile a bundance generally i ncreased from 1989 to 1996 in P ocomoke and Tangier sounds, remaining at a relatively high level through 2001, but generally decreased from 2003 to 2008. This lack of recruitment may explain poor commercial and recreational harvest in recent years. The relatively low abundance of juvenile weakfish since 2003 is similar to that of the early 1990 's, but harvest continues to be exceptionally low, unlike the higher harvest in the early 1990's.

Weakfish otoliths were collected from 71 fish in 2012. Age samples from 2003 2005 w ere c omprised of $45 \%$ or m ore a ge t wo pl us w eakfish, a nd then dr amatically shifted to primarily age one fish from 2006-2011, with 0 to $30 \%$ age two plus fish and no age 3 fish from 2008 to 2011. Age structure expanded to include three year old weakfish in 2012, w ith $46 \%$ of sampled fish be ing a ge two plus, indicating a shift back toward slightly older weakfish (Table 6).

Mortality estimates for 2007 t hrough 2012 could not be calculated be cause of extremely low sample size, while instantaneous total mortality e stimates calculated for 2005 and 2006 were $\mathrm{Z}=1.44$ and $\mathrm{Z}=1.35$, respectively (Table 7). Maryland's lengthbased estimates were similar to the coastal assessment of $Z=1.4$ for cohorts since 1995 (Kahn et al. 2005).

The most recent weakfish Stock Assessment Workshop conducted by ASMFC in 2009 utilized various models to determine natural mortality (M), fishing mortality (F) and current bi omass (NFSC 2009). This assessment i ndicated weakfish bi omass w as extremely low; F was moderate and M was high and increasing (NFSC 2009). The stock was classified as depleted due to high M, not F. The stock assessment confirmed that the low com mercial and recreational weakfish harvest in Maryland, and low abundance in the sampling surveys, is directly related to a coast wide stock decline.

## Summer flounder

Summer flounder pound net survey mean lengths have varied widely from 20042012. Mean total lengths have ranged from the time series high of 374 mm TL in 2005 and 2010 to the time series low of 286 mm TL in 2006 (Table 4). The 2012 mean length of 338 mm TL decreased compared to 2011, it was the ninth lowest of the 20 year time series. This decrease is primarily attributed a greater proportion of juveniles, as indicated by the 1 ength f requency a nalysis be low. Relative stock densities in the 2012 onboard pound ne t s urvey indicated a decrease in the stock and qua lity categories with a corresponding large increase in the preferred category compared to 2011 (Table 8). The 2012 percentage of summer flounder in the preferred category and above was the highest in the 20 year time series. The length frequency distribution from the onboard sampling
was bimodal in 2012 peaking at the 130 to 150 and 410 to 430 mm TL length groups, representing an expansion in smaller sized fish (Figure 9). There was also an increase in the proportion of fish greater than or equal to the 356 mm TL minimum commercial size limit in 2012 ( $60 \%$ ) compared to 2011 ( $51 \%$ ). The number of summer flounder sampled in 2012 was the lowest of the 20 years surveyed (Table 4). Recreational size limits have been adjusted annually, but comparing the onboard pound net survey catches to the 2012 recreational size limit of 432 mm TL indicated a greater proportion of legal fish in the stock during 2012 (28\%) compared to 2011 ( $4 \%$ ).

Maryland's commercial summer flounder harvest totaled 144,580 pounds in 2011, the $5^{\text {th }}$ lowest in the 49 year t ime s eries (Figure 10). The 1 ong-term (1962-2011) commercial harvest average is 412,949 pounds. In recent years the commercial flounder fishery has be en $m$ anaged $b$ y quot $a, w$ ith va rying $r$ egulations a nd $s$ eason $c$ losures $t o$ ensure the quota was not exceeded. The majority of the Maryland commercial flounder harvest comes from the A tlantic O cean and co astal bays (Figures 10 and 11 ). The recreational harvest estimate of $15,347(\operatorname{PSE}=44.8)$ fish caught in 2011 ranked $31^{\text {st }}$ out of the 31 year time series, and declined $39 \%$ from the 2010 estimate of 25,215 ( $\mathrm{PSE}=$ 35.7) fish (Figure 12). The 2011 MRIP recreational release estimate of 472,536 ( $\mathrm{PSE}=$ 23.5) fish was the $18^{\text {th }}$ highest of the 1981-2011 time series, representing a drop back down to 1996 values (Figure 12). This is consistent with the RSD analysis and onboard length frequency distributions, that indicate a de crease in fish greater than the minimum recreational size limit in 2011. The recreational fishery has been subject to increasingly restrictive regulations in the past several years, which most likely reduced harvest rates.

Reported summer flounder charter boat harvest has been variable, but generally
increased to the time series high of 14,371 fish in 2010 from the 2003 low of 1,051 fish (Figure 13). The 2011 h arvest dipped three percent to 14,008 the second highest in the 19 year time series. Linear regression indicated no significant trend between the charter boat catch and the statewide MRIP estimate, the com mercial 1 andings or the for hire inland only MRIP estimate. This is not surprising, since the majority of the commercial harvest occurs in the A tlantic O cean, and the MRIP inland estimate includes both the coastal bays and the Chesapeake Bay, and the charter logs are all from the Chesapeake Bay. The geometric mean index did significantly decline $\left(\mathrm{R}^{2}=0.46, \mathrm{P}=0.0013\right)$ over the entire time series (Figure 14), but has been relatively stable for the past eight years.

A coast w ide stock a ssessment us ing the A ge $S$ tructured Assessment P rogram (ASAP) was conducted in 2008, and updated in 2011(NFSC 2008, Terceiro 2011). The assessment indicated that summer flounder recruitment along the Atlantic coast declined from a peak in 1983 to the time series low in 1988 (NFSC 2008). T he ASAP model estimated recruitment for 2009 at 60 million fish, above the long term mean of 43 million fish (NFSC 2008, T erceiro 2011). T he NMFS coastal as sessment found that F va ried from $\mathrm{F}=1.1$ to $\mathrm{F}=2.0$ from 1982 to 1996 , but has remained below 1.0 since 1996. F was estimated to be 0.22 in 2010, below the threshold, and the estimated 2010 S SB of 132.8 million pounds was slightly above the target level of 132.4 m illion pounds. The NMFS assessment con cluded that summer f lounder stocks were n ot ove rfished, overfishing was not occurring, and the rebuilding target has been met as of 2010 (NFSC 2008, Terceiro 2011).

## Bluefish

Bluefish sampled from t he onboard pound ne ts urvey a veraged 298 mm T L during 2012, an increase from the 2011 mean of 245 mm TL (Table 4). The 2012 mean length was below the 20 year t ime s eries m ean of 302 mm . Ninety-two percent of sampled bluefish were in the $\mathrm{RSD}_{\text {stock }}$ category and $7.9 \%$ were in the $\mathrm{RSD}_{\text {quality }}$ category (Table 9), indicating an increase in larger bluefish compared to 2010. The pound ne t survey length frequency distribution shifted to larger size bluefish in 2012, lengths were mostly distributed between the 190 to 370 mm TL groups with a minor peak in the 230 mm TL group (Figure 15).

The 2005-2007 pound net sampling indicated a small shift to a 1 arger grade of bluefish, although small bluefish still dominated the population. This trend reversed in 2008 through 2011 when larger bluefish became scarce. The 2012 length structure was similar to those of 2005 - 2007. Variable migration patterns into Chesapeake Bay may be responsible for these differences. Crecco (1996) reviewed bluefish angler catches and suggested that the bulk of the stock was di splaced of fshore. Lack of forage and interspecific competition with striped bass were possible reasons for this displacement.

Maryland bl uefish commercial ha rvest decreased by $33 \%$ in 2011 to 70,383 pounds, and remained below the 1929-2011 average of 171,408 pounds (Figure 16). The 2011 catch ranked $57^{\text {th }}$ in the 82 year time series. Total commercial landings fluctuated without $t$ rend from 42,662 to 157,436 pounds from 1993 - 2011 (Figure 16 ). The majority of Maryland's commercial bluefish harvest from 1972 through 1988 came from the C hesapeake Bay. However, Chesapeake B ay cat ches de clined after 1998 while Atlantic Ocean and coastal bay catches remained stable. Recreational harvest estimates
for bluefish were high through most of the 1980 's, but have fluctuated at a lower level since 1991 (Figure 17). The 2011 estimate of 259,286 ( $\mathrm{PSE}=25.6$ ) fish ha rvested decreased slightly compared to 2010 ( 272,764 fish, $\mathrm{PSE}=17.6$ ), and was well below the time series average of 852,601 fish. Estimated recreational releases increased by two and half fold in 2011 to $408,323(\mathrm{PSE}=27.5)$ compared to $2010(161,424$ fish, $\mathrm{PSE}=30.6)$, still lower then the time series mean of 541,923 fish (Figure 17).

Reported bluefish harvest from charter boat logs ranged from 27,667-134,828 fish per year from 1993 to 2011, 2011 harvest declined for the third consecutive year to 30,176 fish (Figure 18). Harvest from charter boat logs generally tracked with state wide MRIP estimates, but regression analysis indicated no s ignificant trend with recreational estimates or commercial la ndings. Two of the 70,182 entries were not used in indices calculations because of excessively high CPA's ( $>300$ ). The geometric mean catch per angler varied in a narrow range from 1993 to 2007, increased to the time series high in 2008, but then declined from 2009 to 2011 (Figure 19).

A st ock assessment of Atlantic c oast bl uefish utilizing the forward pr ojecting catch at a ge (ASAP) model was produced in 2010 , and revised in 2012 (Shepherd and Nieland 2010, NMFS 2012). The as sessment indicated that F was steady at a 1 ow rate since 2000. Recruitment estimated in the ASAP model has remained relatively constant since 2000 at around 22.5 million age- 0 bluefish, with the exception of a relatively large 2006 c ohort e stimated as 35.2 m illion fish. Recruitment during 2009-2011 was be low average (Shepherd and N ieland 2010 , N MFS 2012 ). The m odel indicated that overfishing is not occurring and that the stock is not overfished, but projected spawning stock biomass declines over the next few years due to poor recruitment.

## Atlantic croaker

Atlantic croaker mean length from the onboard pound net survey decreased for the third year to 274 mm TL , and was the third lowest value of the 20 year time series (Table 4). Gill net caught fish were also measured during onboard sampling for the first time in 2012, with a mean length of $296 \mathrm{~mm} \mathrm{TL}(\mathrm{n}=571)$ and a mean weight of $381 \mathrm{~g}(\mathrm{n}$ $=61$ ). Fifty percent of the Atlantic croaker from onboard pound net sampling in 2012 were in the $\mathrm{RSD}_{\text {preferred }}$ category, a decrease o ver 2011. All ot her R SD c ategories increased slightly in 2012 (Table 10 ). The onboard pound ne $t$ length f requency distribution for 2012 indicated an increase in the smaller croaker, but otherwise was very similar to the 2011 distribution, with the primary peak occurring in the 250 and 270 mm length groups (Figure 20). Onboard gill net length frequency peaked in the 270 and 290 mm 1 ength gr oups with $c$ atches dr opping of $q$ uickly $f$ or bot $h s$ maller a nd 1 arge $f$ ish (Figure 21). This could be an indication of net selectivity, or an artifact of the sample being from a single catch (one fisherman on one day).

Atlantic croaker sampled from gill nets in 2012 mean lengths and weights by sex were 295 mm TL and 375 g for females $(\mathrm{n}=47)$ and 308 mm TL and 400 g for males ( n $=14$ ). Gill net samples were $77 \%$ female and $23 \%$ male, but sample size was low, so these pe rcentages may not reflect the actual male to female composition of the gill net harvest. Pound ne ts amples w ere not randomly s elected, t herefore no s ex s pecific analysis was conducted.

During 2011, the Maryland Atlantic croaker total commercial harvest of 704,019 pounds (Chesapeake B ay and Atlantic O cean combined) increased $44 \%$ compared to 2010 (Figure 22). The 2011 harvest was still below the 1929-2011 average of 1,042,700
pounds. The 2011 recreational harvest estimate was 554,206 fish $(\mathrm{PSE}=22.3)$ a $51 \%$ decrease from 2010, the lowest value since 1993, and was below the 1981-2011 average of 756,175 fish (Figure 23). The 2011 recreational release es timate decreased $64 \%$ compared to 2010 (Figure 23), and was well below the 1981-2011 average of 1,241,139 fish.

Reported Atlantic c roaker ha rvest from charter boa ts r anged from 12 7,664$448,789 \mathrm{f}$ ish during the 19 year time p eriod (Figure 24). The c harter boat log book harvest trended with the statewide MRIP estimates $\left(\mathrm{R}^{2}=0.37, \mathrm{P}=0.0055\right)$, but not with the Chesapeake Bay commercial landings or for hire inland only MRIP estimates. The MRIP for hi re i nland only estimates did, ho wever, follow $t$ he $s$ ame $g$ eneral $t$ rend. Twelve of the 51,044 entries were not used to calculate the GM because of CPA values exceeding 200 fish. The geometric mean catch per angler increased significantly $\left(\mathrm{R}^{2}=\right.$ $0.44, \mathrm{P}=0.0021$ ) from 1993 to 2011, with relatively stable values prior to 2004 and generally i ncreased values since 2004 (Figure 25). Following three years of $s$ teadily increasing values, the 2011 GM of 4.66 fish per angler was a de crease from 2010, but was still above the long term mean.

Since 1989, the Atlantic croaker juvenile indices have varied without trend, with the highest values occurring in the late 1990s. This index increased to the third highest value of the 24 year time series for 2008, but fell sharply in 2009 and remained low through 2011 (Figure 26). The 2012 GM increased to 3.8 fish per tow, and was above the 24 year time series mean of 3.4 fish per tow, and was the $7^{\text {th }}$ highest value of the time series. Atlantic croaker recruitment $h$ as been linked to environmental factors including winter temperature in nursery areas (Lankford and Targett 2001, Hare and Able 2007);
prevailing winds, currents and hurricanes during spawning; and larval ingress (Montane and A ustin 2005, Norcross a nd A ustin 1986). Because of these strong e nvironmental influences, high spawning stock biomass may not result in good recruitment.

Ages de rived from pound netc aught Atlantic cr oaker otoliths in 2012 ranged from 0 to $8(\mathrm{n}=255$; Table 11). The number of Atlantic croaker sampled for length in $2012(\mathrm{n}=1,842)$ was applied to an age-length key for 2012 (Table 11). This application indicated that $34 \%$ of the fish were age four, $22 \%$ w ere age two, $22 \%$ w ere a ge three, $10 \%$ were age zero and $6 \%$ were age five. The remaining age groups each accounted for three percent or less of the fish sampled (Table 11). Atlantic croaker greater than six years ol d ha ve be come 1 ess a bundant in $r$ ecent years than int he $m$ id 2000 s . The contribution of strong year classes (1998, 2002, 2006 a nd 2008) to the catch c an also been seen in Table 11. The abundance of age zero fish in the pound net catch in 2012 appears toc orroborate the above av erage juvenile trawl inde x . Instantaneous t otal mortality in 2012 was $Z=0.80$, very similar to $2010(Z=0.81)$, and ended a trend of increasing mortality since the 1999-2012 time series low of 0.30 in 2006 (Table 7).

In 2010, the A SMFC A tlantic Croaker T echnical Committee completed a s tock assessment us ing a s tatistical cat ch at a ge m odel using d ata t hrough 2008 ( ASMFC 2010). The as sessment i ndicated decreasing $F$ and $r$ ising $S$ SB s ince $t$ he late 1980 ' s. Estimated values of F, SSB and biological reference points were too uncertain to be used to de termine s tock status. However, the ratio of F to $\mathrm{F}_{\text {msy }}$ (the F ne eded to pr oduce maximum s ustainable yield) was d eemed reliable and was used to determine $t$ hat overfishing is not occurring. It is not possible to be confident with regard to stock status, particularly a bi omass determination, until the di scards of A tlantic c roaker from the

South Atlantic shrimp trawl fishery can be adequately estimated and incorporated into the stock assessment (ASMFC 2010).

## Spot

Spot mean length from the onboard sampling decreased in 2012 to 179 mm TL $(n=1,508)$, the lowest value of the 18 year time series (Table 4). The onboard sampling length frequency distribution in 2012 shifted to smaller length fish (Figure 27). The 150 and 160 mm TL groups accounted for $64 \%$ of sampled spot. One jumbo spot ( $>254 \mathrm{~mm}$ TL) was pr esent in the 2012 onboard sampling accounting for less than $0.1 \%$ of the sample. Abundance of jumbo spot in the survey have been low for the past several years ( $0-3 \%$ of sample, 2005-2011). This followed good catches in the early part of the decade ( $10 \%$ in $2003,13 \%$ in 2004).

Commercial harvest in 2011 decreased slightly to 552,985 pounds (Figure 28), the $5^{\text {th }}$ highest catch of the 82 year time series. C ommercial harvest peaked in the 1950's with catches nearing 600,000 pounds. Harvest then fell sharply and remained low, except for a few spikes, rebounding to m oderate 1 evels from the m id 1980s through the late 2000s, and returning to near time series high values the past three years. Chesapeake Bay commercial harvest h ad be en fairly steady from 2003-2005 r anging from 66,865 t o 74,722 pounds before declining to 23,500 pounds in 2006. An unusually sharp increase in 2007 and 2009 through 2010 can be at tributed to a large increase in gill net harvest, which accounted for $95 \%$ of the 2007 s pot harvest ( 380,648 pounds), $90 \%$ of the 2009 harvest ( 467,595 pounds), $87 \%$ of the 2010 h arvest ( 507,091 pounds) and $61 \%$ in 2011 ( 388,533 pounds), compared to $43 \%$ of the 2006 harvest ( 16,420 pounds). The reported spot harvest, excluding gill net landings, for 2007 (19,703 pounds ) w as similar to the

2006 non -gill ne $t$ ha rvest of 21,354 pounds . In 2008 gill nets accounted for $48 \%$ of commercial harvest, with an increasing catch in non-gill net fisheries ( 62,934 pounds). The 2009 non-gill net harvest was similar to 2008 ( 52,556 pounds), but the 2011 non-gill net harvest increased and was primarily from fish pots (134,058 pounds, $24 \%$ of total harvest). This would seem to indicate the recent spike in gill net la ndings was due to increased effort directed at spot, likely triggered by market demand and/or the decreased availability of other more desirable species. The increase in fish pot harvest in 2011 is likely a result of charter fishermen with commercial lic enses' reporting s pot caught in pots to use as live bait.

Maryland recreational ha rvest estimates from the MRIP indicated that sot catches since 1981 ha ve been variable (Figure 29). Recreational ha rvest ranged from 300,000 fish in 1988 to $3,800,000 \mathrm{f}$ ish in 1986 and 2007, while the n umber released fluctuated from 200,000 in 1999 to 2,700,000 in 1986 (Figure 29). The 2011 recreational harvest estimate $(912,704$ fish; $\mathrm{PSE}=19)$ decreased $22 \%$ compared to 2010 , dropping well be low the time series mean estimate of $1,630,015$ fish, and marked the $8^{\text {th }}$ lowest value of the 31 year time series. The release estimate of 296,513 fish ( $\mathrm{PSE}=18.6$ ) decreased $74 \%$ compared to 2010 , and was the $4^{\text {th }}$ lowest estimate of the 31 year time series (Figure 29).

Reported s pot charter b oat logbook harvest from 1993 to 2010 ranged from 217,052 to 848,492 fish per year (Figure 30). The 2011 reported harvest was the fourth lowest of the 19 year time series and follows the lowest value in 2010. The charter boat log book harvest did not significantly trend with the MRIP for hire inland only estimates, the C hesapeake Bay commercial 1 andings or statewide MRIP estimates. T his is not
surprising, since charter boat captains sometimes have clients catch spot to use as bait for larger predatory species. MRIP surveys may not accurately account for spot used as bait, while the com mercial ha rvest $t$ ends to be $m$ ore incidental some years and directed in others. Twenty-four of the 44,056 charter $\log$ book entries were not utilized because of greatly i nflated C PA va lues ( $>300$ ). The geometric mean CPA was hi ghest in 1995, stable at a relatively low level from 1999 - 2002 and generally increased from 2002 2007. T he C PA ha s remained a bove a verage from $2008-2011 \mathrm{w}$ ith t he e xception of 2010, which had the second lowest value in the 18 year time series (Figure 31).

Spot juvenile trawl index values from 1989-2012 were quite variable (Figure 32). The 2010 GM value of 104.5 spot per tow was the highest value of the time series, but declined to the second lowest value of the 23 year time series in 2011. The juvenile index increased to 16.4 spot per tow in 2012, just be low the time series mean of 18.3 (Figure 32).

In 2012 age one spot accounted for $60 \%$ of the sample with $39 \%$ being age zero and the remaining $1 \%$ being age two (Table 12). Age one spot dominated the pound net catch from 2007 to 2011, accounting for $75 \%$ to $99 \%$ of sampled fish. During this same time period, age zero and age two fish were present every year, with age zero accounting for $0.4 \%$ to $24.3 \%$ of sampled spot and age two accounting for $0.2 \%$ to $3.3 \%$. Two fish, sampled for 1 ength onl $y$, in bot h 2007 a nd 2011 w ere in length groups four to six centimeters 1 arger t han available M aryland DNR s amples. In bot h cases a ge 1 ength information from spot aged by VMRC were used. These were the only fish in the three and four year old age classes.

In ar elatively short-lived species such as spot, population dynamics and length structure will be greatly influenced by recruitment events. The shift in length frequency, decrease in mean size a nd reduction in percent jumbo s pot observed in 2005 t hrough 2010 could be i ndicative of growth overfishing. Reduced recreational ha rvest and reduced proportion of a ge on e s pot in 2012 a re likely due to the very poor 2011 year class. Commercial harvest may not have been as affected since there app eared to be an increase of spot caught for live bait, many of which may have been age zero. Virginia and N orth C arolina voiced concern ove r decreasing s pot ha rvests in their waters, and ASMFC's spot Plan Review Team continues to monitor catch and biological information to determine if additional management action is necessary. Given the popularity of spot as a recreational finfish, other indicators of $s$ tock status should be de veloped to e nsure production is exceeding harvest and losses due to natural mortality. No stock assessment has been completed for spot; primarily do to lack of necessary data.

## Red Drum

Red dr um ha ve be en e ncountered s poradically $t$ hrough $t$ he 20 years of $t$ he onboard pound ne t s urvey, with 458 be ing m easured in $2012($ Table 4). The number of red drum sampled from the onboard sampling also spiked in 2002 at 177 fish (Table 4); however, none were measured from 1993 to 1998, 2001 or in 2009 and 2010. Red dr um m ean 1 ength from $t$ he 2012 onboard $s$ ampling w as 318 mm T L , indicating the fish were pr imarily juve niles ( most like ly age one fish). Three of the sampled red drum were over the maximum recreational and commercial size limits of 27 and 25 inches respectively, and the remaining 455 were below the 18 inch minimum size limit in place for both sectors.

Maryland commercial fishermen reported harvesting no red drum in 20 11, and only 19 pounds in 2010 (Figure 33). Average harvest from 2004 to 2011 was 27 pounds per year, compared to 700 pounds per year from 1998 to 2003. However, lower harvest since 2003 may not reflect an actual decline in abundance, since more liberal regulations were in effect during previous years. P rior to the regulation change to an $18-25$ inch slot limit with a 5 fish bag limit in 2003, Maryland commercial fishermen were allowed to harvest one fish over 27 inches per day. Most of these fish were much larger than 27 inches which consequently led to higher harvest values by weight.

The MRIP estimated that r ecreational f ishermen did not ha rvest or cat ch and release any red dr um in 2011 (Figure 34). Recreational harvest estimates have b een extremely variable ranging from zero ( 23 of the 31 years in the 1981-2011 time series) to 12,804 fish (in 1986, PSE=67.4). Peak number of red drum releases occurred in 2002 at $18,412 \mathrm{f}$ ish (Figure 34 ). Anecdotal information regarding 2012 r ecreational c atches indicate $j$ uvenile $r$ ed dr um were pl entiful t hroughout m uch of M aryland's p ortion of Chesapeake Bay and its tributaries. Catches were commonly reported on fishing message boards on $t$ he internet and in local news papers. N early all of the reports were of sublegal fish in the 10 to 14 inch range, indicating a strong 2011 year class.

Maryland charter boat captains reported harvesting red drum in every year from 1993-2010, except for 1996. C atches were low for all years, ranging from zero to 99 fish, with a mean of 20 red drum per year (Figure 35). The low reported catch indicated red drum were available in Maryland's portion of Chesapeake Bay, but the low numbers confirm the species limited availability to recreational anglers, as indicated by the annual MRIP estimates. N o annual indices w ere generated because of 1 ow s ample sizes.

Maryland is near the northern limit for red drum and catches of legal size fish would be expected to increase if the stock expands in response to the current A tlantic coast stock recovery plan (ASMFC 2002).

## Black Drum

Black dr um are o nly o ccasionally encountered dur ing the MD D NR onboard pound net sampling, with only one being sampled in 2012 (Table 4). Lengths throughout the time series have ranged from 244 to 1330 mm TL. The one fish measured in 2012 was 997 mm TL. Commercial harvest of black drum was banned for Maryland's portion of Chesapeake B ay in 1999 , but some fish are still harvested along the A tlantic c oast (Figure 36). Recreational harvest and release estimates from 1981 to 2011 have been variable, ranging from zero to over 13,000 fish in 1983 (Figure 3 7). In 20 11, MRIP estimated no black drum were h arvested and $7,971(\mathrm{PSE}=78.8)$ were released by recreational anglers. The harvest e stimates are somewhat $t$ enuous, s ince $t$ he MRIP survey is unlikely to accurately represent a small, short lived seasonal fishery such as the black drum fishery in Maryland.

Examination of the charter boat logs revealed black drum w ere harvested in all years of the 1993-2011 time series, with a mean catch of 407 fish per year (range $=104-$ 905; Figure 38). Charter harvest had no significant trend to either the state wide or inland for hire only MRIP estimates. The geometric mean significantly declined $\left(\mathrm{R}^{2}=0.68, \mathrm{P}<\right.$ 0.001 ) throughout the time series, but did increase slightly in 2009 a nd leveled off in 2010 and 2011 (Figure 39).

## Spanish Mackerel

Spanish mackerel have been measured for FL, TL or bothineach year of the onboard pound n et s ampling. Since 2001 , however, only F L has been taken, to be consistent with data collected by other state and federal agencies. During this time period FL from the onboard sampling has ranged from $208-681 \mathrm{~mm}$. One hundred seven Spanish mackerel were encountered in 2012, with a mean length of 318 mm FL (Table 4). The num ber of $m$ ackerel $m$ easured has been low for most years $w$ ith the largest samples occurring from 2005-2007 (Table 4).

The 2011 commercial harvest of S panish mackerel in Maryland was 5,054 pounds, a $33 \%$ increase from 2010 (3,806 pounds; Figure 40), a nd be low the 1965 t o 2011 mean of 6,359 pounds per year. Commercial harvest was very low from 1965 1986 w ith no c atches greater t han 3,600 pound s i ncluding six years of z ero ha rvest. Commercial harvest has been somewhat more stable since 1987 w ith a peak of 62,688 pounds in 1991. Since 1996, the majority of Spanish mackerel harvest has come from Chesapeake B ay, but during the 1987 - 1995 t ime pe riod Atlantic O cean catches dominated. Recreational harvest estimates peaked in the early to mid 1990's with three years of a pproximately 40,000 fish harvested (Figure 41). This followed a pe riod of seven out of ten annual estimates with zero fish captured. Harvest estimates for 1998 2011 were variable, ranging from $0-20,049$ fish with an average of 8,686 fish taken. In 2011, an estimated $10,544(\mathrm{PSE}=52.6)$ Spanish mackerel were harvested, nearly double the 2010 estimate of 5,580 fish $(\mathrm{PSE}=55.5$, Figure 39). Due to the high PSE values, these estimates are considered tenuous.

Spanish mackerel charter boat harvest from 1993 to 2010 ranged from 563 $10,653 \mathrm{f}$ ish pe r year (Figure 42). The c harter boa t l og book ha rvest di dt rend significantly with the MRIP for hire inland only estimates $\left(\mathrm{R}^{2}=0.58, \mathrm{P}<0.01\right)$ and the statewide MRIP estimates $\left(\mathrm{R}^{2}=0.50, \mathrm{P}<0.01\right)$, but not the Chesapeake Bay commercial landings. The geometric mean CPA varied without trend (Figure 43). It would appear that S panish mackerel a re providing a small and somewhat consistent opportunity for recreational anglers in Chesapeake Bay.

## Spotted Seatrout

Spotted seatrout are rarely encountered during sampling. Eight were measured from the onboard sampling in 2012 with a mean length of 436 mm TL (Table 4). Commercial harvest of spotted seatrout in Maryland averaged 44,921 pounds from 19441954, zero pounds from 1955 - 1990 and 6,497 pounds from 1991-2011 (Figure 44). Reported 2011 harvest was 585 pounds, well below the 1991-2011 mean. Recreational harvest estimates indicated a modest fishery during the mid 1980's and mid 1990's. However, catches became very low to nonexistent from the late 1990's to 2005, with a slight upswing in 2006 before returning to zero in 2007 and 2008. Catches increased in 2009 to 11,680 fish, the highest value since 1998 (Figure 45). The 2010 estimate decreased to 3,146 $(\mathrm{PSE}=71)$ and was similar in $2011(3,058$ fish $\mathrm{PSE}=66)$, but the high PSE values from 2009 to 2011 indicate the MRIP survey does not provide reliable estimates for this species in Maryland.

Spotted seatrout harvest from 2011 charter boats was 1,762 fish. Reported harvest ranged from 224-20,030 fish per year and averaged 4,187 fish per year for the 15 year time series (Figure 46). No harvest was reported from 1993 to 1996, but it is not
clear if spotted seatrout were not reported at that time or none were captured. The charter boat log book harvest did not trend significantly with the MRIP for hire inland only estimates, the statewide MRIP estimates or the Chesapeake Bay commercial landings. The geometric mean CPA varied without significant trend, but has declined the past three years (Figure 47). The recreational spotted seatrout fishery in Chesapeake Bay is prosecuted by a small group of anglers that are likely under-represented in the MRIP estimation design. This is supported by the 2007 and 2008 reported charter harvest values that approximated the time series mean coinciding with zero value estimates by the MRIP.

## Atlantic Menhaden

Mean length for Atlantic menhaden sampled from commercial pound nets in 2012 was 243 mm FL, near the mean of 245 mm FL for the 2004 to 2012 time series (Table 4). Menhaden length frequencies from onboa rd s ampling for 2006 a nd 2007 were ve ry similar and robust compared to 2005 . However, the 2008 length frequency distribution was more con centrated around the mean, with a lower proportion of smaller and larger fish $t$ han $t$ he pr evious $t$ wo years. In $2009 t$ he di stribution e xpanded, but $w$ as $s$ till dominated by larger fish (Figure 48). The 2010 and 2011 length distribution indicated a shift to s maller fish, a nd a m ore even di stribution of lengths. The 2012 di stribution returned to a more truncated distribution similar to 2008 , with $40 \%$ of sampled fish in the 230 mm FL size group.

Atlantic menhaden scale s amples were taken from 375 fish in 2012, but a ges could onl ybe a ssigned to 355 fish (Table 13). After appl ying the annual 1 ength frequencies to the corresponding age length keys, age one was the dominate year-class in

2010 a nd 2011 , accounting f or $43 \%$ and $38 \%$ of pound $n$ et c aught m enhaden, respectively (Table 13). In 2012 a ge t wo menhaden a ccounted for $57 \%$ of pound ne t caught menhaden and age seven fish were present for the first time since aging began in 2005. Menhaden greater then age four made up $2 \%$ to $4.5 \%$ of the population form 2005 to 2012.

Atlantic menhaden commercial harvest in Maryland increased from 7,000 pounds in 1935 to over 8 m illion pounds in 1965 (Figure 49). Commercial ha rvest remained above 3 m illion pounds until 1990 w hen harvest dropped to 1.7 m illion pounds, slowly increased, a nd s piked in 2005 t o a r ecord hi gh of 12.6 m illion pounds. A verage commercial harvest from $1935-2011 \mathrm{w}$ as 4.1 million pounds. The 2011 commercial harvest decreased for the fourth straight year, but was still the $16^{\text {th }}$ highest of the 76 year time series ( 6.9 million pounds), with $95 \%$ of harvest from the Chesapeake Bay (Figure 49).

An update of the ASMFC Atlantic menhaden stock assessment was conducted in 2012 using da ta through 2011 ( ASMFC 2012). T he a ssessment i ndicated that recruitment was generally low and population fecundity declined since the late 1990s. Fishing m ortality i ncreased i n 2010 a nd 2011 a nd t he popul ation i s c urrently experiencing overfishing when compared to the population benchmarks. A mendment 2 of the ASMFC Fisheries Management Plan for Atlantic menhaden is being finalized and will require reductions in harvest to end overfishing and increase the abundance of this important prey species.

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Figure 49. Maryland commercial Atlantic menhaden harvest by area, 1935-2011.

Table 1. Minimum lengths (mm TL) for relative stock density categories.

| Species | Stock | Quality | Preferred | Memorable | Trophy |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Weakfish | 205 | 340 | 420 | 555 | 705 |
| Summer <br> Flounder | 180 | 320 | 400 | 552 | 670 |
| Bluefish | 240 | 430 | 540 | 705 | 885 |
| Atlantic <br> croaker | 125 | 185 | 255 | 305 | 390 |

Table 2. Areas sampled, number of sampling trips, mean water temperature and mean salinity by month for 2012.

| Area | Month | Number of <br> Samples | Mean <br> Water <br> Temp. C | Mean <br> Salinity <br> (ppt) |
| :---: | :---: | :---: | :---: | :---: |
| Point Lookout | May | 1 | 20.5 | 13.2 |
| Central Bay | May | 1 | 24.5 | 12.3 |
| East Bay | May | 1 | 24.3 | 9.2 |
| Point Lookout | June | 2 | 22.7 | 10.2 |
| Central Bay | June | 2 | 24.2 | 9.8 |
| East Bay | June | 1 | 23.6 | 8.1 |
| Point Lookout | July | 2 | 28.0 | 14.2 |
| Central Bay | July | 2 | 27.6 | 12.6 |
| West Bay | July | 3 | 27.3 | 12.4 |
| Point Lookout | August | 2 | 27.0 | 16.2 |
| Central Bay | August | 1 | 27.1 | 15.3 |
| East Bay | August | 1 | 27.1 | 15.1 |
| West Bay | August | 2 | 27.6 | 15.7 |
| Point Lookout | September | 1 | 25.0 | 16.7 |
| Central Bay | September | 1 | 27.1 | 15.1 |
| East Bay | September | 1 | 27.2 | 15.1 |
| West Bay | September | 2 | 27.0 | 15.3 |

Table 3. List of non-target species observed during the 2012 onboard pound net survey.

| Common Name | Scientific Name |
| :--- | :--- |
|  |  |
| American shad | Alosa sapidissima |
| Atlantic cutlassfish | Trichiurus lepturus |
| Atlantic herring | Clupea harengus |
| Butterfish | Peprilus triacanthus |
| Common carp | Cyprinus carpio |
| Cownose ray | Rhinoptera bonasus |
| Crevalle jack | Caranx hippos |
| Florida pompano | Trachinotus carolinus |
| Gizzard shad | Dorosoma cepedianum |
| Harvestfish | Peprilus alepidotus |
| Hogchoker | Trinectes maculates |
| Northern kingfish | Menticirrhus saxatilis |
| Northern puffer | Sphoeroides maculatus |
| Northern searobin | Prionotus carolinus |
| Oyster toadfish | Opsanus tau |
| Silver perch | Bairdiella chrysoura |
| Southern kingfish | Menticirrhus americanus |
| Striped bass | Morone saxatilis |
| Striped burrfish | Chilomycterus schoepfi |
| White perch | Morone americana |

T able 4. Mean length ( mm TL, unless otherwise noted), standard deviation, and sample size of summer migrant fishes from Chesapeake Bay onboard pound net sampling, 1993-2012.

|  | 1993 | 1994 | 1995 | 1996 | 1997\| | 1998 | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Weakfish |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length | 276 | 291 | 306 | 293 | 297 | 337 | 334 | 361 | 334 | 325 | 324 | 273 | 278 | 290 | 275 | 276 | 262 | 253 | 236 | 284 |
| std. dev. | 46 | 50 | 54 | 54 | 39 | 37 | 53 | 83 | 66 | 65 | 68 | 32 | 39 | 30 | 42 | 52 | 22 | 24 | 24 | 48 |
| n | 435 | 642 | 565 | 1431 | 755 | 1234 | 851 | 333 | 76 | 196 | 129 | 326 | 304 | 62 | 61 | 42 | 23 | 47 | 26 | 93 |
| Summer flounder |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length | 347 | 309 | 297 | 335 | 295 | 339 | 325 | 347 | 358 | 324 | 353 | 327 | 374 | 286 | 341 | 347 | 368 | 374 | 359 | 338 |
| std. dev. | 58 | 104 | 62 | 65 | 91 | 53 | 63 | 46 | 50 | 93 | 56 | 101 | 76 | 92 | 66 | 72 | 64 | 84 | 67 | 130 |
| n | 209 | 845 | 1669 | 930 | 818 | 1301 | 1285 | 1565 | 854 | 486 | 759 | 577 | 499 | 1274 | 1056 | 982 | 277 | 197 | 213 | 161 |
| Bluefish |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length | 312 | 316 | 323 | 307 | 330 | 343 | 306 | 303 | 307 | 293 | 320 | 251 | 325 | 311 | 318 | 260 | 265 | 297 | 245 | 298 |
| std. dev. | 75 | 55 | 54 | 50 | 74 | 79 | 65 | 40 | 41 | 45 | 58 | 60 | 92 | 71 | 70 | 41 | 43 | 60 | 48 | 77 |
| n | 45 | 621 | 912 | 619 | 339 | 378 | 288 | 398 | 406 | 592 | 223 | 581 | 841 | 1422 | 1509 | 2676 | 1181 | 493 | 290 | 877 |
| Atlantic croaker |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length | 233 | 259 | 286 | 294 | 301 | 310 | 296 | 302 | 317 | 279 | 287 | 311 | 317 | 304 | 307 | 298 | 320 | 295 | 281 | 274 |
| std. dev. | 35 | 34 | 42 | 31 | 39 | 40 | 54 | 45 | 37 | 73 | 55 | 43 | 48 | 66 | 54 | 62 | 50 | 34 | 31 | 42 |
| n | 471 | 1081 | 974 | 2190 | 1450 | 1057 | 1399 | 2209 | 733 | 771 | 3352 | 1653 | 2398 | 1295 | 2963 | 1532 | 91 | 1970 | 1764 | 1842 |
| Spot |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length | 184 | 207 | 206 | 235 | 190 | 230 | 213 | 230 | 239 | 184 | 216 | 208 | 197 | 191 | 208 | 198 | 185 | 201 | 193 | 179 |
| std. dev. | 28 | 21 | 28 | 28 | 35 | 16 | 25 | 21 | 33 | 36 | 30 | 36 | 37 | 29 | 23 | 21 | 21 | 22 | 18 | 24 |
| n | 309 | 451 | 158 | 275 | 924 | 60 | 572 | 510 | 126 | 681 | 1354 | 882 | 2818 | 2195 | 519 | 1195 | 33 | 51 | 582 | 1508 |
| Spotted Seatrout |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length |  | 448 | 452 |  |  | 541 | 460 |  |  |  |  |  |  |  | 414 | 464 | 262 |  | 361 | 436 |
| std. dev. |  | 86 | 42 |  |  |  | 134 |  |  |  |  |  |  |  | 43 | 72 | 22 |  | 142 | 112 |
| n | 0 | 4 | 6 | 0 | 0 | 1 | 2 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3 | 10 | 23 | 0 | 4 | 8 |
| Black Drum |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length |  | 1106 | 741 | 353 |  | 1074 |  |  |  | 435 | 475 | 780 | 1130 | 1031 | 1144 | 875 | 1147 | 1061 | 978 | 997 |
| std. dev. |  | 175 | 454 | 20 |  | 182 |  |  |  | 190 | 20 | 212 |  | 228 | 95 | 238 | 84 | 345 | 188 |  |
| n | 0 | 2 | 3 | 2 | 0 | 12 | 0 | 0 | 0 | 7 | 4 | 44 | 1 | 8 | 9 | 5 | 13 | 3 | 3 | 1 |

Table 4. Continued.

|  | 1993 | 1994 | 1995 | 1996 | 1997\| | 1998 | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Red Drum |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length |  |  |  |  |  | 302 | 332 | 648 |  | 316 | 506 | 647 | 353 | 366 | 658 | 361 |  |  | 678 | 318 |
| std. dev. |  |  |  |  |  |  | 71 |  |  | 44 |  | 468 |  | 21 | 40 | 57 |  |  | 18 | 71 |
| n | 0 | 0 | 0 | 0 | 0 | 1 | 16 | 1 | 0 | 177 | 1 | 2 | 1 | 16 | 2 | 21 | 0 | 0 | 2 | 458 |
| Spanish Mackerel (Total Length) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length | 261 | 391 | 487 | 481 | 520 | 418 | 468 | 455 |  |  |  |  |  |  |  |  |  |  |  |  |
| std. dev. | 114 | 55 | 38 | 55 |  | 45 | 82 | 66 |  |  |  |  |  |  |  |  |  |  |  |  |
| n | 3 | 78 | 39 | 27 | 1 | 4 | 45 | 35 |  |  |  |  |  |  |  |  |  |  |  |  |
| Spanish Mackerel (Fork Length) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length |  |  | 418 | 401 | 437 | 379 |  | 386 | 406 | 422 | 405 | 391 | 422 | 439 | 436 | 407 | 418 |  |  | 393 |
| std. dev. |  |  | 34 | 62 |  |  |  | 34 | 34 | 81 | 63 | 95 | 33 | 35 | 51 | 59 | 53 |  |  | 74 |
| n |  |  | 44 | 27 | 1 | 1 |  | 49 | 19 | 20 | 11 | 8 | 373 | 445 | 158 | 18 | 7 | 0 | 0 | 107 |
| Menhaden (Fork Length) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mean length |  |  |  |  |  |  |  |  |  |  |  | 262 | 282 | 238 | 243 | 246 | 245 | 232 | 213 | 243 |
| std. dev. |  |  |  |  |  |  |  |  |  |  |  | 28 | 36 | 42 | 41 | 29 | 40 | 36 | 39 | 25 |
| n |  |  |  |  |  |  |  |  |  |  |  | 213 | 1052 | 826 | 854 | 826 | 366 | 836 | 773 | 755 |

Table 5. Relative stock density of weakfish from Chesapeake Bay summer onboard pound net survey, 1993-2012.

| Year | Stock | Quality | Preferred | Memorable | Trophy |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 89 | 10 | 1 | $<1$ |  |
| 1994 | 90 | 9 | 1 |  | $<1$ |
| 1995 | 74 | 23 | 3 |  |  |
| 1996 | 77 | 22 | 1 |  |  |
| 1997 | 90 | 9 | 1 |  |  |
| 1998 | 58 | 39 | 2 | $<1$ |  |
| 1999 | 61 | 33 | 5 | $<1$ |  |
| 2000 | 48 | 29 | 20 | 2 |  |
| 2001 | 58 | 35 | 5 | 1 |  |
| 2002 | 73 | 18 | 8 |  | $<1$ |
| 2003 | 67 | 30 | 2 | $<1$ |  |
| 2004 | 96 | 3 | 1 |  |  |
| 2005 | 94 | 5 | 1 |  |  |
| 2006 | 95 | 5 |  |  |  |
| 2007 | 94 | 3 | 3 |  |  |
| 2008 | 90 | 5 | 5 |  |  |
| 2009 | 100 |  |  |  |  |
| 2010 | 100 |  |  |  |  |
| 2011 | 100 |  |  |  |  |
| 2012 | 87 | 11 | 2 |  |  |

Table 6. Percentage of weakfish by age and year, number of age samples and number of length samples by year, using pound net length and age data 2003-2012.

| Year | Age 1 | Age 2 | Age 3 | Age 4 | \# of Ages | \# of Lengths |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| 2003 | 8.81 | 72.57 | 15.69 | 2.94 | 48 | 129 |
| 2004 | 55.90 | 39.20 | 4.90 |  | 59 | 326 |
| 2005 | 39.80 | 55.20 | 4.80 | 0.30 | 109 | 304 |
| 2006 | 70.10 | 22.20 | 7.60 | 0.10 | 62 | 62 |
| 2007 | 67.80 | 24.20 | 7.90 | 0.10 | 61 | 61 |
| 2008 | 85.71 | 7.14 | 7.14 |  | 41 | 42 |
| 2009 | 77.27 | 22.73 |  |  | 22 | 22 |
| 2010 | 100.00 |  |  |  | 45 | 47 |
| 2011 | 80.77 | 15.38 |  |  | 26 | 27 |
| 2012 | 54.18 | 42.34 | 3.47 |  | 71 | 93 |

Table 7. Weakfish and Atlantic croaker instantaneous total mortality rate estimates ( Z ) from Chesapeake Bay pound net data, 1999-2012.

| Species | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Weakfish | 0.74 | 0.4 | 0.62 | 0.58 | 0.73 | 1.29 | 1.44 | 1.35 | $*$ | $*$ | $*$ | $*$ | $*$ | $*$ |
| Atlantic croaker | 0.45 | 0.46 | 0.36 | 0.36 | 0.52 | 0.42 | 0.35 | 0.30 | 0.37 | 0.37 | 0.52 | 0.67 | 0.81 | 0.80 |

* Insufficient data to calculate 2007-2012 weakfish estimates.

Table 8. Relative stock density of summer flounder from Chesapeake Bay summer onboard pound net survey, 1993-2012.

| Year | Stock | Quality | Preferred | Memorable | Trophy |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 29 | 56 | 16 |  |  |
| 1994 | 24 | 56 | 20 | $<1$ |  |
| 1995 | 68 | 25 | 6 | 1 |  |
| 1996 | 25 | 61 | 13 | 1 |  |
| 1997 | 47 | 39 | 14 |  |  |
| 1998 | 30 | 57 | 12 | $<1$ |  |
| 1999 | 42 | 50 | 8 | $<1$ |  |
| 2000 | 22 | 66 | 12 | $<1$ |  |
| 2001 | 20 | 61 | 19 | $<1$ |  |
| 2002 | 41 | 35 | 24 | $<1$ |  |
| 2003 | 21 | 63 | 15 | $<1$ |  |
| 2004 | 23 | 55 | 21 | 1 |  |
| 2005 | 20 | 46 | 33 | 1 |  |
| 2006 | 57 | 29 | 14 | $<1$ |  |
| 2007 | 40 | 44 | 16 | $<1$ |  |
| 2008 | 31 | 47 | 21 | 1 |  |
| 2009 | 24 | 43 | 32 | $<1$ |  |
| 2010 | 29 | 35 | 34 | 3 |  |
| 2011 | 28 | 47 | 24 | 1 |  |
| 2012 | 19 | 25 | 55 | 1 |  |

Table 9. Relative stock density of bluefish from Chesapeake Bay summer onboard pound net survey, 1993-2012.

| Year | Stock |  |  |  |  |  | Quality |  | Preferred | Memorable ${ }^{\text {Trophy }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 90 | 10 |  |  |  |  |  |  |  |  |  |
| 1994 | 97 | 3 |  |  |  |  |  |  |  |  |  |
| 1995 | 98 | 2 |  |  |  |  |  |  |  |  |  |
| 1996 | 97 | 3 |  |  |  |  |  |  |  |  |  |
| 1997 | 96 | 4 |  |  | $<1$ |  |  |  |  |  |  |
| 1998 | 89 | 6 | 4 |  |  |  |  |  |  |  |  |
| 1999 | 92 | 8 | $<1$ |  |  |  |  |  |  |  |  |
| 2000 | 99 | 1 |  |  |  |  |  |  |  |  |  |
| 2001 | 98 | 2 |  |  |  |  |  |  |  |  |  |
| 2002 | 100 | $<1$ |  |  |  |  |  |  |  |  |  |
| 2003 | 96 | 4 |  |  |  |  |  |  |  |  |  |
| 2004 | 99 | 1 |  |  |  |  |  |  |  |  |  |
| 2005 | 79 | 20 | 1 |  |  |  |  |  |  |  |  |
| 2006 | 95 | 5 | $<1$ |  |  |  |  |  |  |  |  |
| 2007 | 94 | 3 | 3 |  |  |  |  |  |  |  |  |
| 2008 | 99 | 1 |  |  |  |  |  |  |  |  |  |
| 2009 | 100 | $<1$ |  | $<1$ |  |  |  |  |  |  |  |
| 2010 | 98 | 2 | $<1$ |  |  |  |  |  |  |  |  |
| 2011 | 100 |  |  |  |  |  |  |  |  |  |  |
| 2012 | 92 | 8 | $<1$ |  |  |  |  |  |  |  |  |

Table 10. Relative stock density of Atlantic croaker from Chesapeake Bay summer onboard pound net survey, 1993-2012.

| Year | Stock |  | Quality | Preferred | Memorable |  | Trophy |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | 6 | 72 | 19 | 2 |  |  |  |
| 1994 | $<1$ | 48 | 42 | 9 | $<1$ |  |  |
| 1995 | 1 | 21 | 48 | 28 | 2 |  |  |
| 1996 | 0 | 4 | 66 | 29 | 1 |  |  |
| 1997 | 7 | 9 | 32 | 52 | 1 |  |  |
| 1998 | 0 | 7 | 42 | 48 | 3 |  |  |
| 1999 | $<1$ | 28 | 25 | 42 | 4 |  |  |
| 2000 | 0 | 11 | 49 | 35 | 5 |  |  |
| 2001 | 0 | 2 | 38 | 56 | 4 |  |  |
| 2002 | 19 | 14 | 17 | 47 | 2 |  |  |
| 2003 | $<1$ | 43 | 17 | 36 | 3 |  |  |
| 2004 | $<1$ | 3 | 52 | 39 | 5 |  |  |
| 2005 | $<1$ | 11 | 26 | 55 | 7 |  |  |
| 2006 | 1 | 24 | 16 | 51 | 8 |  |  |
| 2007 | 0 | 17 | 37 | 37 | 9 |  |  |
| 2008 | 6 | 21 | 25 | 41 | 6 |  |  |
| 2009 | 0 | 9 | 30 | 52 | 10 |  |  |
| 2010 | 0 | 10 | 53 | 36 | 1 |  |  |
| 2011 | 0 | 18 | 63 | 19 | $<1$ |  |  |
| 2012 | 3 | 25 | 50 | 21 | 1 |  |  |

Table 11. Percentage of Atlantic croaker by age and year, number of age samples and number of length samples by year, using pound net length and age data, 1999-2012.

| Year | Age 0 | Age 1 | Age 2 | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 | Age 8 | Age 9 | Age 10 | Age 11 | Age 12 | Age 13 | \# Aged | \# Measured |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1999 | 0.0 | 34.0 | 22.5 | 3.3 | 9.4 | 4.2 | 16.0 | 6.0 | 4.2 | 0.4 |  |  |  |  | 180 | 1,399 |
| 2000 | 0.0 | 10.1 | 42.5 | 25.1 | 1.0 | 1.4 | 4.9 | 7.4 | 5.3 | 2.2 |  |  |  |  | 145 | 2,209 |
| 2001 | No Data |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 2002 | 18.4 | 4.0 | 10.1 | 8.9 | 29.4 | 24.0 | 1.0 | 0.0 | 3.0 | 0.5 | 0.6 |  |  |  | 66 | 771 |
| 2003 | 0.0 | 15.2 | 38.6 | 1.3 | 12.2 | 26.6 | 3.8 | 0.1 | 0.2 | 0.1 | 0.7 | 0.3 | 1.0 |  | 129 | 3,352 |
| 2004 | 0.0 | 0.6 | 54.9 | 5.0 | 5.4 | 6.9 | 23.3 | 3.1 | 0.0 | 0.2 | 0.0 | 0.6 |  |  | 161 | 1,653 |
| 2005 | 0.0 | 10.1 | 4.8 | 51.5 | 7.6 | 1.5 | 7.3 | 11.4 | 5.6 | 0.0 | 0.1 | 0.1 |  |  | 190 | 2,398 |
| 2006 | 16.7 | 6.3 | 18.1 | 4.8 | 36.8 | 2.3 | 3.2 | 5.0 | 5.2 | 1.8 | 0.0 | 0.0 | 0.0 | 0.1 | 253 | 1,295 |
| 2007 | 0.0 | 11.2 | 14.4 | 30.0 | 8.8 | 27.0 | 1.3 | 1.1 | 1.6 | 3.3 | 1.0 | 0.3 |  |  | 275 | 2,963 |
| 2008 | 5.5 | 7.2 | 28.3 | 14.0 | 19.0 | 4.5 | 17.6 | 1.0 | 0.4 | 0.5 | 1.7 | 0.3 |  |  | 288 | 1,532 |
| 2009 | 0.0 | 30.9 | 8.5 | 37.4 | 11.1 | 7.8 | 1.8 | 2.2 | 0.3 |  |  |  |  |  | 222 | 1,381 |
| 2010 | 0.0 | 1.2 | 25.7 | 8.7 | 36.5 | 15.8 | 9.4 | 0.9 | 1.3 | 0.3 | 0.0 | 0.3 |  |  | 267 | 2,516 |
| 2011 | 0.0 | 0.8 | 17.4 | 48.2 | 11.3 | 16.6 | 3.6 | 1.7 | 0.3 | 0.1 |  |  |  |  | 245 | 1,886 |
| 2012 | 10.2 | 0.9 | 22.5 | 21.8 | 34.1 | 6.5 | 2.8 | 0.9 | 0.3 |  |  |  |  |  | 255 | 1,842 |

Table 12. Percentage of spot by age and year, number of age samples and number of length samples by year, using pound net length and age data, 2007-2012.

| Year | Age 0 | Age 1 | Age 2 | Age 3 | Age 4 | Ages | Lengths |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2007 | 21.26 | 75.03 | 3.32 | 0.00 | 0.39 | 98 | 519 |
| 2008 | 20.77 | 78.62 | 0.61 | 0.00 | 0.00 | 206 | 1201 |
| 2009 | 7.75 | 90.70 | 1.55 | 0.00 | 0.00 | 232 | 614 |
| 2010 | 5.87 | 90.12 | 4.01 | 0.00 | 0.00 | 91 | 300 |
| 2011 | 0.37 | 99.39 | 0.23 | 0.01 | 0.00 | 173 | 582 |
| 2012 | 39.46 | 59.80 | 0.74 | 0.00 | 0.00 | 230 | 1408 |

Table 13. Atlantic menhaden proportion at age in percentage, using pound net length and age data, number of age samples and number of length samples by year, 20052012.

| Year | Age 0 | Age 1 | Age 2 | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 | \# Aged | \# Measured |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2005 |  | 2.74 | 25.86 | 42.61 | 25.64 | 3.15 |  |  | 345 | 1,061 |
| 2006 |  | 40.44 | 28.27 | 18.36 | 9.70 | 2.62 | 0.60 |  | 289 | 826 |
| 2007 |  | 22.64 | 37.44 | 24.70 | 10.72 | 3.95 | 0.55 |  | 379 | 854 |
| 2008 |  | 16.60 | 44.55 | 29.36 | 7.27 | 1.94 | 0.28 |  | 385 | 826 |
| 2009 | 0.40 | 16.79 | 24.92 | 38.04 | 17.15 | 2.72 |  |  | 258 | 512 |
| 2010 |  | 42.98 | 30.61 | 14.93 | 8.26 | 2.50 | 0.60 |  | 388 | 836 |
| 2011 |  | 38.03 | 31.41 | 19.88 | 9.12 | 1.57 |  |  | 392 | 773 |
| 2012 |  | 14.51 | 56.74 | 21.45 | 4.26 | 1.80 | 0.77 | 0.48 | 355 | 755 |

Figure 1. Summer sampling area map for 2012.


Figure 2. Weakfish length frequency distributions from onboard pound net sampling, 2009-2012.





Figure 3. Maryland commercial weakfish harvest by area, 1929-2011.


Figure 4. Maryland commercial weakfish harvest in the Chesapeake Bay, 1955-2011.


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Figure 10. Maryland commercial summer flounder harvest by area, 1962-2011.


Figure 11. Maryland commercial summer flounder harvest in the Chesapeake Bay, 19622011.


Figure 12. Estimated Maryland recreational summer flounder harvest and releases for 1981-2011 (Source: MRIP, 2013).


Figure 13. Summer Flounder statewide MRIP harvest and reported charter boat harvest from Maryland logbooks in numbers, 1993-2011.


Figure 14. Summer flounder geometric mean catch per angler from Maryland charter boat logs, with 95\% confidence intervals, 1993-2011.


Figure 15. Bluefish length frequency distributions from onboard pound net sampling, 2009-2012.





Figure 16. Maryland commercial bluefish harvest by area, 1929-2011.


Figure 17. Estimated Maryland recreational bluefish harvest and releases for 1981-2011 (Source: MRIP, 2013).


Figure 18. Bluefish statewide MRIP harvest in numbers, Maryland reported charter boat harvest in numbers and Maryland commercial harvest in pounds, 1993-2011.


Figure 19. Bluefish geometric mean catch per angler from Maryland charter boat logs, with $95 \%$ confidence intervals, 1993-2011.


Figure 20. Atlantic croaker length frequency distributions from onboard pound net sampling, 2009-2012.





Figure 21. Atlantic croaker length frequency distribution from onboard gill net sampling for 2012.


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Figure 25. Atlantic croaker geometric mean catch per angler from Maryland charter boat logs, with 95\% confidence intervals, 1993-2011.


Figure 26. Maryland juvenile Atlantic croaker geometric mean catch per trawl and 95\% confidence intervals for Maryland's lower Chesapeake Bay, 1989-2012. 1998 data point was omitted for scale (GM $1998=30.05-9.02,+12.72$ ).


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# PROJECT NO. 2 <br> JOB NO 3. <br> TASK NO. 1A 

# SUMMER - FALL STOCK ASSESSMENT AND COMMERCIAL FISHERY MONITORING 

Prepared by Jeffrey Horne

## INTRODUCTION

The primary objective of Project 2, Job 3, Task 1A was to characterize the size and age structures of the 2011 Maryland striped bass (Morone saxatilis) commercial pound net and hook-and-line harvest. The 2011 pound net season ran from 1 J une through 30 November while the commercial hook-and-line fishery was open from 7 June through 8 November. The commercial hook-and-line fishery was closed the entire month of August. These fisheries targeted resident/premigratory striped bass. Harvested fish were sampled at commercial check stations and additional fish were sampled by visiting pound nets throughout the season.

In addition to characterizing the size and age structure of the commercial harvest, data from this survey were used to monitor temporal trends in size-at-age of the harvest. These data also provided the foundation for the construction of the Maryland catch-at-age matrix utilized by the Atlantic States Marine Fisheries Commission (ASMFC) in coastal striped bass stock assessment. Length and age distributions constructed from the 2011 commercial fisheries seasons were used to characterize the length and age structure of the entire 2011 Chesapeake Bay commercial harvest and the majority of the recreational harvest (Fegley 2001).

## METHODS

## Commercial pound net monitoring

Before sampling was implemented at check stations in 2000, fish were sampled directly from pound nets. Between 1993 and 1999, pound net monitoring and accompanying tagging studies were restricted to legal-sized striped bass ( $\geq 457 \mathrm{~mm}$ or 18 inches TL). In 2000, full-net sampling was initiated at pound nets in an effort to quantify the size and age structure of striped bass by-catch. Commercial pound ne t monitoring had been conducted in tandem with a mark-recapture study designed to estimate the total instantaneous fishing mortality rate ( F ) on resident Chesapeake Bay striped bass (Hornick et al. 2005). In 2005, the tagging study was eliminated but striped bass were still s ampled monthly from pound ne ts tocontinue t he characterization of t he resident s tock structure.

From 1993-1999, it was assumed that the size and age structures of striped bass sampled at pound ne ts were representative of the $s$ ize and age $s$ tructures of $s$ triped ba ss 1 anded by the commercial pound net fishery. The validity of this assumption was questioned with the realization that commercial fishermen sometimes removed fish over 650 mm TL from nets prior to Fisheries Service (FS) staff examination, or during the culling process. These larger striped bass are highly marketable, so fishermen prefer to sell them rather than let them be tagged and released. In 2000, potential bias in the tagging study length distributions were ascertained by adding a check station component to the commercial pound net monitoring (MD DNR 2002). This allowed for the direct comparison of the length di stribution of $s$ triped ba ss sampled from pound nets to the length distribution of harvested striped bass sampled at check stations.

Pound net sampling occurred monthly from June through November 2011 (Table 1). The pound nets sampled were not randomly selected, but were chosen according to watermen's schedules and the best chance of attaining fish. During 2011, striped bass were sampled from pound nets in the upper and lower Bay. Whenever possible, all striped bass in each pound net were measured in order to investigate by-catch. Full net sampling was not possible when pound nets contained too many fish to be transferred to FS boats. If a full net could not be sampled, a random sub-sample was taken.

At each net sampled, all striped bass were measured for total length (mm TL), and the presence and category of external anomalies were noted. Scales were removed from three fish per 10 -millimeter length group per month, up to 700 mm TL, and from all striped bass greater than 700 mm TL. Other data recorded included latitude and longitude, date the net was last fished, depth, surface salinity, surface water temperature, air temperature, Secchi depth (m), and whether the net was fully or partially sampled.

## Commercial pound net/hook-and-line fisheries monitoring (check station)

All striped bass harvested in Maryland's commercial striped bass fisheries are required to pass through a MD DNR approved check station (see Project 2, Job 3, Task 5A). Check stations across Maryland were sampled for pound net and hook-and-line harvested fish each month from June through November 2011 (Figure 1). For the pound net fishery, sample targets were established of 100 fish per month from J une through August and 200 fish per m onth for September through November. This monthly allocation reflects consistent historic patterns of harvest levels, which normally increase in the fall to twice summer levels. For the hook-and-line fishery, a sample target of 400 fish per month was established over the six-month season, since historical landings exhibited no clear monthly pattern. Target sample sizes for both fisheries were based on sample sizes and age-
length keys derived from the 1997 and 1998 pound net tagging studies. Check stations were chosen by monitoring their activity and selecting from those landing $8 \%$ or more of the monthly harvest in the previous year. Stations that reported higher harvests were sampled more frequently. This method generally dispersed the sampling effort so that sample sizes were proportional to landings.

Scale samples were removed from two fish per 10-millimeter length group from striped bass less than 650 mm TL and from all striped bass greater than 650 mm TL from pound net and hook-and-line harvested fish. Scales taken from the pound net monitoring survey were combined with check station scales for ageing.

## Analytical Procedures

Scale ages from the pound net and check station surveys were applied to all fish sampled. The number of scales read per length group varied depending on the size of the fish. The decision to apply ages from the pound net fishery to hook-and-line fish was based on the study by Fegley (2001) in which striped bass sampled from pound nets and from commercial hook-and-line check stations were examined for possible differences in length at age. An analysis of covariance (Sokal and Rohlf 1995) test indicated no a ge*gear interaction ( $\mathrm{P}>\mathrm{F}=0.8532$ ). Striped bass harvested by each gear exhibited nearly identical age-length relationships; therefore ages derived from one fishery could be applied to the other. This is not surprising since both fisheries are concurrent within Maryland, and minimum and maximum length size regulations are identical.

Age composition of the pound net and hook-and-line fisheries was estimated via two-stage sampling (Kimura 1977, Quinn and Deriso 1999). In the first stage, total length samples were taken, which were assumed to be a random sample of the commercial harvest. In stage two, a fixed subsample of scales were randomly chosen to be aged. Scales from check station surveys and pound net
monitoring were combined to create the age-length key. Approximately twice as many scales as ages per length group were selected to be read based on the variance of ages per length group (Barker et al. 2004). Target sample sizes were: length group $<300 \mathrm{~mm}=3$ scales per length group; 300-400 $\mathrm{mm}=4$ scales per length group; 400-700 $\mathrm{mm}=5$ scales per length group; $>700 \mathrm{~mm}=10$ scales per length group. In some cases, the actual number of scales aged was limited by the number of samples available per length group.

Year-class was determined by reading acetate impressions of the scales placed in microfiche readers, and age was calculated by subtracting year-class from collection year. The resulting ages were used to construct an age-length key. The catch-at-age for each fishery was cal culated by applying the age-length key to the hook-and-line and pound net length frequencies, and expanding the resulting age distribution to the landings.

In order to examine recruitment into the pound ne $t$ and hook-and-line fisheries, the age structure of the harvest over time was examined. The age structure of the harvest for the 2011 hook-and-line and pound net fisheries was also compared to previous years.

Mean lengths and weights-at-age of striped bass landed in the commercial pound net and hook-and-line fisheries were derived by applying ages to all sampled fish, and weighting the means on the length distribution at each age. Mean lengths and weights at-age were calculated by year-class for the aged sub-sample of fish. Mean lengths-at-age and weights-at-age were also estimated for each year-class using an expansion method. Expanded means were calculated with an age-length key and a probability table which applied ages from the sub-sample of aged fish to all sampled fish. Agespecific length distributions based on the aged sub-sample are often different than the age-specific length distribution based on the entire length sample. Bettoli and Miranda (2001) suggested that the
sub-sample means-at-age are often biased. The two calculation methods would result in equal means only if the length distributions for each age-class were normal, which rarely occurs in these data. Finally, length frequencies from the pound net monitoring and check station samples were examined.

## RESULTS and DISCUSSION

## Pound net monitoring

During the 2011 striped bass pound net study, a total of 2,331 striped bass were sampled from one pound net in the upper Bay and five pound nets in the lower Bay. The six nets were sampled a total of 14 times during the study.

Striped bass sampled from pound nets ranged from 198-861 mm TL, with a mean length of 514 mm TL (Figure 2). In 2011, 32\% of striped bass collected from full net samples were less than the minimum legal size of 18 inches TL, while $25 \%$ of fish from partially sampled nets were sublegal. Mean total lengths of the aged sub-sample from pound nets are presented in Table 2.

Striped bass sampled from pound nets, ranged from 1 to 13 years of age (Table 3, Figure 2). Four year-old fish from the above average 2007 year-class contributed $40 \%$ in 2011; more age 4 fish than in $2010(31 \%)$ and 2009 ( $18 \%$ ). Age 5 fish from the below average 2006 year-class contributed $8 \%$ of the sample, lower than age 5 fish in 2010 ( $21 \%$ ) (Figure 3, Table 3). Age 3 fish contributed $14 \%$ in 2011, which is lower than the contribution in $2010(33 \%)$. Striped bass age 6 and over were more common in 2011, and accounted for $30 \%$ of the sample; more than their contribution in 2009 $(9 \%)$ and $2010(23 \%)$. Fish age 8 and older composed $4 \%$ of the sample in 2011, which was higher than $2009(1 \%)$ and $2010(1 \%)$. Length frequencies of legal sized striped bass sampled at pound nets were almost identical to length distributions from the check stations (Figure 4).

## Hook-and-line check station sampling

A total of 1,431 striped bass were sampled at hook-and-line check stations in 2011. The mean length of sampled striped bass was 554 mm TL. Striped bass sampled from the hook-and-line fishery ranged from 434 to 895 mm TL (Figure 5) and from 3 to 11 years of age (Figure 5).

The length frequency and ages of the sampled fish were applied to the total harvest. Striped bass in the $470-550 \mathrm{~mm}$ length groups accounted for $59 \%$ of the hook-and-line harvest, lower than 2010 ( $69 \%$; Figure 5). Fish $>630 \mathrm{~mm} \mathrm{TL}$ contributed $8 \%$ to the total harvest. As in past years, few large fish were available to the hook-and-line fishery. Striped bass over 700 mm TL were harvested throughout the season, and contributed $3 \%$ to the overall harvest (Figure 6). Historically, these fish have not been available in large numbers during the summer (MDDNR 2002). Approximately $1 \%$ of the harvest was sub-legal ( $<457 \mathrm{~mm} \mathrm{TL}$ ). Mean lengths-at-age and weights-at-age for the 2011 combined hook-and-line and pound net fisheries are shown in Tables 4 and 5.

The 2011 hook-and-line harvest accounted for $23 \%$, by weight, of the Maryland Chesapeake Bay total commercial harvest in 2011 (see Project 2, Job 3, Task 5A). The estimated 2011 catch-atage of the hook-and-line fishery is presented in Table 6. The majority of the harvest was composed of four to seven year-old striped bass (93\%). Striped bass from the 2007 (age 4) and 2005 (age 6) year-classes contributed $48 \%$ and $21 \%$, respectively. Fish from the strong 2003 year-class (age 8) accounted for $3 \%$ of the total, less than in 2010 (11\%). Striped bass from the below average 2006 year-class (age 5) contributed $10 \%$, which was lower than their contribution in 2010 (Figure 7). Fish from the 2004 year-class (age 7) contributed $15 \%$ to the hook-and-line harvest, less than in 2010 (21\%). Striped bass age 8 and older contributed 4\% to the overall harvest in 2011, similar to 2010 (4\%).

## Pound net check station sampling

A total of 1,128 striped bass were sampled at pound net check stations in 2011. Striped bass sampled ranged from 453 to 916 mm TL (Figure 5). Striped bass sampled from the pound net fishery ra nged from 3 to 11 years of age. Striped bass in the $450-550 \mathrm{~mm}$ TL 1 ength groups accounted for $51 \%$ of the 2011 pound net harvest, which is lower than 2010 ( $77 \%$; Figure 5). The contribution of striped bass in the 570-630 mm TL length groups increased from 18\% in 2010 to $32 \%$ in 2011. Fish $>630 \mathrm{~mm}$ TL composed $17 \%$ of the sample, three times that of 2010 (5\%). In general, a number of large fish were available to the 2011 pound net fishery (Figure 6). Mean lengths-at-age and weights-at-age from the 2011 hook-and-line and pound net fisheries combined, are shown in Tables 4 and 5, respectively.

The pound net fishery accounted for $33 \%$, by weight, of the Maryland Chesapeake Bay 2011 commercial harvest (see Proj. 2, Job 3, Task 5A). The estimated 2011 catch-at-age for the pound net fishery is presented in Table 6. Fish age three to six contributed $75 \%$ of the 2011 total pound net harvest. The contribution of eight year-old fish from the 2003 year-class was lower in the pound net harvest in 2011 than in 2010, contributing 7\% to the total harvest (Figure 7). Striped bass age 8 and over composed $10 \%$ of the 2011 harvest, much higher than the contribution in 2010 (2\%). Sub-legal striped bass ( $<457 \mathrm{~mm} \mathrm{TL}$ ) composed $0.1 \%$ of the total pound net harvest.

## Monitoring summary

Striped bass ranging from 457 to 550 mm TL composed $51 \%$ and $59 \%$ of the 2011 pound net and hook-and-line fisheries, respectively. There were more large fish ( $>530 \mathrm{~mm}$ ) harvested in 2011 compared to 2010 ( $71 \%$ for both fisheries; Figure 5). In 2011, 120 fish from pound net monitoring
and 99 fish from check station sampling were aged. Younger fish (age 3 to 6) were abundant, accounting for the majority of the harvest (Figure 7). Length frequencies of legal-sized fish sampled from pound nets and all fish from check stations were almost identical (Figure 4).

The mean lengths of 4,5 , and 6 year-old legal-sized striped bass ( $\geq 457 \mathrm{~mm}$ TL) decreased during the period 1990 to 2000 (Figure 8). Since 2001, there was no apparent trend for mean lengths of striped bass aged 4 to 6 .

An ANOVA with a Duncan's Post Hoc Test (SAS 2006) was performed to compare lengths and weights of striped bass harvested between fisheries and months in 2011. Striped bass were significantly $(\mathrm{P}<0.05)$ longer and heavier from the pound net fishery than the hook-and-line fishery.

During the hook-and-line fishery, the longest and heaviest fish were sampled in June/July and the smallest in September. Striped bass sampled in June/July were significantly longer than fish harvested in September/October/November. No lengths were available for August (season closed). Striped ba ss sampled in June/July were significantly heavier $t$ han $f$ ish ha rvested in September/October. No weights were available for August (season closed) or November (scale malfunction).

In the pound net check station monitoring, the longest and heaviest fish were harvested in October and the smallest in July. Striped bass November and August were similar in length, but significant differences in length were evident in every other month. Striped bass from June and October were significantly heavier than all other months. Striped bass from August and November were significantly heavier than July and September.

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Figure 8. Mean lengths for legal-size striped bass ( $\geq 457 \mathrm{~mm} \mathrm{TL}$ ) by year for $4,5,6$, a nd 7 year-old striped ba ss s ampled from Maryland Chesapeake Bay pound ne ts a nd commercial hook-and-line and pound net check stations, 1990 through 2011. Mean lengths were calculated by using sub-sampled ages only and by expanding ages to sample length frequency before calculating means. The $95 \%$ confidence intervals are shown around points in the s ub-sample da ta s eries. (1990-2007 e dited) Note different scales.

Table 1. Summary of sampling areas, sampling dates, surface temperature, surface salinity and numbers of fish encountered during the 2011 Maryland Chesapeake Bay commercial pound net monitoring survey.

| Month | Area | Number <br> of Nets <br> Sampled | Mean <br> Water <br> Temp ( $\left.{ }^{\circ} \mathbf{C}\right)$ | Mean <br> Salinity <br> $\mathbf{( p p t )}$ | Number <br> of Fish <br> Sampled |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Upper | 1 | 24.0 | 3.8 | 181 |
| June | Middle | - | - | - | - |
|  | Lower | 1 | 25.4 | 9.1 | 132 |
|  | Upper | 1 | 27.3 | 6.3 | 167 |
| July | Middle | - | - | - | - |
|  | Lower | - | - | - | - |
|  | Upper | 1 | 27.3 | 8.4 | 195 |
| August | Middle | - | - | - | - |
|  | Lower | - | - | - | - |
|  | Upper | - | - | - | - |
| September | Middle | - | - | - | - |
|  | Lower | 4 | 23.4 | 10.6 | 428 |
|  | Upper | 1 | 19.2 | 3.9 | 288 |
| October | Middle | - | - | - | - |
|  | Lower | 2 | 14.3 | 8.8 | 406 |
|  | Upper | 1 | 12.5 | 4.5 | 167 |
| November | Middle | - | - | - | - |
|  | Lower | 2 | 11.7 | 8.0 | 367 |

Table 2. Mean length-at-age (mm TL) of striped bass sampled from pound nets in Maryland's Chesapeake Bay, June through November 2011.

| Year-class | Age | n | Mean <br> length <br> (mm TL) | STD | STDERR | LCLM | UCLM |
| :---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: |
| 2010 | 1 | 6 | 232 | 36 | 15 | 195 | 269 |
| 2009 | 2 | 21 | 322 | 52 | 11 | 298 | 346 |
| 2008 | 3 | 13 | 403 | 37 | 10 | 381 | 425 |
| 2007 | 4 | 22 | 488 | 68 | 14 | 458 | 518 |
| 2006 | 5 | 6 | 580 | 103 | 42 | 477 | 683 |
| 2005 | 6 | 8 | 654 | 62 | 22 | 603 | 705 |
| 2004 | 7 | 15 | 668 | 78 | 20 | 625 | 711 |
| 2003 | 8 | 9 | 762 | 59 | 20 | 718 | 806 |
| 2002 | 9 | 7 | 771 | 39 | 15 | 736 | 806 |
| 2001 | 10 | 7 | 793 | 30 | 12 | 766 | 820 |
| 2000 | 11 | 5 | 791 | 57 | 25 | 726 | 856 |
| 1998 | 13 | 1 | 844 | - | - | - | - |

Table 3. Number of striped bass, by age, sampled from pound nets, in Maryland's Chesapeake Bay, June through November 2011.

| Year-class | Age | Pound Net Monitoring |  |
| :---: | :---: | :---: | :---: |
|  |  | Number sampled at age (n) | Percent of Total |
| 2010 | 1 | 13 | 0.54 |
| 2009 | 2 | 174 | 7.48 |
| 2008 | 3 | 318 | 13.65 |
| 2007 | 4 | 935 | 40.10 |
| 2006 | 5 | 197 | 8.45 |
| 2005 | 6 | 344 | 14.75 |
| 2004 | 7 | 248 | 10.62 |
| 2003 | 8 | 66 | 2.82 |
| 2002 | 9 | 17 | 0.73 |
| 2001 | 10 | 14 | 0.62 |
| 2000 | 11 | 6 | 0.24 |
| 1998 | 13 | 1 | 0.02 |
| Total |  | $\mathbf{2 , 3 3 1}$ | $\mathbf{1 0 0 . 0 0}$ |

Table 4. Mean length-at-age ( mm TL ) of legal-size striped bass ( $\geq 457 \mathrm{~mm} \mathrm{TL} / 18 \mathrm{in} \mathrm{TL}$ ) for ages 3-14 sampled from commercial pound net and hook-and-line fisheries in Maryland's Chesapeake Bay, June through November 2011.

| Year-class | Age | n | Mean <br> Length <br> (mm TL) | STD | STDERR | LCLM | UCLM |
| :---: | :---: | ---: | :---: | :---: | :---: | :---: | :---: |
| 2008 | 3 | 3 | 465 | 6 | 3 | 450 | 479 |
| 2007 | 4 | 14 | 507 | 52 | 14 | 476 | 537 |
| 2006 | 5 | 7 | 612 | 34 | 13 | 580 | 643 |
| 2005 | 6 | 17 | 622 | 61 | 15 | 591 | 653 |
| 2004 | 7 | 14 | 690 | 80 | 21 | 644 | 737 |
| 2003 | 8 | 21 | 743 | 80 | 18 | 707 | 780 |
| 2002 | 9 | 9 | 795 | 73 | 24 | 739 | 851 |
| 2001 | 10 | 11 | 828 | 39 | 12 | 802 | 855 |
| 2000 | 11 | 3 | 837 | 65 | 38 | 674 | 999 |

Table 5. Mean weight-at-age (kg) of legal-size striped bass ( $\geq 457 \mathrm{~mm} \mathrm{TL} / 18$ in TL) sampled from commercial pound net and hook-and-line fisheries in Maryland's Chesapeake Bay, June through November 2011. Mean weights are weighted by the sample $n$-at-length in each age.

| Year-Class | Age | n Aged | Weighted Mean <br> weight* (kg) |
| :---: | :---: | ---: | :---: |
| 2008 | 3 | 3 | 0.8 |
| 2007 | 4 | 13 | 1.2 |
| 2006 | 5 | 6 | 2.3 |
| 2005 | 6 | 13 | 2.5 |
| 2004 | 7 | 14 | 3.2 |
| 2003 | 8 | 21 | 4.0 |
| 2002 | 9 | 9 | 5.2 |
| 2001 | 10 | 11 | 5.8 |
| 2000 | 11 | 3 | 6.3 |

[^0]Table 6. Estimated catch-at-age of striped bass landed by Maryland Chesapeake Bay commercial hook-and-line and pound net fisheries, June through November 2011.

| Year-class | Age | Hook and Line |  | Pound Net |  |
| :---: | ---: | ---: | ---: | ---: | ---: |
|  |  | Landings in <br> Pounds of Fish | Percent of <br> Total | Landings in <br> Pounds of Fish | Percent of <br> Total |
| 2009 | 2 | 103 | 0.1 | 0 | 0 |
| 2008 | 3 | 10,746 | 2.4 | 16,094 | 2.5 |
| 2007 | 4 | 211,851 | 48.0 | 277,766 | 42.9 |
| 2006 | 5 | 41,995 | 9.5 | 64,448 | 9.9 |
| 2005 | 6 | 91,389 | 20.7 | 125,169 | 19.3 |
| 2004 | 7 | 65,940 | 14.9 | 101,153 | 15.6 |
| 2003 | 8 | 14,089 | 3.2 | 43,097 | 6.6 |
| 2002 | 9 | 2,622 | 0.6 | 8,981 | 1.4 |
| 2001 | 10 | 1,785 | 0.4 | 7,907 | 1.2 |
| 2000 | 11 | 903 | 0.2 | 3,498 | 0.5 |
| 1999 | 12 | 0 | 0.0 | 0 | 0.0 |
| 1998 | 13 | 0 | 0.1 | 0 | 0.0 |
| 1997 | 14 | 0 | 0.1 | 0 | 0.0 |
| Total* |  | $\mathbf{4 4 1 , 4 2 2}$ | $\mathbf{1 0 0 . 0}$ | $\mathbf{6 4 8 , 1 1 3}$ | $\mathbf{1 0 0 . 0}$ |

[^1]Figure 1. Locations of Chesapeake Bay commercial pound net and hook-and-line check stations sampled from June through November 2011.


Figure 2. Age and length (mm TL) frequencies of striped bass sampled during Maryland Chesapeake Bay pound net monitoring study, June through November 2011.



Figure 3. Age structure of striped bass ( $\geq 457 \mathrm{~mm} \mathrm{TL} / 18 \mathrm{in} \mathrm{TL}$ ) sampled from Maryland Chesapeake Bay commercial pound net monitoring study from 1996 through 2011.


AGE

Figure 3. Continued.




## AGE

Figure 4. Length frequency of striped bass sampled during the 2011 pound net monitoring, pound net check station and hook-and-line check station surveys. All fish were sampled from June through November 2011. Pound net monitoring length frequency is for legal-size fish only ( $\geq 457 \mathrm{~mm} \mathrm{TL} / 18$ in TL).


Figure 5. Age and length frequencies of striped bass sampled from Maryland Chesapeake Bay commercial hook-and-line and pound net check stations, June through November 2011.



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Figure 6. Month-specific length distributions of striped bass sampled from Maryland Chesapeake Bay commercial hook-and-line and pound net fisheries, June through November 2011.

*No fish for August Hook and Line, season was closed for entire month.

Figure 7. Age structure of striped bass sampled from Maryland Chesapeake Bay commercial hook-and-line and pound net check stations, 1999 through 2011. Note-pound net check station sampling began in 2000.


Age

Figure 7. Continued.


Age

Figure 8. Mean lengths for legal-size striped bass ( $\geq 457 \mathrm{~mm} \mathrm{TL}$ ) by year for $4,5,6$, and 7 yearold striped bass sampled from Maryland Chesapeake Bay pound nets and commercial hook-and-line and pound net check stations, 1990 through 2011. Mean lengths were calculated by using sub-sampled ages only and by expanding ages to sample length frequency before calculating means. The $95 \%$ confidence intervals are shown around points in the sub-sample data series. (1990-2007 edited). Note different scales.





Year

# PROJECT NO. 2 <br> JOB NO. 3 <br> TASK NO. 1B 

# WINTER STOCK ASSESSMENT AND COMMERCIAL FISHERY MONITORING 

Prepared by Jeffrey Horne

## INTRODUCTION

The primary objective of Project 2 , J ob 3, Task 1B was to characterize the size and age structure of striped bass (Morone saxatilis) sampled from the December 6, 2011 - February 29, 2012 commercial drift gill net fishery. This fishery targets resident/pre-migratory Chesapeake Bay striped bass and accounts for approximately 40-50\% of the Maryland Chesapeake Bay commercial harvest.

In addition to characterizing the size and age structure of this component of the commercial harvest, these data were used to monitor temporal trends in length and weight-at-age of resident/premigratory striped bass. These data also contributed to the construction of the Maryland catch-at-age matrix utilized in the Atlantic States Marine Fisheries Commission (ASMFC) coastal striped bass stock assessment.

## METHODS

## Data collection procedures

All striped bass harvested in Maryland's commercial striped bass fishery are required to pass through a Maryland Department of Natural Resources (MD DNR) approved check station. Striped bass check stations were sampled for the winter stock assessment according to a stratified random
sampling design. Strata were defined as either high-use, medium-use, or low-use check stations based on landings from the previous year. Individual check stations that processed $8 \%$ or greater of the entire catch were designated as high-use stations, stations that processed between $3 \%$ and $7.9 \%$ of the catch were designated as medium-use, and any station that processed less than $3 \%$ of the catch were designated as low-use. High-use and medium-use stations were sampled at a 3 to 1 ratio; one medium-use station was sampled for every three visits to a high-use station with a sample intensity of one visit per week for the duration of the fishery, or multiple times per week when quota was caught quickly. Low-use sites were not sampled. Days and stations were randomly selected each month, although the results of the random draw were frequently modified because of weather, check station hours, and other logistical constraints. Sampling was distributed as evenly as possible between northern and eastern geographic areas of the Chesapeake Bay. The northern-most check station sampled in this survey was located in Millington, while the southern-most station was located on Hooper's Island (Figure 1).

Monthly sample targets were 1,000 fish in December and 1,250 fish in both January and February, for a total target sample size of 3,500 fish. Sampling at this level provides an accurate representation of both the length and age distributions of the harvest (Fegley et al. 2000). At each check station, attempts were made to measure ( mm TL ) and weigh $(\mathrm{kg})$ a random sample of at least 300 striped bass per visit. On days when fewer than 300 fish were checked in, all individuals were sampled. For fish less than 700 mm TL , scales were taken randomly from two fish per 10 mm length group per visit, but scales were taken from all fish greater than or equal to 700 mm TL.

## Analytical procedures

Age composition of the sample was estimated via two-stage sampling (Kimura 1977, Quinn and Deriso 1999). In the first stage, length and scale samples were taken. These were assumed to be a random sample of the com mercial ha rvest. In stage two, a fixed sub-sample of scales were randomly chosen to be aged. Approximately twice as many scales as ages per length group were selected to be read based on the range of ages per length group (Barker et al. 2004). Target sample sizes of scales to be read were 5 scales per length groups for $400-700 \mathrm{~mm}$ and 10 scales per length group for $>700 \mathrm{~mm}$. In some cases, the actual number of scales aged was limited by the number of samples available per length group.

Ages were assigned to scales by viewing acetate impressions in a microfiche reader. The resulting a ge-length key was a pplied to the sample length-frequency to generate a s ample age distribution. Finally, the age distribution of the total 2011-2012 winter gill net harvest was estimated by applying the sample age distribution to the total reported landings. Because the winter gill net season straddles two calendar years, ages were calculated by subtracting year-class (assigned by scale readers) from the year in which the fishery ended. For example, for the December 2011 - February 2012 gill net season, the year used for age calculations was 2012 .

Mean lengths and weights at-age were calculated by year-class for the aged sub-sample of fish. Mean length-at-age and weight-at-age were al so estimated for each year-class us ing an expansion method (Hoover 2008). Age-specific length distributions based on the aged sub-sample are often different than the age-specific length distribution based on the entire length sample. Bettoli and Miranda (2001) suggest that the sub-sample means-at-age are often biased. Expanded means were calculated with an age-length key and a probability table that applied ages from the sub-sample
of aged fish to all sampled fish. The two calculation methods would result in equal means only if the length distributions for each age-class were normal, which rarely occurs with these data.

To examine recruitment into the winter drift gill net fishery and the age-class structure of the harvest over time, the ex panded age structure of the 2011-2012 harvest was compared to that of previous years beginning with the 1993-1994 gill net season. Trends in growth were examined by plotting actual mean length-at-age and mean weight-at-age of aged sub-samples, with confidence intervals, by year, for individual age-classes. Expanded mean lengths-at-age and weights-at-age were also plotted on the same time series graph for comparison.

## RESULTS and DISCUSSION

The winter dr ift gill ne $t$ com mercial fishery account ed for $45 \%$ of $t$ he $t$ otal $M$ aryland Chesapeake Bay commercial harvest, by weight. A total of4,169 striped bass were sampled and 114 striped bass were aged from the harvest between December 2011 - February 2012. The gill net season was open for 9 days in December, 8 days in January, and 8 days in February due to high catch rates.

Commercial gill nets have been limited to mesh sizes no less than 5 and no greater than 7 inches since the fishery reopened after the 1985-1990 moratorium. As a result, the range in ages of the commercial striped bass drift gill net landings has not fluctuated greatly since the inception of MD DNR check station monitoring during the 1993-1994 gill net season (Figure 2). The majority of fish landed in most years were between 4 and 8 years old. However, the contribution of individual ages to the overall landings has varied between years based on year-class strength. The overall landings of striped bass in this fishery were calculated from the ASMFC compliance report template.

According to the estimated catch-at-age analysis, the 2011-2012 commercial drift gill net harvest consisted primarily of striped bass from the 2007 year-class (age 5; Table 1), which composed $47 \%$ of the total harvest. The 2008 and 2006 year-classes (ages 4 and 6 ) composed an additional $35 \%$ of the total harvest, while ages 8 and older contributed only $2 \%$ to the total. The contribution of fish greater than 8 years old was lower than the 2010-2011 harvest (6\%) and the 2009-2010 harvest (6\%). The youngest fish observed in the 2011-2012 sampled harvest were age 3 .

Mean lengths and weights-at-age of the aged sub-sample and the estimated means from the expansion technique are presented in Tables 2 and 3. Expanded mean lengths and weights-at-age were generally slightly higher for smaller fish and slightly lower for larger fish than sub-sample means. Striped bass were recruited into the 2011-2012 winter gill net fishery at age 3 (2009 yearclass), with an expanded mean length and weight of 489 mm TL and 1.37 kg . The 2007 year-class (age 5) was most commonly observed in the sampled landings with an expanded mean length and weight of 544 mm TL and 1.84 kg , respectively. The expanded mean length and weight of the oldest fish in the aged sub-sample (age 11, 2001 year-class) were 891 mm TL and 9.37 kg , respectively.

The length frequency distributions by check station area are presented in Figure 3. The length frequency distributions were dominated by fish in the 490-610 mm TL range. Sub-legal fish ( $<457 \mathrm{~mm}$ ) composed less than $1 \%$ of the bay-wide sampled harvest.

Time series of sub-sampled and expanded mean lengths and weights for the period 19942012 are shown in Figures 4 and 5 for fish ages 4 through 9, which generally make up $95 \%$ or more of the ha rvest. Mean length-at-age and weight-at-age for a ge 4 a nd 5 s triped bass have been relatively constant. Mean length-at-age and weight-at-age for ages $6,7,8$, and 9 are more variable, likely due to smaller sample sizes or greater range of lengths and weights for each age group.

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Figure 3. Length frequency distribution of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, December 2011 - February 2012.

Figure 4. Mean total lengths (mm TL) of the aged sub-sample, by year, for individual ageclasses of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, 1994-2012 ( $95 \%$ confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. Year refers to the year in which the season ended.

Figure 5. Mean weights (kg) of the aged sub-sample, by year, for individual age-classes of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net fishery, 1994-2012 ( $95 \%$ confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. Year refers to the year in which the season ended.

Table 1. Estimated catch-at-age of striped bass (numbers of fish) landed by the Maryland Chesapeake Bay commercial drift gill net fishery, December 2011-February 2012.

| Year-class | Age | Catch | Percentage <br> of the catch |
| :---: | ---: | ---: | :---: |
| 2009 | 3 | 2,681 | 1 |
| 2008 | 4 | 23,230 | 11 |
| 2007 | 5 | 96,149 | 47 |
| 2006 | 6 | 49,581 | 24 |
| 2005 | 7 | 27,271 | 13 |
| 2004 | 8 | 2,123 | 1 |
| 2003 | 9 | 2,887 | 1 |
| 2002 | 10 | 0 | 0 |
| 2001 | 11 | 49 | 0 |
| Total* |  | $\mathbf{2 0 3 , 9 7 1}$ | $\mathbf{1 0 0}$ |

[^2]Table 2. Mean total lengths (mm TL) by year-class of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, December 2011-February 2012.

| Year-class | Age | n fish <br> aged | Mean TL <br> (mm) of <br> aged sub- <br> sample | Estimated <br> \# at-age <br> in sample | Expanded <br> mean TL <br> (mm) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | 3 | 3 | 463 | 55 | 489 |
| 2008 | 4 | 11 | 482 | 475 | 504 |
| 2007 | 5 | 22 | 555 | 1,965 | 544 |
| 2006 | 6 | 12 | 603 | 1,013 | 569 |
| 2005 | 7 | 25 | 688 | 557 | 597 |
| 2004 | 8 | 22 | 741 | 43 | 713 |
| 2003 | 9 | 18 | 760 | 59 | 682 |
| 2002 | 10 | 0 | - | 0 | - |
| 2001 | 11 | 1 | 891 | 1 | 891 |
| Total* |  | $\mathbf{1 1 4}$ |  | $\mathbf{4 , 1 6 9}$ |  |

[^3]Table 3. Mean weights (kg) by year-class of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, December 2011-February 2012.

| Year-class | Age | n fish <br> aged | Mean <br> weight <br> (kg) of <br> aged sub- <br> sample | Estimated <br> \# at-age <br> in sample | Expanded <br> mean weight <br> (kg) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | 3 | 3 | 1.21 | 55 | 1.37 |
| 2008 | 4 | 11 | 1.28 | 475 | 1.49 |
| 2007 | 5 | 22 | 1.94 | 1,965 | 1.84 |
| 2006 | 6 | 12 | 2.71 | 1,013 | 2.07 |
| 2005 | 7 | 25 | 3.87 | 557 | 2.43 |
| 2004 | 8 | 22 | 4.76 | 43 | 4.28 |
| 2003 | 9 | 18 | 5.39 | 59 | 3.76 |
| 2002 | 10 | 0 | - | 0 | - |
| 2001 | 11 | 1 | 9.37 | 1 | 9.37 |
| Total* |  | $\mathbf{1 1 4}$ |  | $\mathbf{4 , 1 6 9}$ |  |

[^4]Figure 1. Registered Maryland Chesapeake Bay check stations sampled for commercial drift gill net-harvested striped bass, December 2011-February 2012.


Figure 2. Age distribution of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, 1994-2012.


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Figure 2. Continued.


## Age (Years)

Figure 3. Length frequency distributions of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, December 2011-February 2012.


Figure 4. Mean total lengths (mm TL) of the aged sub-sample, by year, for individual ageclasses of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net landings, 1994-2012 ( $95 \%$ confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. Year refers to the year in which the season ended.


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Figure 4. Continued.


## Year

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Figure 5. Mean weights ( kg ) of the aged sub-sample, by year, for individual age-classes of striped bass sampled from the Maryland Chesapeake Bay commercial drift gill net fishery, 1994-2012 ( $95 \%$ confidence intervals are shown around each point).
Expanded means (estimated from entire sample) are also shown. Year refers to the year in which the season ended.




Year
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Figure 5. Continued.




Year

# PROJECT NO. 2 

JOB NO. 3
TASK NO. 1C

# ATLANTIC COAST STOCK ASSESSMENT AND COMMERCIAL HARVEST MONITORING 

Prepared by Amy Batdorf

## INTRODUCTION

The primary objective of Project 2, Job 3, Task 1C was to characterize the size and age structure of commercially harvested striped bass from Maryland's Atlantic coast. Trawls and gill nets were permitted during the Atlantic season, which occurred between November 1, 2011 and April 30, 2012. This fishery was managed with a 24 inch total length (TL) minimum size limit and an annual quota of 126,396 pounds. Although this report covers the November 2011-April 2012 fishing season, the quota is managed by calendar year. Maryland's Atlantic coast fishery is not as large as the Chesapeake Bay commercial fishery and its annual quota comprises only $6 \%$ of Maryland's total commercial harvest quota. Monitoring of the coastal fishery began in 2006 to improve Maryland's catch-at-age and weight-at-age estimates used in the annual compliance report to the Atlantic States Marine Fisheries Commission, as well as the coast-wide stock assessment.

## METHODS

## Data collection procedures

All striped bass commercially harvested in Maryland are required to pass through a Maryland Department of Natural Resources (MD DNR) approved check station. Check stations are typically cooperating fish dealers who report daily landings to MD DNR. A review of 2005-2010 check station activity indicated that $81 \%$ of striped bass harvested
along Maryland's A tlantic c oast pa ssed through t wo c heck stations in Ocean City, Maryland. Consequently, s ampling al ternated between these two check stations as fish came in during the season. Catches were typically intermittent and personnel sampled when fish were ava ilable. A monthly sample $t$ arget of 150 f ish w as established for November, December, and January, because of a previous analysis of check station logs showed that $96 \%$ of the harvest oc curs dur ing these months. Fish were measured (mm TL) and weighed ( kg ) and scales were randomly taken from five fish per 10 mm length group per day for age determination.

## Analytical procedures

Age c omposition of the sample was e stimated via two-stage sampling (Kimura 1977, Quinn and Desiro 1999). In stage one, a random sample of lengths was taken from the total catch from November 2011 through April 2012. For stage two, a sub-sample of scales from Atlantic coast striped bass was aged.

Year-class was determined by reading acetate impressions of the scales placed in microfiche readers. Because the Atlantic coast fishery spans two calendar years, age was calculated by $s$ ubtracting $t$ he as signed year-class $f$ rom $t$ he year in which the $f$ ishery ended. I n the November 2011 -April 2012 Atlantic fishery, the year u sed for a ge calculations w as 2012. These ages were $t$ hen used to construct the age-length ke y (ALK). The resulting ALK was applied to the sample length frequency to generate a sample age distribution for all fish sampled at check stations. The age distribution of the Atlantic c oast ha rvest from November 2011 through A pril 2012 was es timated by applying the sample age distribution to the total landings.

Mean lengths and weights at-age were calculated by year-class for the sub-sample of fish. Mean lengths-at-age and mean weights-at-age were also estimated for each yearclass using an expansion method. Bettoli and Miranda (2001) suggested that age-specific length distributions based on an aged sub-sample are often different than the age-specific
length distribution based on the entire length sample. The two calculation methods (subsample $m$ eans a nd expanded $m$ eans) would $r$ esult in e qual $m$ eans onl $y$ if $t$ he 1 ength distributions $f$ or $e$ ach age-class $w$ ere no rmal, w hich rarely o ccurs in these da ta. Therefore, expanded means were calculated with an ALK and a p robability ta ble that applied ages from the sub-sample of aged fish to all sampled fish.

## RESULTS and DISCUSSION

Sampling at coastal check stations was conducted on twenty-seven days between November 2011 and April 2012. A total of 561 fish were measured and weighed and the ALK was developed from 210 scale samples. This is the largest sample obtained from the Atlantic fishery in the time series. Because this fishery is largely a bycatch fishery, fish were harvested intermittently and are often difficult to intercept at the check stations.

Fish harvested during the 2011-2012 Atlantic c oast fishing season ranged from age 4 (2008 year-class) to age 21 ( 1991 year-class) (Figure 1). Most (72\%) striped bass harvested were ages 7 through 10 (Table 1). Striped bass were recruited into the Atlantic coast fishery as young as age 4 , but due to the 24 inch minimum size limit, few fish younger than age 6 were harvested, which is similar to previous years.

Fourteen year classes w ere r epresented in the s ampled harvest. Based on t he estimated catch-at-age, the most common age harvested dur ing the 2011-2012 A tlantic coast harvest was age 9 (2003 year-class), which represented $34 \%$ of the fishery (Table 1). Large contributions were also made by the 2004 year class (age 8) and the 2005 year class (age 7), which represented $16 \%$ and $13 \%$ of the fishery, respectively.

Striped bass sampled at Atlantic coast check stations during the 2011-2012 season had a mean length of 800 mm TL and mean weight of 5.6 kg . The length distribution of fish harvested in the 2011-2012 season ranged from 610 to 1270 mm TL (Figure 2). The weight distribution of the fish harvested ranged from 2.4 to 22.1 kg .

The sub-sample means-at-age and the expanded means-at-age for both length and weight were very similar (Tables 2 and 3, Figures 3 and 4). In 2012, 210 of the 561 fish (37\%) sampled were aged. Because a high proportion of the total sample was aged, the expanded $m$ ean 1 engths a nd $w$ eights-at-age $w$ ere $s$ imilar $t o$ means of $t$ he ag ed subsample, and generally within the $95 \%$ confidence limits. Recently recruited age 5 fish had an expanded mean length of 657 mm TL and expanded mean weight of 3.1 kg . Age 9 striped bass, the most abundant a ge harvested, had an expanded mean length of 798 mm TL and expanded mean weight of 5.3 kg (Figure 1). Age 8 striped bass, the next most abundant year-class harvested, had an expanded mean length of 770 mm TL and an expanded mean weight of 4.8 kg .

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Quinn, T.J. and R.B. Desiro. 1999. Quantitative Fish Dynamics Oxford University Press.

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Table 2. Sub-sample and expanded mean total lengths (mm TL) by year-class of striped bass sampled from Atlantic coast fishery, November 2011-April 2012. Includes the lower and upper $95 \%$ confidence limits (LCL and UCL, respectively).

Table 3. Sub-sample and expanded mean weights (kg) by year-class of striped bass sampled from Atlantic coast fishery, November 2011-April 2012. Includes the lower and upper $95 \%$ confidence limits (LCL and UCL, respectively).

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Figure 1. Age distribution of striped bass sampled from the Atlantic coast fishery, 2006-2012 seasons.

Figure 2. Length distribution of striped bass sampled from the Atlantic coast fishery, 2006-2012 seasons.

Figure 3. Mean total lengths (mm TL) of the aged sub-sample, by year, for individual age-classes of striped bass (through age 12) sampled from the Maryland Atlantic coast trawl and gill net landings, 2006-2012 (95\% confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. *Note differences in scales on the $y$-axis.

Figure 4. Mean weight (kg) of the aged sub-sample, by year, for individual ageclasses of striped bass (through age 12) sampled from the Maryland Atlantic coast trawl and gill net landings, 2006-2012 ( $95 \%$ confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. *Note differences in scales on the y-axis.

Table 1. Estimated catch-at-age of striped bass (numbers of fish) landed by the Maryland Atlantic coast commercial fishery, November 2011-April 2012.

| Year- <br> Class | Age | Catch | Percent |
| :---: | :---: | :---: | :---: |
| 2008 | 4 | 9 | 0.2 |
| 2007 | 5 | 301 | 5.9 |
| 2006 | 6 | 351 | 6.9 |
| 2005 | 7 | 684 | 13.4 |
| 2004 | 8 | 795 | 15.5 |
| 2003 | 9 | 1729 | 33.8 |
| 2002 | 10 | 462 | 9.0 |
| 2001 | 11 | 404 | 7.9 |
| 2000 | 12 | 193 | 3.8 |
| 1999 | 13 | 81 | 1.6 |
| 1998 | 14 | 20 | 0.4 |
| 1997 | 15 | 61 | 1.2 |
| 1996 | 16 | 18 | 0.4 |
| 1991 | 21 | 9 | 0.2 |
|  | Total | $\mathbf{5 , 1 1 7}$ | $\mathbf{1 0 0}$ |

*Sum of columns may not equal totals due to rounding

Table 2. Sub-sample and expanded mean total lengths (mm TL) by year-class of striped bass sampled from Atlantic coast fishery, November 2011-April 2012. Includes the lower and upper 95\% confidence limits (LCL and UCL, respectively).

| Year- <br> Class | Age | n Fish <br> Aged | Mean TL <br> (mm) of Aged <br> sub-sample | LCL | UCL | Estimated \# <br> at-age in <br> sample | Expanded <br> Mean TL <br> (mm) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 8}$ | 4 | 1 | 617 | --- | --- | 1 | 617 |
| $\mathbf{2 0 0 7}$ | 5 | 12 | 640 | 620 | 660 | 33 | 657 |
| $\mathbf{2 0 0 6}$ | 6 | 9 | 667 | 646 | 689 | 38 | 673 |
| $\mathbf{2 0 0 5}$ | 7 | 19 | 733 | 705 | 761 | 75 | 729 |
| $\mathbf{2 0 0 4}$ | 8 | 28 | 786 | 758 | 814 | 87 | 770 |
| $\mathbf{2 0 0 3}$ | 9 | 64 | 819 | 799 | 839 | 190 | 798 |
| $\mathbf{2 0 0 2}$ | 10 | 23 | 901 | 862 | 941 | 51 | 866 |
| $\mathbf{2 0 0 1}$ | 11 | 24 | 953 | 932 | 974 | 44 | 945 |
| $\mathbf{2 0 0 0}$ | 12 | 13 | 991 | 960 | 1023 | 21 | 980 |
| $\mathbf{1 9 9 9}$ | 13 | 7 | 1004 | 984 | 1024 | 9 | 1002 |
| $\mathbf{1 9 9 8}$ | 14 | 2 | 1060 | 679 | 1441 | 2 | 1056 |
| $\mathbf{1 9 9 7}$ | 15 | 5 | 1034 | 957 | 1111 | 7 | 1021 |
| $\mathbf{1 9 9 6}$ | 16 | 2 | 1078 | 989 | 1167 | 2 | 1079 |
| $\mathbf{1 9 9 1}$ | 21 | 1 | 1260 | --- | --- | 1 | 1260 |
| Total |  | $\mathbf{2 1 0}$ |  |  |  | 561 |  |

Table 3. Sub-sample and expanded mean weights (kg) by year-class of striped bass sampled from Atlantic coast fishery, November 2011-April 2012. Includes the lower and upper $95 \%$ confidence limits (LCL and UCL, respectively).

| Year <br> Class | Age | n Fish <br> Aged | Mean Weight <br> (kg) of Aged <br> sub-sample | LCL | UCL | Estimated \# <br> at-age in <br> sample | Expanded <br> Mean Weight <br> (kg) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 8}$ | 4 | 1 | 2.8 | --- | --- | 1 | 2.8 |
| $\mathbf{2 0 0 7}$ | 5 | 12 | 2.8 | 2.6 | 3.0 | 33 | 3.1 |
| $\mathbf{2 0 0 6}$ | 6 | 9 | 3.1 | 2.9 | 3.3 | 38 | 3.3 |
| $\mathbf{2 0 0 5}$ | 7 | 19 | 4.2 | 3.7 | 4.6 | 75 | 4.1 |
| $\mathbf{2 0 0 4}$ | 8 | 28 | 4.9 | 4.5 | 5.4 | 87 | 4.8 |
| $\mathbf{2 0 0 3}$ | 9 | 64 | 5.8 | 5.3 | 6.2 | 190 | 5.3 |
| $\mathbf{2 0 0 2}$ | 10 | 23 | 7.7 | 6.8 | 8.6 | 51 | 6.8 |
| $\mathbf{2 0 0 1}$ | 11 | 24 | 8.9 | 8.4 | 9.4 | 44 | 8.7 |
| $\mathbf{2 0 0 0}$ | 12 | 13 | 10.0 | 8.9 | 11.1 | 21 | 9.6 |
| $\mathbf{1 9 9 9}$ | 13 | 7 | 10.9 | 9.3 | 12.5 | 9 | 10.7 |
| $\mathbf{1 9 9 8}$ | 14 | 2 | 12.4 | --- | --- | 2 | 11.9 |
| $\mathbf{1 9 9 7}$ | 15 | 5 | 13.1 | 8.7 | 17.4 | 7 | 11.9 |
| $\mathbf{1 9 9 6}$ | 16 | 2 | 13.1 | 8.6 | 17.5 | 2 | 13.0 |
| $\mathbf{1 9 9 1}$ | 21 | 1 | 22.1 | --- | --- | 1 | 22.1 |
| Total |  | $\mathbf{2 1 0}$ |  |  |  | $\mathbf{5 6 1}$ |  |

Figure 1. Age distribution of striped bass sampled from the Atlantic coast fishery, 20062012 seasons.


Figure 2. Length distribution of striped bass sampled from the Atlantic coast fishery, 2006-2012 seasons.


Figure 3. Mean total lengths (mm TL) of the aged sub-sample, by year, for individual age-classes of striped bass (through age 12) sampled from the Maryland Atlantic coast trawl and gill net landings, 2006-2012 (95\% confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. *Note differences in scales on the $y$-axis.




## Season

Figure 3. Continued



Season

Figure 4. Mean weight (kg) of the aged sub-sample, by year, for individual age-classes of striped bass (through age 12) sampled from the Maryland Atlantic coast trawl and gill net landings, 2006-2012 (95\% confidence intervals are shown around each point). Expanded means (estimated from entire sample) are also shown. *Note differences of scale on the $y$-axis.


Figure 4. Continued


# PROJECT NO. 2 <br> JOB NO. 3 <br> TASK NO. 2 <br> \title{ CHARACTERIZATION OF STRIPED BASS SPAWNING STOCKS IN MARYLAND 

}

Prepared by Angela Giuliano and Beth A. Versak

## INTRODUCTION

The primary objective of Project 2, Job 3, Task 2 was to generate estimates of relative abundance-at-age for striped bass in Chesapeake Bay during the 2012 spring spawning season. Since 1985, the Maryland Department of Natural Resources (MD DNR) has e mployed multipanel experimental drift gill nets to monitor the Chesapeake Bay component of the Atlantic coast striped bass population. Because Chesapeake Bay spawners produce up to $90 \%$ of the Atlantic coastal s tock (Richards a nd R ago 1999), i ndices de rived from this e ffort a re i mportant in the coastal stock assessment process. Indices produced from this study are currently used to guide management decisions concerning recreational and commercial striped bass fisheries from North Carolina to Maine.

A secondary objective of Task 2 was to characterize the striped bass spawning population within $t$ he $C$ hesapeake $B$ ay. Length di stribution, a ge $s$ tructure, a verage 1 ength-at-age, a nd percentage of striped bass older than age 8 present on the spawning grounds were examined. In addition, a $n$ Index of Spawning Potential (ISP) for $f$ emale striped bass, an age-independent measure of female spawning biomass within the Chesapeake Bay, was calculated.

## METHODS

## Data Collection Procedures

Multi-panel experimental drift gill nets were deployed in the Potomac River and in the Upper Chesapeake B ay in 2012 (Figure 1). Gill nets were fished 6 da ys p er week, weather permitting, from late March through May. In the Potomac River, sampling was conducted from March 26 to May 7 for a total of 30 sample days. In the Upper Bay, sampling was conducted from March 30 to May 18 with a total of 37 sample days.

Individual n et pa nels were 150 f eet 1 ong , a nd r anged f rom 8.0 t o 11.5 f eet d eep depending on $m$ esh size. $T$ he panels were constructed of multifilament nylon webbing in 3.0, $3.75,4.5,5.25,6.0,6.5,7.0,8.0,9.0$ and 10.0 -inch stretch-mesh. In the Upper Bay, all 10 panels were tied together, end to end, to fish the entire suite of meshes simultaneously. In the Potomac River, because of the design of the fishing boat, the gang of panels was split in half, with two suites of panels ( 5 meshes tied together) fished simultaneously end to end. In both systems, all 10 panels were fished twice daily unless weather prohibited a second set. The order of panels within the suite of nets was randomized with gaps of 5 to 10 feet between each panel. Overall soak times for each panel ranged from 4 to 109 minutes.

Sampling locations were assigned using a stratified random design. The Potomac River and Upper Bay spawning areas were each considered a stratum. One randomly chosen site per day was fished in each spawning area. Sites were chosen from a grid superimposed on a map of each system. The P otomac R iver gr id c onsisted of $40,0.5$-square-mile qua drants, while the upper Bay grid consisted of 31, 1-square-mile quadrants. GPS equipment, buoys, and landmarks were used to locate the appropriate quadrant in the field. O nce in the designated quadrant, air and surface water temperatures, surface salinity, and water clarity (Secchi depth) were measured.

All striped bass captured in the nets were measured for total length (mm TL), sexed by expression of gonadal products, and released. Scales were taken from 2-3 randomly chosen male striped bass per 10 mm length group, per week, for a maximum of 10 s cale samples per length group over the entire season. Scales were also taken from all males over 700 mm TL and from all f emales r egardless of tot al le ngth. Scales were r emoved f rom the le ft s ide of the fish, between t he l ateral 1 ine a nd t he f irst dor sal fin. Additionally, if time a nd fish c ondition permitted, U. S. Fish and Wildlife Service internal anchor tags were applied (Project No. 2, Job No. 3, Task 4).

## Analytical Procedures

## Development of age-length keys

Sex-specific age-length keys (ALKs) were used to develop catch-per-unit-effort (CPUE) estimates. The scale allocation procedure, in use since 2003, designated two sex-specific groups of scales pooled from both the spring gill net sampling and the spring striped bass recreational season creel survey (Project No. 2, Job No. 3, Task 5B; Barker et al., 2003).

## Development of selectivity-corrected CPUEs and variance estimates

CPUEs for individual mesh sizes and length groups were calculated for each spawning area. C PUE was standardized a st he num ber of f ish c aptured in 1000 s quare yards of experimental drift g ill n et per hour. M esh-specific C PUEs w ere calculated by s umming the catch in each length group across days and meshes, and dividing the result by the total effort for each mesh. This $r$ atio of $s$ ums a proach $w$ as a ssumed $t o p r o v i d e t$ he $m$ ost a ccurate characterization of the s pawning popul ation, which exhibits a hi gh de gree of e migration a nd
immigration from the sampling area during the two-month sampling interval. The dynamic state of $t$ he s pawning popul ation precludes obt aining an instantaneous, representative sample on a given day, whereas a sum of the catches absorbs short-term variability and provides a cumulative 'snap-shot' of spawning stock density. In addition, it was necessary to compile catches across the dur ation of $t$ he s urvey in each 1 ength gr oup, s o that $s$ ample sizes w ere 1 arge enough to characterize gill net selectivity.

Sex-specific models ha ve be en us ed since 2000 to de velop selectivity coefficients for female and male fish sampled from the P otomac R iver a nd Upper Bay. M odel building and hypothesis $t$ esting determined $t$ hat uni que physical selectivity characteristics $w$ ere evi dent by sex, but not bya rea (Waller 2000, unpubl ished da ta). T herefore, s ex-specific s electivity coefficients for each mesh and length group were estimated by fitting a skew-normal model to spring data from 1990 to 2000 (Helser et al., 1998).

Sex-specific selectivity coefficients were used to correct the mesh-specific length group CPUE e stimates. T he selectivity-corrected C PUEs w ere $t$ hen av eraged across $m$ eshes and weighted by $t$ he capt ure ef ficiency of $t$ he $m$ esh, $r$ esulting in a ve ctor of selectivity-corrected length group C PUEs f or e ach s pawning a rea and sex. These two sex-specific s electivity coefficients have been used since 2000 .

Sex-specific ALKs were applied to the appropriate vectors of selectivity-corrected length group C PUEs to attain estimates of s electivity-corrected year-class C PUEs. S ex- and areaspecific, selectivity-corrected, year-class C PUEsw ere calculated usingt he s kew-normal selectivity model. These area- and sex-specific estimates of relative abundance were pooled to develop estimates of relative abundance for Maryland's Chesapeake Bay. Before pooling over spawning areas, weights corresponding to the fraction of total spawning habitat encompassed by
each spawning ar ea w ere as signed. The C hoptank $R$ iver ha s not be en $s$ ampled s ince 1996, therefore, values for 1997 to the present were weighted using only the Upper Bay (0.615) and the Potomac River (0.385; Hollis 1967). In order to incorporate Bay-wide indices into the coastal assessment model, 15 age-specific indices were developed, one for each age from age 1 through age 15 -plus.

Confidence 1 imits for the i ndividual sex- and area-specific C PUEs are pr esented. In addition, confidence limits for the pooled age-specific CPUE estimates are produced according to $t$ he methods presented in $C$ ochran (1977), ut ilizing e stimation of va riance $f$ or va lues developed from stratified r andom sampling. D etails of this procedure c an be found in B arker and Sharov (2004).

Finally, a dditional s pawning s tock analyses for C hesapeake Bay s triped bass w ere performed, including:

- Development of daily water and air temperature and catch patterns to examine patterns and relationships;
- Examination of the spawning stock length-at-age (LAA) structure among areas and over time, a nd calculation of confidence i ntervals for sex- and area-specific l ength-at-age ( $\alpha=0.05$ );
- Examination of $t$ rends in $t$ he a ge $c$ omposition of $t$ he $B$ ay spawning s tock a nd the percentage of the female s pawning stock older than age 8, and calculation of the total stock older than age 8 ;
- Development of an index of spawning potential (ISP) for each system by converting the selectivity-corrected 1 ength group C PUE of female striped ba ss ove r 500 mm TLto biomass utilizing the regression equation (Rugolo and Markham 1996):
$\ln$ weight $_{\mathrm{kg}}=2.91 * \ln$ length $_{\mathrm{mm}}-11.08$


## RESULTS AND DISCUSSION

## CPUEs and variance

A total of 624 scales were aged to create the sex-specific ALKs (Table 1). Annual CPUE calculations produced four vectors of selectivity-corrected sex- and age-specific CPUE values. The un-weighted time series data are presented by area in Tables 2-7.

The 2012 un-weighted CPUE for Potomac females (22) ranked fourteenth of 27 years in the time series, below the series average of 27, but was double the value from 2011 (Table 2). The un-weighted CPUE for Potomac males (123) ranked twenty-fourth in the time-series, and well be low the a verage of 433. The three values in that time series lower than 2012 have all occurred within the last seven years. The Upper Bay female CPUE (87) was the highest in the 28 year time series and well above the time series ave rage of 37 (Table 4). The un-weighted CPUE for Upper Bay males (252) was ranked twenty-third in the time series, a decrease from the last several years and well below the time series average of 445 (Table 5). The Choptank River has not been sampled since 1996 (Tables 6 and 7).

Area and sex-specific, weighted C PUE va lues w ere pool ed $f$ or us e in $t$ he a nnual coastwide striped bass stock assessment. These indices are presented in a time series for ages one through $15+$ (Table 8 ). The 2012 selectivity-corrected, total, weighted CPUE (265) w as twenty-seventh in the 28 year time series and well below the time series average of 487 .

Confidence limits were calculated for the pooled and weighted CPUEs (Tables 9 and 10). Confidence limits could not be calculated for the 15+ age group in years when these values are the sum of multiple age-class CPUEs. Coefficients of Variation (CV) of the 2012 age-specific CPUEs were all below 0.20 and indicated a small variance in CPUE. Historically, $80 \%$ of the CV values were less than 0.10 and $89 \%$ were less than 0.25 (Table 11 ). CV values greater than
1.0 were limited to older age-classes sampled during and immediately following the moratorium. The increased variability was likely attributed to small sample sizes associated with those older age-classes when the population size was low.

In both systems, males dominated both the un-weighted and weighted (Tables 12 and 13), total CPUEs. However, in 2012, the female contribution in each system was higher than usual. In the Upper Bay, females made up $26 \%$ of the un-weighted and weighted total CPUEs, and $15 \%$ on the Potomac. H istorically the female contribution is usually less than $10 \%$ in each system. Three year old males from the 2009 year-class contributed $20 \%$ of the total weighted CPUE and $21 \%$ of the un-weighted in 2012 . Potomac River fish only contributed $30 \%$ to the tot al unweighted and $21 \%$ to the weighted CPUEs, unlike in 2011 when they contributed $53 \%$ to the unweighted and $41 \%$ to the weighted CPUEs, respectively.

The 2009 year-class replaced t he 2007 year-class as the 1 argest c ontributor to male CPUE, making up $23 \%$ of the Upper Bay male CPUEs and $36 \%$ of the Potomac male CPUEs. In the P otomac R iver, $75 \%$ of the male CPUEs were made up of fish age 5 and younger. The Upper Bay male CPUEs were more evenly distributed over a wide range of ages.

Female CPUEs were distributed across many year-classes in both systems. Four year old females were again present in the Potomac River, but not in the Upper Bay. In the Upper Bay, female fish age 7 and younger made up only $16 \%$ of the female CPUEs, while on the Potomac River these young females contributed $45 \%$ to the female CPUEs. The $15+$ age group, which includes the record 1996 year-class was the largest contributor (22\%) to the female Upper Bay CPUEs, followed b y a ge 9 f emales from the above a verage 2003 year-class (17\%). In t he Potomac River, the contribution of the $15+$ females to the female CPUEs was lower ( $14 \%$ to unweighted and $13 \%$ to weighted). The highest contribution to female CPUE in the Potomac River
was from age 6 fish from the below average 2006 year-class, which c ontributed $22 \%$ to the female CPUEs.

## Temperature and catch patterns

The NOAA National Climatic Data Center (2012) documented January through April of 2012 as the warmest on record and driest since 1985 for Maryland. Due to the warm weather, both systems started at the earliest date in the time series. In both systems, wide fluctuations in air temperatures were observed, likely due to differences in daily sampling time.

Daily surface water temperatures on the Potomac River ranged from $13.6^{\circ} \mathrm{C}$ to $19.2^{\circ} \mathrm{C}$. The survey started with the water temperature at $16.6^{\circ} \mathrm{C}$, the highest starting temperature in the 27 year time series. While water temperatures increased a few degrees over the cour se of the survey, they were fairly stable throughout. Female CPUE peaked between April 13 and April 20 (Figure 2). This peak in female CPUE corresponds roughly with a peak of male CPUE on April 17, suggesting possible spawning activity. As opposed to previous years when males are present in the survey area after females, male CPUEs were almost nonexistent past April 21 despite the presence of some females still on the spawning grounds. Because the water temperatures at the beginning of the survey were above the $14^{\circ} \mathrm{C}$ needed to initiate spawning (Fay et al. 1983), it is possible that some spawning activity occurred prior to the start of the survey.

Surface water temperatures on $t$ he Upper Bay during the spawning survey ranged from $11.7^{\circ} \mathrm{C}$ to $20.8^{\circ} \mathrm{C}$. Upper Bay water temperatures increased gradually throughout the spawning survey. Water temperatures surpassed $14^{\circ} \mathrm{C}$ on April 17. Peaks in female CPUE oc curred on April 8, 17, and 20 (Figure 3). These dates also had the highest CPUEs for male striped bass in the Upper Bay. These observations suggest spawning activity in early to mid-April. Similar to
the Potomac River, CPUEs for both sexes dropped off after April 21 suggesting the majority of spawning activity had concluded by this date.

## Length composition of the stock

In 2012, 808 male and 172 female striped bass were measured. On the Potomac River, 313 male and 40 female striped bass were sampled; 495 males and 132 females were sampled from the Upper Bay (Figure 4). The mean length of female striped bass in $2012(929 \pm 21 \mathrm{~mm}$ TL) w as 1 arger t han the m ean length of m ale s triped $\operatorname{bass}(578 \pm 12 \mathrm{~mm} \mathrm{TL}, \mathrm{P}<0.0001$ ), consistent with the known biology of the species. Mean lengths a re reported with 2 standard errors.

Mean lengths of male striped bass collected from the Potomac River ( $492 \pm 15 \mathrm{~mm} \mathrm{TL}$ ) and upper Bay ( $613 \pm 16 \mathrm{~mm} \mathrm{TL}$ ) were significantly different $(\mathrm{P}<0.0001)$ in 2012. The majority of males caught on t he Potomac River in 2012 were be tween 390 and 590 mm TL while the Upper Bay male length distribution was much wider and included many more fish between 610 and 830 mm TL (Figure 4).

Male s triped ba ss on t he P otomac r anged f rom 290 to 1138 mm TL. The 1 ength distribution was heavily influenced by the contribution of striped bass from the 2007 through 2010 year-classes. Male s triped bass between 390 and $590 \mathrm{~mm} \mathrm{~T} \mathrm{~L} \mathrm{c} \mathrm{omposed} 71 \%$ of t he Potomac River male catch in 2012 (Figure 4). The unc orrected Potomac male CPUE peaked between 330 and 470 mm T L, r epresenting a combination of the 2008, 2009 and 2010 yearclasses (Figure 5). The selectivity-corrected Potomac male CPUE peaked between 330 and 390 mm T L, increasing the contribution of the younger 2009 a nd 2010 year-classes. This could indicate that the smaller fish are not captured efficiently in the sampling gear

Male striped bass on $t$ he Upper B ay ranged from 268 to 1087 mm TL. Sizes of male Upper Bay fish were ev enly di stributed with one distinct peak between 770 a nd 830 m m TL. This peak coincides with the above average 2003 year-class (Figure 4). Male striped bass CPUE in the Upper Bay was higher across a wide range of sizes, whereas the majority of the Potomac River male C PUE oc curred b etween 290 a nd 650 mm TL (Figure 5). The prominent yearclasses of 2009, 2007, 2005, 2003, a nd 2001 were cl early visible in the selectivity-corrected CPUEs. These year-classes, with the exception of 2009, were all above average.

Female striped bass sampled from the Potomac River and Upper Bay in 2012 were not significantly di fferent in mean total le ngth $(\mathrm{P}=0.84)$. Female s triped bass s ampled from t he Potomac ranged from 468 to 1197 mm TL (mean $=924 \pm 55 \mathrm{~mm} \mathrm{TL}$ ), while females sampled in the Upper Bay ranged from 544 to 1196 mm TL (mean=931 $\pm 22 \mathrm{~mm}$ TL; Figure 4).

There were few discernable peaks in female CPUE by length group the Potomac River in 2012. The CPUE observed in the 470 mm TL length group represents the one 4 year old female caught on the Potomac River. The selectivity-corrected CPUE peaks in the 530 through 730 mm TL length groups are a combination of six and seven year old females (Figure 6). The remainder of the Potomac River female CPUE was distributed over length groups from 870 to 1190 mm TL.

In the Upper B ay, female corrected and unc orrected CPUEs covered a wide r ange of length groups. Application of the selectivity model to the data corrected the catch upward in the extreme ends of the length distribution where few fish were en countered. Large num bers of females were captured in 2012, resulting in a higher than normal CPUEs. The youngest female, in the 550 m m TL length group, was from the 2007 year-class. Peaks in selectivity-corrected CPUEs between 610 and 670 mm TL were composed of fish from the 2005 a nd 2006 year-
classes. The peaks in the larger size groups were a combination of 11 to 19 year old fish from the 2001 through 1993 year-classes.

## Length at age (LAA)

Based on pr evious investigations which indicated no influence of a rea on m ean LAA, samples from the Potomac River, Upper Bay and the spring recreational creel sampling (Project 2, Job 3, T ask 5B) were a gain combined in 2012 to produce separate male and female ALKs (Warner et al., 2006, Warner et al., 2008, Giuliano and Versak 2012).

Age and sex-specific LAA statistics are presented in Tables 14 a nd 15. Small sample sizes at age in both systems precluded testing for differences in LAA relationships in some cases. When year-classes a re small or at the ex tremes in age, sample sizes are too small to a nalyze statistically. This is the case particularly for female striped bass, as they are encountered much less frequently on the spawning grounds. A two-way analysis of variance was performed, where possible, to determine differences in LAA between areas (Upper Bay a nd P otomac). No differences between sample areas were detected in LAA for either sex in 2012 ( $\mathrm{P}>0.05$ ) except for 6 and 14 year old males. Six year old males were significantly longer on the Upper Bay (641 mm TL) than the Potomac ( $572 \mathrm{~mm} \mathrm{TL}, \mathrm{P}=0.05$ ). Fourteen year old males were significantly larger on the Potomac (1138 mm TL) than the Upper Bay ( $991 \mathrm{~mm} \mathrm{TL}, \mathrm{P}=0.03$ ), however the Potomac sample size was just one fish which may not be representative of all 14 year old male fish on the Potomac River.

When comparing LAA between years, only gill net fish were us ed. Male and female LAA has been relatively stable since the mid 1990s (Figures 7 a nd 8). Mean lengths of males were similar in 2011 and 2012 for all ages except for age 7 (ANOVA, $\alpha=0.05, \mathrm{P}=0.003$ ). Mean
lengths of females were similar in 2011 and 2012 for all ages that could be tested except for age 12 (ANOVA, $\alpha=0.05, \mathrm{P}=0.03$ ).

## Age composition of the stock

During the 2012 survey, eighteen age-classes, ranging from 2 to 19 were enc ountered (Tables 14 and 15). Male striped bass ranged from ages 2 to 15 , with ages 8 and 9 fish (2004 and 2003 year-classes) being the most abundant male cohorts. The majority of females were ages 9 to 14, with most of the females collected at age 9 (2003 year-class). The abundance of ages 2 to 5 striped bass in the Maryland Chesapeake Bay spawning stock has been variable since 1985, with clear peaks of abundance corresponding to strong year-classes (Figure 9). In 2012, the largest increase in age-specific CPUE was indicated by the age 11 (2001 year-class) cohort. While all age-specific CPUEs for age 8 and younger showed a decline in 2012, the majority for fish age 9 and older showed an increase. The 1996 year-class has now moved into the 15+ age group, and their contribution is still evident (Figure 9).

In 2012 , the contribution of age $8+$ females to the female spawning stock increased to $80 \%$ (Figure 10). The contribution of females age 8 and older to the spawning stock has been at or above $80 \%$ since 1996 , with the exception of 2011 . Some decline is expected based on the results of the most recent coastwide stock assessment, which showed that female spawning stock biomass has been declining coastwide (ASMFC 2011).

The percentage of the overall sample (males and females combined) age 8 and older has been variable since 1997 (Figure 11). However the 2012 value of $41 \%$ is the highest in the 28 year time series. The percentage of age $8+$ fish is heavily influenced by strong year-classes and shows cyclical variations (Figure 9). While the percentage of age $8+$ females showed a modest
increase in 2012, the sexes-combined sample of age $8+$ fish showed a large increase. This was due to a combination of a large number of older males encountered in the Upper Bay and low recruitment to the spawning stock of three year-classes since 2005.

Historically, C hesapeake B ay es timates of ISP, expressed as bi omass, have followed trends similar to the coastal estimates. Recent estimates of spawning stock biomass (SSB) for coastal females have shown a decline over the past several years (ASMFC 2011). The MD DNR estimate of ISP generated from the upper Bay has been variable, but in 2012 the ISP value of 799 was the highest on record, well above the time-series average of 301 (Table 16, Figure 12). The 2012 Potomac River female ISP increased slightly to 150 , but was still well below the time series average of 231.

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Table 2. Estimates of selectivity-corrected age-class CPUE by year for female striped bass captured in the Potomac River during the 1985 - 2012 spawning stock surveys. CPUE is $s$ tandardized as the num ber of fish captured in 1000 s quare yards of experimental drift gill net per hour. The Potomac River was not sampled in 1994.

Table 3. Estimates of selectivity-corrected age-class CPUE by year for male striped bass captured in the Potomac River during the 1985 - 2012 spawning stock surveys. CPUE is standardized as the num ber of fish captured in 1000 s quare yards of experimental drift gill net per hour. The Potomac River was not sampled in 1994.

Table 4. Estimates of selectivity-corrected age-class CPUE by year for female striped bass captured inthe Upper Bay dur ing the 1985 - 2012 spawning stock surveys. CPUE is $s$ tandardized as $t$ he num ber of fish $c$ aptured in 1000 s quare yards of experimental drift gill net per hour.

Table 5. Estimates of selectivity-corrected age-class CPUE by year for male striped bass captured inthe Upper Bay dur ing the 1985 - 2012 spawning stock surveys. CPUE is $s$ tandardized as the num ber of fish captured in 1000 s quare yards of experimental drift gill net per hour.

Table 6. Estimates of selectivity-corrected age-class CPUE by year for female striped bass captured in the Choptank River during the 1985 - 1996 spawning stock surveys. CPUE is standardized as the num ber of fish captured in 1000 s quare yards of experimental dr ift g ill n et per hour. T he C hoptank R iver w as not s ampled in 1995, and has not been sampled since 1996.

Table 7. Estimates of selectivity-corrected age-class CPUE by year for male striped bass captured in the Choptank River during the 1985 - 1996 spawning stock surveys. CPUE is standardized as the num ber of fish captured in 1000 s quare yards of experimental dr ift g ill n et per hour. T he C hoptank R iver w as not s ampled in 1995, and has not been sampled since 1996.

Table 8. Mean values of the annual, pooled, weighted, age-specific CPUEs (1985-2012) for the Maryland Chesapeake Bay striped bass spawning stock. CPUE is reported as the number of fish captured in 1000 square yards of net per hour.

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Table 12. Un-weighted striped bass catch per unit effort (CPUE) by year-class, late March through M ay 2012. V alues ar e pr esented by sex, area, and percent of t otal. CPUE is number of fish per hour in 1000 yards of experimental drift net.

Table 13. Striped bass catch per unit ef fort (CPUE) by year-class, w eighted by s pawning area, late March through May 2012. Values are presented as percent of total, sexspecific, and area-specific CPUE. CPUE is number of fish per hour in 1000 yards of experimental drift net.

Table 14. Mean length-at-age ( mm TL) statistics for the aged sub-sample of male striped bass collected in the Potomac River and the Upper Bay, and areas combined, late March through May 2012.

Table 15. Mean length-at-age ( mm TL) statistics 0 for the aged sub-sample of female striped bass collected in the Potomac River and the Upper Bay, and areas combined, late March through May 2012.

Table 16. Index of s pawning bi omass by year, for female striped $b a 5 \$ 0 \mathrm{~mm}$ TL sampled from s pawning a reas of the C hesapeake B ay dur ing M arch, A pril a nd May since 1985. T he index is selectivity-corrected CPUE converted to biomass $(\mathrm{kg})$ using parameters from a length-weight regression.

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Figure 1. Drift gill net sampling locations in spawning areas of the Upper Chesapeake Bay and the Potomac River, late March - May 2012.

Figure 2. Daily effort-corrected catch of female and male striped bass, with surface water and a ir temperatures in the s pawning reach of t he P otomac R iver, late M arch through May 2012. E ffort is standardized as 1000 square yards of experimental drift gill net per hour. Note different scales.

Figure 3. Daily effort-corrected catch of female and male striped bass, with surface water and air temperatures in the spawning reach of the Upper Chesapeake Bay, late March through M ay 2012. E ffort is standardized a s 1000 s quare yards of experimental drift gill net per hour. Note different scales.

Figure 4. Length frequency of male and female striped bass from the spawning areas of the Upper Chesapeake Bay and Potomac River, late March through May 2012.

Figure 5. Length group C PUE ( uncorrected and corrected for $g$ ear s electivity) of $m$ ale striped bass collected from spawning areas of the Upper Bay and Potomac River, late March - May 2012. C PUE is the number of fish captured per hour in 1000 square yards of experimental drift gill net. Note different scales.

Figure 6. Length group CPUE (uncorrected and corrected for gear selectivity) of female striped bass collected from spawning areas of the Upper Bay and Potomac River, late March - May 2012. CPUE is the number of fish captured per hour in 1000 square yards of experimental drift gill net.

Figure 7. Mean length (mm TL) by year for individual ages of male striped bass sampled from spawning areas of the Potomac River and Upper Chesapeake Bay during late March through May, 1985-2012. Error bars are $\pm 1$ standard error (SE). Note the Potomac River was not sampled in 1994. *Note different scales.

Figure 8. Mean length (mm TL) by year for individual ages of female striped bass sampled from spawning areas of the Potomac River and Upper Chesapeake Bay during late March through May, 1985-2012. Error bars are $\pm 1$ standard error (SE). Note the Potomac River was not sampled in 1994. *Note different scales.

Figure 9. Maryland Chesapeake Bay spawning stock indices used in the coastal assessment. These are selectivity-corrected estimates of CPUE by year for ages 2 through 15plus. A reas and sexes are pooled, although the contribution of sexes is shown in the stacked bars. Note different scales.

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Figure 10. Percentage ( selectivity-corrected CPUE) of female striped bass that were age 8 and older sampled from experimental drift gill nets set in spawning reaches of the Potomac R iver, C hoptank R iver a nd the Upper Chesapeake B ay, late M arch through M ay, 1985-2012 (Choptank R iver to 1996). E ffort is standardized as 1000 square yards of net per hour. Area-specific indices were weighted based on the relative size of the spawning areas before area-specific indices were pooled.

Figure 11. Percentage ( selectivity-corrected CPUE) of male and female striped bass age 8 and over sampled from experimental drift gill nets set in spawning reaches of the Potomac R iver, C hoptank R iver a nd the Upper Chesapeake B ay, late M arch through M ay, 1985-2012 (Choptank R iver to 1996). E ffort is standardized a s 1000 square yards of net per hour. Area-specific indices were weighted based on the relative size of the spawning areas before area-specific indices were pooled.

Figure 12. Biomass (kg) of female striped bass greater than or equal to 500 mm TL collected from experimental drift gill nets fished in two spawning a reas of the Maryland Chesapeake B ay dur ing late M arch through May, 1985-2012. T he i ndex is corrected for gear selectivity, and bootstrap $95 \%$ confidence intervals are shown around each point.

Table 1. Number of scales aged per sex, area, and survey, by length group (mm TL), in 2012.

|  | MALES |  |  |  | FEMALES |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Length group (mm) | Upper Bay | Potomac River | Creel | Male <br> Total | Upper Bay | Potomac River | Creel | Female Total |
| 270 | 1 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 290 | 0 | 2 | 0 | 2 | 0 | 0 | 0 | 0 |
| 310 | 2 | 3 | 0 | 5 | 0 | 0 | 0 | 0 |
| 330 | 3 | 3 | 0 | 6 | 0 | 0 | 0 | 0 |
| 350 | 3 | 3 | 0 | 6 | 0 | 0 | 0 | 0 |
| 370 | 3 | 3 | 0 | 6 | 0 | 0 | 0 | 0 |
| 390 | 3 | 3 | 0 | 6 | 0 | 0 | 0 | 0 |
| 410 | 3 | 3 | 0 | 6 | 0 | 0 | 0 | 0 |
| 430 | 3 | 3 | 0 | 6 | 0 | 0 | 1 | 1 |
| 450 | 3 | 3 | 0 | 6 | 0 | 0 | 6 | 6 |
| 470 | 3 | 3 | 1 | 7 | 0 | 1 | 8 | 9 |
| 490 | 3 | 3 | 0 | 6 | 0 | 0 | 9 | 9 |
| 510 | 3 | 3 | 0 | 6 | 0 | 0 | 9 | 9 |
| 530 | 3 | 3 | 0 | 6 | 0 | 1 | 9 | 10 |
| 550 | 3 | 3 | 0 | 6 | 1 | 0 | 7 | 8 |
| 570 | 6 | 5 | 0 | 11 | 0 | 0 | 10 | 10 |
| 590 | 5 | 5 | 0 | 10 | 2 | 0 | 5 | 7 |
| 610 | 5 | 5 | 0 | 10 | 0 | 0 | 6 | 6 |
| 630 | 5 | 5 | 0 | 10 | 3 | 0 | 5 | 8 |
| 650 | 8 | 2 | 0 | 10 | 0 | 1 | 2 | 3 |
| 670 | 10 | 0 | 0 | 10 | 3 | 1 | 4 | 8 |
| 690 | 9 | 1 | 0 | 10 | 2 | 1 | 4 | 7 |
| 710 | 10 | 0 | 5 | 15 | 1 | 1 | 4 | 6 |
| 730 | 10 | 0 | 5 | 15 | 2 | 1 | 1 | 4 |
| 750 | 8 | 0 | 3 | 11 | 0 | 0 | 0 | 0 |
| 770 | 8 | 2 | 5 | 15 | 1 | 0 | 0 | 1 |
| 790 | 10 | 1 | 4 | 15 | 2 | 0 | 2 | 4 |
| 810 | 8 | 3 | 5 | 16 | 1 | 0 | 5 | 6 |
| 830 | 6 | 4 | 5 | 15 | 4 | 0 | 4 | 8 |
| 850 | 10 | 0 | 1 | 11 | 5 | 0 | 10 | 15 |
| 870 | 5 | 3 | 4 | 12 | 5 | 3 | 7 | 15 |
| 890 | 6 | 2 | 2 | 10 | 9 | 1 | 5 | 15 |
| 910 | 2 | 0 | 1 | 3 | 7 | 3 | 5 | 15 |
| 930 | 8 | 1 | 0 | 9 | 10 | 0 | 5 | 15 |
| 950 | 1 | 0 | 0 | 1 | 6 | 4 | 5 | 15 |
| 970 | 5 | 1 | 1 | 7 | 8 | 3 | 2 | 13 |
| 990 | 2 | 0 | 0 | 2 | 9 | 2 | 4 | 15 |
| 1010 | 0 | 3 | 0 | 3 | 9 | 4 | 0 | 13 |
| 1030 | 4 | 0 | 0 | 4 | 3 | 1 | 0 | 4 |
| 1050 | 3 | 0 | 0 | 3 | 8 | 2 | 1 | 11 |
| 1070 | 1 | 0 | 0 | 1 | 1 | 2 | 1 | 4 |
| 1090 | 1 | 0 | 0 | 1 | 7 | 1 | 1 | 9 |
| 1110 | 0 | 0 | 0 | 0 | 1 | 2 | 0 | 3 |
| 1130 | 0 | 1 | 0 | 1 | 2 | 1 | 0 | 3 |
| 1150 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 1 |
| 1170 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 1 |
| 1190 | 0 | 0 | 0 | 0 | 3 | 2 | 0 | 5 |
| 1210 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| Total | 195 | 85 | 42 | 322 | 117 | 38 | 147 | 302 |

Table 2. Estimates of selectivity-corrected age-class CPUE by year for female striped bass captured in the Potomac River during the 1985-2012 spawning stock surveys. CPUE is standardized as the number of fish captured in 1000 square yards of experimental drift gill net per hour. The Potomac River was not sampled in 1994.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | $15+$ | Total |
| 1985 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | 0.5 | 0.2 | 0.0 | 0.2 | 0.1 | 0.1 | 0.0 | 0.5 | 0.0 | 0.6 | 2 |
| 1986 | 0.0 | 0.0 | 1.0 | 7.3 | 0.7 | 0.0 | 0.4 | 0.4 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 10 |
| 1987 | 0.0 | 0.0 | 0.0 | 2.9 | 6.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.8 | 0.0 | 10 |
| 1988 | 0.0 | 0.0 | 0.0 | 1.7 | 2.4 | 5.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.9 | 10 |
| 1989 | 0.0 | 0.0 | 0.0 | 0.0 | 6.9 | 4.7 | 4.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 16 |
| 1990 | 0.0 | 0.0 | 0.0 | 0.0 | 1.6 | 3.7 | 3.5 | 1.7 | 0.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 11 |
| 1991 | 0.0 | 0.0 | 0.0 | 0.0 | 0.6 | 0.6 | 1.5 | 2.0 | 6.6 | 0.3 | 1.8 | 0.0 | 0.0 | 0.0 | 0.6 | 14 |
| 1992 | 0.0 | 0.0 | 0.0 | 2.6 | 6.4 | 6.7 | 8.7 | $\begin{array}{r} 11 . \\ 4 \end{array}$ | 8.2 | 8.7 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 53 |
| 1993 | 0.0 | 0.0 | 0.0 | 1.0 | 8.2 | 7.7 | 9.4 | $\begin{array}{r} 15 . \\ 2 \end{array}$ | $\begin{array}{r} 14 . \\ 3 \end{array}$ | 8.6 | 4.3 | 0.0 | 0.0 | 0.0 | 0.0 | 69 |
| 1994 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1995 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 3.1 | 4.6 | 4.8 | 4.6 | 6.6 | 5.5 | 5.0 | 0.7 | 0.0 | 0.0 | 35 |
| 1996 | 0.0 | 0.0 | 0.0 | 0.0 | 0.8 | 0.2 | 3.9 | 7.1 | 6.8 | 8.8 | 5.4 | 8.1 | 3.3 | 0.0 | 0.0 | 45 |
| 1997 | 0.0 | 0.0 | 0.0 | 3.1 | 0.5 | 4.0 | 3.0 | 5.3 | 9.2 | $\begin{array}{r} 10 . \\ 2 \end{array}$ | 4.2 | 4.8 | 1.4 | 1.5 | 0.0 | 47 |
| 1998 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.8 | 0.3 | 1.0 | 3.2 | 2.7 | 4.4 | 4.6 | 1.6 | 0.7 | 0.0 | 19 |
| 1999 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 2.1 | 3.7 | 4.2 | 4.8 | 2.0 | 6.4 | 2.6 | 0.6 | 0.0 | 0.3 | 27 |
| 2000 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 7.4 | 1.4 | 2.4 | 7.8 | 1.2 | 1.4 | 5.1 | 0.0 | 27 |
| 2001 | 0.0 | 0.0 | 0.0 | 1.0 | 0.0 | 0.0 | 2.9 | 4.6 | 7.2 | 4.0 | 4.3 | 3.0 | 5.2 | 0.0 | 0.0 | 32 |
| 2002 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 1.0 | 3.1 | $\begin{array}{r} 12 . \\ 3 \\ \hline \end{array}$ | 5.9 | 5.5 | 2.7 | 6.0 | 1.8 | 2.2 | 40 |
| 2003 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.5 | 1.8 | 3.5 | 2.8 | 1.6 | 0.3 | 1.5 | 0.0 | 12 |
| 2004 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 1.6 | 2.8 | $\begin{array}{r} 13 . \\ 5 \end{array}$ | 6.3 | 8.6 | $\begin{array}{r} 11 . \\ 6 \\ \hline \end{array}$ | 6.6 | 3.5 | 4.8 | 1.3 | 61 |
| 2005 | 0.0 | 0.0 | 0.0 | 0.0 | 1.9 | 0.0 | 1.6 | 0.6 | 2.7 | 2.5 | 4.6 | 4.1 | 1.7 | 0.8 | 2.3 | 23 |
| 2006 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.2 | 0.8 | 6.3 | 9.2 | 4.1 | 5.1 | 9.6 | 2.3 | 6.5 | 44 |
| 2007 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | 0.4 | 0.9 | 1.4 | 3.2 | 7.5 | 4.5 | 1.4 | 3.8 | 3.2 | 26 |
| 2008 | 0.0 | 0.0 | 0.0 | 0.4 | 0.4 | 0.0 | 0.9 | 0.1 | 0.4 | 1.8 | 2.4 | 4.9 | 1.2 | 1.2 | 1.4 | 15 |
| 2009 | 0.0 | 0.0 | 0.3 | 0.0 | 0.5 | 0.5 | 0.3 | 2.6 | 4.3 | 1.9 | 2.3 | 1.9 | 4.6 | 1.2 | 1.4 | 22 |


| 2010 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | 2.1 | 2.3 | 0.7 | 1.5 | 2.2 | 5.9 | 4.1 | 19 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2011 | 0.0 | 0.0 | 0.1 | 0.8 | 0.4 | 0.0 | 0.0 | 0.9 | 0.4 | 2.0 | 1.1 | 1.1 | 1.1 | 0.4 | 2.6 | 11 |
| 2012 | 0.0 | 0.0 | 0.0 | 1.0 | 1.4 | 4.7 | 2.6 | 1.1 | 1.6 | 1.0 | 1.6 | 1.8 | 0.8 | 1.0 | 3.0 | 22 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 27 |

Table 3. Estimates of selectivity-corrected age-class CPUE by year for male striped bass captured in the Potomac River during the 1985-2012 spawning stock surveys. CPUE is standardized as the number of fish captured in 1000 square yards of experimental drift gill net per hour. The Potomac River was not sampled in 1994.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 + | Total |
| 1985 | 0.0 | $\begin{array}{r} 285 . \\ 3 \end{array}$ | $\begin{array}{r} 517 . \\ 6 \end{array}$ | 80.6 | 10.5 | 0.7 | 1.4 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 896 |
| 1986 | 0.0 | $\begin{array}{r} 241 . \\ 5 \end{array}$ | $\begin{array}{r} 375 . \\ 9 \end{array}$ | $\begin{array}{r} 531 . \\ \hline 2 \\ \hline \end{array}$ | 8.2 | 8.2 | 0.6 | 0.7 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 116 |
| 1987 | 0.0 | $\begin{array}{r} 144 . \\ 5 \\ \hline \end{array}$ | $\begin{array}{r} 283 . \\ 5 \\ \hline \end{array}$ | $\begin{array}{r} 174 . \\ 6 \end{array}$ | $\begin{array}{r} 220 . \\ 8 \\ \hline \end{array}$ | 3.6 | 1.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.6 | 829 |
| 1988 | 0.0 | 18.2 | $\begin{array}{r} 107 . \\ \hline \end{array}$ | 63.8 | 75.9 | $\begin{array}{r} 81 . \\ 2 \\ \hline \end{array}$ | 0.4 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 347 |
| 1989 | 0.0 | 51.9 | $\begin{array}{r} 240 . \\ 9 \end{array}$ | $\begin{array}{r} 134 . \\ 5 \end{array}$ | 39.1 | $\begin{array}{r} 55 . \\ \hline \end{array}$ | $\begin{array}{r} 21 . \\ 8 \end{array}$ | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 543 |
| 1990 | 0.0 | $114 .$ | $\begin{array}{r} 351 . \\ 8 \end{array}$ | $\begin{array}{r} 172 . \\ 8 \end{array}$ | 73.8 | $\begin{array}{r}28 . \\ 3 \\ \hline\end{array}$ | $\begin{array}{r} 33 . \\ 8 \end{array}$ | $\begin{array}{r} 26 . \\ 6 \\ \hline \end{array}$ | 1.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 803 |
| 1991 | 0.0 | 19.9 | 91.2 | 96.6 | 49.7 | 37. 8 | 28. 7 | $\begin{array}{r} 22 . \\ 3 \end{array}$ | 6.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 352 |
| 1992 | 0.3 | 36.3 | $\begin{array}{r} 202 . \\ 4 \end{array}$ | $\begin{array}{r} 148 . \\ 9 \end{array}$ | 97.6 | 73. 0 | $\begin{array}{r} 39 . \\ \hline 1 \end{array}$ | $\begin{array}{r} 19 . \\ 0 \end{array}$ | 6.1 | 0.8 | 8.4 | 0.0 | 0.0 | 0.0 | 0.0 | 632 |
| 1993 | 0.0 | 30.4 | 141. | $\begin{array}{r} 133 . \\ 9 \end{array}$ | $\begin{array}{r} 101 . \\ 4 \end{array}$ | $\begin{array}{r} 83 . \\ \hline 7 \end{array}$ | $\begin{array}{r} 62 . \\ 6 \end{array}$ | $\begin{array}{r} 43 . \\ 6 \end{array}$ | $\begin{array}{r} 21 . \\ 9 \end{array}$ | 1.8 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 621 |
| 1994 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1995 | 0.0 | 9.1 | $\begin{array}{r} 143 . \\ 9 \end{array}$ | 61.1 | 18.7 | $\begin{array}{r} 20 . \\ 4 \end{array}$ | $\begin{array}{r} 25 . \\ 3 \end{array}$ | $\begin{array}{r} 32 . \\ 2 \end{array}$ | $\begin{array}{r} 11 . \\ \hline \end{array}$ | $\begin{array}{r} 10 . \\ 7 \end{array}$ | 0.1 | 0.0 | 0.8 | 0.0 | 0.0 | 334 |
| 1996 | 0.0 | 0.0 | $\begin{array}{r} 230 . \\ 6 \end{array}$ | $\begin{array}{r} 172 . \\ \hline 9 \end{array}$ | 24.8 | $\begin{array}{r}26 . \\ 8 \\ \hline\end{array}$ | 17. | $\begin{array}{r} 22 . \\ \hline \end{array}$ | $\begin{array}{r} 19 . \\ \hline \end{array}$ | 3.6 | 0.6 | 0.8 | 0.0 | 0.0 | 0.0 | 520 |
| 1997 | 0.0 | 49.5 | 54.3 | $\begin{array}{r} 112 . \\ \hline 9 \end{array}$ | 95.7 | $\begin{array}{r} 12 . \\ \hline \end{array}$ | 5.7 | $\begin{array}{r} \hline 10 . \\ 8 \end{array}$ | $\begin{array}{r} 17 . \\ \hline \end{array}$ | $\begin{array}{r} 13 . \\ 6 \\ \hline \end{array}$ | 2.2 | 2.6 | 0.0 | 0.0 | 0.0 | 377 |
| 1998 | 0.0 | 72.9 | $\begin{array}{r} 200 . \\ 7 \end{array}$ | 29.8 | $\begin{array}{r} 128 . \\ 9 \end{array}$ | $\begin{array}{r}49 . \\ 8 \\ \hline\end{array}$ | $\begin{array}{r} 16 . \\ 9 \end{array}$ | $\begin{array}{r} 11 . \\ 7 \end{array}$ | 4.3 | 9.0 | 8.6 | 5.0 | 2.9 | 0.5 | 0.0 | 541 |
| 1999 | 0.0 | 9.9 | $\begin{array}{r} 316 . \\ 9 \end{array}$ | $151 .$ | $\begin{array}{r} 103 . \\ 6 \end{array}$ | $\begin{array}{r} 65 . \\ \hline \end{array}$ | $\begin{array}{r} 19 . \\ 1 \end{array}$ | $\begin{array}{r} 10 . \\ 3 \end{array}$ | 6.9 | 3.8 | 4.4 | 3.1 | 1.9 | 0.0 | 0.0 | 696 |
| 2000 | 0.0 | 1.9 | 42.2 | $\begin{array}{r} 136 . \\ 8 \end{array}$ | 48.5 | $18 .$ | $\begin{array}{r} 14 . \\ 8 \end{array}$ | 9.8 | 5.5 | 0.0 | 0.1 | 3.7 | 0.1 | 0.4 | 0.9 | 283 |
| 2001 | 0.0 | 10.6 | 36.1 | 43.5 | 33.8 | 12. | 8.9 | 7.8 | 4.8 | 1.7 | 2.2 | 4.0 | 0.8 | 0.6 | 0.0 | 167 |


|  |  |  |  |  |  | 6 |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2002 | 0.0 | 27.2 | 75.4 | 48.7 | 52.4 | $\begin{array}{r} 23 . \\ 0 \end{array}$ | $\begin{array}{r} 20 . \\ 9 \end{array}$ | 7.9 | 2.3 | 3.4 | 2.2 | 1.6 | 2.0 | 0.0 | 0.6 | 268 |
| 2003 | 0.0 | 12.6 | 79.0 | 39.6 | 24.5 | $\begin{array}{r} 31 . \\ 6 \end{array}$ | $\begin{array}{r} 22 . \\ 5 \end{array}$ | $\begin{array}{r} 10 . \\ 0 \end{array}$ | 7.0 | 9.5 | 3.2 | 3.7 | 5.8 | 0.2 | 0.2 | 249 |
| 2004 | 0.0 | 10.5 | $\begin{array}{r} 148 . \\ 8 \end{array}$ | 90.4 | 25.9 | 17. | $\begin{array}{r} 19 . \\ 5 \end{array}$ | $\begin{array}{r} 17 . \\ \hline \end{array}$ | 8.4 | 8.1 | 11. | 1.8 | 1.1 | 1.6 | 1.6 | 364 |
| 2005 | 0.0 | 10.9 | 11.0 | 14.9 | 16.3 | 4.7 | 4.5 | 3.6 | 4.1 | 3.1 | 1.9 | 1.2 | 0.0 | 0.0 | 0.0 | 76 |
| 2006 | 0.0 | 8.3 | $\begin{array}{r} 127 . \\ 1 \end{array}$ | 20.7 | 33.5 | $\begin{array}{r} 14 . \\ 5 \end{array}$ | 6.3 | 6.9 | 8.2 | 9.1 | 7.4 | 4.7 | 0.6 | 0.4 | 0.0 | 248 |
| 2007 | 0.0 | 10.4 | 16.6 | 37.1 | 5.3 | 5.6 | 4.3 | 2.1 | 2.6 | 2.8 | 5.4 | 1.0 | 0.8 | 2.0 | 0.1 | 96 |
| 2008 | 0.0 | 6.1 | 35.8 | 20.1 | 12.0 | 1.7 | 1.8 | 2.3 | 1.1 | 1.2 | 1.3 | 2.5 | 0.4 | 0.0 | 0.2 | 86 |
| 2009 | 0.0 | 35.2 | 35.9 | $\begin{array}{r} 116 . \\ 5 \end{array}$ | 23.1 | $\begin{array}{r} 56 . \\ \hline \end{array}$ | 9.1 | $\begin{array}{r} 10 . \\ 5 \end{array}$ | $\begin{array}{r} 10 . \\ 5 \end{array}$ | 2.8 | 3.8 | 2.6 | 3.7 | 0.6 | 0.6 | 312 |
| 2010 | 0.0 | 3.2 | $\begin{array}{r} 104 . \\ 9 \end{array}$ | 58.0 | 49.2 | $\begin{array}{r} 29 . \\ 7 \end{array}$ | $\begin{array}{r} 23 . \\ 9 \end{array}$ | 1.7 | 6.8 | 3.6 | 0.9 | 1.2 | 1.3 | 0.6 | 0.4 | 285 |
| 2011 | 0.0 | 27.6 | 95.7 | $\begin{array}{r} 164 . \\ 4 \end{array}$ | 51.2 | $\begin{array}{r} 54 . \\ 4 \\ \hline \end{array}$ | $\begin{array}{r} 29 . \\ 6 \\ \hline \end{array}$ | $\begin{array}{r} 24 . \\ \hline \end{array}$ | 6.2 | 5.2 | 6.1 | 4.1 | 4.9 | 2.1 | 5.3 | 481 |
| 2012 | 0.0 | 19.0 | 44.4 | 15.1 | 13.9 | 6.4 | 6.0 | 4.8 | 4.1 | 1.4 | 2.1 | 1.3 | 0.6 | 4.1 | 0.0 | 123 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 433 |

Table 4. Estimates of selectivity-corrected age-class CPUE by year for female striped bass captured in the Upper Bay during the 1985-2012 spawning stock surveys. CPUE is standardized as the number of fish captured in 1000 square yards of experimental drift gill net per hour.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | $15+$ | Total |
| 1985 | 0.0 | 0.0 | 0.8 | 0.0 | 0.3 | 0.1 | 0.5 | 0.0 | 0.1 | 0.0 | 0.0 | 0.2 | 0.0 | 0.0 | 0.3 | 2 |
| 1986 | 0.0 | 0.0 | 0.3 | $24 .$ | 0.0 | 0.0 | 0.5 | 0.5 | 3.8 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.3 | 30 |
| 1987 | 0.0 | 0.0 | 0.0 | 3.1 | $\begin{array}{r} 26 . \\ 8 \end{array}$ | 0.0 | 0.0 | 2.7 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 8.8 | 8.5 | 50 |
| 1988 | 0.0 | 0.0 | 4.2 | 8.8 | 6.5 | $\begin{array}{r} 31 . \\ 7 \end{array}$ | 1.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 52 |
| 1989 | 0.0 | 0.0 | 1.2 | 1.8 | 6.2 | 3.9 | 9.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 22 |
| 1990 | 0.0 | 0.0 | 0.0 | 0.3 | 0.0 | 0.3 | 1.8 | 5.3 | 0.0 | 0.0 | 0.0 | 0.9 | 0.6 | 0.0 | 0.0 | 9 |
| 1991 | 0.0 | 0.0 | 0.0 | 0.5 | 3.2 | 0.5 | 2.3 | 3.1 | 2.2 | 0.0 | 1.2 | 0.0 | 0.0 | 0.0 | 1.2 | 14 |
| 1992 | 0.0 | 0.0 | 0.2 | 4.4 | 3.5 | 5.6 | 4.4 | 4.9 | 4.3 | 4.2 | 0.3 | 0.0 | 0.5 | 1.1 | 0.4 | 34 |
| 1993 | 0.0 | 0.0 | 0.0 | 3.0 | 5.1 | 2.0 | 4.0 | 4.8 | 4.0 | 3.9 | 2.0 | 1.3 | 2.3 | 2.1 | 0.0 | 35 |
| 1994 | 0.0 | 0.0 | 0.0 | 0.4 | 0.8 | 3.0 | 1.3 | 2.9 | 1.5 | 2.9 | 1.1 | 0.0 | 0.0 | 0.0 | 0.0 | 14 |
| 1995 | 0.0 | 0.0 | 0.0 | 0.0 | 1.7 | $\begin{array}{r} 20 . \\ 2 \\ \hline \end{array}$ | $\begin{array}{r} 19 . \\ 5 \\ \hline \end{array}$ | 7.7 | $\begin{array}{r} 11 . \\ \hline 2 \\ \hline \end{array}$ | 5.2 | 5.7 | 2.0 | 7.0 | 0.0 | 0.0 | 80 |
| 1996 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 1.3 | $11 .$ | $\begin{array}{r} 10 . \\ 2 \end{array}$ | 6.4 | 5.4 | 7.0 | 1.8 | 0.0 | 0.0 | 0.0 | 43 |
| 1997 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 1.9 | $\begin{array}{r} 10 . \\ 9 \end{array}$ | $\begin{array}{r} 17 . \\ 9 \\ \hline \end{array}$ | 1.6 | 0.0 | 0.7 | 0.5 | 0.0 | 0.0 | 33 |
| 1998 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.7 | 5.0 | 2.6 | 5.2 | 1.3 | 1.3 | 0.0 | 0.0 | 0.5 | 17 |
| 1999 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 2.8 | 0.0 | 1.7 | 6.7 | 3.2 | 0.7 | 0.9 | 0.0 | 3.5 | 0.0 | 19 |
| 2000 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 2.2 | 3.3 | 1.0 | 3.0 | 5.9 | 2.5 | 5.7 | 0.1 | 0.3 | 0.0 | 24 |
| 2001 | 0.0 | 0.0 | 0.0 | 0.0 | 0.5 | 2.1 | 4.6 | $\begin{array}{r} 13 . \\ 5 \\ \hline \end{array}$ | 5.6 | 5.8 | 7.5 | 5.0 | 1.4 | 1.5 | 0.3 | 48 |
| 2002 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 6.9 | 1.1 | 3.1 | 9.0 | 2.6 | 2.3 | 2.0 | 1.6 | 0.8 | 0.0 | 29 |
| 2003 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 1.7 | 7.0 | 8.5 | 8.9 | $\begin{array}{r} 16 . \\ 8 \end{array}$ | $12 .$ | 4.3 | 3.9 | 2.6 | 0.0 | 66 |
| 2004 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.3 | 2.2 | 7.9 | $\begin{array}{r} 11 . \\ \hline 0 \end{array}$ | 7.2 | 9.4 | 3.0 | 1.5 | 0.5 | 3.0 | 46 |
| 2005 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.2 | 1.4 | 3.3 | 7.9 | 9.0 | 10. | 9.5 | 3.4 | 1.2 | 4.8 | 51 |


| 2006 | 0.0 | 0.0 | 0.0 | 0.0 | 2.8 | 4.2 | 3.1 | 0.3 | 4.3 | 6.2 | 3.2 | 5.4 | 7.4 | 1.8 | 5.9 | 45 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2007 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.5 | 3.4 | 2.8 | 4.3 | 5.5 | 11. | 5.0 | 1.3 | 3.8 | 7.1 | 45 |
| 2008 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.5 | 1.8 | 2.6 | 4.2 | 3.6 | 7.8 | 2.1 | 0.8 | 1.7 | 25 |
| 2009 | 0.0 | 0.0 | 0.0 | 0.0 | 3.2 | 3.8 | 0.2 | 2.9 | 8.5 | 2.8 | 6.6 | 4.8 | 10. | 3.8 | 5.1 | 52 |
| 2010 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 2.3 | 1.3 | 2.2 | 2.7 | 1.4 | 2.0 | 2.1 | 6.6 | 6.3 | 27 |
| 2011 | 0.0 | 0.0 | 0.0 | 4.9 | 2.0 | 1.2 | 1.3 | 6.4 | 1.3 | 2.5 | 1.2 | 1.0 | 2.1 | 1.2 | 2.2 | 27 |
| 2012 | 0.0 | 0.0 | 0.0 | 0.0 | 1.5 | 6.8 | 6.2 | 6.4 | 15. | 5.8 | 8.8 | 9.3 | 4.5 | 3.8 | 19. | 87 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 37 |

Table 5. Estimates of selectivity-corrected age-class CPUE by year for male striped bass captured in the Upper Bay during the 19852012 spawning stock surveys. CPUE is standardized as the number of fish captured in 1000 square yards of experimental drift gill net per hour.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 + | Total |
| 1985 | 0.0 | 47.5 | $\begin{array}{r} 148 . \\ 8 \\ \hline \end{array}$ | 1.9 | 0.0 | 0.8 | 0.5 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 199 |
| 1986 | 0.0 | $\begin{array}{r} 219 . \\ 0 \end{array}$ | $\begin{array}{r} 192 . \\ 3 \end{array}$ | $\begin{array}{r} 450 . \\ 8 \end{array}$ | 0.4 | 3.4 | 2.2 | 3.8 | 1.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 1.2 | 874 |
| 1987 | 0.0 | $\begin{array}{r} 131 . \\ 7 \end{array}$ | $\begin{array}{r} 231 . \\ 0 \end{array}$ | 68.1 | $\begin{array}{r} 138 . \\ 8 \end{array}$ | 0.0 | 2.1 | 4.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 576 |
| 1988 | 0.0 | 52.1 | 38.0 | 61.6 | 37.8 | 36.8 | 0.6 | 0.0 | 0.0 | 7.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 234 |
| 1989 | 0.0 | 8.1 | $\begin{array}{r} 102 . \\ 3 \end{array}$ | 17.4 | 21.1 | 26.9 | $\begin{array}{r} 16 . \\ 6 \end{array}$ | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 192 |
| 1990 | 0.0 | 56.7 | 28.4 | 92.8 | 20.1 | 24.9 | $\begin{array}{r} 22 . \\ 9 \end{array}$ | $\begin{array}{r} 16 . \\ 8 \end{array}$ | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 263 |
| 1991 | 0.0 | 84.1 | $\begin{array}{r} 254 . \\ 9 \end{array}$ | 36.8 | 40.9 | 11.3 | $\begin{array}{r} 16 . \\ 0 \\ \hline \end{array}$ | 9.5 | 4.3 | 0.1 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 458 |
| 1992 | 0.0 | 22.5 | $\begin{array}{r} 193 . \\ 9 \end{array}$ | $\begin{array}{r} 150 . \\ 1 \end{array}$ | 19.4 | 52.9 | 27. | $\begin{array}{r} 19 . \\ 1 \end{array}$ | 7.5 | 0.5 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 494 |
| 1993 | 0.0 | 30.6 | $\begin{array}{r} 126 . \\ 2 \end{array}$ | $\begin{array}{r} 149 . \\ 1 \end{array}$ | 63.0 | 16.3 | 27. | 9.9 | 7.5 | 0.5 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 430 |
| 1994 | 0.0 | 25.4 | 54.5 | 96.3 | $\begin{array}{r} 101 . \\ 8 \\ \hline \end{array}$ | 43.2 | $\begin{array}{r} 14 . \\ 5 \end{array}$ | $\begin{array}{r} 26 . \\ 8 \\ \hline \end{array}$ | 6.4 | 2.1 | 0.3 | 0.0 | 0.0 | 0.0 | 0.0 | 371 |
| 1995 | 0.0 | 79.0 | $\begin{array}{r} 108 . \\ 4 \end{array}$ | 75.8 | 89.8 | 52.9 | 30. | $\begin{array}{r} 11 . \\ 6 \\ \hline \end{array}$ | $\begin{array}{r} 12 . \\ 4 \end{array}$ | 3.7 | 7.2 | 0.9 | 0.0 | 0.0 | 0.0 | 471 |
| 1996 | 0.0 | 6.2 | $\begin{array}{r} 433 . \\ 5 \\ \hline \end{array}$ | 57.6 | 23.3 | 86.2 | 59. | $\begin{array}{r} 34 . \\ \hline \end{array}$ | $\begin{array}{r} 29 . \\ 0 \\ \hline \end{array}$ | $\begin{array}{r} 11 . \\ \hline 8 \\ \hline \end{array}$ | 12. | 0.0 | 0.6 | 0.0 | 0.0 | 753 |
| 1997 | 0.0 | 28.9 | 38.8 | $\begin{array}{r} 155 . \\ 5 \end{array}$ | 15.4 | 23.9 | 23. | $\begin{array}{r} 15 . \\ 0 \end{array}$ | 8.9 | 2.0 | 12. 1 | 0.0 | 0.7 | 0.0 | 0.0 | 325 |
| 1998 | 0.0 | 13.0 | $\begin{array}{r} 106 . \\ 6 \end{array}$ | 34.6 | $\begin{array}{r} 162 . \\ 0 \end{array}$ | 20.9 | $\begin{array}{r} 10 . \\ 0 \end{array}$ | $\begin{array}{r} 17 . \\ 1 \\ \hline \end{array}$ | $\begin{array}{r} 20 . \\ 9 \end{array}$ | $\begin{array}{r} 11 . \\ 9 \\ \hline \end{array}$ | 5.4 | 8.7 | 0.0 | 0.0 | 0.0 | 411 |
| 1999 | 0.0 | 7.7 | 81.8 | 33.6 | 30.4 | 14.6 | 4.8 | 0.6 | 4.7 | 1.6 | 0.4 | 0.2 | 0.3 | 0.0 | 0.0 | 181 |
| 2000 | 0.0 | 22.2 | 64.6 | 83.6 | 47.7 | 80.4 | 28. 0 | 10. | 6.1 | 6.2 | 3.9 | 3.3 | 1.4 | 0.4 | 0.3 | 359 |
| 2001 | 0.0 | 1.4 | 40.9 | 70.2 | 64.9 | 27.6 | 35. | $\begin{array}{r} 33 . \\ 0 \end{array}$ | 5.8 | 10. | 3.5 | 0.4 | 0.5 | 0.0 | 0.4 | 294 |


| 2002 | 0.0 | $\begin{array}{r} 120 . \\ 7 \\ \hline \end{array}$ | 19.1 | 34.1 | $\begin{array}{r} 106 . \\ 7 \end{array}$ | 48.2 | $\begin{array}{r} 42 . \\ 2 \\ \hline \end{array}$ | $\begin{array}{r} 43 . \\ \hline \end{array}$ | $\begin{array}{r} 20 . \\ 1 \end{array}$ | 5.2 | 2.4 | 1.1 | 1.9 | 0.0 | 0.0 | 445 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2003 | 0.0 | 17.7 | 131. <br> 9 | 62.1 | 42.2 | 89.8 | $\begin{array}{r} \hline 62 . \\ \hline \end{array}$ | $\begin{array}{r} 29 . \\ 7 \\ \hline \end{array}$ | $\begin{array}{r} 29 . \\ \hline \end{array}$ | $\begin{array}{r} 22 . \\ \hline \end{array}$ | 8.1 | 4.0 | 2.4 | 0.4 | 0.4 | 503 |
| 2004 | 0.0 | 40.3 | $\begin{array}{r} 221 . \\ \hline \end{array}$ | $\begin{array}{r} 140 . \\ 5 \end{array}$ | 52.7 | 44.0 | 56. | $\begin{array}{r} \hline 49 . \\ 7 \end{array}$ | $\begin{array}{r} 28 . \\ \hline \end{array}$ | $\begin{array}{r} 20 . \\ 0 \end{array}$ | $\begin{array}{r} 13 . \\ 7 \\ \hline \end{array}$ | 2.6 | 2.5 | 1.4 | 0.0 | 673 |
| 2005 | 0.0 | $\begin{array}{r} 100 . \\ 6 \end{array}$ | $\begin{array}{r} 161 . \\ 8 \end{array}$ | $\begin{array}{r} 110 . \\ 2 \end{array}$ | $\begin{array}{r} 145 . \\ 9 \end{array}$ | 36.3 | $\begin{array}{r} 36 . \\ 8 \end{array}$ | $\begin{array}{r} 29 . \\ \hline \end{array}$ | $\begin{array}{r} 32 . \\ 5 \end{array}$ | 20. | $\begin{array}{r} 14 . \\ 2 \end{array}$ | 5.7 | 0.3 | 0.0 | 0.0 | 694 |
| 2006 | 0.0 | 7.0 | $\begin{array}{r} 339 . \\ 9 \end{array}$ | 52.2 | 53.6 | 34.3 | $\begin{array}{r} 16 . \\ 9 \end{array}$ | $15$ | $\begin{array}{r} 16 . \\ 6 \end{array}$ | $\begin{array}{r} 17 . \\ \hline \end{array}$ | $\begin{array}{r} 11 . \\ \hline 0 \end{array}$ | 6.3 | 1.3 | 1.0 | 0.0 | 573 |
| 2007 | 0.0 | 6.3 | 26.2 | $\begin{array}{r} 100 . \\ 4 \end{array}$ | 20.9 | 20.8 | $\begin{array}{r} 15 . \\ \hline \end{array}$ | 7.3 | 7.8 | 7.1 | 6.5 | 4.5 | 2.2 | 1.4 | 0.2 | 227 |
| 2008 | 0.0 | 1.5 | $\begin{array}{r} 117 . \\ 5 \end{array}$ | $\begin{array}{r} 163 . \\ 5 \end{array}$ | $\begin{array}{r} 175 . \\ 0 \end{array}$ | 26.4 | $\begin{array}{r} 35 . \\ 2 \end{array}$ | $\begin{array}{r} 28 . \\ 8 \end{array}$ | $\begin{array}{r} \hline 14 . \\ 8 \end{array}$ | $13 .$ | $\begin{array}{r} 10 . \\ 4 \end{array}$ | $\begin{array}{r} 10 . \\ 3 \end{array}$ | $\begin{array}{r} 18 . \\ 7 \end{array}$ | 3.8 | 3.2 | 623 |
| 2009 | 0.0 | 43.2 | 45.7 | $\begin{array}{r} 175 . \\ 9 \end{array}$ | 66.0 | $\begin{array}{r} 185 . \\ 1 \end{array}$ | $\begin{array}{r} 28 . \\ 3 \end{array}$ | $\begin{array}{r} 25 \\ 7 \end{array}$ | $\begin{array}{r} 32 . \\ 9 \end{array}$ | 8.8 | $\begin{array}{r} 15 . \\ 4 \end{array}$ | $\begin{array}{r} 12 . \\ \hline \end{array}$ | $\begin{array}{r} 22 . \\ 3 \\ \hline \end{array}$ | 2.9 | 1.5 | 666 |
| 2010 | 0.0 | 10.2 | $\begin{array}{r} 177 . \\ 8 \end{array}$ | 45.6 | 74.8 | 63.6 | $72 .$ | 8.4 | $\begin{array}{r} \hline 14 . \\ 8 \end{array}$ | $\begin{array}{r} 10 . \\ \hline \end{array}$ | 4.1 | 4.7 | 5.4 | 5.4 | 22. | 520 |
| 2011 | 0.0 | 20.1 | 59.2 | 92.8 | 39.5 | 57.9 | $\begin{array}{r} 42 . \\ 0 \end{array}$ | $\begin{array}{r} 50 . \\ 7 \end{array}$ | $\begin{array}{r} 10 . \\ 9 \end{array}$ | 7.9 | 7.0 | 8.5 | 0.7 | 4.2 | 8.3 | 410 |
| 2012 | 0.0 | 12.8 | 56.8 | 27.7 | 27.5 | 15.3 | $\begin{array}{r} 26 . \\ 0 \end{array}$ | $\begin{array}{r} 26 . \\ \hline \end{array}$ | 21. | 4.8 | 15. | $\begin{array}{r} 10 . \\ 8 \end{array}$ | 1.7 | 4.0 | 0.7 | 252 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 445 |

Table 6. Estimates of selectivity-corrected age-class CPUE by year for female striped bass captured in the Choptank River during the 1985-1996 s pawning s tock s urveys. C PUE is standardized a st he nu mber of $f$ ish c aptured in 1000 s quare yards of experimental drift gill net per hour. The Choptank River was not sampled in 1995, and has not been sampled since 1996.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15+ | Total |
| 1985 | 0 | 0.0 | 0.0 | 0.0 | 2.2 | 0.8 | 2.9 | 0.8 | 1.0 | 0.4 | 0.0 | 0.6 | 1.3 | 0.5 | 1.0 | 12 |
| 1986 | 0 | 0.0 | 0.0 | $\begin{array}{r} 12 . \\ 8 \end{array}$ | 1.9 | 1.0 | 1.6 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | 0.2 | 0.5 | 18 |
| 1987 | 0 | 0.0 | 0.0 | 6.8 | $\begin{array}{r} 20 . \\ 7 \end{array}$ | 3.3 | 0.6 | 0.0 | 5.6 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.5 | 38 |
| 1988 | 0 | 0.0 | 0.0 | 9.2 | $\begin{array}{r} 10 . \\ 8 \end{array}$ | $\begin{array}{r} 16 . \\ 4 \end{array}$ | 3.2 | 0.0 | 1.0 | 1.0 | 0.0 | 0.0 | 0.0 | 0.7 | 0.4 | 43 |
| 1989 | 0 | 0.0 | 0.0 | 17. | 31. | 22. | 39. | 3.0 | 0.5 | 0.6 | 0.0 | 0.0 | 0.5 | 0.0 | 0.0 | 115 |
| 1990 | 0 | 0.0 | 0.0 | 0.0 | $15$ | $24 .$ | $\begin{array}{r} 15 . \\ 9 \\ \hline \end{array}$ | $\begin{array}{r} 40 . \\ 7 \end{array}$ | 3.1 | 3.0 | 0.0 | 0.0 | 4.7 | 2.5 | 4.4 | 114 |
| 1991 | 0 | 0.0 | 0.0 | 1.3 | 0.8 | 22. 9 | 23. | 15. | $\begin{array}{r}32 . \\ 9 \\ \hline\end{array}$ | 4.8 | 3.4 | 0.0 | 14. | 14. 1 | 5.1 | 138 |
| 1992 | 0 | 0.0 | 1.0 | 0.0 | 1.4 | 9.9 | 28. | $\begin{array}{r} 18 . \\ 7 \end{array}$ | $\begin{array}{r} \hline 19 . \\ 0 \end{array}$ | $15 .$ | 0.0 | 0.0 | $16 .$ | 3.4 | 0.0 | 113 |
| 1993 | 0 | 0.0 | 0.0 | 3.0 | 0.0 | 5.4 | 15. | 30. 1 | 23. 5 | 19. | 8.2 | 1.6 | 2.8 | 5.6 | 2.8 | 117 |
| 1994 | 0 | 0.0 | 0.0 | 0.0 | 7.5 | 7.1 | 8.8 | 7.7 | $\begin{array}{r} 31 . \\ 3 \end{array}$ | 6.1 | 4.0 | 0.0 | 0.0 | 0.0 | 0.0 | 73 |
| 1995 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1996 | 0 | 0.0 | 0.0 | 0.0 | 6.9 | 26. | 38. | 37. 0 | 36. | 37. | 21. | 8.7 | 1.1 | 0.0 | 0.0 | 214 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 90 |

Table 7. Estimates of selectivity-corrected age-class CPE by year for male striped bass captured in the Choptank River during the 1985-1996 spawning stock surveys. CPUE is standardized as the number of fish captured in 1000 square yards of experimental drift gill net per hour. The Choptank River was not sampled in 1995, and has not been sampled since 1996.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15+ | Total |
| 1985 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | $\begin{array}{r} 162 . \\ 2 \end{array}$ | $\begin{array}{r} 594 . \\ 7 \\ \hline \end{array}$ | 23.9 | 7.3 | 4.8 | 10.0 | 0.0 | 3.5 | 0.0 | 0.0 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | 0 5 | 0.0 | 0 | 807 |
| 1986 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | $\begin{array}{r} 290 . \\ 2 \end{array}$ | $\begin{array}{r} 172 . \\ 6 \end{array}$ | 393.9 | 12.0 | 6.1 | 1.6 | 1.2 | 0.0 | 0.0 | 0.0 | 0. | 0. 6 | 0.0 | 0 | 878 |
| 1987 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | $\begin{array}{r} 223 . \\ 3 \end{array}$ | $\begin{array}{r} 262 . \\ 0 \end{array}$ | 79.0 | $\begin{array}{r} 156 . \\ 4 \end{array}$ | 9.6 | 0.7 | 1.2 | 0.4 | 0.0 | 0.0 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 0 7 | 0.0 | 0 | 733 |
| 1988 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 27.0 | $\begin{array}{r} 223 . \\ 3 \end{array}$ | 114.6 | 53.5 | $\begin{array}{r} 111 . \\ 5 \end{array}$ | 4.7 | 0.0 | 0.0 | 1.4 | 0.0 | 0. 0 | 0. 0 | 0.0 | 0 | 536 |
| 1989 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | $\begin{array}{r} 228 . \\ 5 \end{array}$ | 58.1 | 466.1 | $\begin{array}{r} 278 . \\ 6 \end{array}$ | $\begin{array}{r} 191 . \\ 9 \end{array}$ | $\begin{array}{r} 173 . \\ 9 \end{array}$ | 1.1 | 1.1 | 0.0 | 0.0 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 0. 0 | 0.0 | 0 | 139 |
| 1990 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | 59.5 | $\begin{array}{r} 280 . \\ 4 \end{array}$ | 36.3 | $\begin{array}{r} 198 . \\ 1 \end{array}$ | $\begin{array}{r} 165 . \\ 8 \end{array}$ | 75.9 | $\begin{array}{r} 116 . \\ 9 \end{array}$ | 5.0 | 0.0 | 2.3 | 0. 0 | $\begin{array}{r} 4 . \\ 3 \\ \hline \end{array}$ | 0.0 | 0 | 944 |
| 1991 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | $\begin{array}{r} 410 . \\ 4 \end{array}$ | $\begin{array}{r} 174 . \\ 9 \end{array}$ | 112.2 | 62.1 | $\begin{array}{r} 115 . \\ 6 \end{array}$ | 79.8 | 55.5 | $\begin{array}{r} 18 . \\ 2 \\ \hline \end{array}$ | 0.6 | 0.0 | 0. 0 | 0. 0 | 0.0 | 0 | 102 9 |
| 1992 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 16.2 | $\begin{array}{r} 733 . \\ 0 \end{array}$ | 135.2 | $\begin{array}{r} 168 . \\ 4 \end{array}$ | $\begin{array}{r} 141 . \\ 9 \end{array}$ | $\begin{array}{r} 136 . \\ 4 \end{array}$ | 81.2 | 23. | 10. | 0.0 | 0. 0 | 0. 0 | 11. | 0 | 145 7 |
| 1993 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | $\begin{array}{r} 291 . \\ 3 \end{array}$ | $\begin{array}{r} 128 . \\ 8 \end{array}$ | $\begin{array}{r} 1156 . \\ \hline 4 \end{array}$ | $\begin{array}{r} 193 . \\ 5 \end{array}$ | $\begin{array}{r} 158 . \\ 8 \end{array}$ | $\begin{array}{r} 161 . \\ 5 \end{array}$ | $\begin{array}{r} 147 . \\ 3 \end{array}$ | $\begin{array}{r} 45 . \\ 9 \\ \hline \end{array}$ | $\begin{array}{r} 11 . \\ 3 \end{array}$ | 3.5 | 0. 0 | 0. 0 | 0.0 | 0 | 229 8 |
| 1994 | 0. | $\begin{array}{r} 112 . \\ 8 \end{array}$ | $\begin{array}{r} 463 . \\ 3 \end{array}$ | 99.5 | $\begin{array}{r} 835 . \\ 2 \end{array}$ | $\begin{array}{r} 270 . \\ 9 \end{array}$ | $\begin{array}{r} 139 . \\ 4 \end{array}$ | $\begin{array}{r} 188 . \\ 5 \end{array}$ | $\begin{array}{r} 54 . \\ 9 \\ \hline \end{array}$ | 9.2 | 7.6 | 8. 3 | 0. 9 | 0.0 | 0 | 219 1 |
| 1995 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1996 | 0. 0 | 7.8 | 682. 2 | 106.0 | 280. | $\begin{array}{r}171 . \\ 5 \\ \hline\end{array}$ | 334. 1 | 91.1 | 85. | 11. | 23. | 0. 0 | 0. 0 | 0.0 | 0 | 179 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 127 |

Table 8. Mean values of the annual, pooled, weighted, age-specific CPUEs (1985-2012) for the Maryland Chesapeake Bay striped bass spawning stock. CPUE is reported as the number of fish captured in 1000 square yards of net per hour.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| YEAR | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15+ | Sum |
| 1985 | $\begin{array}{r} 0 . \\ 0 \end{array}$ | $\begin{array}{r} 140 . \\ 5 \end{array}$ | $\begin{array}{r} 305 . \\ 5 \end{array}$ | 31.9 | 4.8 | 1.3 | 2.2 | 0.0 | 0.4 | 0.1 | 0.0 | 0.4 | 0.3 | 0.0 | 0.7 | 488 |
| 1986 | $\begin{array}{r} 0 . \\ 0 \end{array}$ | $\begin{array}{r} 230 . \\ 2 \end{array}$ | $\begin{array}{r} 261 . \\ 1 \end{array}$ | $\begin{array}{r} 497 . \\ 6 \end{array}$ | 4.0 | 5.3 | 2.0 | 2.9 | 2.8 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.9 | 100 7 |
| 1987 | $\begin{array}{r} \hline 0 . \\ 0 \\ \hline \end{array}$ | $\begin{array}{r} 142 . \\ 2 \end{array}$ | $\begin{array}{r} 258 . \\ 0 \end{array}$ | $\begin{array}{r} 115 . \\ 1 \end{array}$ | $\begin{array}{r} 176 . \\ 1 \end{array}$ | 17.9 | 2.2 | 2.6 | 0.2 | 0.0 | 0.0 | 0.0 | 0.0 | 0.3 | 0.3 | 715 |
| 1988 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 40.8 | 77.6 | 71.3 | 57.0 | 74.6 | 1.3 | 0.0 | 0.0 | 4.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.3 | 327 |
| 1989 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 33.1 | $154 .$ | 80.5 | 45.5 | 48.8 | $\begin{array}{r} 32 . \\ 9 \end{array}$ | 0.2 | 0.1 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 396 |
| 1990 | $\begin{array}{r} 0 . \\ 0 \end{array}$ | 78.1 | $\begin{array}{r} 158 . \\ 1 \end{array}$ | $\begin{array}{r} 120 . \\ 4 \end{array}$ | 48.3 | 34.3 | 32. | $\begin{array}{r} 29 . \\ 8 \\ \hline \end{array}$ | 0.9 | 0.1 | 0.1 | 0.5 | 0.7 | 0.1 | 0.2 | 504 |
| 1991 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 73.4 | $\begin{array}{r} 191 . \\ 9 \end{array}$ | 62.2 | 47.1 | 26.7 | $\begin{array}{r} 26 . \\ 0 \end{array}$ | $\begin{array}{r} 19 . \\ 2 \end{array}$ | $\begin{array}{r} 10 . \\ 6 \end{array}$ | 0.4 | 1.5 | 0.0 | 0.6 | 0.6 | 1.1 | 461 |
| 1992 | 0. 1 | 27.4 | $\begin{array}{r} 221 . \\ 1 \end{array}$ | $\begin{array}{r} 153 . \\ 5 \end{array}$ | 58.6 | 69.9 | $\begin{array}{r} 42 . \\ 9 \end{array}$ | $\begin{array}{r} 29 . \\ 1 \end{array}$ | $13 .$ | 7.0 | 3.3 | 0.0 | 0.9 | 1.2 | 0.2 | 629 |
| 1993 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 41.0 | $\begin{array}{r} 132 . \\ 0 \end{array}$ | $\begin{array}{r} 187 . \\ 2 \end{array}$ | 88.2 | 51.0 | $\begin{array}{r} 51 . \\ 9 \end{array}$ | $\begin{array}{r} 37 . \\ \hline \end{array}$ | $\begin{array}{r} 22 . \\ 6 \end{array}$ | 7.4 | 3.1 | 0.8 | 1.4 | 1.4 | 0.1 | 625 |
| 1994 | 0. 0 | 26.8 | $\begin{array}{r} 103 . \\ 5 \end{array}$ | 98.0 | $\begin{array}{r} 117 . \\ \hline \end{array}$ | 59.5 | 34. | $\begin{array}{r} 42 . \\ 9 \end{array}$ | $\begin{array}{r} 17 . \\ 6 \\ \hline \end{array}$ | 8.6 | 3.1 | 1.3 | 0.3 | 0.0 | 0.0 | 513 |
| 1995 | $\begin{array}{r} 0 . \\ 0 \end{array}$ | 50.0 | $\begin{array}{r} 117 . \\ 2 \end{array}$ | 68.4 | 60.9 | 51.6 | $\begin{array}{r} 40 . \\ 0 \end{array}$ | $\begin{array}{r} 25 . \\ 0 \end{array}$ | $\begin{array}{r} 19 . \\ 7 \\ \hline \end{array}$ | $\begin{array}{r} 11 . \\ 6 \\ \hline \end{array}$ | 9.6 | 3.5 | 4.6 | 0.0 | 0.0 | 462 |
| 1996 | 0. 0 | 4.0 | $\begin{array}{r} 368 . \\ 3 \end{array}$ | $\begin{array}{r} 102 . \\ 2 \end{array}$ | 34.7 | 69.5 | 64. 4 | $\begin{array}{r} 42 . \\ 3 \end{array}$ | $\begin{array}{r} 35 . \\ 4 \end{array}$ | $\begin{array}{r} 16 . \\ 7 \end{array}$ | $\begin{array}{r} 15 . \\ 2 \end{array}$ | 4.7 | 1.6 | 0.0 | 0.0 | 759 |
| 1997 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 36.8 | 44.8 | $\begin{array}{r} 140 . \\ 3 \end{array}$ | 46.5 | 20.9 | $\begin{array}{r} 18 . \\ 9 \end{array}$ | $\begin{array}{r} 22 . \\ 1 \end{array}$ | $\begin{array}{r} 26 . \\ 6 \end{array}$ | $\begin{array}{r} 11 . \\ 4 \\ \hline \end{array}$ | 9.9 | 3.3 | 1.2 | 0.6 | 0.0 | 387 |
| 1998 | 0. 0 | 36.1 | $\begin{array}{r} 142 . \\ 8 \end{array}$ | 32.7 | $\begin{array}{r} 149 . \\ 3 \end{array}$ | 32.3 | 13. 2 | 18. 5 | $\begin{array}{r} 17 . \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 15 . \\ 0 \end{array}$ | 9.1 | 9.9 | 1.7 | 0.4 | 0.3 | 479 |
| 1999 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 8.6 | $\begin{array}{r} 172 . \\ 4 \end{array}$ | 78.9 | 58.6 | 36.7 | $\begin{array}{r} 11 . \\ 7 \end{array}$ | 7.0 | $\begin{array}{r} 11 . \\ 5 \end{array}$ | 5.2 | 4.8 | 2.8 | 1.1 | 2.1 | 0.1 | 397 |
| 2000 | $\begin{array}{r} 0 . \\ 0 \end{array}$ | 14.4 | 55.9 | $\begin{array}{r} 104 . \\ 1 \end{array}$ | 48.0 | 57.7 | $\begin{array}{r} 25 . \\ 0 \end{array}$ | $13 .$ | 8.3 | 8.3 | 7.0 | 7.4 | 1.5 | 2.5 | 0.5 | 352 |
| 2001 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 4.9 | 39.1 | 60.3 | 53.2 | 23.1 | 29. | $\begin{array}{r} 33 . \\ 3 \end{array}$ | $11 .$ | $\begin{array}{r} 12 . \\ 1 \end{array}$ | 9.3 | 6.1 | 3.5 | 1.2 | 0.4 | 283 |


| 2002 | $\begin{array}{r} 0 . \\ 0 \\ \hline \end{array}$ | 84.6 | 40.8 | 39.7 | 85.8 | 42.7 | $\begin{array}{r} 35 . \\ 0 \\ \hline \end{array}$ | $\begin{array}{r} 33 . \\ 1 \end{array}$ | $\begin{array}{r} 23 . \\ 5 \end{array}$ | 8.4 | 5.8 | 3.6 | 5.2 | 1.2 | 0.4 | 400 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2003 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 15.7 | $\begin{array}{r} 111 . \\ 5 \end{array}$ | 53.4 | 35.4 | 68.4 | $\begin{array}{r} 51 . \\ 6 \\ \hline \end{array}$ | $\begin{array}{r} 27 . \\ 6 \end{array}$ | $\begin{array}{r} 26 . \\ \hline \end{array}$ | $\begin{array}{r} 29 . \\ \hline \end{array}$ | 14. | 7.2 | 6.1 | 2.5 | 0.3 | 455 |
| 2004 | $\begin{array}{r} \hline 0 . \\ 0 \\ \hline \end{array}$ | 28.8 | $\begin{array}{r} 193 . \\ 2 \\ \hline \end{array}$ | $\begin{array}{r} 121 . \\ \hline \end{array}$ | 42.4 | 34.6 | 44. $4$ | $\begin{array}{r} \hline 47 . \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 30 . \\ \hline \end{array}$ | $\begin{array}{r} 23 . \\ \hline \end{array}$ | $\begin{array}{r} 23 . \\ \hline \end{array}$ | 6.7 | 4.2 | 3.7 | 2.6 | 611 |
| 2005 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 66.0 | $\begin{array}{r} 103 . \\ 6 \end{array}$ | 73.5 | 96.6 | 24.3 | 25. | $\begin{array}{r} \hline 21 . \\ \hline \end{array}$ | $\begin{array}{r} 27 . \\ 5 \end{array}$ | $\begin{array}{r} 20 . \\ 4 \\ \hline \end{array}$ | 17. | 11. | 3.0 | 1.0 | 3.8 | 496 |
| 2006 | $\begin{array}{r} \hline 0 . \\ 0 \\ \hline \end{array}$ | 7.5 | $\begin{array}{r} 257 . \\ \hline 9 \end{array}$ | 40.1 | 47.6 | 29.2 | 14. $8$ | $\begin{array}{r} 12 . \\ 7 \\ \hline \end{array}$ | $\begin{array}{r} 18 . \\ 4 \\ \hline \end{array}$ | $\begin{array}{r} 21 . \\ 6 \\ \hline \end{array}$ | $\begin{array}{r} 13 . \\ \hline \end{array}$ | 11. 0 | 9.3 | 2.7 | 6.1 | 492 |
| 2007 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 7.9 | 22.5 | 76.0 | 14.9 | 15.3 | $\begin{array}{r} 13 . \\ 5 \\ \hline \end{array}$ | 7.4 | 9.0 | 10. | $16 .$ | 8.0 | 3.0 | 5.4 | 5.3 | 214 |
| 2008 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 3.3 | 86.0 | $\begin{array}{r} 108 . \\ 4 \end{array}$ | $\begin{array}{r} 112 . \\ 3 \end{array}$ | 16.9 | $\begin{array}{r} 23 . \\ 0 \end{array}$ | $\begin{array}{r} 19 . \\ 7 \end{array}$ | $11 .$ | $\begin{array}{r} 12 . \\ 0 \end{array}$ | 10. | $\begin{array}{r} 14 . \\ 0 \end{array}$ | 13. | 3.3 | 3.6 | 437 |
| 2009 | 0. | 40.1 | 42.1 | 153. | 51.6 | 138. | 21. | 22. | $\begin{array}{r} 31 . \\ 2 \\ \hline \end{array}$ | 9.0 | 15. 8 | 12. | 23. | 4.8 | 4.8 | 570 |
| 2010 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 7.5 | $\begin{array}{r} 149 . \\ 7 \end{array}$ | 50.4 | 65.0 | 50.5 | $\begin{array}{r} 54 . \\ \hline 9 \end{array}$ | 6.7 | $13 .$ | $\begin{array}{r} 10 . \\ 2 \end{array}$ | 4.0 | 5.1 | 5.9 | 9.9 | $\begin{array}{r} 19 . \\ \hline 4 \end{array}$ | 453 |
| 2011 | $\begin{array}{r} \hline 0 . \\ 0 \\ \hline \end{array}$ | 23.0 | 73.3 | $\begin{array}{r} 123 . \\ 7 \end{array}$ | 45.4 | 57.3 | 38. | $\begin{array}{r} 44 . \\ 9 \end{array}$ | $\begin{array}{r} 10 . \\ 1 \end{array}$ | 9.1 | 7.9 | 7.8 | 4.0 | 4.3 | 9.5 | 458 |
| 2012 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 15.2 | 52.0 | 23.2 | 23.7 | 17.8 | $\begin{array}{r} 23 . \\ 1 \end{array}$ | $\begin{array}{r} 22 . \\ 6 \end{array}$ | $\begin{array}{r} \hline 25 . \\ 0 \end{array}$ | 7.4 | $16 .$ | $13 .$ | 4.4 | 6.7 | $\begin{array}{r} 13 . \\ \hline \end{array}$ | 265 |
| Average |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 487 |

Table 9. Lower confidence limits (95\%) of the annual, pooled, weighted, age-specific CPUEs (1985-2012) for the Maryland Chesapeake Bay striped bass spawning stock. CPUE is reported as the number of fish captured in 1000 square yards of net per hour.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15+ |
| 1985 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | $\begin{array}{r} 127 . \\ 3 \end{array}$ | $\begin{array}{r} 277 . \\ 1 \end{array}$ | 28.8 | 4.2 | 1.0 | 0.9 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | * |
| 1986 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | $\begin{array}{r} 214 . \\ 2 \end{array}$ | $\begin{array}{r} 245 . \\ 6 \end{array}$ | $\begin{array}{r} 464 . \\ 6 \end{array}$ | 3.6 | 4.8 | 1.7 | 2.7 | 1.8 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | * |
| 1987 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | $\begin{array}{r} 130 . \\ 4 \end{array}$ | $\begin{array}{r} 245 . \\ 1 \end{array}$ | $\begin{array}{r} 110 . \\ 6 \end{array}$ | $\begin{array}{r} 167 . \\ 8 \end{array}$ | 12.1 | 0.0 | 2.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | * |
| 1988 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 36.2 | 69.3 | 65.8 | 53.8 | 68.0 | 0.1 | 0.0 | 0.0 | 0.7 | 0.0 | 0.0 | 0.0 | 0.0 | * |
| 1989 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 24.7 | $\begin{array}{r} 148 . \\ 0 \end{array}$ | 66.1 | 35.5 | 41.5 | $\begin{array}{r} 24 . \\ 8 \end{array}$ | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | * |
| 1990 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 65.6 | $\begin{array}{r} 148 . \\ 3 \end{array}$ | $\begin{array}{r} 116 . \\ 3 \end{array}$ | 42.3 | 28.9 | $\begin{array}{r} 29 . \\ 4 \end{array}$ | $\begin{array}{r} 23 . \\ 9 \end{array}$ | 0.4 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | * |
| 1991 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 57.0 | $\begin{array}{r} 182 . \\ 6 \end{array}$ | 58.6 | 44.8 | 22.6 | $\begin{array}{r} 22 . \\ 4 \end{array}$ | $\begin{array}{r} 16 . \\ 5 \end{array}$ | 5.4 | 0.0 | 0.6 | 0.0 | 0.0 | 0.0 | 0.0 |
| 1992 | $\begin{gathered} \hline 0 . \\ 1 \\ \hline \end{gathered}$ | 23.0 | $\begin{array}{r} 206 . \\ 8 \end{array}$ | $\begin{array}{r} 145 . \\ 6 \end{array}$ | 54.6 | 65.7 | $\begin{array}{r} 38 . \\ 7 \end{array}$ | $\begin{array}{r} 26 . \\ 1 \end{array}$ | $\begin{array}{r} 11 . \\ 0 \end{array}$ | 4.1 | 2.3 | 0.0 | 0.0 | 0.0 | * |
| 1993 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 30.5 | $\begin{array}{r} 125 . \\ 3 \end{array}$ | $\begin{array}{r} 159 . \\ 4 \end{array}$ | 83.6 | 47.7 | 47. | $\begin{array}{r} 31 . \\ 7 \end{array}$ | $\begin{array}{r} 18 . \\ 1 \end{array}$ | 3.8 | 1.7 | 0.0 | 0.0 | 0.0 | * |
| 1994 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 21.7 | 89.3 | 94.5 | 96.8 | 52.9 | $\begin{array}{r} 31 . \\ \hline \end{array}$ | $\begin{array}{r} 38 . \\ 7 \end{array}$ | $\begin{array}{r} 12 . \\ 5 \end{array}$ | 7.5 | 2.3 | 1.0 | 0.3 | 0.0 | * |
| 1995 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 45.8 | $\begin{array}{r} 114 . \\ 5 \end{array}$ | 66.4 | 59.3 | 49.6 | 38. | 24. 1 | $\begin{array}{r} 18 . \\ 7 \\ \hline \end{array}$ | $\begin{array}{r} 11 . \\ 0 \end{array}$ | 9.2 | 3.2 | 1.9 | 0.0 | * |
| 1996 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | 0.0 | $347 .$ | 98.2 | 26.3 | 65.2 | $\begin{array}{r} 57 . \\ \hline \end{array}$ | $\begin{array}{r} \hline 37 . \\ \hline \end{array}$ | $\begin{array}{r} 30 . \\ 4 \end{array}$ | $\begin{array}{r} 10 . \\ 3 \end{array}$ | $\begin{array}{r} 10 . \\ 3 \end{array}$ | 3.1 | 1.1 | 0.0 | 0.0 |
| 1997 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 35.9 | 43.5 | $\begin{array}{r} 136 . \\ 8 \end{array}$ | 44.9 | 20.3 | 18. | 20. | $\begin{array}{r} 21 . \\ 9 \end{array}$ | $\begin{array}{r} 10 . \\ 7 \end{array}$ | 6.3 | 3.0 | 1.1 | 0.5 | 0.0 |
| 1998 | $\begin{gathered} \hline 0 . \\ 0 \end{gathered}$ | 35.7 | $\begin{array}{r} 138 . \\ 9 \end{array}$ | 31.4 | $\begin{array}{r} 144 . \\ 5 \end{array}$ | 31.6 | $11 .$ | 17. | $\begin{array}{r} 16 . \\ 7 \end{array}$ | $\begin{array}{r} 14 . \\ 3 \\ \hline \end{array}$ | 8.7 | 8.8 | 1.2 | 0.3 | 0.2 |
| 1999 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 6.9 | $\begin{array}{r} 168 . \\ 6 \end{array}$ | 76.5 | 56.8 | 35.5 | $\begin{array}{r} \hline 11 . \\ \hline \end{array}$ | 6.6 | 10 3 | 4.6 | 4.4 | 2.5 | 1.1 | 0.5 | 0.1 |
| 2000 | $\begin{gathered} 0 . \\ 0 \end{gathered}$ | 13.5 | 53.7 | $\begin{array}{r} 101 . \\ 8 \end{array}$ | 46.7 | 55.8 | $\begin{array}{r} 23 . \\ 4 \end{array}$ | $\begin{array}{r} 13 . \\ 2 \\ \hline \end{array}$ | 7.9 | 7.6 | 6.5 | 5.5 | 1.4 | 1.2 | 0.5 |
| 2001 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 4.4 | 37.6 | 58.6 | 51.7 | 22.1 | 28. 2 | 32. | 11. 0 | 11. | 8.7 | 5.3 | 3.0 | 0.8 | 0.4 |


| 2002 | $\begin{array}{r} 0 . \\ 0 \end{array}$ | 75.7 | 39.3 | 38.8 | 83.3 | 40.4 | 33. | 32. | 22. | 7.4 | 5.4 | 3.3 | 3.7 | 0.3 | * |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2003 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 14.4 | $\begin{array}{r} 107 . \\ 5 \end{array}$ | 51.8 | 34.2 | 65.8 | $\begin{array}{r} \hline 49 . \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 26 . \\ 7 \end{array}$ | $\begin{array}{r} 25 . \\ 5 \end{array}$ | $\begin{array}{r} 26 . \\ 7 \end{array}$ | $13 .$ | 6.3 | 5.1 | 1.5 | 0.3 |
| 2004 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 22.8 | $\begin{array}{r} 188 . \\ 7 \end{array}$ | $\begin{array}{r} 118 . \\ 3 \end{array}$ | 41.1 | 33.3 | $\begin{array}{r} \hline 43 . \\ 3 \end{array}$ | $\begin{array}{r} 45 . \\ 5 \end{array}$ | $\begin{array}{r} \hline 28 . \\ 0 \end{array}$ | 22. | $\begin{array}{r} 21 . \\ 8 \end{array}$ | 6.1 | 3.8 | 3.2 | * |
| 2005 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 62.8 | 98.9 | 71.0 | 92.8 | 23.3 | 24. | 21. | $\begin{array}{r} 26 . \\ 4 \end{array}$ | 19. | 16. $4$ | $10 .$ | 2.6 | 0.9 | * |
| 2006 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 6.4 | $\begin{array}{r} 242 . \\ 1 \end{array}$ | 38.4 | 45.6 | 27.6 | $\begin{array}{r} 14 . \\ 2 \end{array}$ | $\begin{array}{r} 12 . \\ 3 \end{array}$ | $\begin{array}{r} 17 . \\ 2 \end{array}$ | $\begin{array}{r} 20 . \\ 0 \end{array}$ | $\begin{array}{r} 12 . \\ 1 \end{array}$ | 9.8 | 7.2 | 2.2 | * |
| 2007 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 6.9 | 21.4 | 74.0 | 14.5 | 14.9 | $\begin{array}{r} 12 . \\ 5 \end{array}$ | 6.2 | 8.0 | 9.3 | $\begin{array}{r} 13 . \\ \hline \end{array}$ | 7.0 | 2.8 | 3.9 | * |
| 2008 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 2.8 | 82.1 | $\begin{array}{r} 104 . \\ 0 \end{array}$ | $\begin{array}{r} 106 . \\ 8 \end{array}$ | 16.2 | $\begin{array}{r} \hline 22 . \\ 0 \end{array}$ | $\begin{array}{r} 18 . \\ 7 \end{array}$ | $\begin{array}{r} 10 . \\ 7 \end{array}$ | $11 .$ | 9.3 | 12. | 6.8 | 2.9 | * |
| 2009 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 38.5 | 40.6 | $\begin{array}{r} 148 . \\ 4 \end{array}$ | 49.8 | $\begin{array}{r} 133 . \\ 1 \end{array}$ | $\begin{array}{r} 20 . \\ 5 \end{array}$ | $\begin{array}{r} 21 . \\ 9 \\ \hline \end{array}$ | $\begin{array}{r} 29 . \\ 3 \end{array}$ | 8.5 | 15. | $10 .$ | $20 .$ | 4.3 | * |
| 2010 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 7.0 | $\begin{array}{r} 144 . \\ 8 \end{array}$ | 49.2 | 63.3 | 49.0 | $53 .$ | 6.2 | $13 .$ | 9.7 | 3.8 | 4.8 | 5.6 | 8.8 | * |
| 2011 | $\begin{array}{r} \hline 0 . \\ 0 \end{array}$ | 22.0 | 71.1 | $\begin{array}{r} 120 . \\ 2 \end{array}$ | 43.8 | 55.2 | $\begin{array}{r} 37 . \\ 1 \\ \hline \end{array}$ | $\begin{array}{r} 43 . \\ \hline \end{array}$ | 9.8 | 8.8 | 7.6 | 5.5 | 3.5 | 3.8 | * |
| 2012 | 0. | 14.2 | 50.2 | 22.4 | 22.8 | 16.7 | 22. | 20. | 23. | 6.9 | 15. | 9.2 | 3.8 | 5.5 | * |

* Notes: Shadings note negative values that have been changed to zero. Confidence intervals could not be calculated for age $15+$ when more than one age class was present in the group.

Table 10. Upper confidence limits ( $95 \%$ ) of the annual, pooled, weighted, age-specific CPUEs (1985-2012) for the Maryland Chesapeake Bay striped bass spawning stock. CPUE is reported as the number of fish captured in 1000 square yards of net per hour.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15+ |
| 1985 | 0.0 | $\begin{array}{r} 153 . \\ 6 \end{array}$ | $\begin{array}{r} 334 . \\ 0 \end{array}$ | 35.1 | 5.4 | 1.6 | 3.4 | 0.2 | 2.6 | 0.2 | 0.1 | 0.8 | 0.6 | 0.1 | * |
| 1986 | 0.0 | $\begin{array}{r} 246 . \\ 2 \end{array}$ | $\begin{array}{r} 276 . \\ 6 \end{array}$ | $\begin{array}{r} 530 . \\ 6 \end{array}$ | 4.5 | 5.8 | 2.4 | 3.2 | 3.8 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | * |
| 1987 | 0.0 | $\begin{array}{r} 154 . \\ 0 \\ \hline \end{array}$ | $\begin{array}{r} 270 . \\ \hline \end{array}$ | $\begin{array}{r} 119 . \\ 6 \end{array}$ | $\begin{array}{r} 184 . \\ 5 \end{array}$ | 23.7 | 5.4 | 2.8 | 2.3 | 0.0 | 0.0 | 0.0 | 0.0 | 0.5 | * |
| 1988 | 0.0 | 45.3 | 86.0 | 76.8 | 60.2 | 81.1 | 2.5 | 1.0 | 1.1 | 8.0 | 0.0 | 0.0 | 0.0 | 0.1 | * |
| 1989 | 0.0 | 41.6 | $\begin{array}{r} 161 . \\ 4 \end{array}$ | 95.0 | 55.5 | 56.0 | $\begin{array}{r} 41 . \\ 0 \end{array}$ | 0.6 | 0.1 | 0.2 | 0.0 | 0.0 | 0.1 | 0.0 | * |
| 1990 | 0.0 | 90.5 | $\begin{array}{r} 168 . \\ 0 \end{array}$ | $\begin{array}{r} 124 . \\ 5 \end{array}$ | 54.3 | 39.6 | $\begin{array}{r} 34 . \\ 7 \end{array}$ | $\begin{array}{r} 35 . \\ 7 \end{array}$ | 1.3 | 0.5 | 0.3 | 1.0 | 5.3 | 1.7 | * |
| 1991 | 0.0 | 89.8 | $\begin{array}{r} 201 . \\ 2 \end{array}$ | 65.8 | 49.4 | 30.8 | $\begin{array}{r} 29 . \\ 6 \end{array}$ | $\begin{array}{r} 21 . \\ 8 \\ \hline \end{array}$ | $\begin{array}{r} 15 . \\ 8 \\ \hline \end{array}$ | 1.2 | 2.3 | 0.0 | 6.3 | 5.4 | 2.9 |
| 1992 | 0.3 | 31.8 | $\begin{array}{r} 235 . \\ 4 \end{array}$ | $\begin{array}{r} 161 . \\ 4 \\ \hline \end{array}$ | 62.7 | 74.1 | 47. 1 | $\begin{array}{r} 32 . \\ 0 \end{array}$ | $\begin{array}{r} 16 . \\ 3 \end{array}$ | 10. | 4.2 | 0.0 | 7.3 | 8.9 | * |
| 1993 | 0.0 | 51.4 | $\begin{array}{r} 138 . \\ 7 \end{array}$ | $\begin{array}{r} 215 . \\ 1 \end{array}$ | 92.9 | 54.2 | 56. | 42. | 27. | 11. | 4.5 | 1.7 | 2.8 | 7.6 | * |
| 1994 | 0.0 | 32.0 | 117. 8 | $\begin{array}{r} 101 . \\ 5 \end{array}$ | $\begin{array}{r} 138 . \\ 9 \end{array}$ | 66.1 | $\begin{array}{r} 36 . \\ \hline \end{array}$ | $\begin{array}{r} \hline 47 . \\ 0 \\ \hline \end{array}$ | $\begin{array}{r} 22 . \\ \hline \end{array}$ | 9.6 | 3.8 | 1.5 | 0.3 | 0.0 | * |
| 1995 | 0.0 | 54.2 | $\begin{array}{r} 120 . \\ 0 \end{array}$ | 70.3 | 62.5 | 53.5 | 41. 5 | $\begin{array}{r} 25 . \\ \hline 9 \end{array}$ | $\begin{array}{r} 20 . \\ 6 \end{array}$ | $\begin{array}{r} 12 . \\ \hline \end{array}$ | $\begin{array}{r} 10 . \\ 1 \end{array}$ | 3.8 | 7.2 | 0.0 | * |
| 1996 | 0.0 | 10.8 | $\begin{array}{r} 389 . \\ 5 \\ \hline \end{array}$ | $\begin{array}{r} 106 . \\ 1 \end{array}$ | 43.2 | 73.9 | $\begin{array}{r} 71 . \\ 5 \\ \hline \end{array}$ | $\begin{array}{r} 46 . \\ 6 \\ \hline \end{array}$ | $\begin{array}{r} 40 . \\ 4 \end{array}$ | $\begin{array}{r} 23 . \\ \hline \end{array}$ | $\begin{array}{r} 20 . \\ 1 \end{array}$ | 6.3 | 2.2 | 0.0 | 0.0 |
| 1997 | 0.0 | 37.8 | 46.1 | $\begin{array}{r} 143 . \\ 9 \end{array}$ | 48.2 | 21.6 | $\begin{array}{r}19 \\ 7 \\ \hline\end{array}$ | 23. | $\begin{array}{r}31 . \\ 2 \\ \hline\end{array}$ | 12. | $\begin{array}{r} 13 . \\ 6 \\ \hline \end{array}$ | 3.6 | 1.3 | 0.6 | 0.0 |
| 1998 | 0.0 | 36.4 | $\begin{array}{r} 146 . \\ 7 \end{array}$ | 34.1 | $\begin{array}{r} 154 . \\ 0 \end{array}$ | 33.0 | $\begin{array}{r} 15 . \\ 1 \\ \hline \end{array}$ | $\begin{array}{r} 19 . \\ 4 \end{array}$ | 17. | 15. | 9.5 | 11. | 2.2 | 0.5 | 0.4 |
| 1999 | 0.0 | 10.3 | $\begin{array}{r} 176 . \\ 2 \end{array}$ | 81.3 | 60.4 | 37.9 | $\begin{array}{r} 12 . \\ \hline \end{array}$ | 7.4 | $\begin{array}{r} 12 . \\ 7 \end{array}$ | 5.7 | 5.3 | 3.1 | 1.2 | 3.8 | 0.2 |
| 2000 | 0.0 | 15.2 | 58.2 | $\begin{array}{r} 106 . \\ 4 \end{array}$ | 49.2 | 59.7 | $\begin{array}{r} 26 . \\ 5 \end{array}$ | $\begin{array}{r} 14 . \\ 4 \\ \hline \end{array}$ | 8.6 | 9.0 | 7.4 | 9.3 | 1.6 | 3.8 | 0.6 |
| 2001 | 0.0 | 5.4 | 40.5 | 61.9 | 54.6 | 24.2 | 30 0 | 34. | $\begin{array}{r}12 . \\ 1 \\ \hline\end{array}$ | 12. | 9.8 | 6.8 | 4.0 | 1.6 | 0.5 |


| 2002 | 0.0 | 93.6 | 42.3 | 40.7 | 88.3 | 45.0 | $\begin{array}{r} 36 . \\ 2 \end{array}$ | $\begin{array}{r} 33 . \\ 9 \end{array}$ | $\begin{array}{r} 25 . \\ 0 \end{array}$ | 9.3 | 6.2 | 3.9 | 6.7 | 2.1 | * |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2003 | 0.0 | 17.1 | $\begin{array}{r} 115 . \\ 5 \end{array}$ | 55.1 | 36.6 | 71.0 | $\begin{array}{r} \hline 54 . \\ 0 \end{array}$ | $\begin{array}{r} 28 . \\ 5 \end{array}$ | $\begin{array}{r} 28 . \\ 0 \end{array}$ | $\begin{array}{r} 31 . \\ 4 \end{array}$ | $\begin{array}{r} 16 . \\ 2 \end{array}$ | 8.1 | 7.2 | 3.5 | 0.4 |
| 2004 | 0.0 | 34.9 | 197 7 | $\begin{array}{r} 124 . \\ 0 \end{array}$ | 43.7 | 35.9 | $\begin{array}{r} 45 . \\ \hline 4 \end{array}$ | 49 0 | $\begin{array}{r} \hline 32 . \\ 2 \\ \hline \end{array}$ | $\begin{array}{r} \hline 24 . \\ 0 \end{array}$ | $\begin{array}{r} 24 . \\ 3 \end{array}$ | 7.3 | 4.7 | 4.2 | * |
| 2005 | 0.0 | 69.2 | $\begin{array}{r} 108 . \\ 4 \end{array}$ | 76.0 | $\begin{array}{r} 100 . \\ 5 \end{array}$ | 25.2 | $\begin{array}{r} \hline 26 . \\ 8 \end{array}$ | $\begin{array}{r} 22 . \\ 5 \end{array}$ | $\begin{array}{r} 28 . \\ 5 \end{array}$ | $\begin{array}{r} 21 . \\ 5 \end{array}$ | $\begin{array}{r} 18 . \\ 5 \end{array}$ | $\begin{array}{r} 12 . \\ 5 \end{array}$ | 3.3 | 1.2 | * |
| 2006 | 0.0 | 8.6 | $\begin{array}{r} 273 . \\ 7 \\ \hline \end{array}$ | 41.7 | 49.5 | 30.9 | 15. | $13 .$ | 19. | 23. | 14. | $\begin{array}{r} 12 . \\ 2 \end{array}$ | 11. 3 | 3.2 | * |
| 2007 | 0.0 | 8.9 | 23.6 | 78.1 | 15.3 | 15.7 | $\begin{array}{r} 14 . \\ 4 \\ \hline \end{array}$ | 8.5 | $\begin{array}{r} 10 . \\ 1 \end{array}$ | $\begin{array}{r} 10 . \\ 8 \end{array}$ | $\begin{array}{r} 18 . \\ \hline 8 \end{array}$ | 8.9 | 3.3 | 7.0 | * |
| 2008 | 0.0 | 3.7 | 90.0 | $\begin{array}{r} 112 . \\ 8 \end{array}$ | $\begin{array}{r} 117 . \\ 9 \end{array}$ | 17.6 | 24. | $\begin{array}{r} \hline 20 . \\ 7 \end{array}$ | 11. | 12. | 10. | 15. | 20. | 3.6 | * |
| 2009 | 0.0 | 41.7 | 43.6 | $\begin{array}{r} 157 . \\ 6 \\ \hline \end{array}$ | 53.5 | $\begin{array}{r} 143 . \\ 3 \end{array}$ | $\begin{array}{r} 21 . \\ 8 \end{array}$ | $\begin{array}{r} 23 . \\ \hline \end{array}$ | $\begin{array}{r} 33 . \\ \hline \end{array}$ | 9.4 | $\begin{array}{r} 16 . \\ 7 \end{array}$ | $\begin{array}{r} 13 . \\ 5 \end{array}$ | $\begin{array}{r} 26 . \\ \hline \end{array}$ | 5.3 | * |
| 2010 | 0.0 | 8.0 | $\begin{array}{r} 154 . \\ 6 \end{array}$ | 51.6 | 66.6 | 52.0 | $\begin{array}{r} 56 . \\ 7 \end{array}$ | 7.2 | $\begin{array}{r} 14 . \\ 5 \end{array}$ | $\begin{array}{r} 10 . \\ 7 \end{array}$ | 4.1 | 5.4 | 6.2 | 11. | * |
| 2011 | 0.0 | 24.0 | 75.6 | $\begin{array}{r} 127 . \\ 3 \end{array}$ | 46.9 | 59.4 | $\begin{array}{r} 39 . \\ 0 \end{array}$ | $\begin{array}{r} 46 . \\ 8 \\ \hline \end{array}$ | $\begin{array}{r} 10 \\ 3 \end{array}$ | 9.5 | 8.1 | $\begin{array}{r} 10 . \\ 2 \end{array}$ | 4.6 | 4.8 | * |
| 2012 | 0.0 | 16.2 | 53.8 | 24.0 | 24.6 | 19.0 | $24 .$ | 24. | 26. | 7.9 | 17. | 17. | 4.9 | 8.0 | * |

* Note: Confidence intervals could not be calculated for age 15+ when more than one age class was present in the group.

Table 11. Coefficients of Variation of the annual, pooled, weighted, age-specific CPUEs (1985-2012) for the Maryland Chesapeake Bay striped bass spawning stock.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15+ |
| 1985 | 0 | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | $\begin{array}{r} 0.1 \\ 1 \end{array}$ | $\begin{array}{r} 0.2 \\ 8 \end{array}$ | 2.1 6 | 2.50 | $\begin{array}{r} 1.0 \\ 4 \end{array}$ | $\begin{array}{r} 0.2 \\ \hline 9 \end{array}$ | 0.5 8 | 0.6 4 | 2.1 4 | * |
| 1986 | 0 | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ \hline 9 \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | 0.18 | 0 | 0 | 0 | 0.2 8 | 2.6 2 | * |
| 1987 | 0 | $\begin{array}{r} 0.0 \\ 4 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \\ \hline \end{array}$ | $\begin{array}{r} 0.1 \\ 6 \\ \hline \end{array}$ | 0.7 6 | 0.0 5 | 4.32 | 0 | 0 | 0 | 0.3 4 | 0.3 6 | * |
| 1988 | 0 | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ 4 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ \hline 4 \end{array}$ | $\begin{array}{r} 0.4 \\ 5 \end{array}$ | 0.0 0 | $\begin{array}{r} 13.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.4 \\ 2 \end{array}$ | 0 | 0 | 0 | 1.1 0 | * |
| 1989 | 0 | $\begin{array}{r} 0.1 \\ 3 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 9 \end{array}$ | $\begin{array}{r} \hline 0.1 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 7 \end{array}$ | $\begin{array}{r} 0.1 \\ 2 \end{array}$ | 1.1 7 | 0.29 | $\begin{array}{r} 2.9 \\ 2 \end{array}$ | 0 | 0 | 1.3 1 | 0 | * |
| 1990 | 0 | $\begin{array}{r} 0.0 \\ 8 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | $\begin{array}{r} 0.0 \\ 8 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 4 \end{array}$ | 0.1 0 | 0.28 | $\begin{array}{r} 1.5 \\ 1 \\ \hline \end{array}$ | $\begin{array}{r} 1.0 \\ 7 \end{array}$ | $\begin{array}{r} 0.4 \\ 9 \end{array}$ | 3.1 8 | 7.8 5 | * |
| 1991 | 0 | $\begin{array}{r} 0.1 \\ \hline \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 8 \end{array}$ | 0.0 7 | 0.0 7 | 0.25 | $\begin{array}{r} \hline 0.9 \\ 6 \end{array}$ | $\begin{array}{r} 0.2 \\ \hline 9 \end{array}$ | 0 | 5.1 0 | 4.2 9 | 0.8 2 |
| 1992 | $\begin{array}{r} 0.7 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 8 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ 5 \\ \hline \end{array}$ | 0.10 | $\begin{array}{r} 0.2 \\ 1 \\ \hline \end{array}$ | $\begin{array}{r} \hline 0.1 \\ 4 \end{array}$ | 0 | 3.3 8 | 3.1 6 | * |
| 1993 | 0 | $\begin{array}{r} 0.1 \\ 3 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 7 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | 0.0 5 | 0.0 7 | 0.10 | $\begin{array}{r} 0.2 \\ \hline 4 \end{array}$ | $\begin{array}{r} 0.2 \\ 3 \end{array}$ | 0.5 4 | 0.4 9 | 2.1 9 | * |
| 1994 | 0 | $\begin{array}{r} \hline 0.1 \\ 0 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ \hline 7 \end{array}$ | $\begin{array}{r} 1.0 \\ \hline 2 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 9 \end{array}$ | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | 0.0 4 | 0.0 5 | 0.15 | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | $\begin{array}{r} \hline 0.1 \\ 3 \end{array}$ | $\begin{array}{r} 0.1 \\ 1 \end{array}$ | 0.0 6 | 0 | * |
| 1995 | 0 | $\begin{array}{r} 0.0 \\ 4 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | 0.0 2 | 0.02 | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | 0.0 2 | 0.0 4 | 0.2 9 | 0 | * |
| 1996 | 0 | $\begin{array}{r} 0.8 \\ 7 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.1 \\ 2 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 3 \\ \hline \end{array}$ | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | 0.0 5 | 0.07 | $\begin{array}{r} 0.1 \\ \hline 9 \end{array}$ | $\begin{array}{r} \hline 0.1 \\ 6 \end{array}$ | $\begin{array}{r} 0.1 \\ 7 \end{array}$ | 0.1 6 | 0 | 0 |
| 1997 | 0 | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | 0.0 2 | 0.0 4 | 0.09 | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | 0.1 8 | 0.0 5 | 0.0 5 | 0.0 7 | 0 |
| 1998 | 0 | $\begin{array}{r} \hline 0.0 \\ 0 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ \hline 2 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 2 \end{array}$ | 0.0 1 | 0.0 7 | 0.0 2 | 0.02 | 0.0 2 | 0.0 2 | 0.0 5 | 0.1 5 | 0.1 1 | 0.2 1 |
| 1999 | 0 | $\begin{array}{r} 0.1 \\ 0 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | 0.0 2 | 0.0 2 | 0.0 2 | 0.0 3 | 0.05 | 0.0 6 | 0.0 5 | 0.0 6 | 0.0 2 | 0 | 0.1 9 |
| 2000 | 0 | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | 0.0 2 | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | 0.0 1 | 0.0 2 | 0.0 3 | 0.0 2 | 0.02 | 0.0 4 | 0.0 3 | 0.1 3 | 0.0 3 | 0.2 6 | 0.0 2 |
| 2001 | 0 | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} \hline 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | 0.0 1 | 0.0 2 | 0.0 2 | 0.0 2 | 0.02 | 0.0 3 | 0.0 3 | 0.0 6 | 0.0 7 | 0.1 8 | 0.0 3 |


| 2002 | 0 | $\begin{array}{r} 0.0 \\ 5 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | $\begin{array}{r} 0.0 \\ 3 \end{array}$ | $\begin{array}{r} 0.0 \\ 2 \end{array}$ | $\begin{array}{r} 0.0 \\ 1 \end{array}$ | 0.03 | $\begin{array}{r} 0.0 \\ 6 \end{array}$ | 0.0 3 | 0.0 4 | 0.1 4 | 0.3 7 | * |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.2 | 0.0 |
| 2003 | 0 | 4 | 2 | 2 | 2 | 2 | 2 | 2 | 0.02 | 4 | 5 | 6 | 9 | 0 | 4 |
|  |  | 0.1 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  |
| 2004 | 0 | 0 | 1 | 1 | 2 | 2 | 1 | 2 | 0.03 | 2 | 3 | 4 | 6 | 7 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  |
| 2005 | 0 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 0.02 | 3 | 3 | 5 | 6 | 7 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.1 | 0.0 |  |
| 2006 | 0 | 7 | 3 | 2 | 2 | 3 | 2 | 2 | 0.03 | 4 | 4 | 6 | 1 | 9 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 |  |
| 2007 | 0 | 6 | 2 | 1 | 1 | 1 | 3 | 8 | 0.06 | 4 | 9 | 6 | 4 | 4 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.2 | 0.0 |  |
| 2008 | 0 | 7 | 2 | 2 | 2 | 2 | 2 | 2 | 0.02 | 3 | 4 | 5 | 5 | 5 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  |
| 2009 | 0 | 2 | 2 | , | 2 | 2 | 2 | 2 | 0.03 | 3 | 3 | 6 | 6 | 5 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |  |
| 2010 | 0 | 3 | 2 | 1 | 1 | 1 | 2 | 4 | 0.02 | 2 | 2 | 3 | 3 | 6 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.01 | 0.0 | 0.0 | 0.1 | 0.0 | 0.0 |  |
| 2011 | 0 | 2 | 2 | 1 | 2 | 2 | 1 | 2 |  | 2 | 2 | 5 | 7 | 6 | * |
|  |  | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.04 | 0.0 | 0.0 | 0.1 | 0.0 | 0.1 |  |
| 2012 | 0 | 3 | 2 | 2 | 2 | 3 | 2 | 4 |  | 3 | 3 | 6 | 7 | 0 | * |

* Note: CV values $>1.00$ are noted by shadings. CVs could not be calculated for age $15+$ when more than one age class was present in the group.

Table 12. Un-weighted striped bass catch per unit effort (CPUE) by year-class, late March through May 2012. Values are presented by sex, area, and percent of total. CPUE is number of fish per hour in 1000 yards of experimental drift net.

| Year-class | Age | Pooled Unweighted CPUE | $\begin{aligned} & \text { \% of } \\ & \text { Total } \end{aligned}$ | Females |  | Males |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Potomac | Upper Bay | Potomac | Upper Bay |
| 2011 | 1 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |
| 2010 | 2 | 31.8 | 6.6 | 0.0 | 0.0 | 19.0 | 12.8 |
| 2009 | 3 | 101.2 | 20.9 | 0.0 | 0.0 | 44.4 | 56.8 |
| 2008 | 4 | 43.7 | 9.0 | 1.0 | 0.0 | 15.1 | 27.7 |
| 2007 | 5 | 44.3 | 9.1 | 1.4 | 1.5 | 13.9 | 27.5 |
| 2006 | 6 | 33.2 | 6.8 | 4.7 | 6.8 | 6.4 | 15.3 |
| 2005 | 7 | 40.8 | 8.4 | 2.6 | 6.2 | 6.0 | 26.0 |
| 2004 | 8 | 39.0 | 8.0 | 1.1 | 6.4 | 4.8 | 26.7 |
| 2003 | 9 | 42.9 | 8.8 | 1.6 | $\begin{array}{r} 15 . \\ 4 \\ \hline \end{array}$ | 4.1 | 21.8 |
| 2002 | 10 | 13.0 | 2.7 | 1.0 | 5.8 | 1.4 | 4.8 |
| 2001 | 11 | 28.3 | 5.8 | 1.6 | 8.8 | 2.1 | 15.8 |
| 2000 | 12 | 23.2 | 4.8 | 1.8 | 9.3 | 1.3 | 10.8 |
| 1999 | 13 | 7.6 | 1.6 | 0.8 | 4.5 | 0.6 | 1.7 |
| 1998 | 14 | 12.9 | 2.7 | 1.0 | 3.8 | 4.1 | 4.0 |
| $\leq 1997$ | 15+ | 23.0 | 4.7 | 3.0 | 19. | 0.0 | 0.7 |
| Total |  | 484.7 |  | 21.6 | 87. 5 | $\begin{array}{r} 123 . \\ 2 \end{array}$ | $\begin{array}{r} 252 . \\ \hline \end{array}$ |
| \% of Total |  |  |  | 4 | 18 | 25 | 52 |
| \% of Sex |  |  |  | 20 | 80 | 33 | 67 |
| \% of System |  |  |  | 15 | 26 | 85 | 74 |

Table 13. Striped bass catch per unit effort (CPUE) by year-class, weighted by spawning area*, late March through May 2012. Values are presented as percent of total, sex-specific, and area-specific C PUE. C PUE is nu mber of $f$ ish pe $r$ hour in 1000 yards of experimental drift net.

| Year-class | Age | Pooled Weighted CPUE | $\%$ of <br> Total | Females |  | Males |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Potomac | Upper Bay | Potomac | Upper Bay |
| 2011 | 1 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |
| 2010 | 2 | 15.2 | 5.7 | 0.0 | 0.0 | 7.3 | 7.8 |
| 2009 | 3 | 52.0 | 19.6 | 0.0 | 0.0 | 17.1 | 34.9 |
| 2008 | 4 | 23.2 | 8.8 | 0.4 | 0.0 | 5.8 | 17.0 |
| 2007 | 5 | 23.7 | 9.0 | 0.5 | 0.9 | 5.4 | 16.9 |
| 2006 | 6 | 17.8 | 6.7 | 1.8 | 4.2 | 2.5 | 9.4 |
| 2005 | 7 | 23.1 | 8.7 | 1.0 | 3.8 | 2.3 | 16.0 |
| 2004 | 8 | 22.6 | 8.5 | 0.4 | 3.9 | 1.8 | 16.4 |
| 2003 | 9 | 25.0 | 9.5 | 0.6 | 9.4 | 1.6 | 13.4 |
| 2002 | 10 | 7.4 | 2.8 | 0.4 | 3.6 | 0.5 | 3.0 |
| 2001 | 11 | 16.5 | 6.2 | 0.6 | 5.4 | 0.8 | 9.7 |
| 2000 | 12 | 13.6 | 5.1 | 0.7 | 5.7 | 0.5 | 6.7 |
| 1999 | 13 | 4.4 | 1.6 | 0.3 | 2.7 | 0.2 | 1.0 |
| 1998 | 14 | 6.7 | 2.5 | 0.4 | 2.3 | 1.6 | 2.5 |
| $\leq 1997$ | 15+ | 13.4 | 5.1 | 1.1 | 11.8 | 0.0 | 0.5 |
| Total |  | 264.7 |  | 8.3 | 53.8 | 47.5 | 155.1 |
| \% of Total |  |  |  | 3 | 20 | 18 | 59 |
| \% of Sex |  |  |  | 13 | 87 | 23 | 77 |
| \% of System |  |  |  | 15 | 26 | 85 | 74 |

* Spawning area weights used: Potomac (0.385); Upper Bay (0.615).

Table 14. Mean length-at-age (mm TL) statistics for male striped bass collected in the Potomac River and the Upper Bay, and areas combined, late March through May 2012.

| YEAR- <br> CLASS | AGE | AREA | N | MEAN | LCL | UCL | SD | SE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2010 | 2 | POTOMAC | 9 | 326 | 301 | 350 | 32 | 11 |
|  |  | UPPER | 2 | 320 | 244 | 396 | 8 | 6 |
|  |  | COMBINED | 11 | 325 | 305 | 344 | 28 | 9 |
| 2009 | 3 | POTOMAC | 16 | 400 | 379 | 420 | 39 | 10 |
|  |  | UPPER | 15 | 376 | 353 | 398 | 40 | 10 |
|  |  | COMBINED | 31 | 388 | 373 | 403 | 41 | 7 |
| 2008 | 4 | POTOMAC | 4 | 463 | 336 | 590 | 80 | 40 |
|  |  | UPPER | 13 | 446 | 403 | 489 | 71 | 20 |
|  |  | COMBINED | 17 | 450 | 414 | 486 | 71 | 17 |
| 2007 | 5 | POTOMAC | 13 | 535 | 510 | 561 | 43 | 12 |
|  |  | UPPER | 9 | 536 | 489 | 582 | 61 | 20 |
|  |  | COMBINED | 22 | 536 | 514 | 558 | 50 | 11 |
| 2006 | 6 | POTOMAC | 10 | 572 | 537 | 607 | 49 | 15 |
|  |  | UPPER | 8 | 641 | 568 | 714 | 87 | 31 |
|  |  | COMBINED | 18 | 603 | 565 | 640 | 75 | 18 |
| 2005 | 7 | POTOMAC | 9 | 618 | 597 | 638 | 27 | 9 |
|  |  | UPPER | 27 | 646 | 618 | 674 | 70 | 14 |
|  |  | COMBINED | 36 | 639 | 617 | 660 | 63 | 11 |
| 2004 | 8 | POTOMAC | 4 | 672 | 544 | 800 | 81 | 40 |
|  |  | UPPER | 43 | 748 | 729 | 772 | 79 | 12 |
|  |  | COMBINED | 47 | 742 | 718 | 767 | 81 | 12 |
| 2003 | 9 | POTOMAC | 7 | 824 | 791 | 857 | 36 | 14 |
|  |  | UPPER | 43 | 777 | 754 | 800 | 74 | 11 |
|  |  | COMBINED | 50 | 784 | 763 | 804 | 71 | 10 |
| 2002 | 10 | POTOMAC | 3 | 855 | 781 | 929 | 30 | 17 |
|  |  | UPPER | 8 | 850 | 790 | 910 | 72 | 25 |
|  |  | COMBINED | 11 | 851 | 810 | 893 | 61 | 19 |
| 2001 | 11 | POTOMAC | 4 | 839 | 776 | 901 | 39 | 20 |
|  |  | UPPER | 11 | 917 | 861 | 974 | 84 | 25 |
|  |  | COMBINED | 15 | 896 | 851 | 942 | 82 | 21 |
| 2000 | 12 | POTOMAC | 3 | 974 | 862 | 1086 | 45 | 26 |
|  |  | UPPER | 5 | 984 | 897 | 1072 | 70 | 31 |
|  |  | COMBINED | 8 | 981 | 932 | 1029 | 59 | 21 |
| 1999 | 13 | POTOMAC | 2 | 992 | 782 | 1201 | 23 | 17 |
|  |  | UPPER | , | 1042 | - | - | - | - |
|  |  | COMBINED | 3 | 1008 | 925 | 1092 | 34 | 19 |
| 1998 | 14 | POTOMAC | 1 | 1138 | - | - | - | - |
|  |  | UPPER | 9 | 991 | 950 | 1032 | 54 | 18 |
|  |  | COMBINED | 10 | 1006 | 957 | 1055 | 69 | 22 |
| 1997 | 15 | POTOMAC | 0 | - | - | - | - | - |
|  |  | UPPER | 1 | 1043 | - | - | - | - |
|  |  | COMBINED | 1 | 1043 | - | - | - | - |

Table 15. Mean length-at-age (mm TL) statistics for female striped bass collected in the Potomac River and the Upper Bay, and areas combined, late March through May 2012.

| YEAR- <br> CLASS | AGE | AREA | N | MEAN | LCL | UCL | SD | SE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2008 | 4 | $\begin{aligned} & \hline \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & 1 \\ & 0 \\ & 1 \\ & \hline \end{aligned}$ | $\begin{gathered} 468 \\ - \\ 468 \end{gathered}$ |  |  |  | - |
| 2007 | 5 | $\begin{gathered} \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \end{gathered}$ | $\begin{aligned} & 0 \\ & 1 \\ & 1 \end{aligned}$ | $\begin{aligned} & 544 \\ & 544 \end{aligned}$ |  |  |  |  |
| 2006 | 6 | $\begin{aligned} & \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & 4 \\ & 3 \\ & 7 \\ & \hline \end{aligned}$ | $\begin{aligned} & 647 \\ & 655 \\ & 651 \end{aligned}$ | $\begin{aligned} & 527 \\ & 510 \\ & 592 \\ & \hline \end{aligned}$ | $\begin{aligned} & 767 \\ & 801 \\ & 709 \\ & \hline \end{aligned}$ | $\begin{aligned} & 75 \\ & 59 \\ & 63 \\ & \hline \end{aligned}$ | $\begin{aligned} & 38 \\ & 34 \\ & 24 \end{aligned}$ |
| 2005 | 7 | $\begin{aligned} & \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & 2 \\ & 6 \\ & 8 \end{aligned}$ | $\begin{aligned} & 697 \\ & 657 \\ & 667 \end{aligned}$ | $\begin{aligned} & 290 \\ & 620 \\ & 635 \\ & \hline \end{aligned}$ | $\begin{gathered} 1104 \\ 695 \\ 700 \\ \hline \end{gathered}$ | $\begin{aligned} & 45 \\ & 36 \\ & 39 \\ & \hline \end{aligned}$ | $\begin{aligned} & 32 \\ & 14 \\ & 14 \end{aligned}$ |
| 2004 | 8 | $\begin{aligned} & \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & 0 \\ & 6 \\ & 6 \\ & \hline \end{aligned}$ | $\begin{aligned} & 794 \\ & 794 \\ & \hline \end{aligned}$ | 711 <br> 711 | $\begin{aligned} & 876 \\ & 876 \\ & \hline \end{aligned}$ | $\begin{aligned} & 79 \\ & 79 \\ & \hline \end{aligned}$ | $\begin{aligned} & 32 \\ & 32 \\ & \hline \end{aligned}$ |
| 2003 | 9 | $\begin{gathered} \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \\ \hline \end{gathered}$ | $\begin{gathered} 5 \\ 32 \\ 37 \\ \hline \end{gathered}$ | $\begin{aligned} & 908 \\ & 866 \\ & 872 \\ & \hline \end{aligned}$ | $\begin{aligned} & 861 \\ & 838 \\ & 847 \\ & \hline \end{aligned}$ | $\begin{aligned} & 955 \\ & 895 \\ & 898 \\ & \hline \end{aligned}$ | $\begin{aligned} & 38 \\ & 80 \\ & 76 \\ & \hline \end{aligned}$ | $\begin{aligned} & 17 \\ & 14 \\ & 13 \\ & \hline \end{aligned}$ |
| 2002 | 10 | $\begin{gathered} \hline \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \end{gathered}$ | $\begin{gathered} \hline 4 \\ 18 \\ 22 \end{gathered}$ | $\begin{aligned} & \hline 928 \\ & 939 \\ & 937 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 854 \\ & 917 \\ & 918 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline 1002 \\ 961 \\ 957 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 47 \\ & 44 \\ & 44 \\ & \hline \end{aligned}$ | $\begin{gathered} 23 \\ 10 \\ 9 \end{gathered}$ |
| 2001 | 11 | $\begin{gathered} \hline \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \end{gathered}$ | $\begin{gathered} \hline 8 \\ 12 \\ 20 \end{gathered}$ | $\begin{aligned} & 964 \\ & 966 \\ & 965 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 919 \\ & 941 \\ & 945 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline 1008 \\ 990 \\ 985 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 53 \\ & 38 \\ & 43 \\ & \hline \end{aligned}$ | $\begin{aligned} & 19 \\ & 11 \\ & 10 \end{aligned}$ |
| 2000 | 12 | $\begin{gathered} \hline \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \\ \hline \end{gathered}$ | $\begin{gathered} \hline 4 \\ 13 \\ 17 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 1031 \\ & 1031 \\ & 1031 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline 960 \\ 999 \\ 1005 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 1102 \\ & 1062 \\ & 1056 \\ & \hline \end{aligned}$ | $\begin{array}{r} \hline 45 \\ 53 \\ 50 \\ \hline \end{array}$ | $\begin{aligned} & 22 \\ & 15 \\ & 12 \end{aligned}$ |
| 1999 | 13 | $\begin{gathered} \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \end{gathered}$ | $\begin{aligned} & \hline 0 \\ & 9 \\ & 9 \\ & \hline \end{aligned}$ | $\begin{gathered} - \\ 1040 \\ 1040 \\ \hline \end{gathered}$ | $\begin{aligned} & 1002 \\ & 1002 \end{aligned}$ | $\begin{aligned} & 1079 \\ & 1079 \end{aligned}$ | $\begin{array}{r} 50 \\ 50 \\ \hline \end{array}$ | $\begin{gathered} \hline- \\ 17 \\ 17 \\ \hline \end{gathered}$ |
| 1998 | 14 | $\begin{gathered} \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \end{gathered}$ | $\begin{gathered} \hline 3 \\ 7 \\ 10 \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 1081 \\ & 1039 \\ & 1052 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline 830 \\ 1001 \\ 1008 \\ \hline \end{gathered}$ | $\begin{aligned} & 1331 \\ & 1078 \\ & 1096 \end{aligned}$ | $\begin{gathered} \hline 101 \\ 41 \\ 62 \\ \hline \end{gathered}$ | $\begin{aligned} & 58 \\ & 16 \\ & 19 \end{aligned}$ |
| 1997 | 15 | $\begin{gathered} \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \end{gathered}$ | $\begin{aligned} & 2 \\ & 4 \\ & 6 \end{aligned}$ | $\begin{aligned} & 1090 \\ & 1078 \\ & 1082 \\ & \hline \end{aligned}$ | $\begin{gathered} 931 \\ 945 \\ 1013 \end{gathered}$ | $\begin{aligned} & 1248 \\ & 1211 \\ & 1150 \\ & \hline \end{aligned}$ | $\begin{aligned} & 18 \\ & 83 \\ & 65 \\ & \hline \end{aligned}$ | $\begin{aligned} & 13 \\ & 42 \\ & 27 \\ & \hline \end{aligned}$ |
| 1996 | 16 | $\begin{gathered} \hline \text { POTOMAC } \\ \text { UPPER } \\ \text { COMBINED } \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 4 \\ & 3 \\ & 7 \\ & \hline \end{aligned}$ | $\begin{aligned} & 1080 \\ & 1119 \\ & 1097 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline 988 \\ 981 \\ 1045 \\ \hline \end{gathered}$ | $\begin{aligned} & 1172 \\ & 1257 \\ & 1148 \\ & \hline \end{aligned}$ | $\begin{aligned} & 58 \\ & 56 \\ & 56 \\ & \hline \end{aligned}$ | $\begin{aligned} & 29 \\ & 32 \\ & 21 \\ & \hline \end{aligned}$ |
| 1995 | 17 | $\begin{aligned} & \hline \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & 1 \\ & 0 \\ & 1 \\ & \hline \end{aligned}$ | $\begin{gathered} 1197 \\ - \\ 1197 \end{gathered}$ |  |  |  | - |
| 1994 | 18 | $\begin{aligned} & \hline \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & 0 \\ & 2 \\ & 2 \end{aligned}$ | $\begin{gathered} - \\ 1158 \\ 1158 \\ \hline \end{gathered}$ | $\begin{gathered} - \\ 695 \\ 695 \\ \hline \end{gathered}$ | $\begin{aligned} & 1622 \\ & 1622 \end{aligned}$ | $\begin{aligned} & 52 \\ & 52 \end{aligned}$ | $\begin{aligned} & 37 \\ & 37 \end{aligned}$ |
| 1993 | 19 | $\begin{aligned} & \hline \text { POTOMAC } \\ & \text { UPPER } \\ & \text { COMBINED } \end{aligned}$ | $\begin{aligned} & \hline 0 \\ & 1 \\ & 1 \\ & \hline \end{aligned}$ | $\begin{gathered} \hline- \\ 1196 \\ 1196 \\ \hline \end{gathered}$ |  |  |  |  |

Table 16. Index of spawning biomass by year, for female striped bass $\geq 500 \mathrm{~mm}$ TL sampled from s pawning a reas of the C hesapeake Bay during M arch, April and May s ince 1985. T he index is selectivity-corrected CPUE converted to biomass (kg) using parameters from a lengthweight regression.

| Year | Upper Bay | Potomac River |
| :---: | :---: | :---: |
| 1985 | 64.93 | 25.90 |
| 1986 | 151.95 | 45.70 |
| 1987 | 400.49 | 88.84 |
| 1988 | 250.32 | 63.60 |
| 1989 | 120.29 | 80.54 |
| 1990 | 98.42 | 62.52 |
| 1991 | 109.38 | 138.65 |
| 1992 | 274.95 | 379.35 |
| 1993 | 278.52 | 420.88 |
| 1994 | 87.26 | Not Sampled |
| 1995 | 547.66 | 293.77 |
| 1996 | 347.87 | 391.57 |
| 1997 | 240.42 | 362.33 |
| 1998 | 155.86 | 226.78 |
| 1999 | 168.44 | 280.82 |
| 2000 | 192.75 | 325.22 |
| 2001 | 479.14 | 272.49 |
| 2002 | 276.46 | 398.94 |
| 2003 | 563.41 | 118.46 |
| 2004 | 376.19 | 530.23 |
| 2005 | 469.68 | 195.80 |
| 2006 | 406.22 | 458.23 |
| 2007 | 418.54 | 263.27 |
| 2008 | 228.60 | 162.78 |
| 2009 | 482.52 | 189.77 |
| 2010 | 279.71 | 212.79 |
| 2011 | 167.56 | 105.43 |
| 2012 | 799.21 | 149.96 |
| Average | 301.31 | 231.28 |

Figure 1. Drift gill net sampling locations in spawning areas of the Upper Chesapeake Bay and the Potomac River, late March - May 2012.


Figure 2. Daily effort-corrected catch of female and male striped bass, with surface water and air temperatures in the spawning reach of the Potomac River, late March through May 2012. Effort is standardized as 1000 square yards of experimental gill net per hour. Note different scales.
Females

Date

Date
$\square$

Figure 3. Daily effort-corrected catch of female and male striped bass, with surface water and air temperatures in the spawning reach of the Upper Chesapeake Bay, late March through May 2012. Effort is standardized as 1000 square yards of experimental drift gill net per hour. Note different scales.



Date
$\square$

Figure 4. Length frequency of male and female striped bass from the spawning areas of the Upper Chesapeake Bay and Potomac River, late March through May 2012. Note different scales.



Figure 5. Length group CPUE (uncorrected and corrected for gear selectivity) of male striped bass collected from spawning areas of the Upper Bay and Potomac River, late March May 2012. CPUE is the number of fish captured per hour in 1000 square yards of experimental drift net.


Length group (mm)


Length group (mm)

Figure 6. Length group CPUE (uncorrected and corrected for gear selectivity) of female striped bass collected from spawning areas of the Upper Bay and Potomac River, late March May 2012. CPUE is the number of fish captured per hour in 1000 square yards of experimental drift net.



Length group (mm)

Figure 7. Mean length (mm TL) by year for individual ages of male striped bass sampled from spawning areas of the Potomac River and Upper Chesapeake Bay during late March through May, 1985-2012. Error bars are $\pm 2$ standard errors (SE). The Potomac River was not sampled in 1994. *Note difference in scales on y-axis.


Figure 7. Continued.






Figure 8. Mean length (mm TL) by year for individual ages of female striped bass sampled from spawning areas of the Potomac River and Upper Chesapeake Bay during late March through May, 1985-2012. Error bars are $\pm 2$ standard errors (SE). Note the Potomac River was not sampled in 1994. *Note difference in scales on $y$-axis.


Figure 8. Continued.


Figure 9. Maryland Chesapeake Bay spawning stock indices used in the coastal assessment. These are selectivity-corrected estimates of CPUE by year for ages 2 through 15-plus. Areas and sexes are pooled, although the contribution of sexes is shown in the stacked bars. Note different scales.








Year
II-242

Figure 9. Continued.







## Year

Figure 10. Percentage (selectivity-corrected CPUE) of female striped bass that were age 8 and older sampled from experimental drift gill nets set in spawning reaches of the Potomac River, Choptank River and the Upper Chesapeake Bay, late March through May, 1985-2012 (Choptank River to 1996). Effort is standardized as 1000 square yards of net per hour. Areaspecific indices were weighted based on the relative size of the spawning areas before area-specific indices were pooled.*

*Weights for spawning areas (1985-1996): Upper Bay=0.59; Potomac River=0.37; Choptank River=0.04.
(1997 - Present): Upper Bay=0.615; Potomac River=0.385 (Hollis 1967).

Figure 11. Percentage (selectivity-corrected CPUE) of male and female striped bass age 8 and over sampled from experimental drift gill nets set in spawning reaches of the Potomac River, Choptank River and the Upper Chesapeake Bay, late March through May, 1985-2012 (Choptank River to 1996). Effort is standardized as 1000 square yards of net per hour. Areaspecific indices were weighted based on the relative size of the spawning areas before area-specific indices were pooled.*

*Weights for spawning areas (1985-1996): Upper Bay=0.59; Potomac River=0.37; Choptank River=0.04.
(1997-Present): Upper Bay=0.615; Potomac River=0.385; (Hollis 1967).

Figure 12. Biomass (kg) of female striped bass greater than or equal to 500 mm TL collected from experimental drift gill nets fished in two spawning areas of the Maryland Chesapeake Bay during late March through May, 1985-2012. The index is corrected for gear selectivity, and bootstrap $95 \%$ confidence intervals are shown around each point.


# PROJECT NO. 2 

JOB NO. 3
TASK NO. 3

# MARYLAND JUVENILE STRIPED BASS SURVEY 

Prepared by Eric Q. Durell

## INTRODUCTION

The primary objective of Project 2, Job 3, Task 3 was to document annual year-class success for young-of-the-year (YOY) striped bass (Morone saxatilis) in Chesapeake Bay. Annual indices of relative abundance provide an early indicator of future a dult stock recruitment (Schaefer 1972; Goodyear 1985) and document annual variation and long-term trends in abundance and distribution.

## METHODS

## Sample Area and Intensity

Juvenile indices were derived from sampling at 22 fixed stations within Maryland's portion of the Chesapeake B ay (Table 1, Figure 1). Sample sites were divided a mong four of the major spawning and nursery areas; seven each in the Potomac River and Head of Bay areas and four each in the Nanticoke and Choptank rivers.

Stations have been sampled continuously since 1954, with changes in some station locations. Recent e rosion at $t$ he $W$ orton $C$ reek $s$ ite ( site \#11) in the $H$ ead of $B$ ay ar ea prompted $t$ he establishment of an auxiliary site directly across the creek called Handy Point (site \#164). Handy Point will be assessed as an eventual replacement for Worton Creek.

From 1954 to 1961, Maryland's juvenile surveys included inconsistent stations and rounds.

Sample sizes ranged from 34 to 46 . Indices derived for this period include only stations which are consistent with subsequent years. In 1962, stations were standardized and a second sample round was added for a total of 88 samples. A third sample round, added in 1966, increased sample size to 132.

Sites were sampled monthly, with rounds ( sampling e xcursions) occurring during July (Round I), August (Round II), and September (Round III). Replicate seine hauls, a minimum of thirty minutes apart, were taken at each site in each sample round. This protocol produced a total of 132 samples from which Bay-wide means were calculated.

Auxiliary stations have been sampled on an inconsistent basis and were not included in survey indices. These data enhance geographical coverage in rivers with permanent stations or provide information from other river systems. They are also useful for replacement of permanent stations when necessary. Replicate hauls at auxiliary stations were discontinued in 1992 to conserve time and allow increased geographical coverage of spawning areas. Auxiliary stations were sampled at the Head of Bay (Susquehanna Flats and one downstream station) and the Patuxent River (Table 1, Figure 1).

## Sample Protocol

A $30.5-\mathrm{m} \times 1.24-\mathrm{m}$ bagless beach seine of untreated $6.4-\mathrm{mm}$ bar mesh was set by hand. One end was held on shore while the other was fully stretched perpendicular from the beach and swept with the current. Ideally, the area swept was equivalent to a $729 \mathrm{~m}^{2}$ quadrant. When depths of $1.6-\mathrm{m}$ or greater were encountered, the offshore end was deployed along this depth contour. An estimate of distance from the beach to this depth was recorded.

Striped bass and selected other species were separated into 0 and $1+$ age groupings. Ages were assigned from length-frequencies and verified through scale examination. Age 0 fish were
measured (mm total length) from a random sample of up to 30 individuals per site and round. All other finfish were identified to species and counted.

Additional data were collected at each site and sample round. These included: time of first haul, maximum distance from shore, weather, maximum depth, surface water temperature $\left({ }^{\circ} \mathrm{C}\right)$, tide stage, surface salinity (ppt), pr imary a nd secondary bot tom substrates, a nd submerged aquatic vegetation within the sample area (ranked by quartiles). Dissolved oxygen (DO), pH , and turbidity (Secchi disk) were added in 1997. All data were entered and archived in Statistical Analysis System (SAS) databases (SAS 1990).

## Estimators

The most commonly referenced striped bass 'juvenile index' is the arithmetic mean (AM). The AM has been used to predict harvest in New York waters (Schaefer 1972). Goodyear (1985) validated this index as a predictor of harvest in the Chesapeake Bay. The AM is an unbiased estimator of the mean regardless of the underlying frequency distribution (McConnaughey and Conquest 1992). The AM, however, is sensitive to high sample values (Sokol and Rolhf 1981). Additionally, detection of significant differences between annual ar ithmetic means is often not possible due to high variances (Heimbuch et al. 1983; Wilson and Wiesburg 1991).

The geometric mean (GM) was adopted by the Atlantic States Marine Fisheries Commission (ASMFC) Striped Bass Technical Committee in 1992 as the preferred index of relative abundance to model $s$ tock status. $T$ he GM is calculated from the $\log _{e}(x+1)$ transformation, where $x$ is an individual seine haul catch. One is added to all catches in order to transform zero catches, because the log of 0 does not exist (Ricker 1975). Since the $\log _{\mathrm{e}}$-transformation stabilizes the variance of catches (Richards 1992) the GM estimate is more precise than the AM and is not as sensitive to a single large sample value. It is almost always lower than the AM (Ricker 1975). The GM is
presented with $95 \%$ confidence intervals (CIs) which are calculated as antilog $\left(\log _{\mathrm{e}}(\mathrm{x}+1)\right.$ mean $\pm 2$ standard errors), and provide a visual depiction of sample variability.

A third estimator, the proportion of positive hauls (PPHL), is the ratio of hauls containing juvenile striped bass to total hauls. Because the PPHL is based on the binomial distribution, it is very robust to bias a nd sampling e rror and greatly reduces variances (Green 1979). Its use as supplementary information is appropriate since seine estimates are often neither normally nor lognormally distributed (Richards 1992).

Comparison of these three indices is one method of assessing their accuracy. Similar trends among indices create more certainty that indices reflect actual changes in population abundance. Greatly diverging trends may identify error in one or more of the indices.

Bay-wide annual indices are compared to the target period average (TPA). The TPA is the average of indices from 1959 through 1972. These years have been suggested as a period of stable biomass and general stock he alth (ASMFC 1989) and "an a ppropriate stock rebuilding target" (Gibson 1993). The TPA provides a fixed reference representing an average index produced by a healthy population. A fixed reference is an advantage over a time-series average that is revised annually and may be significantly biased by long-term trends in annual indices.

Differences among annual means were tested with an analysis of variance (GLM; SAS 1990) on the $\log _{e}(\mathrm{x}+1)$ transformed data. Means were considered significant at the $\mathrm{p}=0.05$ level. Duncan's multiple range test was used to differentiate means.

## RESULTS

## Bay-wide Means

A total of 117 YOY striped bass was collected at permanent stations in 2012, with individual samples yielding between 0 and 12 fish. The $\mathrm{AM}(0.9)$ and $\mathrm{GM}(0.49)$ were both the lowest in their respective time-series (Table 2 and 3, Figures 2 and 3). The PPHL was 0.35 , indicating that $35 \%$ of samples produced juvenile striped bass (Table 4, Figure 4).

A one-way analysis of variance (ANOVA) performed on the $\log _{\mathrm{e}}$-transformed catch values indicated significant differences among annual means (ANOVA: $\mathrm{P}<0.0001$ ) (SAS 1990). Duncan's multiple range test $(\mathrm{p}=0.05)$ found that the $2012 \log _{\mathrm{e}}$-mean was significantly lower than 46 years of the time-series, and indiscernible from the seven lowest years.

## System Means

Head of Bay - In 42 samples, 28 juveniles were collected at the Head of Bay sites for an AM of 0.7, less than the time-series average (11.7) and the TPA of 17.3 (Table 2, Figure 5). The GM of 0.44 was also below the time-series average (5.55) and TPA (7.27) (Table 3, Figure 6). Differences in annual $\log _{\mathrm{e}}$-means were significant (ANOVA: $\mathrm{P}<0.0001$ ). Duncan's multiple range test $(\mathrm{p}=0.05)$ found the 2012 Head of Bay $\log _{\mathrm{e}}$-mean significantly less than 44 years of the time-series, and indiscernible from the smallest 11 year-classes of the time-series.

Potomac River - A total 72 juveniles was collected in 42 samples on the Potomac River. The AM of 1.7 was less than the TPA (9.2) and the time-series average (8.3) (Table 2, Figure 5). The GM of 0.95 was also less than the time-series average (3.62) and TPA (3.93) (Table 3, Figure 7). Analysis of va riance of $\log _{\mathrm{e}}$-means indicated significant di fferences am ong y ears (ANOVA: $\mathrm{P}<0.0001$ ). D uncan's multiple range test ( $\mathrm{p}=0.05$ ) r anked the 2012 Potomac R iver year-class significantly smaller than 26 years, and not significantly different than the 29 other years of the time-
series.
Choptank River - A total of 3 juveniles was collected in 24 Choptank River samples. The AM of 0.1 was lower than the time-series average of 21.6 and the TPA of 10.8 (Table 2, Figure 5). The GM of 0.08 was also lower than its time-series average (8.12) and TPA (5.00) (Table 3, Figure 8). Differences among years were significant (ANOVA: $\mathrm{P}<0.0001$ ). Duncan's multiple range test $(\mathrm{p}=0.05) \mathrm{r}$ anked t he 2012 Choptank $R$ iver year-class $s$ ignificantly s maller tha n 41 years, a nd indiscernible from 14 years of the time series.

Nanticoke River - A total of 14 juveniles was collected in 24 samples on the Nanticoke River. The AM of 0.6 was below the time-series average (8.4) and TPA (8.6) (Table 2, Figure 5). The GM of 0.37 was also less than its time-series average (3.76) and TPA (3.12) (Table 3, Figure 9). The Nanticoke River also exhibited significant differences among years (ANOVA: $\mathrm{P}<0.0001$ ). Duncan's multiple range test $(\mathrm{p}=0.05)$ ranked the 2012 index significantly smaller than 27 years of the time-series. The 2012 index was statistically indiscernible from the remaining 28 years of the time-series.

## Auxiliary Indices

At the Head of Bay auxiliary sites, 39 juveniles were caught in 21 samples, resulting in an AM of 1.9 and a GM of 0.71 . Both indices were less than their respective time-series averages (Table 5).

On the Patuxent River, one YOY striped bass was caught in 18 samples for an AM of 0.1 and a GM of 0.04. Both Patuxent River indices were less than their respective time-series averages and medians (Table 5).

## DISCUSSION

By all measures, striped bass recruitment in Maryland's Chesapeake Bay was very poor in 2012. The bay-wide AM and GM indices were both the lowest in the history of the survey (Tables 2 and 3). Duncan's multiple range test $(\mathrm{p}=0.05)$ found the $2012 \log _{\mathrm{e}}$-mean was indiscernible from the seven smallest year-classes on record (1959, 1980, 1981, 1983, 1985, 1988 and 1990). YOY striped bass occurred in only $35 \%$ of the samples ( $\mathrm{PPHL}=0.35$ ), another indication of a small year-class and the lowest observed since 1959 (Table 4, Figure 4).

Recruitment was below average in all individual systems. The 2012 year-class was among the smallest ever recorded in the Head of Bay ( $5^{\text {th }}$ percentile), Choptank River (lowest on record), and N anticoke R iver ( $2^{\text {nd }}$ percentile) as m easured by g eometric m eans. The P otomac R iver performed slightly better with a GM at the $21^{\text {st }}$ percentile of the time-series.

High variability in annual spawning success is a hallmark of striped bass populations, which are known for producing occasional dominant year-classes under optimal spawning conditions. The disparity in spawning success between 2011 (among the best years on record) and 2012 (the worst year on record) may be attributable to differing weather conditions during the spawning season in those years. Ulanowicz and Polgar (1980) speculated that high variability in annual recruitment is due primarily to extrinsic environmental factors. Boynton et al (1977) noted that recruitment may not be limited by $t$ he num ber of $s$ triped ba ss on $t$ he $P$ otomac $R$ iver $s$ pawning $g$ rounds, a nd demonstrated that dominant year-classes were associated with colder than normal winters and higher than normal spring river flows. Consistent with these hypotheses, temperature and precipitation in the months before and during the 2011 and 2012 spawns were markedly different according to the NOAA National Climatic Data Center (NCDC). The NCDC (2012) ranked the period January-April 2011, a year of high recruitment, colder and wetter than normal. The NCDC ranked January-April

2012, a year of very low recruitment, the warmest on record and the driest since 1985. Furthermore, the pattern of recruitment success in 2011 followed by subsequent recruitment failure in 2012 was also apparent $\mathrm{i} n$ ot her a nadromous s pecies doc umented by t he M DDNR seine survey (http://dnr.maryland.gov/fisheries/juvindex/). This points to the influence of extrinsic environmental factors that were not conducive to the success of anadromous spawning behavior in general in 2012.

## RELATIONSHIP OF AGE 0 TO AGE 1 INDICES

## INTRODUCTION

Indices of age 1 (yearling) striped bass (Table 6) developed from the Maryland juvenile striped bass surveys were tested for relationship to YOY indices by year-class. Previous analysis yielded a significant relationship with age 0 indices explaining $73 \%(\mathrm{P} \leq 0.001)$ of the variability in age 1 i ndices one $y$ ear 1 ater (MD D NR 1994). T he $s$ trength of $t$ his $r$ elationship 1 ed to the incorporation of the age 1 index into coastal stock assessment models by the ASMFC Striped Bass Technical Committee. The utility of age 1 indices as a potential fishery independent verification of the YOY index also makes this relationship of interest.

## METHODS

Age 1 indices were developed from the Maryland beach seine data (Table 6). Size ranges were used to determine catch of age 1 fish from records prior to 1991. Since 1991, striped bass have been separated into 0,1 and $2+$ age groups in the recorded data. Age groups were assigned by length-frequencies and later confirmed through direct examination of scales. Annual indices were computed as arithmetic means of $\log$ transformed catch values $\left[\log _{e}(c a t c h+1)\right]$. Regression analysis was used to test the relationship between age 0 and subsequent age 1 mean catch per haul.

## RESULTS AND DISCUSSION

The relationship of a ge-0 t o s ubsequent a ge- 1 relative abunda nce w as significant and explained $61 \%$ of the variability $\left(\mathrm{r}^{2}=0.606, \mathrm{p} \leq 0.001\right)$ in the age 1 indices (Figure 10). The equation that best described this relationship was: $\mathrm{C}_{1}=(0.18916)\left(\mathrm{C}_{0}\right)-0.07263$, where $\mathrm{C}_{1}$ is the age 1 index and $C_{0}$ is the age 0 index. While still significant, the model has lost predictive power since 1994 when $r^{2}=0.73$. The addition of quadratic and cubic terms yielded even poorer fits.

This year's actual index of age 1 striped bass ( 0.30 ) was less than the index of 0.37 predicted by the regression analysis. Examination of residuals (Figure 11) shows that this regression equation can be used to predict subsequent yearling striped bass abundance with reasonable certainty in the case of small and average sized year-classes. Estimates of future abundance of age 1 striped bass are less reliable for dominant year-classes such as 2011. L ower than expected abundance of age 1 striped bass $m$ ay be an indication of de nsity-dependent processes ope rating at hi gh 1 evels of abundance, such as cannibalism, increased competition for food, increased spatial distribution, or overwintering mortality. H igher t han expected a bundance of age 1 s triped ba ss m ay identify particularly good conditions that enhanced survival.

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Figure 10. Relationship between age 0 and subsequent age 1 striped bass indices.
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Table 1. Maryland juvenile striped bass survey sample sites.

| Site | River or | Area or |
| :--- | :--- | :--- |
| Number | Creek | Nearest Land Mark |

## HEAD-OF-CHESAPEAKE BAY SYSTEM

| $* 58$ | Susquehanna Flats | North side Spoil Island, 1.9 miles south of Tyding's Park |
| ---: | :--- | :--- |
| $* 130$ | Susquehanna Flats | North side of Plum Point |
| *144 | Susquehanna Flats | Tyding's Estate, west shore of flats |
| *132 | Susquehanna Flats | 0.2 miles east of Poplar Point |
| * 59 | Northeast River | Carpenter Point, K.O.A. Campground beach |
| 3 | Northeast River | Elk Neck State Park beach |
| 4 | Elk River | Welch Point, Elk River side |
| 5 | Elk River | Hyland Point Light |
| 115 | Bohemia River | Parlor Point |
| 160 | Sassafras River | Sassafras N.R.M.A., opposite Ordinary Point |
| 10 | Sassafras River | Howell Point, 500 yards east of point |
| 11 | Worton Creek | Mouth of Tim's Creek, west shore |
| *164 | Worton Creek | Handy Point, 0.3 miles west of Green Point Wharf |
| $* 88$ | Chesapeake Bay | Beach at Tolchester Yacht Club |

## POTOMAC RIVER SYSTEM

139
50
51
52
163
56
55

Potomac River
Potomac River
Hallowing Point, VA
Indian Head, old boat basin
Potomac River Liverpool Point, south side of pier
Potomac River Blossom Point, mouth of Nanjemoy Creek
Potomac River Aqualand Marina
Potomac River St. George Island, south end of bridge
Wicomico River Rock Point

[^5]Table 1. Continued.

| Site | River or | Area or |
| :--- | :--- | :--- |
| Number | Creek | Nearest Land Mark |

## CHOPTANK RIVER SYSTEM

| 2 | Tuckahoe Creek | Northeast side near mouth |
| ---: | :--- | :--- |
| 148 | Choptank River | North side of Jamaica Point |
| 161 | Choptank River | Dickinson Bay, 0.5 miles from Howell Point |
| 29 | Choptank River | Castle Haven, northeast side |

## NANTICOKE RIVER SYSTEM

36 Nanticoke River Sharptown, pulpwood pier

Nanticoke River
Nanticoke River
Nanticoke River Tyaskin Beach

## PATUXENT RIVER SYSTEM

* 85
* 86
* 91
* 92
* 106
* 90

Patuxent River
Patuxent River
Patuxent River
Patuxent River
Patuxent River
Patuxent River

Selby Landing
Nottingham, Windsor Farm
Milltown Landing
Eagle Harbor
Sheridan Point
Peterson Point

[^6]Table 2. Maryland juvenile striped bass survey arithmetic mean catch per haul at permanent sites.

| Year | Head-of-Bay | Potomac River | Choptank River | Nanticoke River | Bay-wide |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1954 | 0.9 | 5.2 | 1.2 | 25.1 | 5.2 |
| 1955 | 4.4 | 5.7 | 12.5 | 5.9 | 5.5 |
| 1956 | 33.9 | 6.2 | 9.8 | 8.2 | 15.2 |
| 1957 | 5.4 | 2.5 | 2.1 | 1.3 | 2.9 |
| 1958 | 28.2 | 8.4 | 19.5 | 22.5 | 19.3 |
| 1959 | 1.9 | 1.6 | 0.1 | 1.8 | 1.4 |
| 1960 | 9.3 | 4.3 | 9.0 | 4.7 | 7.1 |
| 1961 | 22.1 | 25.8 | 6.0 | 1.5 | 17.0 |
| 1962 | 11.4 | 19.7 | 6.1 | 6.6 | 12.2 |
| 1963 | 6.1 | 1.1 | 5.4 | 4.1 | 4.0 |
| 1964 | 31.0 | 29.1 | 10.6 | 13.3 | 23.5 |
| 1965 | 2.2 | 3.4 | 9.5 | 21.6 | 7.4 |
| 1966 | 32.3 | 10.5 | 13.6 | 3.3 | 16.7 |
| 1967 | 17.4 | 1.9 | 5.3 | 4.1 | 7.8 |
| 1968 | 13.1 | 0.7 | 6.3 | 9.0 | 7.2 |
| 1969 | 26.6 | 0.2 | 4.8 | 6.2 | 10.5 |
| 1970 | 33.1 | 20.1 | 57.2 | 17.1 | 30.4 |
| 1971 | 23.7 | 8.5 | 6.3 | 2.0 | 11.8 |
| 1972 | 12.1 | 1.9 | 11.0 | 25.0 | 11.0 |
| 1973 | 24.5 | 2.1 | 1.3 | 1.1 | 8.9 |
| 1974 | 19.9 | 1.5 | 15.3 | 3.9 | 10.1 |
| 1975 | 7.6 | 7.8 | 4.7 | 5.2 | 6.7 |
| 1976 | 9.9 | 3.2 | 2.4 | 1.7 | 4.9 |
| 1977 | 12.1 | 1.9 | 1.2 | 1.0 | 4.8 |
| 1978 | 12.5 | 7.9 | 6.0 | 4.8 | 8.5 |
| 1979 | 8.3 | 2.2 | 2.8 | 0.9 | 4.0 |
| 1980 | 2.3 | 2.2 | 1.0 | 1.8 | 2.0 |
| 1981 | 0.3 | 1.4 | 1.3 | 2.4 | 1.2 |
| 1982 | 5.5 | 10.0 | 13.0 | 6.2 | 8.4 |
| 1983 | 1.2 | 2.0 | 0.9 | 1.0 | 1.4 |
| 1984 | 6.1 | 4.7 | 2.8 | 1.5 | 4.2 |
| 1985 | 0.3 | 5.6 | 3.7 | 2.1 | 2.9 |
| 1986 | 1.6 | 9.9 | 0.5 | 2.2 | 4.1 |
| 1987 | 1.3 | 6.4 | 12.1 | 2.5 | 4.8 |
| 1988 | 7.3 | 0.4 | 0.7 | 0.4 | 2.7 |
| 1989 | 19.4 | 2.2 | 97.8 | 2.9 | 25.2 |
| 1990 | 3.8 | 0.6 | 3.1 | 0.9 | 2.1 |
| 1991 | 3.9 | 2.5 | 12.2 | 1.1 | 4.4 |

Table 2. Continued.

| Year | Head-of-Bay | Potomac <br> River | Choptank <br> River | Nanticoke <br> River | Bay-wide |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1992 | 1.3 | 22.1 | 4.3 | 4.3 | 9.0 |
| 1993 | 23.0 | 36.4 | 105.5 | 9.3 | 39.8 |
| 1994 | 23.4 | 3.9 | 19.3 | 21.5 | 16.1 |
| 1995 | 4.4 | 8.7 | 17.7 | 10.4 | 9.3 |
| 1996 | 25.0 | 48.5 | 154.4 | 43.7 | 59.4 |
| 1997 | 8.3 | 10.6 | 7.3 | 3.5 | 8.0 |
| 1998 | 8.3 | 10.8 | 32.6 | 3.8 | 12.7 |
| 1999 | 3.1 | 15.7 | 48.2 | 18.7 | 18.1 |
| 2000 | 13.3 | 7.8 | 21.2 | 17.6 | 13.8 |
| 2001 | 13.4 | 7.8 | 201.9 | 40.1 | 50.8 |
| 2002 | 3.1 | 7.0 | 0.7 | 7.8 | 4.7 |
| 2003 | 28.4 | 23.6 | 41.8 | 8.7 | 25.8 |
| 2004 | 7.8 | 4.0 | 22.8 | 19.5 | 11.4 |
| 2005 | 13.2 | 10.3 | 55.2 | 1.5 | 17.8 |
| 2006 | 1.5 | 6.7 | 5.8 | 3.2 | 4.3 |
| 2007 | 20.2 | 4.9 | 14.3 | 15.4 | 13.4 |
| 2008 | 5.9 | 3.3 | 0.5 | 1.0 | 3.2 |
| 2009 | 6.8 | 7.8 | 11.3 | 6.5 | 7.9 |
| 2010 | 7.3 | 5.7 | 3.3 | 4.6 | 5.6 |
| 2011 | 10.3 | 12.8 | 125.7 | 24.3 | 34.6 |
| 2012 | 0.7 | 1.7 | 0.1 | 0.6 | 0.9 |
|  |  |  |  |  |  |
| Average | 11.7 | 8.3 | 21.6 | 8.4 | 11.8 |
| TPA* | 17.3 | 9.2 | 10.8 | 8.6 | 12.0 |

* TPA (target period average) is the average from 1959 through 1972.

Table 3. Maryland juvenile striped bass survey geometric mean catch per haul at permanent sites.

| Year | Head-of-Bay | Potomac River | Choptank River | Nanticoke River | Bay-wide |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1955 | 1.49 | 3.78 | 2.36 | 2.26 | 2.26 |
| 1956 | 6.88 | 4.50 | 6.22 | 5.29 | 5.29 |
| 1957 | 1.92 | 1.78 | 1.16 | 1.40 | 1.40 |
| 1958 | 22.07 | 3.93 | 11.01 | 11.12 | 11.12 |
| 1959 | 0.95 | 0.61 | 0.09 | 0.59 | 0.59 |
| 1960 | 3.18 | 2.44 | 4.31 | 3.01 | 3.01 |
| 1961 | 7.46 | 12.82 | 5.40 | 6.61 | 6.61 |
| 1962 | 3.73 | 6.70 | 3.14 | 4.25 | 4.25 |
| 1963 | 3.01 | 0.54 | 2.01 | 1.61 | 1.61 |
| 1964 | 15.41 | 9.15 | 4.92 | 9.04 | 9.04 |
| 1965 | 0.76 | 0.92 | 2.18 | 1.56 | 1.56 |
| 1966 | 15.89 | 4.95 | 5.52 | 6.24 | 6.24 |
| 1967 | 3.92 | 1.03 | 2.80 | 2.28 | 2.28 |
| 1968 | 6.13 | 0.39 | 3.85 | 2.69 | 2.69 |
| 1969 | 12.21 | 0.12 | 2.55 | 2.81 | 2.81 |
| 1970 | 13.71 | 10.97 | 25.41 | 12.48 | 12.48 |
| 1971 | 10.45 | 3.48 | 2.51 | 4.02 | 4.02 |
| 1972 | 4.95 | 0.96 | 5.36 | 3.26 | 3.26 |
| 1973 | 11.92 | 1.10 | 0.43 | 2.33 | 2.33 |
| 1974 | 6.79 | 0.66 | 3.55 | 2.62 | 2.62 |
| 1975 | 2.34 | 3.56 | 2.71 | 2.81 | 2.81 |
| 1976 | 2.70 | 1.46 | 0.89 | 1.58 | 1.58 |
| 1977 | 4.99 | 0.78 | 0.81 | 1.61 | 1.61 |
| 1978 | 6.51 | 3.33 | 2.65 | 3.75 | 3.75 |
| 1979 | 4.56 | 1.15 | 1.12 | 1.73 | 1.73 |
| 1980 | 1.43 | 1.04 | 0.58 | 1.01 | 1.01 |
| 1981 | 0.17 | 0.68 | 0.84 | 0.59 | 0.59 |
| 1982 | 2.98 | 3.50 | 5.68 | 3.54 | 3.54 |
| 1983 | 0.61 | 0.62 | 0.64 | 0.61 | 0.61 |
| 1984 | 2.23 | 1.42 | 2.13 | 0.81 | 1.64 |
| 1985 | 0.19 | 1.45 | 1.78 | 0.94 | 0.91 |
| 1986 | 0.90 | 3.09 | 0.32 | 1.24 | 1.34 |
| 1987 | 0.16 | 3.01 | 3.06 | 1.36 | 1.46 |
| 1988 | 2.25 | 0.22 | 0.40 | 0.28 | 0.73 |
| 1989 | 8.54 | 1.15 | 28.10 | 1.94 | 4.87 |
| 1990 | 2.20 | 0.38 | 1.34 | 0.56 | 1.03 |
| 1991 | 1.99 | 0.84 | 4.42 | 0.52 | 1.52 |

Table 3. Continued.

| Year | Head-of-Bay | Potomac <br> River | Choptank <br> River | Nanticoke <br> River | Bay-wide |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1992 | 0.87 | 6.00 | 2.07 | 1.72 | 2.34 |
| 1993 | 15.00 | 15.96 | 27.87 | 4.56 | 13.97 |
| 1994 | 12.88 | 2.01 | 7.71 | 9.06 | 6.40 |
| 1995 | 2.85 | 4.47 | 9.96 | 3.76 | 4.41 |
| 1996 | 15.00 | 13.60 | 33.29 | 19.13 | 17.61 |
| 1997 | 6.15 | 3.67 | 3.95 | 1.74 | 3.91 |
| 1998 | 4.32 | 4.42 | 21.10 | 2.74 | 5.50 |
| 1999 | 1.91 | 5.84 | 20.01 | 5.52 | 5.34 |
| 2000 | 8.84 | 3.52 | 12.53 | 10.86 | 7.42 |
| 2001 | 7.15 | 5.01 | 86.71 | 20.31 | 12.57 |
| 2002 | 1.35 | 3.95 | 0.38 | 4.89 | 2.20 |
| 2003 | 11.89 | 12.81 | 20.56 | 3.25 | 10.83 |
| 2004 | 4.17 | 2.36 | 9.52 | 9.65 | 4.85 |
| 2005 | 8.48 | 7.92 | 16.81 | 1.07 | 6.91 |
| 2006 | 0.95 | 2.42 | 2.81 | 1.65 | 1.78 |
| 2007 | 8.21 | 2.20 | 7.87 | 5.41 | 5.12 |
| 2008 | 2.33 | 1.40 | 0.34 | 0.73 | 1.26 |
| 2009 | 2.85 | 3.75 | 6.61 | 4.18 | 3.92 |
| 2010 | 2.90 | 2.17 | 2.23 | 2.96 | 2.54 |
| 2011 | 5.79 | 7.18 | 26.14 | 12.99 | 9.57 |
| 2012 | 0.44 | 0.95 | 0.08 | 0.37 | 0.49 |
|  |  |  |  |  |  |
| Average | 5.55 | 3.62 | 8.12 | 3.76 | 4.22 |
| TPA* | 7.27 | 3.93 | 5.00 | 3.12 | 4.32 |

* TPA (target period average) is the average from 1959 through 1972.

Table 4. Maryland Chesapeake Bay arithmetic mean (AM) and log mean with coefficients of variation (CV), proportion of positive hauls (PPHL) with 95\% confidence intervals (CI), and number of seine hauls (n) for juvenile striped bass.

| Year | AM | CV (\%) <br> of AM | Log <br> Mean | CV (\%) of <br> Log Mean | PPHL | Low <br> CI | High <br> CI | n |
| :---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1957 | 2.9 | 205.5 | 0.87 | 100.72 | 0.66 | 0.52 | 0.80 | 44 |
| 1958 | 19.3 | 94.2 | 2.50 | 48.56 | 0.89 | 0.79 | 0.99 | 36 |
| 1959 | 1.4 | 198.3 | 0.47 | 171.23 | 0.30 | 0.14 | 0.45 | 34 |
| 1960 | 7.1 | 149.2 | 1.39 | 86.32 | 0.72 | 0.58 | 0.87 | 36 |
| 1961 | 17.0 | 183.3 | 2.03 | 61.04 | 0.96 | 0.90 | 1.02 | 46 |
| 1962 | 12.2 | 160.8 | 1.66 | 82.85 | 0.75 | 0.66 | 0.84 | 88 |
| 1963 | 4.0 | 182.6 | 0.96 | 111.85 | 0.56 | 0.45 | 0.66 | 88 |
| 1964 | 23.5 | 162.3 | 2.31 | 60.35 | 0.90 | 0.83 | 0.96 | 88 |
| 1965 | 7.4 | 247.7 | 0.94 | 140.06 | 0.47 | 0.36 | 0.57 | 88 |
| 1966 | 16.7 | 184.8 | 1.98 | 67.16 | 0.86 | 0.80 | 0.92 | 132 |
| 1967 | 7.8 | 263.9 | 1.19 | 100.40 | 0.69 | 0.61 | 0.77 | 132 |
| 1968 | 7.2 | 175.3 | 1.31 | 94.10 | 0.65 | 0.57 | 0.73 | 132 |
| 1969 | 10.5 | 224.0 | 1.34 | 104.40 | 0.62 | 0.54 | 0.70 | 132 |
| 1970 | 30.4 | 157.5 | 2.60 | 52.73 | 0.95 | 0.91 | 0.99 | 132 |
| 1971 | 11.8 | 187.0 | 1.61 | 80.43 | 0.81 | 0.74 | 0.88 | 132 |
| 1972 | 11.0 | 250.8 | 1.45 | 91.54 | 0.72 | 0.64 | 0.80 | 132 |
| 1973 | 8.9 | 229.2 | 1.20 | 110.90 | 0.61 | 0.53 | 0.70 | 132 |
| 1974 | 10.1 | 261.9 | 1.29 | 102.42 | 0.65 | 0.57 | 0.74 | 132 |
| 1975 | 6.7 | 152.2 | 1.34 | 86.76 | 0.73 | 0.66 | 0.81 | 132 |
| 1976 | 4.9 | 279.4 | 0.95 | 113.88 | 0.60 | 0.51 | 0.68 | 132 |
| 1977 | 4.8 | 236.4 | 1.96 | 113.00 | 0.62 | 0.54 | 0.70 | 132 |
| 1978 | 8.5 | 145.6 | 1.56 | 77.24 | 0.77 | 0.69 | 0.84 | 132 |
| 1979 | 4.0 | 182.1 | 1.00 | 100.24 | 0.66 | 0.58 | 0.74 | 132 |
| 1980 | 2.0 | 174.8 | 0.70 | 114.68 | 0.54 | 0.45 | 0.62 | 132 |
| 1981 | 1.2 | 228.2 | 0.46 | 150.34 | 0.39 | 0.30 | 0.47 | 132 |
| 1982 | 8.4 | 160.1 | 1.51 | 79.73 | 0.76 | 0.68 | 0.83 | 132 |
| 1983 | 1.4 | 268.0 | 0.48 | 152.37 | 0.38 | 0.30 | 0.46 | 132 |
| 1984 | 4.2 | 228.2 | 0.97 | 106.58 | 0.65 | 0.57 | 0.73 | 132 |
| 1985 | 2.9 | 253.0 | 0.65 | 152.02 | 0.42 | 0.33 | 0.50 | 132 |
| 1986 | 4.1 | 272.2 | 0.85 | 121.40 | 0.55 | 0.47 | 0.64 | 132 |
| 1987 | 4.8 | 262.1 | 0.90 | 124.54 | 0.51 | 0.42 | 0.59 | 132 |
| 1988 | 2.7 | 313.8 | 0.55 | 170.46 | 0.37 | 0.29 | 0.45 | 132 |
| 1989 | 25.2 | 309.1 | 1.77 | 90.18 | 0.75 | 0.68 | 0.82 | 132 |
| 1990 | 2.1 | 174.8 | 0.71 | 120.74 | 0.49 | 0.41 | 0.58 | 132 |
| 1991 | 4.4 | 203.8 | 0.93 | 120.27 | 0.52 | 0.43 | 0.60 | 132 |
|  |  |  |  |  |  |  |  |  |

Table 4. Continued.

| Year | AM | CV (\%) <br> of AM | Log <br> Mean | CV (\%) of <br> Log Mean | PPHL | Low <br> CI | High <br> CI | n |
| :---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1992 | 9.0 | 267.0 | 1.20 | 105.19 | 0.67 | 0.59 | 0.75 | 132 |
| 1993 | 39.8 | 279.1 | 2.71 | 49.53 | 0.96 | 0.93 | 0.99 | 132 |
| 1994 | 16.1 | 150.4 | 2.00 | 66.96 | 0.84 | 0.78 | 0.90 | 132 |
| 1995 | 9.3 | 153.3 | 1.69 | 66.42 | 0.86 | 0.80 | 0.92 | 132 |
| 1996 | 59.4 | 369.2 | 2.92 | 45.50 | 0.99 | 0.96 | 1.00 | 132 |
| 1997 | 8.0 | 135.6 | 1.59 | 70.98 | 0.80 | 0.74 | 0.87 | 132 |
| 1998 | 12.7 | 164.8 | 1.87 | 65.72 | 0.86 | 0.78 | 0.92 | 132 |
| 1999 | 18.1 | 208.4 | 1.85 | 77.45 | 0.80 | 0.75 | 0.88 | 132 |
| 2000 | 13.8 | 120.8 | 2.13 | 53.69 | 0.91 | 0.86 | 0.96 | 132 |
| 2001 | 50.8 | 308.9 | 2.61 | 57.22 | 0.92 | 0.88 | 0.97 | 132 |
| 2002 | 4.7 | 141.3 | 1.16 | 91.89 | 0.67 | 0.59 | 0.75 | 132 |
| 2003 | 25.8 | 136.9 | 2.47 | 55.42 | 0.92 | 0.88 | 0.97 | 132 |
| 2004 | 11.4 | 177.8 | 1.77 | 67.01 | 0.87 | 0.81 | 0.93 | 132 |
| 2005 | 17.8 | 237.3 | 2.07 | 59.12 | 0.90 | 0.86 | 0.95 | 132 |
| 2006 | 4.3 | 178.6 | 1.02 | 103.67 | 0.59 | 0.51 | 0.67 | 132 |
| 2007 | 13.4 | 177.3 | 1.81 | 71.92 | 0.83 | 0.76 | 0.89 | 132 |
| 2008 | 3.2 | 213.1 | 0.81 | 119.32 | 0.54 | 0.45 | 0.62 | 132 |
| 2009 | 7.9 | 154.3 | 1.59 | 66.66 | 0.86 | 0.80 | 0.92 | 132 |
| 2010 | 5.6 | 175.0 | 1.26 | 82.49 | 0.77 | 0.69 | 0.84 | 132 |
| 2011 | 34.6 | 580.4 | 2.36 | 51.94 | 0.93 | 0.89 | 0.97 | 132 |
| 2012 | 0.9 | 197.5 | 0.40 | 152.53 | 0.35 | 0.27 | 0.43 | 132 |
|  |  |  |  |  |  |  |  |  |
| Average | 12.1 | 212.6 | 1.46 | 92.48 | 0.71 | 0.63 | 0.78 |  |
| TPA* | 12.0 | 194.8 | 1.52 | 93.18 | 0.71 | 0.62 | 0.80 |  |

* TPA (target period average) is the average from 1959 through 1972.

Table 5. Maryland juvenile striped bass survey arithmetic (AM) and geometric (GM) mean catch per haul and number of seine hauls per year (n) for auxiliary sample sites.

|  | Patuxent River |  |  | Head of Bay |  |  |
| :---: | ---: | ---: | ---: | ---: | ---: | :---: |
| Year | AM | GM | n | AM | GM | n |
| 1983 | 0.06 | 0.04 | 18 | 0.58 | 0.33 | 12 |
| 1984 | 0.61 | 0.39 | 18 | 0.92 | 0.43 | 12 |
| 1985 | 3.17 | 1.95 | 18 | 1.00 | 0.24 | 12 |
| 1986 | 2.44 | 1.17 | 18 | 0.92 | 0.54 | 12 |
| 1987 | 2.94 | 0.94 | 17 | 0.33 | 0.26 | 9 |
| 1988 | 0.59 | 0.40 | 17 | 1.62 | 1.07 | 21 |
| 1989 | 1.39 | 0.92 | 18 | 10.43 | 1.91 | 21 |
| 1990 | 0.28 | 0.17 | 18 | 4.95 | 2.24 | 21 |
| 1991 | 0.94 | 0.53 | 18 | 2.15 | 0.98 | 20 |
| 1992 | 9.50 | 1.85 | 18 | 0.50 | 0.26 | 20 |
| 1993 | 104.30 | 47.18 | 18 | 28.00 | 11.11 | 21 |
| 1994 | 4.10 | 2.82 | 18 | 6.30 | 2.31 | 21 |
| 1995 | 7.28 | 3.46 | 18 | 2.95 | 1.15 | 21 |
| 1996 | 420.39 | 58.11 | 18 | 12.40 | 4.69 | 20 |
| 1997 | 7.33 | 2.72 | 18 | 2.70 | 2.18 | 20 |
| 1998 | 13.22 | 7.58 | 18 | 2.94 | 1.51 | 16 |
| 1999 | 7.28 | 5.39 | 18 | 3.62 | 2.13 | 13 |
| 2000 | 9.67 | 5.03 | 18 | 8.60 | 5.68 | 15 |
| 2001 | 17.28 | 10.01 | 18 | 19.47 | 6.62 | 15 |
| 2002 | 1.22 | 0.69 | 18 | 1.00 | 0.42 | 15 |
| 2003 | 61.11 | 22.17 | 18 | 16.06 | 11.79 | 16 |
| 2004 | 2.11 | 1.29 | 18 | 7.73 | 4.40 | 15 |
| 2005 | 8.94 | 3.91 | 18 | 5.53 | 4.35 | 15 |
| 2006 | 1.00 | 0.66 | 18 | 0.67 | 0.31 | 15 |
| 2007 | 15.22 | 6.07 | 18 | 5.33 | 2.72 | 15 |
| 2008 | 0.33 | 0.24 | 18 | 3.47 | 2.02 | 15 |
| 2009 | 3.00 | 1.87 | 18 | 2.13 | 1.14 | 15 |
|  |  |  |  |  |  |  |

Table 5. Continued.

|  | Patuxent River |  |  | Head of Bay |  |  |
| :---: | ---: | ---: | ---: | ---: | ---: | :---: |
| Year | AM | GM | $\mathbf{n}$ | AM | GM | n |
| 2010 | 3.33 | 2.49 | 18 | 3.67 | 1.45 | 15 |
| 2011 | 42.5 | 13.41 | 18 | 12.29 | 5.75 | 21 |
| 2012 | 0.06 | 0.04 | 18 | 1.86 | 0.71 | 21 |
|  |  |  |  |  |  |  |
| Average | 25.05 | 6.78 |  | 5.67 | 2.69 |  |
| Median | 3.25 | 1.91 |  | 3.21 | 1.71 |  |

Table 6. Log mean catch per haul of age 0 and age 1 striped bass by year-class.

| Year-class | Age 0 | Age 1 |
| :---: | :---: | :---: |
| 1957 | 0.87 | 0.08 |
| 1958 | 2.50 | 0.45 |
| 1959 | 0.47 | 0.07 |
| 1960 | 1.39 | 0.14 |
| 1961 | 2.03 | 0.39 |
| 1962 | 1.66 | 0.19 |
| 1963 | 0.96 | 0.07 |
| 1964 | 2.31 | 0.29 |
| 1965 | 0.94 | 0.19 |
| 1966 | 1.98 | 0.14 |
| 1967 | 1.19 | 0.20 |
| 1968 | 1.31 | 0.19 |
| 1969 | 1.34 | 0.10 |
| 1970 | 2.60 | 0.74 |
| 1971 | 1.61 | 0.37 |
| 1972 | 1.45 | 0.35 |
| 1973 | 1.20 | 0.21 |
| 1974 | 1.29 | 0.20 |
| 1975 | 1.32 | 0.12 |
| 1976 | 0.95 | 0.05 |
| 1977 | 0.96 | 0.16 |
| 1978 | 1.56 | 0.26 |
| 1979 | 1.00 | 0.16 |
| 1980 | 0.70 | 0.02 |
| 1981 | 0.46 | 0.02 |
| 1982 | 1.51 | 0.28 |
| 1983 | 0.48 | 0.00 |
| 1984 | 0.97 | 0.14 |
| 1985 | 0.65 | 0.03 |
| 1986 | 0.85 | 0.05 |
| 1987 | 0.90 | 0.06 |
| 1988 | 0.55 | 0.14 |
| 1989 | 1.77 | 0.28 |
| 1990 | 0.71 | 0.17 |
| 1991 | 0.93 | 0.11 |
| 1992 | 1.20 | 0.18 |
| 1993 | 2.71 | 0.56 |

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Table 6. Continued.

| Year-class | Age 0 | Age 1 |
| :---: | :---: | :---: |
| 1994 | 2.00 | 0.12 |
| 1995 | 1.69 | 0.07 |
| 1996 | 2.92 | 0.23 |
| 1997 | 1.59 | 0.16 |
| 1998 | 1.87 | 0.31 |
| 1999 | 1.85 | 0.23 |
| 2000 | 2.13 | 0.28 |
| 2001 | 2.61 | 0.58 |
| 2002 | 1.16 | 0.07 |
| 2003 | 2.47 | 0.55 |
| 2004 | 1.77 | 0.25 |
| 2005 | 2.07 | 0.25 |
| 2006 | 1.02 | 0.07 |
| 2007 | 1.81 | 0.27 |
| 2008 | 0.81 | 0.11 |
| 2009 | 1.59 | 0.16 |
| 2010 | 1.26 | 0.02 |
| 2011 | 2.36 | 0.30 |
| 2012 | 0.40 | $\mathrm{~N} / \mathrm{A}$ |

Figure 1. Maryland Chesapeake Bay juvenile striped bass survey site locations.


Figure 2. Maryland Chesapeake Bay arithmetic mean (AM) catch per haul and $95 \%$ confidence intervals ( $\pm 2$ SE) for juvenile striped bass with target period average (TPA).


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Figure 3. Maryland Chesapeake Bay geometric mean (GM) catch per haul and $95 \%$ confidence intervals ( $\pm 2$ SE) for juvenile striped bass with target period average (TPA).


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Figure 4. Maryland Chesapeake Bay juvenile striped bass indices. Arithmetic mean (AM), scaled geometric mean (GM)*, and proportion of positive hauls (PPHL) as percent.


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Figure 5. Arithmetic mean (AM) catch per haul by system for juvenile striped bass. Note different scales.


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Figure 6. Head of Bay geometric mean (GM) catch per haul and $95 \%$ confidence intervals ( $\pm 2 \mathrm{SE}$ ) for juvenile striped bass with target period average (TPA).


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Figure 7. Potomac River geometric mean (GM) catch per haul and $95 \%$ confidence intervals ( $\pm 2$ SE) for juvenile striped bass with target period average (TPA).


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Figure 8. Choptank River geometric mean (GM) catch per haul and $95 \%$ confidence intervals ( $\pm 2 \mathrm{SE}$ ) for juvenile striped bass with target period average (TPA).


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Figure 9. Nanticoke River geometric mean (GM) catch per haul and $95 \%$ confidence intervals ( $\pm 2 \mathrm{SE}$ ) for juvenile striped bass with target period average (TPA).


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Figure 10. Relationship between age 0 and subsequent age 1 striped bass indices.


Figure 11. Residuals of age 1 and age 0 striped bass regression.



# PROJECT NO. 2 

JOB NO. 3
TASK NO. 4

## STRIPED BASS TAGGING

Prepared by Beth A. Versak

## INTRODUCTION

The primary objective of Project 2, Job 3, Task 4 was to summarize all striped bass tagging activities in Maryland's portion of the Chesapeake Bay and the North Carolina cooperative tagging cruise, during the time period of summer 2011 through spring 2012. The Maryland Department of Natural Resources (MD DNR) and partnering agencies tagged striped bass as part of the United States Fish and Wildlife Service's (USFWS) Cooperative Coastal Striped Bass Tagging Program. Fish were tagged from the Chesapeake Bay resident/pre-migratory and spawning stocks, and from the Atlantic coastal stock. S ubsequently, tag num bers a nd associated fish attribute data were forwarded to the USFWS, with the captor providing recovery information directly to the USFWS. These data are used to evaluate stock dynamics (mortality rates, survival rates, growth rates, etc.) of Atlantic coast striped bass stocks.

## METHODS

## Sampling procedures

During late March through May 20 12, a fishery-independent spawning stock study was conducted, in which tags were applied to fish captured with experimental multi-panel drift gill nets in the upper Chesapeake Bay and the Potomac River (see Project 2, Job 3, Task 2) (Figure 1). Fish sampled during this study were measured for total length (TL) to the nearest millimeter ( mm ) and
examined for sex, maturation stage and external anomalies. Internal anchor tags were applied to healthy fish, regardless of size, and scale s amples were col lected from as ub-sample for age determination. Scales were taken from two to three male fish per week per 10-mm length group, up to 700 mm TL. No more than 10 scale samples per 10-mm length group were taken over the course of the survey. Scale samples were taken from all males over 700 mm TL and all female fish. Tagging stopped when water temperatures exceeded $70^{\circ} \mathrm{F}$.

The 2012 cooperative tagging cruise was again conducted on a sportfishing vessel and fish were captured via hook and line. Sampling was conducted on only one day, February 16, 2012, by staff from the USFWS and the North Carolina Division of Marine Fisheries (NC DMF), with support from MD DNR. The goal of this year's sampling was to tag coastal migratory striped bass wintering in the Atlantic Ocean off northeastern North Carolina and/or southeastern Virginia (state and federal waters). Up to seven lines containing custom-made tandem parachute rigs were trolled from the 40 foot sportfishing vessel, Smokin Gun II, at 2.5 to 3.5 knots, in depths of 50 to 75 feet ( 15 to 23 m ). Vigorous fish with no external anomalies were measured for total length to the nearest millimeter ( mm TL) and tagged immediately after being landed in the boat. Scales were taken from the first five striped bass per 10-mm TL group from 400-800 mm TL, and from all striped bass less than 400 mm TL and greater than 800 mm TL.

## Taqging procedures

For all surveys, internal anchor tags, supplied by the USFWS, were inserted through an incision made in the left ventral side of healthy fish, slightly behind and below the tip of the pectoral fin. This small, shallow incision was made with a \#12 curved scalpel after removing a few scales from the tag area. The incision was angled anteriorly through the musculature, encouraging the
incision to fold together and the tag streamer to lie back along the fish's side. The tag anchor was then pushed through the remaining muscle tissue and peritoneum into the body cavity and checked for retention.

## Analytical Procedures

Survival rates from fish tagged during the spring in Maryland were estimated using two approaches, all based on historic release and recovery data. During the most recent ASMFC stock assessment, the instantaneous rates-catch and release (IRCR) model became the primary model utilized. The IRCR method employs an age-independent form of the IRCR model developed in Jiang et al. (2007) to estimate survival, fishing mortality and natural mortality. The candidate models run in the IRCR model are similar in structure to the models used in Program MARK. Additional details on the methodologies can be found in the latest stock assessment report (ASMFC 2011).

Previously, Program MARK w as us ed to e stimate s urvival us ing tag-recovery m odels (Brownie et al. 1985) and subsequent extensions of those models. Estimates of survival and recovery w ere cal culated by fitting a s et of candi date models, chosen "a priori" and based on knowledge of the biology of the species, to the observed release and recovery data (Brownie et al. 1985; Burnham et al. 1995). Further details on Program MARK methodologies can be found in Versak (2007). Survival was converted to total mortality, and a constant value of natural mortality $(M=0.15)$ was subtracted to obtain an estimate of fishing mortality. It is be lieved that natural mortality in Chesapeake Bay is increasing (ASMFC 2011). Thus, the use of a constant value for M became a weakness of the MARK method.

For all methods, the recovery year began on the first day of tagging in the time series (March
28) and continued until March 27 of the following year. Since survival and F estimates for fish released in spring 2012 will not be completed until after March 27, 2013, these estimates will not appear in this report.

Tag release and return data from spring male fish, $\geq 457 \mathrm{~mm}$ TL and $<711 \mathrm{~mm} \mathrm{TL}$ ( $18-28$ inches TL), were used to develop the 2011-2012 estimate of F for Chesapeake Bay (unpublished data). Male fish 18 to 28 inches are generally accepted to compose the Chesapeake Bay resident stock, while 1 arger fish are pr edominantly coa stal migrants. R elease and recapture da ta from Maryland and Virginia (provided by Virginia Institute of Marine Science) were combined to produce a B aywide es timate of F. Similar to the coastwide me thods, the IRCR model was utilized to calculate the Chesapeake Bay F. Further details on the methodologies can be found in the latest stock assessment report (ASMFC 2011).

Estimates of survival, fishing mortality and recovery rates for the North Carolina cooperative tagging cruise data were calculated using the same methods as Maryland's spring tagging data. If the 2012 cruise data are used in the upcoming assessment, the calculations will be conducted by the USFWS.

For each study, t-tests were used to test for significant differences between the mean lengths of striped bass that were tagged and all striped bass measured for total length (SAS 1990). This was done to determine if the tagged fish were representative of $t$ he ent ire $s$ ample. Lengths were considered different at $\mathrm{P}<0.05$.

## RESULTS AND DISCUSSION

## Spring taqging

The spring sampling component monitored the size and sex characteristics of striped bass spawning in the Potomac River and the upper Chesapeake Bay. Sampling occurred between March 26, 2012 and May 18, 2012. A total of 983 striped bass were sampled and 682 (69\%) were tagged as part of this long-term survey (Table 1). In 2012, fewer striped bass were captured in the survey than normal, which resulted in a higher proportion of fish being tagged than in previous years. However, there were still occasions when large samples were caught in a short period of time, which required fish to spend a considerable amount of time submerged in the gill net or on the boat, thereby increasing the potential for mortality. In these cases, biologists measured all fish but were only able to tag a sub-sample. Typically, these large concentrations of fish were of a smaller size and captured in small mesh panels. Larger fish were encountered less frequently, and therefore a higher proportion was tagged. This resulted in a significantly greater mean length of tagged fish than the mean length of all fish sampled. Mean total length of striped bass tagged during spring 2012 ( 660 mm TL ) was significantly greater $(\mathrm{P}<0.05)$ than that of the sampled population (630 mm TL) (Figure 2).

Tag releases and recaptures from both Maryland and Virginia's sampling (combined spring 2011 data) were used to estimate an instantaneous fishing mortality rate (F) for the 2011-2012 recreational, charter boat, and commercial fisheries for the entire Chesapeake Bay. Fishing mortality estimates from the two analysis methods were below the target $\mathrm{F}=0.27$ set by ASMFC (unpublished data).

Estimates of survival and fishing mortality for the 2012 Chesapeake Bay spawning stock, as well as the resident stock, will be presented in the next report of the ASMFC Striped Bass Tagging

Subcommittee. Stock assessments are currently being conducted every two years.

## North Carolina cooperative tagging cruise

Although a different gear was used, the primary objective of the cooperative tagging cruise remained to apply tags to as many striped bass as possible. In 2012, only six striped bass were captured and all were tagged during the cruise (Table 2). Because the sample size was so low, scales were taken from all striped bass captured, regardless of total length.

The mean length of all fish captured and tagged on the 2012 cruise was 905 mm TL. This length was significantly larger than the mean total length for the 2011 cruise ( 810 mm TL total sampled and tagged; $\mathrm{P}<0.0001$ ). Although the sample size was small, it is not uncommon for the mean lengths to vary from year to year. Funding has been secured to conduct the 2013 c ruise onboard a research trawler, as well as a sportfishing charter vessel, to ensure that gear comparison studies are done.

Estimates of survival and fishing mortality based on fish tagged in the 2012 North Carolina study will likely not be calculated due to small sample sizes.

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Figure 2. Length frequencies of striped bass measured and tagged during the spring in Chesapeake Bay.

Table 1. Summary of USFWS internal anchor tags applied to striped bass in Maryland's portion of Chesapeake Bay and Potomac River, late March - May 2012.

| System | Inclusive <br> Release Dates | Total Fish <br> Sampled | Total Fish <br> Tagged | Approximate Tag <br> Sequences ${ }^{\text {a }}$ |
| :---: | :---: | :---: | :---: | :---: |
| Potomac River | $3 / 26 / 12-5 / 7 / 12$ | 354 | 229 | $521806-522000$ <br> $524211-524247$ |
| Upper Chesapeake Bay | $3 / 30 / 12-5 / 18 / 12$ | 629 | 453 | $518001-518456$ |
| Spring spawning survey totals: |  |  |  |  |

${ }^{\text {a }}$ Not all tags in reported sequences were applied; some tags were lost, destroyed, or applied out of order.
${ }^{\mathrm{b}}$ Total sampled includes one USFWS recapture.
${ }^{\mathrm{c}}$ Total sampled includes two fish with no total length recorded.

Table 2. Summary of U SFWS internal a nchor t ags a pplied to s triped ba ss dur ing the 2012 SEAMAP cooperative tagging cruise.

| System | Inclusive <br> Release Dates | Total Fish <br> Sampled | Total Fish <br> Tagged | Approximate Tag <br> Sequences |
| :---: | :---: | :---: | :---: | :---: |
| Nearshore Atlantic Ocean <br> (Near VA-NC line) | $2 / 16 / 12$ | 6 | 6 | $561083-561088$ |
| Cooperative tagging cruise totals: |  |  |  | 6 |
| 6 |  |  |  |  |

Figure 1. Tagging locations in spawning areas of the Upper Chesapeake Bay and the Potomac River, late March - May 2012.


Figure 2. Length frequencies of striped bass measured and tagged during the spring in Chesapeake Bay.


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# PROJECT NO. 2 

JOB NO. 3
TASK NO. 5A

## COMMERCIAL FISHERY HARVEST MONITORING

Prepared by Amy Batdorf

## INTRODUCTION

The pr imary objectives of Project 2, Job 3, Task 5A were to quantify the commercial striped ba ss harvest in 2011 and describe the harvest monitoring conducted bythe Maryland Department of N atural Resources (MD D NR). M D D NR changed $t$ he or ganization of $i$ ts commercial quot a s ystem from a seasonal to a calendar year s ystem in 1999. M aryland completed its twenty-second year of commercial fishing under the quota system since the striped bass fishing $m$ oratorium was lifted in 1990. The com mercial fishery received $42.5 \%$ of $t$ he state's total Chesapeake Bay striped bass quota. The 2011 commercial quota for the Chesapeake Bay and its tributaries was $1,963,873$ pounds, a $7 \%$ decrease from 2010, with an 18 to 36 inch total length (TL) slot limit. There was a separate quota of 126,396 pounds, with a 24 -inch (TL) minimum size for the state's jurisdictional waters off the Atlantic coast.

The Chesapeake Bay commercial quota was further divided by gear type (Table 1). The hook-and-line a nd drift gill ne $t$ fisheries were combined and allotted $75 \%$ of the commercial quota. $T$ he pound ne $t$ and ha ul seine $f$ isheries were allotted $t$ he $r$ emaining $25 \%$. When $t$ he allotted quota for a fishery (gear type) was not landed, it was transferred to another commercial fishery.

Each fishery was managed with specific seasons that could be modified by MD DNR as necessary. T he hook -and-line f ishery w as op en f rom June 7 t hrough N ovember 30, 2011 , Monday through $T$ hursday onl $y$. The pound ne $t \mathrm{f}$ ishery was ope n from June 1 t hrough November 30, 2011, Monday through Saturday. The haul seine fishery was open from June 7 through November 30, 2011, Monday through Friday. The Chesapeake Bay drift gill net season was split, with the first s egment from J anuary 1 through F ebruary 28,2011 and t he second
segment from December 1 through December 31, 2011, Monday through Friday. The Atlantic coast fishery consisted of two gear types, drift gill net and trawl. Both gear types were permitted during the Atlantic season, which occurred in two segments: January 1 through April 30, 2011 and November 1 through December 31, 2011, Monday through Friday.

Commercial harvest data for striped bass can be used as a general measure of stock size (Schaefer 1972, Goodyear 1985). Catch per unit effort (CPUE) data have traditionally been used more widely out side of the Chesapeake B ay as an indicator of stock abundance (Ricker 1975, Cowx 1991). Catch and effort data provide useful information regarding the various components of a fishery and group patterns of use for the fisheries resource. Catch data col lected from the check station reports and effort data from the monthly fishing reports (MFR) for striped bass fishermen were analyzed with the primary objective of presenting a post-moratoria summary of baseline data on commercial catch and CPUE.

## METHODS

In July 2008, commercial finfish license hol ders w ere not ified by MD DNR that participation in the striped bass fishery required a declaration of intent to fish using a specified legal ge ar. A deadline of A ugust 31 was es tablished for receipt of declaration; this process is repeated for every year in which the license holder intends to fish. MD DNR charged a fee to participants ba sed upon the $t$ ype of 1 icense he ld. P articipants w ho he ld an Unlimited Tidal Fishing License (TFL) were required to pay $\$ 300$. Participants who held an Unlimited Finfish Harvester License (FIN) w ere to pay $\$ 100$ a nd the Hook-and-Line only License (HLI) were required to pay $\$ 37.50$ Daily allocations were established to distribute harvest over as many days as was practical, in an effort to avoid flooding the market (Table 1). Individual allocations were printed on each striped bass permit issued by MD DNR.

All com mercially ha rvested striped bass were required to be tagged by the fishermen prior to landing with colored, serial numbered, tamper evident tags inserted in the mouth of the fish and out through the ope rculum. These tags could ve rify $t$ he ha rvester and easily i dentify
legally harvested fish to the public and law enforcement. Each harvest day and prior to sale, all tagged striped $b$ ass w ere required to pa ss through a MD DNR a pproved commercial fishery check station. Fish dealers distributed throughout the state volunteered to act as check stations (Figure 1). C heck station employees, acting as representatives of MD DNR, were responsible for counting, weighing and verifying that all fish were tagged. C heck s tations a lso recorded harvest data on the individual fisherman's striped bass permit. Each morning following a harvest day, the check station was required to telephone MD DNR and report the total pounds of striped bass che cked the previous day (Figures 2, 3). These reports allowed MD DNR to monitor the fisheries' daily reported progress towards their respective quotas. Check stations were required to keep daily written logs detailing the activity of each fisherman, which were returned weekly by mail to MD DNR. Individual fishermen were then required to return their striped bass permit to MD DNR at the end of the season.

In addition, individual fishermen were required to report their striped bass harvest on a monthly fishing report (MFR). MFRs were required to be returned by the $10^{\text {th }}$ of the following month on a monthly basis, regardless of fishing activity. Fishermen who did not return a MFR were considered late. The na mes of those individuals with late reports appeared on the "Late Reports" list on the commercial fisheries website. If the report is still not received by DNR 50 days after the report due date, the licensee received an official violation. Two or more of ficial violations for any of the report types in a 12 m onth period may result in a license suspension. The following information w as c ompiled from each c ommercial fisherman's M FR: D ay of Month, NOAA F ishing A rea, Gear C ode, Quantity of G ear, Duration, N umber of S ets, T rip Length (hours), Number of Crew, and Pounds (by species). CPUE estimates for each gear type were derived by dividing total pounds landed by each gear by the number of reported trips from the MFRs.

The pounds of $s$ triped ba ss presented in $t$ his $r$ eport were $s$ upplied $b y t$ he Data Management and Quota Monitoring Program of the MD DNR Fisheries Service. Prior to 2001, the pounds landed were determined using the MFRs. Due to delays in submission of the MFRs
and the time necessary to enter the data, there would often appear to be discrepancies between the MFRs, check station log sheets, a nd daily check station telephone reports. Since 2001, in order to avoid these issues and have more timely data, the pounds landed have come from the daily check station telephone reports and the weekly check station log sheets. However, all three data $s$ ources are $g$ enerally cor roborative a nd $t$ he change in data s ource $r$ eported here $w$ as considered to have no appreciable effect on the results and conclusions.

## RESULTS AND DISCUSSION

On $t$ he C hesapeake Bay and i ts t ributaries, $1,955,072$ pounds of s triped ba ss w ere harvested in 2011, 8,801 pounds under the 2011 quota. The estimated number of fish landed was 520,772 (Table 2). The Chesapeake drift gill net fishery landed $44 \%$ of the total landings by weight, followed by the pound net fishery at $33 \%$. The hook-and-line fishery contributed $23 \%$ of the total landings and less than $1 \%$ of fish were harvested by the haul seine fishery.

Maryland's Atlantic coast landings were estimated at 2,072 striped bass, weighing 21,401 pounds (Table 2). The drift gill net fishery made up $87 \%$ of the Atlantic harvest, by weight, with the remainder from the trawl fishery.

## Comparisons of Average Weight

The average weight of fish harvested was calculated using two methods. The first was by dividing the total weight of landings by the number of fish reported in the weekly check station $\log$ sheets. The second method involved direct sampling of striped bass at check stations by MD DNR biologists to characterize the harvest of commercial fisheries by measuring and weighing a sub-sample of fish (Project 2, Job 3, Tasks 1A, 1B, and 1C, in this report).

The mean weight per fish of striped bass harvested in Chesapeake Bay, regardless of gear type, was 4.00 pounds when calculated from the check station log sheets and 4.17 pounds when measured by biologists (Table 3), an increase from the 2010 s eason. Mean weights by specific gear type ranged from 3.54 to 3.97 pounds from check station log sheets, and were 3.56 to 4.44 pounds when measured by biologists. The largest striped ba ss landed in the Chesapeake B ay were taken by the drift gill net fishery. The average weight of fish harvested by gill net was 3.97 pounds when calculated using the $\log$ sheet data and 4.44 pounds when calculated using the MD DNR measurements.

Striped bass were al so sampled at A tlantic coast check stations to characterize co astal harvest, although sample size was small (Project 2, J ob 3, Task 1C, this report). Striped bass sampled from the Atlantic coast fisheries by MD DNR biologists averaged 14.95 pounds (Table 3). The av erage weight cal culated from the check station $\log$ sheets $w$ as 10.33 pounds. Fish caught in the Atlantic trawl fishery averaged 16.87 pounds according to MD DNR estimates, and were larger on a verage than those caught in the gill net fishery ( 14.60 pounds). The average weights of fish from the Atlantic trawl and gill net fisheries, as calculated from check station log sheets, were 14.54 and 9.90 pounds, respectively.

## Commercial Harvest Trends

Since the moratorium was lifted in 1990, striped bass harvests and quotas have become relatively consistent in the Chesapeake Bay (Table 4, Figure 4). The majority of the commercial striped bass ha rvest in Chesapeake B ay has hi storically be en by drift gill net. Since the late

1990s, however, an increasing portion of the harvest has come from the pound net and hook-andline $f$ isheries. The hook -and-line fishery $g$ enerally harvests $t$ he $l$ east of $t$ he $t$ hree major Chesapeake Bay gears. The pound net fishery harvest increased through the early 1990s and by 1998 averaged approximately 600,000 pounds of striped bass harvested per year between 19982011.

Similar to the C hesapeake B ay fisheries, the Atlantic ha rvest has increased since the moratorium was lifted in 1990 and the fishery harvests nearly $100 \%$ of its quota; with a decline in harvest for the 2009-2011 seasons (Figure 5). In almost all years since 1990, the Atlantic trawl fishery harvest has been greater that the Atlantic drift gill net harvest with the exception of 2010 and 2011 where the gill net harvest was larger than the trawl harvest (Table 5, Figure 5). Though the Atlantic drift gill net fishery harvested very little initially after the moratorium was lifted, the harvest be gan to increase in 1994, 1 ikely due to increased interest in the fishery and increased abundance of the stock.

## Commercial CPUE Trends

Weight harvested by year and gear type was taken from check station $\log$ sheets. The number of fishing trips in which striped bass were landed was determined from the MFRs (Table 2). The pounds landed were divided by the number of trips to calculate an estimate of CPUE.

The pound net fishery CPUE w as 390 pounds pertrip, the same as last season. The Chesapeake B ay drift gill net fishery CPUE was 397 pounds pertrip, an $11 \%$ decrease from 2010 CPUE. The hook-and-line fishery CPUE was 224 pounds per trip, a $16 \%$ increase from the previous y ear (Table 5 , F igure 6). With $t$ he e xception of 2004, $t$ he hook -and-line $f$ ishery continues to have the lowest CPUE of all the Chesapeake Bay fisheries. Over the past five years, the gill net fishery had the highest average CPUE value ( 365 lbs per trip), followed closely by the pound net fishery ( 351 lbs per trip) and the hook-and-line fishery (206 lbs per trip) (Table 6, Figure 6).

The Atlantic trawl fishery CPUE was 187 pounds per trip in 2011, a $63 \%$ drop from the 2010 CPUE and significantly be low the twenty-two year average of 546 pounds per trip. The 2011 CPUE for the Atlantic drift gill net fishery was 155 pounds per trip, below the twenty-two year average of 196 pounds per trip (Table 6, Figure 7).

In general, all C hesapeake Bay com mercial striped bass f isheries have ex hibited positive trends in CPUE estimates since the lifting of the mor atorium in 1990 (Figure 6). The Atlantic drift gill net fishery has been variable with a downward trend since 2009. The Atlantic trawl fishery has also been variable, with several spikes in harvest in 1995 and from 2006-2009.

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Figure 4. Maryland's Chesapeake Bay striped bass total harvest (thousands of pounds) per calendar year by commercial gear type, 1990 to 2011.

Figure 5. Maryland's Atlantic gill net and trawl fisheries total striped bass harvest (thousands of pounds) per calendar year by commercial gear type, 1990-2011.

Figure 6. Maryland's Chesapeake Bay striped bass catch (pounds) per trip (CPUE) by commercial gear type, 1990-2011. Trips were determined as days fished when striped bass catch was reported.

Figure 7. Maryland's Atlantic gill net and trawl fisheries striped bass catch (pounds) per trip (CPUE), 1990-2011. Trips were determined as days fished when striped bass catch was reported.

Table 1. Striped bass commercial regulations by gear type for the 2011 calendar year.

| Area | Gear Type | Annual Quota (pounds) | Number of Participants | Trip Limit | $\begin{aligned} & \text { Minimum } \\ & \text { Size } \end{aligned}$ | Reporting Requirement |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Bay and Tributaries | Pound Net | 490,968 | 222 | single permit holders: $800 \mathrm{lbs} /$ day; multiple permit holders $1,600 \mathrm{lbs} /$ day | $\begin{gathered} \text { 18-36 in TL } \\ \text { slot } \end{gathered}$ | Monthly Harvest Report |
|  | Haul Seine | included in Pound Net | 3 | $750 \mathrm{lbs} /$ license/day; 1,250 lbs/license/net/season | $\begin{gathered} \text { 18-36 in TL } \\ \text { slot } \end{gathered}$ | Monthly Harvest Report |
|  | Hook-andLine | 589,162 | 149 | $500 \mathrm{lbs} /$ license/day; 1,500 lbs/license/week; max 4 people/boat; 2 crew/licensee | $\begin{gathered} \text { 18-36 in TL } \\ \text { slot } \end{gathered}$ | Monthly Harvest Report |
|  | Gill Net | 883,743 | 761 | $300 \mathrm{lbs} /$ licensee/day; max 4 licenses/boat | $\begin{gathered} \text { 18-36 in TL } \\ \text { slot } \end{gathered}$ | Monthly Harvest Report |
| Total Bay Quota |  | 1,963,873 |  |  |  |  |
| Atlantic Coast | Atlantic Trawl | 126,396 | 40 | $1,950 \mathrm{lbs} /$ license/season for both Atlantic gears | 24 in TL min | Monthly Harvest Report |
|  | Atlantic Gill Net | included in Trawl | 46 |  |  |  |
| Total Maryland Quota |  | 2,090,269 |  |  |  |  |

Table 2. Summary of striped bass commercial harvest statistics by gear type for the 2011 calendar year.

| Area | Gear Type | Pounds ${ }^{1}$ | Estimated ${ }^{1}$ <br> Number <br> of Fish | Trips ${ }^{2}$ |
| :---: | :---: | :---: | :---: | :---: |
| Chesapeake Bay ${ }^{3}$ | Haul Seine | 1,135 | 326 | 3 |
|  | Pound Net | 646,978 | 177,592 | 1,661 |
|  | Hook-and-Line | 441,422 | 124,841 | 1,972 |
|  | Gill Net | 865,537 | 218,013 | 2,180 |
|  | Chesapeake Total Harvest | 1,955,072 | 520,772 | 5,816 |
| Atlantic Coast | Atlantic Trawl | 2,806 | 193 | 15 |
|  | Atlantic Gill Net | 18,595 | 1,879 | 120 |
|  | Atlantic Total Harvest | 21,401 | 2,072 | 135 |
| Maryland Totals |  | 1,976,473 | 522,844 | 5,951 |

1. Data from check station log sheets.
2. Trips were determined as days fished when striped bass catch was reported on MFRs.
3. Includes all Maryland Chesapeake Bay and tributaries, except main stem Potomac River.

Table 3. Striped bass average weight (lbs) by gear type for the 2011 calendar year. Average weights calculated by MD DNR biologists include $95 \%$ confidence intervals.

| Area | Gear Type | Average Weight <br> from Check <br> Station Logs <br> (pounds) | Average Weight from <br> Biological Sampling <br> (pounds) $^{2}$ | Sample <br> Size from <br> Biological <br> Sampling |
| :---: | :---: | :---: | :---: | :---: |
|  | Haul Seine | N/A | N/A | N/A |
|  | Pound Net | 3.64 | $4.03(3.90-4.17)$ | 1,104 |
|  | Hook-and-Line | 3.54 | $3.56(3.48-3.65)$ | 1,328 |
|  | Gill Net | 3.97 | $4.44(4.39-4.49)$ | 3,441 |
|  | Chesapeake <br> Total Harvest | $\mathbf{4 . 0 0}$ | $\mathbf{4 . 1 7}(\mathbf{4 . 1 2 - 4 . 2 1 )}$ | 5,873 |
| Atlantic Coast | Trawl | 14.54 | $16.87(13.56-20.18)$ | 3 |
|  | Gill Net | 9.90 | $14.60(13.72-15.49)$ | 175 |
|  | Atlantic Total <br> Harvest | $\mathbf{1 0 . 3 3}$ | $\mathbf{1 4 . 9 5}(\mathbf{1 4 . 0 5 - 1 5 . 8 5 )}$ | $\mathbf{2 0 7}$ |

1. Data from check station log sheets, pounds divided by the number of fish reported.
2. Data from check station sampling by MD DNR biologists, all months combined.
3. Includes all Maryland Chesapeake Bay and tributaries, except main stem Potomac River.

Table 4. Pounds of striped bass harvested by commercial gear type, 1990 to 2011.

| Year | Hook-and- <br> Line | Pound Net | Drift Gill Net | Atlantic Gill <br> Net | Atlantic <br> Trawl |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{1 9 9 0}$ | 700 | 1,533 | 130,947 | 83 | 4,843 |
| $\mathbf{1 9 9 1}$ | 2,307 | 37,062 | 331,911 | 1,426 | 14,202 |
| $\mathbf{1 9 9 2}$ | 7,919 | 157,627 | 609,197 | 422 | 17,348 |
| $\mathbf{1 9 9 3}$ | 8,188 | 181,215 | 647,063 | 127 | 3,938 |
| $\mathbf{1 9 9 4}$ | 51,948 | 227,502 | 831,823 | 3,085 | 15,066 |
| $\mathbf{1 9 9 5}$ | 29,135 | 290,284 | 869,585 | 10,464 | 71,587 |
| $\mathbf{1 9 9 6}$ | 54,038 | 336,887 | $1,186,447$ | 23,894 | 38,688 |
| $\mathbf{1 9 9 7}$ | 367,287 | 467,217 | $1,216,686$ | 28,764 | 55,792 |
| $\mathbf{1 9 9 8}$ | 536,809 | 613,122 | 721,987 | 36,404 | 51,824 |
| $\mathbf{1 9 9 9}$ | 790,262 | 667,842 | $1,087,123$ | 24,590 | 51,955 |
| $\mathbf{2 0 0 0}$ | 747,256 | 462,086 | $1,001,304$ | 40,806 | 66,968 |
| $\mathbf{2 0 0 1}$ | 398,695 | 647,990 | 586,892 | 20,660 | 71,156 |
| $\mathbf{2 0 0 2}$ | 359,344 | 470,828 | 901,407 | 21,086 | 68,300 |
| $\mathbf{2 0 0 3}$ | 372,551 | 602,748 | 744,790 | 24,256 | 73,893 |
| $\mathbf{2 0 0 4}$ | 355,629 | 507,140 | 921,317 | 27,697 | 87,756 |
| $\mathbf{2 0 0 5}$ | 283,803 | 513,519 | $1,211,365$ | 12,897 | 33,974 |
| $\mathbf{2 0 0 6}$ | 514,019 | 672,614 | 929,540 | 45,710 | 45,383 |
| $\mathbf{2 0 0 7}$ | 643,598 | 528,683 | $1,068,304$ | 38,619 | 74,172 |
| $\mathbf{2 0 0 8}$ | 432,139 | 559,087 | $1,216,581$ | 37,117 | 80,888 |
| $\mathbf{2 0 0 9}$ | 650,207 | 566,898 | $1,050,188$ | 32,937 | 94,390 |
| $\mathbf{2 0 1 0}$ | 519,117 | 650,628 | 934,742 | 28,467 | 16,335 |
| $\mathbf{2 0 1 1}$ | 441,422 | 646,978 | 865,537 | 18,595 | 2,806 |

Table 5. Striped bass average catch per trip (CPUE) in pounds by commercial gear type, 1990 to 2011.

| Year | Hook-and- <br> Line | Pound Net | Drift Gill Net | Atlantic Gill <br> Net | Atlantic <br> Trawl |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{1 9 9 0}$ | 25 | 81 | 76 | 21 | 161 |
| $\mathbf{1 9 9 1}$ | 77 | 96 | 84 | 65 | 254 |
| $\mathbf{1 9 9 2}$ | 70 | 130 | 114 | 84 | 271 |
| $\mathbf{1 9 9 3}$ | 52 | 207 | 125 | 25 | 188 |
| $\mathbf{1 9 9 4}$ | 108 | 248 | 139 | 129 | 284 |
| $\mathbf{1 9 9 5}$ | 71 | 220 | 156 | 75 | 994 |
| $\mathbf{1 9 9 6}$ | 85 | 210 | 188 | 151 | 407 |
| $\mathbf{1 9 9 7}$ | 145 | 252 | 228 | 215 | 465 |
| $\mathbf{1 9 9 8}$ | 164 | 273 | 218 | 217 | 381 |
| $\mathbf{1 9 9 9}$ | 151 | 273 | 293 | 167 | 416 |
| $\mathbf{2 0 0 0}$ | 160 | 225 | 276 | 281 | 485 |
| $\mathbf{2 0 0 1}$ | 154 | 231 | 202 | 356 | 416 |
| $\mathbf{2 0 0 2}$ | 178 | 208 | 252 | 248 | 382 |
| $\mathbf{2 0 0 3}$ | 205 | 266 | 292 | 240 | 582 |
| $\mathbf{2 0 0 4}$ | 170 | 162 | 285 | 148 | 636 |
| $\mathbf{2 0 0 5}$ | 168 | 200 | 324 | 143 | 336 |
| $\mathbf{2 0 0 6}$ | 251 | 360 | 340 | 315 | 873 |
| $\mathbf{2 0 0 7}$ | 201 | 322 | 359 | 327 | 1325 |
| $\mathbf{2 0 0 8}$ | 205 | 303 | 298 | 383 | 1108 |
| $\mathbf{2 0 0 9}$ | 206 | 351 | 324 | 326 | 1348 |
| $\mathbf{2 0 1 0}$ | 193 | 391 | 448 | 235 | 511 |
| $\mathbf{2 0 1 1}$ | 224 | 390 | 397 | 155 | 187 |

Figure 1. Map of the 2011 Maryland authorized commercial striped bass check stations.


Figure 2. Maryland's Chesapeake Bay pound net and hook-and-line fisheries cumulative striped bass landings from check stations daily call-in reports, June-November 2011.


Figure 3. Maryland's Chesapeake Bay gill net and the Atlantic trawl and gill net fisheries (combined) cumulative striped bass landings from check stations' daily call-in reports, January-December 2011. Note different scales.



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Figure 7. M aryland's A tlantic gill net and trawl fisheries striped bass cat ch (pounds) pertrip (CPUE), 1990-2011. Trips were determined as days fished when striped bass catch was reported.


PROJECT NO. 2<br>JOB NO. 3<br>TASK NO. 5B

# CHARACTERIZATION OF THE STRIPED BASS <br> SPRING RECREATIONAL SEASON <br> AND SPAWNING STOCK IN MARYLAND 

Prepared by Angela Giuliano

## INTRODUCTION

The primary objective of Project 2, Job 3, Task 5B was to characterize the size, age and sex composition of striped bass (Morone saxatilis) s ampled from the 2012 spring recreational season, which be gan on Saturday, April 21 and c ontinued through M ay 15. The s econdary objective was to conduct a dockside creel survey to characterize the a ngler popul ation. D ata collected includes catch and demographic information.

A por tion of $t$ he A tlantic $m$ igratory $s$ triped ba ss $s$ tock $r$ eturns $t o C$ hesapeake $B$ ay annually in the spring to spawn in the various tributaries (Pearson 1938; Merriman 1941; Tresselt 1952; Raney 1952; Raney 1957; Chapoton and Sykes 1961; Dovel 1971; Dovel and Edmunds 1971; Kernehan et al. 1981.). Mansueti and Hollis (1963) reported that the spawning season runs from A pril through J une. After spawning, migratory striped bass leave the tributaries and exit the B ay totheir s ummer f eeding grounds int he A tlantic O cean. Water t emperatures can significantly i nfluence $t$ he ha rvest of m igratory striped bass in any o ne year, with coastal migrants remaining in Chesapeake Bay longer during cool springs (Jones and Sharov 2003). In some years, ripe, pre-spawn females have been captured as late as the end of June and early July (Pearson 1938; Raney 1952; Vladykov and W allace 1952). Increasing water temperatures tend to trigger mi grations out of the B ay and nor thward a long the A tlantic c oast (Merriman 1941;

Raney 1952; Vladykov and Wallace 1952).
Estimates indicate that in the mid-1970s, over $90 \%$ of the coastal striped bass harvested from s outhern Maine to Cape H atteras w ere fish spawned in C hesapeake B ay (Berggren a nd Lieberman 1978; S etzler et al.1980; F ay et a l. 1983). Consequently, s pawning s uccess a nd young-of-year s urvival in the $C$ hesapeake $B$ ay and its tributaries have a significant ef fect on subsequent s triped b ass stock s ize and catch from N orth C arolina to Maine ( Raney 1952; Mansueti 1961; Alperin 1966; Schaefer 1972; Austin and Custer 1977; Fay et al. 1983).

Maryland's post-moratorium spring striped bass season targets coastal migrant fish in the main stem of Chesapeake Bay. The first season opened in 1991 with a 16 -day season, 36 -inch minimum size, a nd a one fish pe r s eason c reel 1 imit (Speir e tal.1999). Spring season regulations ha ve become pr ogressively more liberal s ince 1991 as stock abunda nce i ncreased (Table 1). The 2012 season was 25 days long (April 21 - May 15), with a one fish ( $\geq 28$ inches) per person, perday, creel limit. Fishing was permitted in Chesapeake Bay from B rewerton Channel to the Maryland - Virginia line, excluding all bays and tributaries (Figure 1).

The $M$ aryland Department of $N$ atural $R$ esources (MD D NR) $S$ triped Bass Program initiated a dockside creel survey for the spring fishery in 2002. The main objectives are:

1. Develop at ime s eries of $r$ elative abund ance oft he $C$ hesapeake $B$ ay s pawning $s$ tock harvested during the spring trophy fishery,
2. Determine the sex ratio and spawning condition of harvested fish,
3. Characterize length and weight of harvested fish,
4. Characterize the age-distribution of harvested fish, and
5. Collect scales and otoliths to supplement MD DNR age-length keys and for an ongoing ageing validation study of older fish.

## METHODS

A dockside creel survey was conducted at le ast two days per week at high-use charter boat marinas (Table 2A) with effort focused on collecting biological data on the catch. Because of the half-day structure of some charter trips, charter boats returned in two waves. Return times depended on how fast customers reached the creel daily limit. Charter boats sometimes caught their limit and returned to the dock as early as 10:00 AM. In 2012, many trips did not return to the dock until noon or later while trying to catch their daily creel limit. Sites were not chosen by a true random draw. Biologists arrived at a chosen site between 9:00 and 10:00 AM to intercept the first wave of returning boats. If it be came apparent that fishing activity from that site was minimal (i.e. most charter boats were tied up at the dock), biologists moved to the nearest site in search of higher fishing activity.

Biologists a lternated between five major charter fishing por ts in 20 12: S olomons Island/Calvert M arina, Solomons Island/Beacon M arina, Kentmorr M arina, Chesapeake Beach/Rod \& Reel, and Deale/Happy Harbor (Table 2A). Preference was given to high-use sites to ensure the target of 60 fish per week would be sampled. Geographic coverage was spread out as much a s pos sible between the middle and lower Bay. Biological data were collected from charter boat harvest. Interviews with anglers from charter boats w ere eliminated in 2008 to allow staff more time to survey private boat anglers. Charter boat fishing activity is adequately characterized through the mandated charter logbook system. Charter boat mates, however, were asked how long lines were in the water so that catch rates could be calculated.

A separate creel survey was conducted at public boat ramps to specifically target private boat and shore anglers. Access sites were randomly selected from a list of five public boat ramps (Table 2B). Sites were categorized as high- or medium-use based on the ex periences of creel
interviewers in previous years. High- and medium-use sites were given relative weights of 2:1 for a probability-based random draw. Low-use sites have not been sampled since 2008. Public boat ramps were visited on one randomly selected weekday and one randomly selected weekend day per week. Interviewers were stationed at two sites per selected day and they remained onsite from 10:00 AM-3:00 PM or until 20 trips were intercepted, whichever came first. If no boat trailers were present and no shore anglers were encountered within 2.5 hour s, the sampling day was c oncluded and the site w as c haracterized as having no f ishing a ctivity. Private boat and shore anglers were only interviewed after their trip was completed.

## Biological Data Collection

Biologists a pproached mates of cha rter bo ats and requested $p$ ermission to collect da ta from the catch (Table 3). Total length (mm TL) and weight ( kg ) were measured. The season sampling target for collecting scales was 12 scale samples per 10 mm length group up to 1000 mm T L, for each sex. Scales were col lected from eve ry fish greater than 1000 mm TL . A portion of $t$ hese $s$ cale samples $w$ as used $t$ os upplement s cales c ollected dur ing t he spring spawning stock gill net survey (Project No. 2, Job No. 3, Task No. 2) for the construction of a combined spring a ge-length key. The number of scales read from the creel survey has varied between years. In 2012,85 scale samples were read. The age structure of fish sampled by the creel survey was estimated using the combined spring age-length key.

The season sampling target for otoliths was 2 fish per 10 mm length group greater than or equal to 800 mm TL , for each sex. Otoliths were extracted by using a hacksaw to make a vertical cut from the top of the head above the margin of the pre-operculum down to a level above the eye socket. A second cut was made horizontally from the front of the head above the
eye until it intersected the first cut, exposing the brain. The brain was removed ca refully to expose the sagittal otoliths, which lie below and behind the brain. Otoliths were removed with tweezers and stored dry in labeled plastic vials for later processing.

Spawning condition was determined based on de scriptions of gonad maturity presented by Snyder (1983). Spawning condition was coded as pre-spawn, post-spawn or unknown, and sex was coded as male, female or unknown. "Unknown" for sex or spawning condition refers to fish that were not examined internally, or were not identified with certainty. Ovaries that were swollen and either or ange colored (early phase) or green colored (late phase) indicated a prespawn female. Shrunken ovaries of a darker coloration indicated post-spawn females. Pre- and post-spawn males were more difficult to distinguish. To verify sex and spawning condition of males, pressure was applied to the abdomen to judge the amount of milt expelled, and an incision was made in the abdomen for internal inspection. Those fish yielding large amounts of milt were determined to be pre-spawn. Male fish with flaccid abdomens or that produced only a s mall amount of milt were considered post-spawn.

## Calculation of Harvest and Catch Rates

Survey personnel interviewed private boat and shore anglers to obtain information from which to de velop e stimates of Harvest Per Trip (HPT), Harvest Per A ngler (HPA), Catch Per Trip (CPT), and Catch Per Hour (CPH) (T able 4). The interview que stions a re pr ovided in Appendix I. HPT was defined as the number of fish kept (harvested) for each trip. HPA was calculated by dividing the number of fish $h$ arvested on a trip by the number of anglers in the fishing party. CPT was defined as number of fish kept (harvest), plus number of fish released, for each trip. CPH was calculated by dividing the total catch by the number of hours fished for
each trip.
HPT, HPA and CPT were also calculated from charter boat $\log$ data. CPH was calculated using the charter boat log data and the average duration of charter boat trips from mate interview data. Charter boat captains are required to submit logbooks to MD DNR indicating the days and areas $f$ ished, a nd num bers of $s$ triped ba ss c aught a nd r eleased. In cases w here a c aptain combined data from multiple trips into one log entry, those data were excluded, so only single trip entries were analyzed. Approximately 20\% of the logbook data has been excluded each year using this criterion, but sample sizes have still exceeded 1,000 trips per year. In 2012, $26 \%$ of the logbook data was excluded.

The analysis of charter boat catch rates used a subset of data to include only fishing that occurred in a reas specified in the MD D NR regulations dur ing the spring season (Figure 1). Data from the fisheries in the Susquehanna Flats area were therefore excluded from this analysis.

## RESULTS AND DISCUSSION

The number of private and charter of boats intercepted, number of anglers interviewed, and numbers of striped bass examined each year are presented in Table 5A. In 2012, 172 private boat trips were intercepted for interviews. Fish were sampled from 37 intercepted charter trips (Table 5 B). No shore a nglers with completed trips were intercepted du ring the s pring trophy season. Fishing activity during the spring season was highest in the middle Bay, specifically the region between the Chesapeake Bay Bridge and the mouth of the Patuxent River.

## BIOLOGICAL DATA

## Length and Weight

## Length distribution

The minimum size limit for the 2012 spring striped bass season was 28 inches ( 711 mm ) TL. Lengths ranged from 690 mm TL to 1096 mm T . The c atch was dom inated by fish between 800 and 900 mm TL ( 31 to 35 inches, Figure 2). The majority of fish were smaller in 2011 as demonstrated by a length frequency skewed to the right.

## Mean length

In 2012 , the mean length for all fish ( 863 mm TL ) was significantly s maller than that observed in any year of the survey except 2007 when there was a slot limit (Table 6A, Figure 3). The mean length of females ( 885 mm TL) was greater than the mean length of males ( 795 mm TL), which is typical of the biology of the species. The mean total length of the females was significantly smaller than that observed in 2006 and 2008-2010 but similar to other years. Mean length of males in 2012 was statistically similar to all other years of the survey except for 2002, 2005-2006, and 2008-2009.

The mean daily lengths of female striped bass harvested in 2012 showed no trend as the season progressed (Figure 4). This is in contrast to mean daily length data for 2002 and 2011 and other studies, when larger females were caught e arlier in the season (Goshorn et al.1992, Barker et al. 2003).

## Mean weight

The mean weight of fish sampled in $2012(6.7 \mathrm{~kg})$ was significantly s maller than that observed in all years of the survey except for 2002, 2005, 2007, and 2011 (Table 6B). Based on $95 \%$ c onfidence intervals, the mean weight of females ( 7.2 kg ) w as significantly s maller than

2006 and 2008-2010 but statistically similar in all other years (Figure 5). The mean weight of males $(5.3 \mathrm{~kg})$ in 2012 was the low est in the time series but was statistically similar to those observed in all other study years, except in 2005, 2006, and 2008. The mean weight of females ( 7.2 kg ) was greater than the mean weight of males ( 5.3 kg ), consistent with data from previous years. Females tend to grow larger than males, and most striped bass over $13.6 \mathrm{~kg}(30.0 \mathrm{lb})$ are females (Bigelow and Schroeder 1953).

## Age Structure

The age distribution of striped bass from the sampled harvest in 2012 ranged from 5 to 17 years old (Figure 6). Most fish harvested were between 8 and 11 years ol d. The 2003 (9 years ol din 2012 ) a nd 2004 ( 8 years ol d) year-classes w ere the most frequently obs erved cohorts, each constituting $50 \%$ and $17 \%$ of the sampled harvest, respectively. The strong 2003 year-class has increased annually in the harvest since 2008 and dominated the 2012 harvest with the pr oportion ne arly doubling s ince 1 ast year. The record 1996 y ear-class ( 16 years ol d in 2012), which dominated catches in 2005,2006 , and 2008 , constituted just $0.4 \%$ of the sample harvest.

## Sex Ratio

The da ta i ncluded $t$ hree de signations for $s$ ex: female, $m$ ale a nd unknow $n$. As in past years, the 2012 spring season harvest was dom inated by female striped bass (Table 7A). Sex ratios ( $\%$ of females in the harvest) were calculated us ing three methods: 1 ) including fish of unknown sex in total, 2) using only known-sex fish, and 3) assuming that the unknown fish were female (Table 7B).

Calculation method did not affect the proportion of females in the sampled harvest as there were no fish of unknown sex in 2012. Females constituted $75 \%$ of the sampled harvest. This is one of the lowest proportions of females harvested in the time series, though similar to 2008 and 2009.

## Spawning Condition

## Percent pre-spawn females

The need to unde rstand spawning condition of the female portion of the c atch helped initiate this study in 2002. Goshorn et al. (1992) studied the spawning condition of large female striped bass in the upp er C hesapeake Bay s pawning area dur ing the 1982-1991 s pawning seasons. Their results suggested that most large females spawn before mid-May in the upper Chesapeake B ay s pawning a rea, indicating a hi gh pot ential to ha rvest gravid females in the spring fishery during the first two weeks of May. Data from the 2012 creel survey indicated that $30 \%$ of the females caught between April 21 and May 15 were in pre-spawn condition (Table 8 ). This pe rcentage is lower $t$ han the av erage of the pa st nine years and one of $t$ he $l$ owest percentages in the time series suggesting that most spawning activity was complete prior to the start of the spring season.

## Daily spawning condition of females

Although the percentage of pre-spawn female striped bass appears to increase throughout the s urvey (Figure 7 ), sample s izes w ere v ery s mall. The pe rcent of pre-spawn females harvested ranged from $32 \%$ to $100 \%$ on a ny given day. Sample sizes of female striped bass ranged from 55 female fish on the first day of sampling to zero female fish towards the end of the trophy season (mean=14 fish, median=7 fish). The peak seen on M ay 10 in Figure 7 was
based on j ust two sampled f ish, bot h of w hich were pre-spawn. T he t hree s ample da ys surrounding $t$ his da te (May 7, 11, a nd 14) consisted of 19 f emale f ish, all in post-spawn condition. The low numbers of female fish encountered, especially towards the end of the trophy season, s uggests that s pawning may ha ve o ccurred e arly in A pril prior to the ope ning of the spring fishing season and that the larger migratory fish had already returned to the ocean. This hypothesis is supported by the s pring s pawning s tock survey (Project 2, J ob 3, Task 2) which showed that few fish remained on the spawning grounds past April 21, the opening date of the 2012 spring trophy season.

## CATCH RATES AND FISHING EFFORT

## Harvest Per Trip Unit Effort

Charter boat activity can be accurately characterized from existing reporting methods so no interviews of charter boat anglers were conducted in 2012. Because of increased focus on improving our understanding of private boat fishing effort, all trips intercepted in 2012 for interviews were private boat trips. Creel survey interview data were used to obtain harvest rate estimates for private vessels. Harvest per trip (HPT) was calculated from charter boat logbooks and creel survey interviews using only fish kept during each trip.

The mean HPT in 2012 according to charter boat logbooks was 4.0 fish per trip, the statistically lowest value in the time series (Table 9A). Mean HPT from private boat interviews ( 0.5 fish per trip) was much lower than HPT from charter boats and the lowest private boat HPT in the time series. Though it was statistically similar to the mean private boat HPT from 2002 and 2006-2008, it was significantly less than all other years.

Mean harvest per angler, per trip (HPA) was calculated by dividing the total number of
fish kept on a vessel by the number of people in the fishing party. HPA from charter boat logbook data in 2012 was 0.6 fish per person, significantly lower than all other years (Table 9B). HPA for private anglers, calculated from interview data, was 0.2 fish per person. While the 2012 HPA number is one of the lowest values in the time series and significantly lower than most years, the value is statistically similar to values from 2006-2008 (Table 9B).

## Catch Per Unit Effort

In all years, charter bo ats cau ght m ore f ish per t rip and per hour than pr ivate boa ts (Tables 10 A and 10 B ). The higher charter boat catch rates are likely attributable to the greater level of experience o ft he cha rter bo at c aptains. A lso, charter captains ar ein constant communication amongst themselves, enabling them to better track daily movements and feeding patterns of migratory striped bass and consistently operate near larger aggregations of fish.

In 2012 , private boats caught an average of 0.8 fish per trip, while charter boats caught 4.8 fish per trip. While the 2012 private boat catch per trip was similar to many past years, the charter logbook mean catch per trip was the lowest in the ten year time series. The private boat CPH was 0.2 fish per hour while charter boats had a CPH of 0.9 fish per hour. The 2012 private boat catch per hour was similar to all years except 2004 and the charter boat mean catch per hour was significantly lower than every year other than 2002.

## Mean Daily Catch Per Hour

Anecdotal information from anglers and charter boat captains in most years indicates a decrease in catch rates during the latter portion of the spring season. In 2012, many captains in the lower portion of Maryland's Chesapeake Bay canceled trips towards the end of the season because of the lack of fish. Interview data showed that mean daily CPH declined slightly over
time in some years, but has generally varied without trend since 2002 (Figure 8). Though there were not enough observations to make a definitive conclusion, it appears that daily CPH in 2012 varied without trend. CPH values have de creased since 2007 due to the lack of ch arter boat interview data. Comparing 2008-2012, however, it appears that the 2012 daily CPH values are generally lower than the other years.

## Angler Characterization

States of residence
In 2012, 172 private boat trips were intercepted for interviews and 447 anglers were interviewed during the period April 21-May 15 (Table 5A and Table 5B). Twelve states of residence were represented in 2012 (Table 11). Most anglers were from Maryland (85\%), Virginia (6\%), and Pennsylvania (5\%), similar to previous years.

## Proportion of License Exempt Anglers

Under cu rrent 1 icense $r$ egulations, a pe rson c an purchase a boa t 1 icense which allows anyone aboard the boat to fish without purchasing an individual Maryland tidal fishing license. This cr eates a pot entially significant, but indeterminate amount of unlicensed fishing effort. Consequently, a question was added to the dockside creel survey in 2008 to determine how many anglers on each boat were license-exempt by virtue of the boat license or other reason in order to determine the amount of license-exempt effort during the spring striped bass season. In 2012, there were on average 2.6 anglers per boat and of these anglers, 1.3 were license-exempt (Table 12). These results are remarkably consistent with previous years.

## Number of Lines Fished

In order to determine fishing effort, the number of lines fished was asked in the creel survey in 2006 and 2010-2012. In 2006, six lines were fished on average per private boat and the maximum number encountered on a boat was 15 . In 2012, the average number of lines fished per private boat was seven and ranged from two to 18 lines (Table 13). This was more lines, on average, than in 2006 ( 6 lines) but less than 2010 and 2011. In addition, the range of the number of lines fished was smaller (3-15 lines) in 2006.

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Table 1. History of MD DNR-Fisheries Service regulations for Maryland striped bass spring trophy seasons, 1991-2012.

| Year | Open <br> Season | Min Size <br> Limit (In.) | Bag Limit (\# Fish) | Open Fishing Area |
| :---: | :---: | :---: | :---: | :---: |
| 1991 | 5/11-5/27 | 36 | 1 per person, per season, with permit | Main stem Chesapeake Bay, Annapolis Bay Bridge-VA State line |
| 1992 | 5/01-5/31 | 36 | 1 per person, per season, with permit | Main stem Chesapeake Bay, Annapolis Bay Bridge-VA State line |
| 1993 | 5/01-5/31 | 36 | 1 per person, per season | Main stem Chesapeake Bay, Annapolis Bay Bridge-VA State line |
| 1994 | 5/01-5/31 | 34 | 1 per person, per day, 3 per season | Main stem Chesapeake Bay, Annapolis Bay Bridge-VA State line |
| 1995 | 4/28-5/31 | 32 | 1 per person, per day, 5 per season | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 1996 | 4/26-5/31 | 32 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 1997 | 4/25-5/31 | 32 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 1998 | 4/24-5/31 | 32 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 1999 | 4/23-5/31 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2000 | 4/25-5/31 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2001 | 4/20-5/31 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2002 | 4/20-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2003 | 4/19-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2004 | 4/17-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2005 | 4/16-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2006 | 4/15-5/15 | 33 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2007 | 4/21-5/15 | $\begin{gathered} 28-35 \text { or } \\ \text { larger than } 41 \end{gathered}$ | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2008 | 4/19-5/13 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2009 | 4/18-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2010 | 4/17-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2011 | 4/16-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |
| 2012 | 4/21-5/15 | 28 | 1 per person, per day | Main stem Chesapeake Bay, Brewerton Channel-VA State line |

Table 2A. Survey sites for the Maryland striped bass spring season dockside creel survey, 20022012. Sites are listed in a clockwise direction around Maryland's section of the Chesapeake Bay.

| Region | Site Name | Site Number |
| :--- | :--- | :---: |
| Eastern Shore-Upper Bay | Rock Hall | 01 |
| Eastern Shore-Middle Bay | Matapeake Boat Ramp | 02 |
| Eastern Shore-Middle Bay | Kent Island Marina/Hemingway’s | 15 |
| Eastern Shore-Middle Bay | Kentmorr Marina | 03 |
| Eastern Shore-Middle Bay | Queen Anne Marina | 04 |
| Eastern Shore-Middle Bay | Knapps Narrows Marina | 13 |
| Eastern Shore-Middle Bay | Tilghman Is./Harrison' s | 05 |
| Western Shore-Lower Bay | Pt. Lookout State Park | 16 |
| Western Shore-Lower Bay | Solomons Island Boat Ramp | 17 |
| Western Shore-Lower Bay | Solomons Island/Harbor Marina | 18 |
| Western Shore-Lower Bay | Solomons Island/Beacon Marina | 19 |
| Western Shore-Lower Bay | Solomons Island/Bunky's Charter Boats | 06 |
| Western Shore-Lower Bay | Solomons /Calvert Marina | 07 |
| Western Shore-Middle Bay | Breezy Point Fishing Center and Ramp | 08 |
| Western Shore-Middle Bay | Chesapeake Beach/Rod \& Reel | 09 |
| Western Shore-Middle Bay | Herrington Harbor South | 14 |
| Western Shore-Middle Bay | Deale/Happy Harbor | 10 |
| Western Shore-Middle Bay | South River | 12 |
| Western Shore-Upper Bay | Sandy Pt. State Park Boat Ramp and Beach | 11 |

Table 2B. Survey sites for the Maryland striped bass spring angler-intercept survey, 2012.

| Relative Use | Access Intercept Site |
| :--- | :--- |
| High | Sandy Pt. State Park Boat Ramp and Beach |
|  | Solomons Island Boat Ramp |
|  | Matapeake Boat Ramp |
|  | Breezy Point Fishing Center and Ramp |
|  | Chesapeake Beach Boat Ramp |

Table 3. Biological data collected by the Maryland striped bass spring season creel survey, 2012.

| Measurement or Test | Units or Categories |
| :--- | :--- |
| Total length (TL) | to nearest millimeter (mm) |
| Weight | kilograms $(\mathrm{kg})$ to the nearest tenth |
| Sex | male, female, unknown |
| Spawning condition | pre-spawn, post-spawn, unknown |

Table 4. Angler and catch information collected by the Maryland striped bass spring season creel survey, 2012.

| Angler and Catch Data Collected |
| :--- |
| Number of hours fished |
| Fishing type: private boat or shore |
| Number of anglers on boat |
| Area fished: upper, middle, lower |
| Number of lines fished |
| Number of fish kept |
| Number of fish released |
| Number of anglers license exempt |
| State of residence |

Table 5A. Numbers of trips intercepted, anglers interviewed, and fish examined by the Maryland striped bass spring season creel survey, through May 15.

| Year | Trips Intercepted | Anglers Interviewed | Fish Examined |
| :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 187 | 458 | 503 |
| $\mathbf{2 0 0 3}$ | 181 | 332 | 478 |
| $\mathbf{2 0 0 4}$ | 138 | 178 | 462 |
| $\mathbf{2 0 0 5}$ | 54 | 93 | 275 |
| $\mathbf{2 0 0 6}$ | 139 | 344 | 464 |
| $\mathbf{2 0 0 7}$ | 542 | 809 | 301 |
| $\mathbf{2 0 0 8}$ | 305 | 329 | 200 |
| $\mathbf{2 0 0 9}$ | 303 | 747 | 216 |
| $\mathbf{2 0 1 0}$ | 238 | 601 | 263 |
| $\mathbf{2 0 1 1}$ | 362 | 824 | 234 |
| $\mathbf{2 0 1 2}$ | 209 | 447 | 130 |

Table 5B. Number of trips, by type (fishing mode), intercepted by the Maryland striped bass spring season creel survey, through May 15.

| Year | Charter Boat | Private Boat | Shore | Not Specified | Total |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 140 | 45 | 0 | 2 | 187 |
| $\mathbf{2 0 0 3}$ | 114 | 65 | 0 | 2 | 181 |
| $\mathbf{2 0 0 4}$ | 88 | 42 | 1 | 7 | 138 |
| $\mathbf{2 0 0 5}$ | 53 | 1 | 0 | 0 | 54 |
| $\mathbf{2 0 0 6}$ | 101 | 28 | 10 | 0 | 139 |
| $\mathbf{2 0 0 7}$ | 50 | 483 | 9 | 0 | 542 |
| $\mathbf{2 0 0 8}$ | 34 | 265 | 6 | 0 | 305 |
| $\mathbf{2 0 0 9}$ | 27 | 275 | 1 | 0 | 303 |
| $\mathbf{2 0 1 0}$ | 45 | 193 | 0 | 0 | 238 |
| $\mathbf{2 0 1 1}$ | 63 | 299 | 0 | 0 | 362 |
| $\mathbf{2 0 1 2}$ | 37 | 172 | 0 | 0 | 209 |

Table 6A. Mean lengths of striped bass (mm TL) with $95 \%$ confidence limits sampled by the Maryland striped bass spring season creel survey, through May 15.

| Year | TL (mm) - All fish | TL (mm) - Females | TL (mm) - Males |
| :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | $\mathbf{8 8 7}(879-894)$ | $\mathbf{8 9 5}(886-903)$ | $\mathbf{8 4 6}(828-864)$ |
| 2003 | $\mathbf{8 9 4}(885-903)$ | $\mathbf{8 9 9}(889-909)$ | $\mathbf{8 3 4}(813-864)$ |
| 2004 | $\mathbf{8 8 9}(881-897)$ | $\mathbf{8 9 6}(886-903)$ | $\mathbf{8 2 7}(810-845)$ |
| 2005 | $\mathbf{8 9 3}(885-902)$ | $\mathbf{8 9 8}(888-907)$ | $\mathbf{8 6 7}(852-883)$ |
| 2006 | $\mathbf{9 2 3}(917-930)$ | $\mathbf{9 2 9}(922-936)$ | $\mathbf{8 8 6}(875-897)$ |
| 2007 | $\mathbf{8 6 1}(852-871)$ | $\mathbf{8 6 9}(858-881)$ | $\mathbf{8 2 7}(806-848)$ |
| 2008 | $\mathbf{9 2 0}(910-931)$ | $\mathbf{9 3 3}(922-944)$ | $\mathbf{8 7 7}(853-900)$ |
| 2009 | $\mathbf{9 1 3}(902-925)$ | $\mathbf{9 3 0}(917-942)$ | $\mathbf{8 6 0}(836-883)$ |
| $\mathbf{2 0 1 0}$ | $\mathbf{9 1 3}(902-924)$ | $\mathbf{9 3 2}(921-944)$ | $\mathbf{8 3 3}(812-855)$ |
| $\mathbf{2 0 1 1}$ | $\mathbf{8 9 0}(880-901)$ | $\mathbf{9 0 6}(895-917)$ | $\mathbf{8 2 9}(808-851)$ |
| $\mathbf{2 0 1 2}$ | $\mathbf{8 6 3}(849-876)$ | $\mathbf{8 8 5}(872-899)$ | $\mathbf{7 9 5}(771-818)$ |

Table 6B. Mean weights of striped bass ( kg ) with $95 \%$ confidence limits sampled by the Maryland striped bass spring season creel survey, through May 15.

| Year | Mean Weight (kg) <br> All fish | Mean Weight (kg) <br> Females | Mean Weight (kg) <br> Males |
| :---: | :---: | :---: | :---: |
| 2002 | $7.3(7.1-7.5)$ | $7.4(7.2-7.6)$ | $\mathbf{6 . 1}(5.7-6.4)$ |
| 2003 | $7.6(7.3-7.9)$ | $7.7(7.3-8.0)$ | $5.9(5.2-6.6)$ |
| 2004 | $7.6(7.4-7.8)$ | $7.8(7.5-8.0)$ | $5.9(5.5-6.4)$ |
| 2005 | $7.3(7.1-7.6)$ | $7.5(7.2-7.8)$ | $\mathbf{6 . 4}(6.0-6.7)$ |
| 2006 | $\mathbf{8 . 1}(7.9-8.4)$ | $\mathbf{8 . 3}(8.0-8.5)$ | $\mathbf{6 . 7}(6.4-7.1)$ |
| 2007 | $\mathbf{6 . 8}(6.4-7.1)$ | $\mathbf{7 . 1}(6.7-7.5)$ | $5.7(5.2-6.1)$ |
| 2008 | $7.8(7.5-8.1)$ | $\mathbf{8 . 2}(7.8-8.5)$ | $\mathbf{6 . 7}(6.1-7.2)$ |
| 2009 | $7.9(7.6-8.2)$ | $\mathbf{8 . 3}(8.0-8.7)$ | $\mathbf{6 . 4}(5.8-6.9)$ |
| 2010 | $\mathbf{7 . 8}(7.5-8.1)$ | $\mathbf{8 . 3}(8.0-8.6)$ | $5.7(5.2-6.1)$ |
| 2011 | $7.3(7.0-7.6)$ | $\mathbf{7 . 7}(7.4-8.0)$ | $5.6(5.1-6.1)$ |
| 2012 | $\mathbf{6 . 7}(6.4-7.1)$ | $\mathbf{7 . 2}(6.9-7.6)$ | $\mathbf{5 . 3}(4.7-5.8)$ |

Table 7A. Number of female (F), male (M), and unknown (U) sex striped bass sampled by the Maryland striped bass spring season creel survey, through May 15.

| Year | $\mathbf{F}$ | $\mathbf{M}$ | $\mathbf{U}$ | Total <br> (nclude U) | Total <br> (Exclude U) | F <br> (Assume U were female) |
| :---: | :---: | :---: | ---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 342 | 70 | 92 | 504 | 412 | 434 |
| $\mathbf{2 0 0 3}$ | 404 | 37 | 39 | 480 | 441 | 443 |
| $\mathbf{2 0 0 4}$ | 406 | 45 | 11 | 462 | 451 | 417 |
| $\mathbf{2 0 0 5}$ | 233 | 39 | 3 | 275 | 272 | 236 |
| $\mathbf{2 0 0 6}$ | 393 | 63 | 8 | 464 | 456 | 401 |
| $\mathbf{2 0 0 7}$ | 242 | 49 | 10 | 301 | 291 | 252 |
| $\mathbf{2 0 0 8}$ | 155 | 45 | 0 | 200 | 200 | 155 |
| $\mathbf{2 0 0 9}$ | 166 | 48 | 2 | 216 | 214 | 168 |
| $\mathbf{2 0 1 0}$ | 212 | 50 | 1 | 263 | 262 | 213 |
| $\mathbf{2 0 1 1}$ | 186 | 48 | 0 | 234 | 234 | 186 |
| $\mathbf{2 0 1 2}$ | 98 | 32 | 0 | 130 | 130 | 98 |

Table 7B. Percent females, using three different calculation methods, sampled by the Maryland striped bass spring season creel survey, through May 15.

| Year | \%F <br> (Include U) | \%F <br> (Exclude U) | \%F <br> (Assume U were Female) |
| :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 68 | 83 | 86 |
| $\mathbf{2 0 0 3}$ | 84 | 92 | 92 |
| $\mathbf{2 0 0 4}$ | 88 | 90 | 90 |
| $\mathbf{2 0 0 5}$ | 85 | 86 | 86 |
| $\mathbf{2 0 0 6}$ | 85 | 86 | 86 |
| $\mathbf{2 0 0 7}$ | 80 | 83 | 84 |
| $\mathbf{2 0 0 8}$ | 78 | 78 | 78 |
| $\mathbf{2 0 0 9}$ | 77 | 78 | 78 |
| $\mathbf{2 0 1 0}$ | 81 | 81 | 81 |
| $\mathbf{2 0 1 1}$ | 79 | 79 | 79 |
| $\mathbf{2 0 1 2}$ | 75 | 75 | 75 |
| Mean | $\mathbf{8 0}$ | $\mathbf{8 3}$ | $\mathbf{8 3}$ |

Table 8. Spawning condition of the female portion of catch, sampled by the Maryland striped bass spring season creel survey, through May 15. Females of unknown spawning condition are excluded.

|  | Pre-spawn Females |  | Post-spawn Females |  |
| :--- | :---: | :---: | :---: | :---: |
| Year | $\mathbf{n}$ | $\mathbf{\%}$ | $\mathbf{n}$ | $\mathbf{\%}$ |
| $\mathbf{2 0 0 2}$ | 150 | 45 | 181 | 55 |
| $\mathbf{2 0 0 3}$ | 231 | 58 | 168 | 42 |
| $\mathbf{2 0 0 4}$ | 222 | 55 | 180 | 45 |
| $\mathbf{2 0 0 5}$ | 144 | 63 | 85 | 37 |
| $\mathbf{2 0 0 6}$ | 162 | 41 | 231 | 59 |
| $\mathbf{2 0 0 7}$ | 142 | 59 | 97 | 41 |
| $\mathbf{2 0 0 8}$ | 47 | 30 | 108 | 70 |
| $\mathbf{2 0 0 9} \boldsymbol{2 0 1 0}$ | 81 | 49 | 83 | 50 |
| $\mathbf{2 0 1 1}$ | 62 | 29 | 150 | 71 |
| $\mathbf{2 0 1 2}$ | 79 | 42 | 107 | 58 |
| Mean | $\mathbf{1 2 3}$ | 30 | 69 | 70 |

*Two female fish ( $1 \%$ of females sampled) were of unknown spawning condition.
Table 9A. Mean harvest of striped bass per trip (HPT), with $95 \%$ confidence limits, calculated from Maryland charter boat logbooks and spring season creel survey interview data, through May 15.

| Year | Charter <br> Logbook <br> Trips (n) | Charter <br> Logbook <br> Mean HPT | Private <br> Creel Int. <br> Trips (n) | Private <br> Creel Int. <br> Mean HPT |
| :---: | :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 1,424 | $4.7(4.6-4.8)$ | 44 | $\mathbf{1 . 1}(0.6-1.4)$ |
| $\mathbf{2 0 0 3}$ | 1,393 | $\mathbf{5 . 7}(5.6-5.8)$ | 64 | $\mathbf{1 . 1}(0.7-1.4)$ |
| $\mathbf{2 0 0 4}$ | 1,591 | $\mathbf{5 . 4}(5.3-5.5)$ | 42 | $\mathbf{2 . 2}(1.7-2.8)$ |
| $\mathbf{2 0 0 5}$ | 1,965 | $\mathbf{5 . 5}(5.4-5.6)$ | 1 | $\mathbf{0 . 0}$ |
| $\mathbf{2 0 0 6}$ | 1,934 | $\mathbf{5 . 3}(5.2-5.4)$ | 28 | $\mathbf{1 . 4}(0.6-2.1)$ |
| $\mathbf{2 0 0 7}$ | 1,607 | $\mathbf{4 . 3}(4.2-4.4)$ | 483 | $\mathbf{0 . 7}(0.6-0.8)$ |
| $\mathbf{2 0 0 8}$ | 1,755 | $\mathbf{4 . 9}(4.8-5.1)$ | 260 | $\mathbf{0 . 6}(0.5-0.7)$ |
| $\mathbf{2 0 0 9}$ | 1,849 | $\mathbf{5 . 0}(4.9-5.1)$ | 275 | $\mathbf{0 . 9}(0.7-1.0)$ |
| $\mathbf{2 0 1 0}$ | 1,986 | $\mathbf{4 . 8}(4.7-4.9)$ | 193 | $\mathbf{1 . 1}(0.9-1.3)$ |
| $\mathbf{2 0 1 1}$ | 1,660 | $\mathbf{4 . 8}(4.7-4.9)$ | 298 | $\mathbf{0 . 9}(0.7-1.0)$ |
| $\mathbf{2 0 1 2}$ | 1,127 | $\mathbf{4 . 0}(3.8-4.1)$ | 172 | $\mathbf{0 . 5}(0.3-0.6)$ |

Table 9B. Mean harvest of striped bass per angler, per trip (HPA), with $95 \%$ confidence limits, calculated from Maryland charter boat logbooks and spring season creel survey interview data, through May 15.

| Year | Charter <br> Logbook <br> Trips (n) | Charter <br> Logbook <br> Mean HPA | Private <br> Creel Int. <br> Trips (n) | Private <br> Creel Int. <br> Mean HPA |
| :---: | :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 1,424 | $\mathbf{0 . 7 8}(0.76-0.79)$ | 43 | $\mathbf{0 . 4}(0.3-0.6)$ |
| $\mathbf{2 0 0 3}$ | 1,393 | $\mathbf{0 . 9 3}(0.92-0.94)$ | 64 | $\mathbf{0 . 4}(0.3-0.5)$ |
| $\mathbf{2 0 0 4}$ | 1,591 | $\mathbf{0 . 8 8}(0.86-0.89)$ | 42 | $\mathbf{0 . 7}(0.5-0.8)$ |
| $\mathbf{2 0 0 5}$ | 1,965 | $\mathbf{0 . 8 8}(0.87-0.89)$ | 1 | $\mathbf{0 . 0}$ |
| $\mathbf{2 0 0 6}$ | 1,934 | $\mathbf{0 . 8 6}(0.87-0.85)$ | 27 | $\mathbf{0 . 5}(0.2-0.7)$ |
| $\mathbf{2 0 0 7}$ | 1,607 | $\mathbf{0 . 6 9}(0.68-0.71)$ | 483 | $\mathbf{0 . 3}(0.2-0.3)$ |
| $\mathbf{2 0 0 8}$ | 1,755 | $\mathbf{0 . 7 9}(0.78-0.81)$ | 260 | $\mathbf{0 . 2}(0.2-0.3)$ |
| $\mathbf{2 0 0 9}$ | 1,849 | $\mathbf{0 . 8 1}(0.80-0.82)$ | 275 | $\mathbf{0 . 3}(0.3-0.4)$ |
| $\mathbf{2 0 1 0}$ | 1,986 | $\mathbf{0 . 7 6}(0.75-0.77)$ | 193 | $\mathbf{0 . 4}(0.3-0.5)$ |
| $\mathbf{2 0 1 1}$ | 1,660 | $\mathbf{0 . 7 8}(0.77-0.80)$ | 298 | $\mathbf{0 . 3}(0.3-0.3)$ |
| $\mathbf{2 0 1 2}$ | 1,127 | $\mathbf{0 . 6 4}(0.62-0.66)$ | 172 | $\mathbf{0 . 2}(0.1-0.2)$ |

Table 10A. Private boat mean catch, effort, and catch per hour, with $95 \%$ confidence limits, from the Maryland striped bass spring season creel survey interview data, through May 15. Catch is defined as number of fish harvested plus number of fish released.

| Year | $\mathbf{n}$ | Mean catch/trip | Mean hours/trip | Mean catch/hour |
| :---: | :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 2}$ | 41 | $\mathbf{1 . 6}(0.9-2.4)$ | $4.9(4.3-5.5)$ | $\mathbf{0 . 3}(0.2-0.5)$ |
| $\mathbf{2 0 0 3}$ | 63 | $\mathbf{1 . 8}(0.9-2.8)$ | $5.4(4.8-6.0)$ | $\mathbf{0 . 5}(0.2-0.7)$ |
| $\mathbf{2 0 0 4}$ | 42 | $\mathbf{3 . 5}(2.0-4.9)$ | $\mathbf{4 . 6}(3.8-5.3)$ | $\mathbf{1 . 0}(0.6-1.4)$ |
| $\mathbf{2 0 0 5}$ | 1 | $\mathbf{0 . 0}$ | $\mathbf{2 . 5}$ | $\mathbf{0 . 0}$ |
| $\mathbf{2 0 0 6}$ | 28 | $\mathbf{2 . 3}(1.1-3.5)$ | $\mathbf{4 . 9}(4.2-5.7)$ | $\mathbf{0 . 7}(0.3-1.1)$ |
| $\mathbf{2 0 0 7}$ | 483 | $\mathbf{1 . 6}(1.2-2.0)$ | $5.0(4.9-5.1)$ | $\mathbf{0 . 3}(0.2-0.4)$ |
| $\mathbf{2 0 0 8}$ | 260 | $\mathbf{1 . 0}(0.7-1.3)$ | $\mathbf{4 . 5}(4.2-4.7)$ | $\mathbf{0 . 3}(0.2-0.4)$ |
| $\mathbf{2 0 0 9}$ | 275 | $\mathbf{1 . 6}(1.0-2.1)$ | $4.7(4.5-4.8)$ | $\mathbf{0 . 4}(0.2-0.5)$ |
| $\mathbf{2 0 1 0}$ | 193 | $\mathbf{1 . 6}(1.2-2.0)$ | $\mathbf{4 . 7}(4.5-4.9)$ | $\mathbf{0 . 4}(0.3-0.5)$ |
| $\mathbf{2 0 1 1}$ | 298 | $\mathbf{1 . 2}(1.0-1.4)$ | $\mathbf{4 . 4}(4.2-4.6)$ | $\mathbf{0 . 3}(0.2-0.4)$ |
| $\mathbf{2 0 1 2}$ | 172 | $\mathbf{0 . 8}(0.5-1.1)$ | $\mathbf{4 . 8}(4.6-5.1)$ | $\mathbf{0 . 2}(0.1-0.3)$ |

Table 10B. Charter boat mean catch, effort, and catch per hour, with $95 \%$ confidence limits, calculated from logbook data, through May 15. Catch is defined as number of fish harvested plus number of fish released. Mean hours per trip are from creel survey interview data until 2009 where the mean hours per trip are from mate interviews.

| Year | $\mathbf{n}$ | Mean catch/trip | Mean hours/trip <br> (From creel interview data) | Mean <br> catch/hour |
| :---: | :---: | :---: | :---: | :---: |
| 2002 | 1,487 | $5.5(5.4-5.7)$ | $5.5(5.3-5.7)$ | $\mathbf{1 . 0}(0.9-1.1)$ |
| 2003 | 1,420 | $7.3(7.0-7.6)$ | $\mathbf{4 . 0}(3.7-4.4)$ | $\mathbf{1 . 8}(1.7-1.9)$ |
| 2004 | 1,629 | $7.4(7.0-7.7)$ | $\mathbf{4 . 0}(3.6-4.4)$ | $\mathbf{1 . 8}(1.7-1.9)$ |
| 2005 | 1,994 | $\mathbf{6 . 9}(6.6-7.1)$ | $\mathbf{3 . 1}(2.6-3.5)$ | $\mathbf{2 . 2}(2.1-2.3)$ |
| 2006 | 1,990 | $\mathbf{8 . 0}(7.7-8.2)$ | $\mathbf{3 . 6}(3.2-3.9)$ | $\mathbf{2 . 2}(2.1-2.3)$ |
| 2007 | 1,793 | $\mathbf{8 . 1}(7.8-8.4)$ | $\mathbf{4 . 6}(4.1-5.0)$ | $\mathbf{1 . 8}(1.7-1.8)$ |
| 2008 | 1,755 | $\mathbf{6 . 4}(6.2-6.6)$ | $\mathrm{N} / \mathrm{A}$ | $\mathrm{N} / \mathrm{A}$ |
| $\mathbf{2 0 0 9}$ | 1,849 | $\mathbf{6 . 0}(5.9-6.2)$ | $\mathbf{3 . 4}(2.9-4.0)$ | $\mathbf{1 . 8}(1.7-1.8)$ |
| $\mathbf{2 0 1 0}$ | 1,986 | $5.7(5.5-5.8)$ | $\mathbf{4 . 4}(4.0-4.9)$ | $\mathbf{1 . 3}(1.2-1.3)$ |
| $\mathbf{2 0 1 1}$ | 1,660 | $\mathbf{5 . 7}(5.5-5.8)$ | $\mathbf{4 . 2}(3.5-4.9)$ | $\mathbf{1 . 3}(1.3-1.4)$ |
| $\mathbf{2 0 1 2}$ | 1,127 | $\mathbf{4 . 8}(4.6-5.0)$ | $\mathbf{5 . 5}(4.9-6.1)$ | $\mathbf{0 . 9}(0.8-0.9)$ |

Table 11. State of residence and number of anglers interviewed by the Maryland striped bass spring season creel survey, through May 15.

| State of <br> residence | $\mathbf{2 0 0 2}$ | $\mathbf{2 0 0 3}$ | $\mathbf{2 0 0 4}$ | $\mathbf{2 0 0 5}$ | $\mathbf{2 0 0 6}$ | $\mathbf{2 0 0 7}$ | $\mathbf{2 0 0 8}$ | $\mathbf{2 0 0 9}$ | $\mathbf{2 0 1 0}$ | $\mathbf{2 0 1 1}$ | $\mathbf{2 0 1 2}$ |
| :---: | :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| AL | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| AZ | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| CA | 1 | 0 | 1 | 0 | 0 | 2 | 0 | 0 | 0 | 0 | 0 |
| CO | 0 | 0 | 1 | 0 | 1 | 1 | 0 | 0 | 1 | 0 | 1 |
| DC | 6 | 1 | 1 | 0 | 1 | 2 | 1 | 0 | 6 | 1 | 0 |
| DE | 6 | 7 | 3 | 0 | 9 | 8 | 1 | 0 | 3 | 1 | 2 |
| FL | 0 | 0 | 1 | 1 | 2 | 0 | 1 | 0 | 3 | 1 | 0 |
| GA | 1 | 1 | 0 | 2 | 2 | 0 | 0 | 0 | 0 | 0 | 0 |
| IL | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| KY | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| KS | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| MA | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 1 | 0 | 0 |
| MD | 353 | 260 | 107 | 66 | 227 | 679 | 266 | 651 | 482 | 491 | 381 |
| MI | 1 | 0 | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 0 |
| MN | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 4 | 0 |
| MT | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 2 | 0 | 0 |
| NC | 0 | 2 | 0 | 1 | 0 | 1 | 1 | 0 | 0 | 0 | 3 |
| NJ | 2 | 2 | 6 | 0 | 3 | 2 | 4 | 0 | 0 | 1 | 3 |
| NY | 4 | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 1 | 1 | 0 |
| OH | 0 | 0 | 0 | 0 | 0 | 3 | 1 | 0 | 1 | 0 | 1 |
| PA | 27 | 19 | 17 | 4 | 22 | 32 | 16 | 46 | 18 | 19 | 23 |
| RI | 2 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 |
| SC | 0 | 0 | 1 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| TN | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 |
| TX | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 |
| VA | 48 | 31 | 30 | 13 | 56 | 71 | 29 | 44 | 42 | 23 | 26 |
| WA | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| WI | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| WV | 0 | 1 | 0 | 2 | 6 | 3 | 2 | 4 | 4 | 0 | 4 |
| Outside U.S. | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| Unknown | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 36 | 0 | 0 |

Table 12. The average number of anglers and average number of unlicensed anglers, per boat, with $95 \%$ confidence intervals, from the 2008-2012 Maryland striped bass spring season creel survey interview data.

| Year | Number of Trips <br> Interviewed | Average Number of <br> Anglers per Boat | Average Number of <br> Unlicensed Anglers per Boat |
| :---: | :---: | :---: | :---: |
| 2008 | 261 | $\mathbf{2 . 8}(2.7-2.9)$ | $\mathbf{1 . 5}(1.3-1.6)$ |
| 2009 | 276 | $\mathbf{2 . 7}(2.6-2.8)$ | $\mathbf{1 . 3}(1.2-1.5)$ |
| 2010 | 193 | $\mathbf{2 . 8}(2.6-2.9)$ | $\mathbf{1 . 4}(1.2-1.5)$ |
| 2011 | 298 | $\mathbf{2 . 7}(2.6-2.9)$ | $\mathbf{1 . 5}(1.3-1.6)$ |
| 2012 | 172 | $\mathbf{2 . 6}(2.4-2.8)$ | $\mathbf{1 . 3}(1.1-1.5)$ |

Table 13. Number of lines fished by private boats.

| Year | Minimum | Maximum | Mean |
| :---: | :---: | :---: | :---: |
| $\mathbf{2 0 0 6}$ | 3 | 15 | 6 |
| $\mathbf{2 0 1 0}$ | 1 | 19 | 8 |
| $\mathbf{2 0 1 1}$ | 2 | 22 | 8 |
| $\mathbf{2 0 1 2}$ | 2 | 18 | 7 |

Figure 1. MD DNR map showing legal open and closed striped bass fishing areas in Chesapeake Bay during the spring season, April 21-May 15, 2012.


* Note: The text on the map refers to the dates catch and release fishing is allowed on the Susquehanna Flats prior to the area closure May 4-15, not the dates the spring trophy fishery is open.

Figure 2. Length distribution of striped bass sampled by year, during the Maryland striped bass spring season creel survey, through May 15.


Figure 2. Continued.


Figure 3. Mean length of striped bass (mm TL) with $95 \%$ confidence intervals, sampled by the Maryland striped bass spring season creel survey, through May 15.


Figure 4. Mean daily length of female striped bass with $95 \%$ confidence intervals, sampled by the Maryland striped bass spring season creel survey, through May 15.


Figure 5. Mean weight of striped bass (kg) with $95 \%$ confidence intervals, sampled by the Maryland striped bass spring season creel survey, through May 15.


Figure 6. Age distribution of striped bass sampled by the Maryland striped bass spring season creel survey, through May 15.


Figure 6. Continued.


Figure 7. Daily percent of female striped bass in pre-spawn condition sampled by the Maryland striped bass spring season creel survey, through May 15.


Figure 7. Continued.



Figure 8. Daily mean catch per hour (CPH) of striped bass with $95 \%$ confidence intervals, calculated from angler interview data collected by the Maryland striped bass spring season creel survey, through May 15. Note different scale since 2008.


## APPENDIX I

INTERVIEW FORMAT AND QUESTIONS
MARYLAND STRIPED BASS SPRING SEASON CREEL SURVEY MARYLAND DEPARTMENT OF NATURAL RESOURCES-FISHERIES SERVICE
1.) How many anglers were on your boat today?
2.) How many striped bass were kept by your party?
3.) How many striped bass were released by your party?
4.) How many hours did you fish today? (Line in until Lines out)
5.) How many lines were you fishing?
6.) Where did you spend most of your time fishing today? $\mathbf{U}, \mathbf{M}$, or $\mathbf{L}$ Bay: Upper Bay = above Bay Bridge, Middle Bay = Bay Bridge to Cove Pt., Lower Bay = Cove Pt. to MD/VA line at Smith Pt.
7.) What is your state of residence?
8.) a. Do you have a boat license?
b. How many anglers in your party were fishing under the boat license? (Or, how many anglers in the party have their own individual licenses?)

# PROJECT NO. 2 

## JOB NO. 4

# INTER-GOVERNMENT COORDINATION 

Prepared by Harry T. Hornick and Eric Q. Durell

The objective of Job 4 was to document and summarize participation of Survey personnel in various research and management forums regarding fifteen resident and migratory finfish species found in Maryland's Chesapeake Bay. W ith the pa ssage of t he A tlantic C oastal F isheries Cooperative Management Act, various management entities such as the A tlantic States Marine Fisheries C ommission (ASMFC), the M id-Atlantic M igratory F ish C ouncil (MAMFC), the Chesapeake B ay L iving R esources S ubcommittee ( CBLRS), the Potomac River Fisheries Commission (PRFC), a nd the $S$ usquehanna $R$ iver A nadromous $F$ ish $R$ estoration $C$ ooperative (SRAFRAC), require current stock assessment information in order to assess management measures. The Survey staff also participated in ASMFC, US Fish and Wildlife Service (USFWS) and National Marine Fisheries Service (NMFS) fishery research and management forums.

Direct participation by Survey personnel as representatives to various management entities provided effective representation of Maryland interests through the development, implementation and refinement of management options for Maryland as well as coastal fisheries management plans. In addition, survey information was used to formulate management plans for thirteen finfish species as well as providing evidence of compliance with state and federal regulations. A summary of this participation and contributions is presented below.

## Atlantic menhaden:

Project staff provided Atlantic menhaden data utilized for stock assessments, FMP's and shared coastal management activities with ASMFC, NMFS, USFWS and various academic institutions.

## Alosines:

Project staff attended SRAFRC meetings as Maryland representatives to discuss American shad and river herring stock status, restoration, and management in the Susquehanna River.

ASMFC Technical Committee representative attended the American shad Technical Committee meetings to approve the annual state compliance report, examine the current population abundance estimates and discuss the ocean and river-specific fisheries, and prepared the Annual American Shad Status Compliance Report for Maryland.

## Bluefish:

The ASMFC Bluefish Technical Committee representative provided Chesapeake Bay juvenile bluefish data to the ASMFC and the Mid-Atlantic Fishery Management Council.

ASMFC Technical Committee representative prepared the Annual Bluefish Status Compliance Report for Maryland.

## Red Drum:

ASMFC Technical Committee representative prepared the Annual Red Drum Status Compliance Report for Maryland.

## Weakfish:

ASMFC Weakfish Technical Committee representative for Maryland attended annual Weakfish Technical Committee meetings and prepared the ASMFC Annual Weakfish Status Compliance report

## Striped Bass:

Project staff served on the A SMFC Striped Bass Tagging Sub Committee, the Interstate Tagging C ommittee, the ASMFC Bluefish Technical C ommittee, and as M aryland representatives to the Potomac River Fisheries Commission (PRFC) Finfish Advisory Board.

Project staff served as Maryland alternate representatives to the A SMFC S triped Bass Scientific and Statistical Committee, the Striped Bass Stock Assessment Subcommittee, and produced Maryland's Annual Striped Bass Compliance Report.

## Striped Bass Data Sharing and Web Page Development

To augment data sharing e fforts, Striped B ass S tock A ssessment program staff in 2002 developed a web page within the MD DNR web site presenting historic Juvenile Striped Bass Survey (Job 3) results. This effort has enabled the public to access SBSA program data directly. The web page, http://www.dnr.maryland.gov/fisheries/Pages/striped-bass/juvenile-index.aspx, is updated a nnually in October.

Monthly individual visits to the Juvenile Striped Bass Survey web page by individual IP address for the period July 2012 to January 2013 are provided in Table 1. Because of a change in MD DNR Information Technology Service policy and data management, and incorporation of a new server, web site visit statistics from January 27, 2012 to July 12, 2012 were not available.

An increase in volume in October 2012 coincided with publication of the juvenile survey results in the media and advertisement on the main Fisheries Service page. Many large or complex data requests are still handled directly by Striped Bass Stock Assessment Program staff. However, the web page has saved staff a considerable amount of time answering basic and redundant data requests.

Table 1. Monthly visits to the Juvenile Striped Bass Survey web page, July 13, 2012 January 12, 2013.

| Date | Visits |
| :--- | ---: |
| July 13, 2012 -Aug. 12, 2012 | 260 |
| Aug. 13, 2012 - Sept. 12, 2012 | 275 |
| Sept. 13, 2012 - Oct. 12, 2012 | 447 |
| Oct.13, 2012 -Nov. 12, 2012 | 249 |
| Nov. 13, 2012-Dec. 12, 2012 | 187 |
| Dec. 12, 2012-Jan. 12, 2013 | 175 |
| TOTAL | 1593 |

Project staff also provided Maryland striped bass data and biological samples such as scale and finfish samples, to other state, federal, private and academic researchers. These included the National M arine F isheries S ervice (NMFS), US Fish and Wildlife S ervice (USFWS), Duke University, the University of Maryland, University of Massachusetts, Virginia Institute of Marine Sciences, Georgetown University, the Pennsylvania State University, Syracuse University, and State management agencies from Delaware, Massachusetts, New York and Virginia. For the past contract year, (November 1, 2011 through October 31, 2012) the following specific requests for information have been accommodated:
-Mr. A.C. Carpenter, Potomac River Fisheries Commission (PRFC).
Provision of striped bass juvenile survey data commercial harvest regulations.
-Ms. Emily Argo, Duke University (PRFC).
Provision of biological samples and data from the Juvenile Striped Bass Survey.
-Dr. Robert Aguilar, Smithsonian Environmental Research Center (SERC). Provided biological samples and data from the Juvenile Striped Bass Survey.
-Atlantic States Marine Fisheries Commission (ASMFC).
Provision of striped bass juvenile index data; updated striped bass fishery regulations; striped bass commercial fishery data, striped bass spawning stock CPUE data; current striped bass commercial fishery data; results f rom f ishery de pendent m onitoring pr ograms, a nd age/length keys developed from results of fishery monitoring programs.
-Dr. Trevor Avery, Dept. of Biology, Acadia University, Nova Scotia, Canada. Provided striped bass juveniles and the striped bass juvenile index data set
-Mr. Jim Cummins, Pennsylvania Fish and Boat Commission.
Provided American Shad data from the Juvenile Striped Bass Survey.
-Ms. Cassie Gurbiz, University of Maryland, Horn Point Laboratory. Provided striped bass data from the Juvenile Striped Bass Survey.

- Maryland Charterboat Association (MCA)

Provision of striped ba ss fishery regulations, striped ba ss recreational, and charter boat harvest data.
-Interstate Commission for the Potomac River Basin,( ICPRB). Provision of current striped bass recreational, charter, and commercial fishery data, and American shad and striped bass juvenile survey data.
-Dr. Matthew Hamilton, Georgetown University.
Provision of juvenile striped bass biological samples for genetic research and abundance indices.
-Dr. John Harrison, The Pennsylvania State University.
Provision of striped bass juvenile survey data and striped bass recreational and commercial fishery data.
-Mr. Ken Hastings.
Provided striped bass commercial fishery monitoring information, striped bass recreational survey data, and ASMFC Striped Bass Compliance Report information.
-Dr. Desmond Kahn, Delaware Division of Fish and Wildlife.
Provision of historic Striped Bass Juvenile Survey data.

- National Marine Fisheries Service, NOAA, Chesapeake Bay Program Staff.

Provision of results from fishery dependent monitoring programs, striped bass juvenile index data, and Atlantic menhaden juvenile survey data.
-Mr. Rob O’Reilly, Virginia Marine Resources Commission.
Provision of c urrent a nd hi storical striped ba ss c ommercial fishery data; Striped bass Voluntary Angler Survey data, results of fishery dependent monitoring programs and striped bass juvenile survey data.
-Mr. Jason Schaffler, Old Dominion University.
Provision of juvenile Atlantic menhaden biological samples and abundance indices.
-Dr. Amy Schueller, NMFS, SEFSC.
Provision of historic data from the Juvenile Striped Bass Survey
-Ms. Sara Turner, Syracuse University.
Provision of biological samples and data from the Juvenile Striped Bass Survey
-University of Maryland (U MD - CEES), Chesapeake Biological Laboratory and Horn Point Environmental Laboratory.
Provided six (6) staff and students with current striped bass juvenile index data, American shad juvenile index data, recreational and commercial landings data, and biological samples.
-Ms. Allison Watts, Virginia Marine Resources Commission..
Provision of data from the Juvenile Striped Bass Survey, MD Volunteer Angler Survey, and commercial fishery monitoring data.
-The Interjurisdictional Project also provided related biological information and reports to thirty three (33) additional scientists, students and concerned stakeholders.


[^0]:    * Mean weights-at-age were calculated based on the age-length key and length and weight measurements of individual fish.

[^1]:    * Sum of columns may not equal totals due to rounding.

[^2]:    * Sum of columns may not equal totals due to rounding.

[^3]:    * Sum of columns may not equal totals due to rounding.

[^4]:    * Sum of columns may not equal totals due to rounding.

[^5]:    * Indicates auxiliary seining site

[^6]:    * Indicates auxiliary seining site

